## MSAN 504 — Probability/Statistics — Summer 2016 Homework Five

- 1. Let  $X_1, ... X_n$  be independent exponential random variables with **different** means  $1/\lambda_1, 1/\lambda_2, ..., 1/\lambda_n$ . Find the cumulative distribution function of  $Z_1 = \min\{X_1, ..., X_n\}$  and  $Z_2 = \max\{X_1, ..., X_n\}$ . (**Hint:** Use the definition of the cumulative distribution function!)
- 2. Let  $X_1,...X_n$  be independent geometric random variables with common success parameter p. Find the cumulative distribution function of  $Z_1 = \min\{X_1,...,X_n\}$  and  $Z_2 = \max\{X_1,...,X_n\}$ .
- 3. Let X and Y be continuous random variables with joint probability density function given by f(x,y) = x + y if  $0 \le x \le 1$  and  $0 \le y \le 1$  and 0 otherwise. Calculate  $f_{X|Y}(x|y)$ , compute  $\mathbb{P}(0.25 < X < 0.5|Y = 0.2)$ , and then determine  $\mathbb{E}[X|Y = y]$ .
- 4. Let the joint probability density of X and Y be bivariate normal. For what values of  $\alpha$  is the variance of  $\alpha X + Y$  minimized?
- 5. Let X be the height of a man and Y the height of his daughter (both in inches). Suppose that the joint probability density function of X and Y is bivariate normal with the following parameters:  $\mu_X = 71$ ,  $\mu_Y = 60$ ,  $\sigma_X = 3$ ,  $\sigma_Y = 2.7$ , and  $\rho = 0.45$ . Find the probability that the height of the daughter, of a man who is 70 inches tall, is at least 59 inches.
- 6. Let **X** be a multivariate Gaussian random vector with mean vector  $\boldsymbol{\mu} = [2, -3, 1]$  and variance-covariance matrix  $\boldsymbol{\Sigma}$ , which is given by

$$\left(\begin{array}{ccc}
1 & 1 & 1 \\
1 & 3 & 2 \\
1 & 2 & 2
\end{array}\right)$$

Suppose that  $Y = 3X_1 - 2X_2 + X_3$ , where  $X_i$  is the  $i^{th}$  component of **X**. What is the distribution of Y?

- 7. Let  $X_1, ..., X_n$  be a random sample from a normal distribution with mean  $\mu$  and variance  $\sigma^2$ . Solve the MLE problem simultaneously for both  $\mu$  and  $\sigma$ .
- 8. Show that the two components of your answer to problem #7, i.e.,  $\hat{\mu}$  and  $\hat{\sigma}^2$ , are independent. (Arguably, your answer to problem #9 will immediately solve #8, but in case you are unable to do problem #9, there is another way to do problem #8.)
- 9. Use the theorem we stated at the end of class on Tuesday to determine the asymptotic distribution of the vector  $(\hat{\mu}, \hat{\sigma}^2)$ . You may need to look up a clean version of the multidimensional verson of this theorem. **Hint:** In this case, the inverse of the Fisher information

will be the inverse of a particular matrix. The form of that matrix is given by the picture below! You'll need to compute a matrix of second-order partial derivatives of the log-likelihood function, take the expected values of the entries in that matrix after plugging in the random vector  $\mathbf{X}$ , multiply by -1, and then do a matrix inversion.

## **THEOREM 14.4** Cramér-Rao Lower Bound

Assuming that the density of  $y_i$  satisfies the regularity conditions R1–R3, the asymptotic variance of a consistent and asymptotically normally distributed estimator of the parameter vector  $\theta_0$  will always be at least as large as

$$[\mathbf{I}(\boldsymbol{\theta}_0)]^{-1} = \left(-E_0 \left[ \frac{\partial^2 \ln L(\boldsymbol{\theta}_0)}{\partial \boldsymbol{\theta}_0 \, \partial \boldsymbol{\theta}_0'} \right] \right)^{-1} = \left(E_0 \left[ \left( \frac{\partial \ln L(\boldsymbol{\theta}_0)}{\partial \boldsymbol{\theta}_0} \right) \left( \frac{\partial \ln L(\boldsymbol{\theta}_0)}{\partial \boldsymbol{\theta}_0} \right)' \right] \right)^{-1}.$$