

5. Bivariate and vector random variables

- If we have a collection of random variables Y_1, Y_2, \dots, Y_n we can gather them together into a vector random variable \mathbf{Y} .
- Suppose that, for each $i = 1, \dots, n$ we have $E[Y_i] = \mu_i$. Then, we write $E[\mathbf{Y}] = \boldsymbol{\mu}$ for $\boldsymbol{\mu} = (\mu_1, \dots, \mu_n)$.
- Now, write $\text{Cov}(Y_i, Y_j) = V_{ij}$ for $i \neq j$ and $\text{Var}(Y_i) = \text{Cov}(Y_i, Y_i) = V_{ii}$. We call $\mathbb{V} = [V_{ij}]_{n \times n}$ the **variance-covariance matrix** for \mathbf{Y} .
- We can also call \mathbb{V} the **covariance matrix** or, more simply, just the **variance matrix**. We write $\mathbb{V} = \text{Var}(\mathbf{Y})$.

Example. Let $\boldsymbol{\epsilon} = (\epsilon_1, \epsilon_2, \dots, \epsilon_n)$ be a vector consisting of n independent random variables, each with mean zero and variance σ^2 . This is a common model for **measurement error** on n measurements. We have

$$E[\boldsymbol{\epsilon}] = \mathbf{0}, \quad \text{Var}(\boldsymbol{\epsilon}) = \sigma^2 \mathbb{I}$$

where $\mathbf{0} = (0, \dots, 0)$ and \mathbb{I} is the $n \times n$ identity matrix. The off-diagonal entries of $\text{Var}(\boldsymbol{\epsilon})$ are zero since $\text{Cov}(\epsilon_i, \epsilon_j) = 0$ for $i \neq j$. For measurement error models, we break our usual rule of using upper case letters for random variables.

Example. A population version of the linear model

- First recall the sample version, which is

$$(LM3) \quad \mathbf{y} = \mathbb{X} \mathbf{b} + \mathbf{e},$$

where \mathbf{y} is the measured response, \mathbb{X} is an $n \times p$ matrix of explanatory variables, \mathbf{b} is chosen by least squares, and \mathbf{e} is the resulting vector of residuals.

- We want to build a random vector \mathbf{Y} that provides a population model for the data \mathbf{y} . We write this as

$$(LM6) \quad \mathbf{Y} = \mathbb{X} \boldsymbol{\beta} + \boldsymbol{\epsilon}$$

where \mathbb{X} is the same explanatory matrix as in (LM3), $\boldsymbol{\beta} = (\beta_1, \dots, \beta_p)$ is an unknown coefficient vector (we don't know the true population coefficient!) and $\boldsymbol{\epsilon}$ is measurement error with $\mathbf{E}[\boldsymbol{\epsilon}] = \mathbf{0}$ and $\text{Var}(\boldsymbol{\epsilon}) = \sigma^2 \mathbb{I}$.

- Our model (LM6) asserts that the process which generated the response data \mathbf{y} was like drawing a random vector \mathbf{Y} constructed using a random measurement error model with known matrix \mathbb{X} for some fixed but unknown value of $\boldsymbol{\beta}$.

Mean and variance of the least squares estimate for

$$\mathbf{Y} = \mathbb{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$$

- Recall that the main purpose of having a probability model is so that we can investigate the chance variation due to picking the sample.
- Recall that for (LM3), the least squares estimate is $\mathbf{b} = (\mathbb{X}^T \mathbb{X})^{-1} \mathbb{X}^T \mathbf{y}$.
- This is a **statistic**, which means a function of the data and not a random variable. We cannot properly talk about the mean and variance of \mathbf{b} .
- We can work out the mean and variance of $(\mathbb{X}^T \mathbb{X})^{-1} \mathbb{X}^T \mathbf{Y}$, as long as we know how to work out the mean and variance of linear combinations.
- As long as $\mathbf{Y} = \mathbb{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$ is a good **probability model** for the relationship between the response variable \mathbf{y} and the explanatory variable \mathbb{X} , calculations done with this model may be useful.

Mean of a linear combination, in matrix form

- The linear property of expectation lets us take expectation through a summation sign, and we get

$$\mathbb{E} \left[\sum_{j=1}^n a_{ij} Y_j \right] = \sum_{j=1}^n a_{ij} \mathbb{E}[Y_j].$$

- In matrix form, with $\mathbb{A} = [a_{ij}]$, this is

$$\mathbb{E}[\mathbb{A}\mathbf{Y}] = \mathbb{A}\mathbb{E}[\mathbf{Y}].$$

Example. For $\mathbf{Y} = \mathbb{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$,

$$\mathbb{E}[\mathbf{Y}] = \mathbb{X}\boldsymbol{\beta} + \mathbb{E}[\boldsymbol{\epsilon}] = \mathbb{X}\boldsymbol{\beta}$$

Example. $\hat{\boldsymbol{\beta}} = (\mathbb{X}^T \mathbb{X})^{-1} \mathbb{X}^T \mathbf{Y}$, we have