A useful variance-stabilizing transformation - ln y.

The data in https://netfiles.uiuc.edu/stepanov/www/Initech.csv are the salaries, y, and years of experience, x, for a sample of 50 Initech employees.

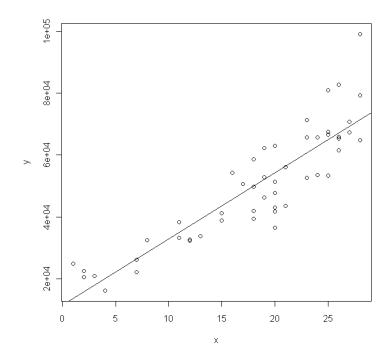
> Initech

	X	У
1	7	26075
2	28	79370
3	23	65726
4	18	41983
5	19	62308
6	15	41154
7	24	53610
8	13	33697
9	2	22444
10	8	32562
11	20	43076
12	21	56000
13	18	58667
14	7	22210
15	2	20521
16	18	49727
17	11	33233
18	21	43628
19	4	16105
20	24	65644
21	20	63022
22	20	47780
23	15	38853
24	25	66537
25	25	67447
26	28	64785
27	26	61581
28	27	70678
29	20	51301
30	18	39346

31 1 24833 32 26 65929 33 20 41721

First, consider the first-order model $Y = \beta_0 + \beta_1 x_1 + \epsilon$.

- $> fit1 = lm(y \sim x)$
- > plot(x, y)
- > abline(fit1\$coefficients)



> summary(fit1)

Call:

 $lm(formula = y \sim x)$

Residuals:

Min 1Q Median 3Q Max -17665.6 -5497.7 -725.7 4667.3 27812.9

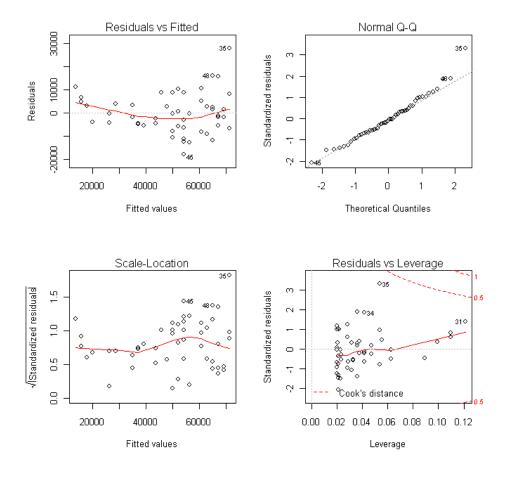
Coefficients:

Estimate Std. Error t value Pr(>|t|)
(Intercept) 11369.4 3160.2 3.598 0.000757 ***
x 2141.3 160.8 13.314 < 2e-16 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 8642 on 48 degrees of freedom Multiple R-Squared: 0.7869, Adjusted R-squared: 0.7825 F-statistic: 177.3 on 1 and 48 DF, p-value: < 2.2e-16

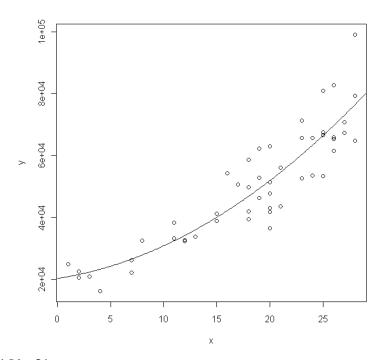
> par(mfrow=c(2,2)) > plot(fit1)



Consider the second-order model $Y = \beta_0 + \beta_1 x_1 + \beta_2 x_1^2 + \epsilon$.

```
> fit2 = lm(y \sim x + I(x^2))
```

- > plot(x, y)
- > xx = seq(0,30,by=0.1)
- > yy = fit2\$coeff[1] + fit2\$coeff[2]*xx + fit2\$coeff[3]*xx^2
- > lines(xx,yy)



> summary(fit2)

Call:

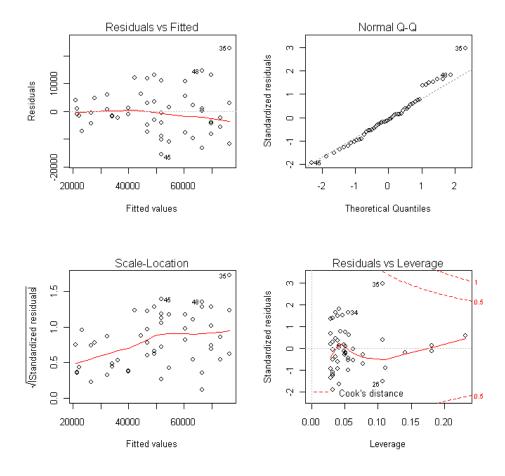
$$lm(formula = y \sim x + I(x^2))$$

Residuals:

Coefficients:

Residual standard error: 8123 on 47 degrees of freedom Multiple R-Squared: 0.8157, Adjusted R-squared: 0.8078 F-statistic: 104 on 2 and 47 DF, p-value: < 2.2e-16

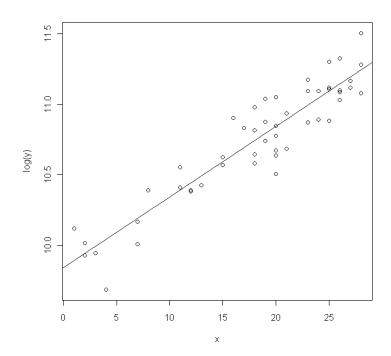
> plot(fit2)



A useful **variance-stabilizing transformation** – ln y.

Consider the model $\ln Y = \beta_0 + \beta_1 x_1 + \epsilon$.

- $> fit3 = lm(log(y) \sim x)$
- > plot(x, log(y))
- > abline(fit3\$coefficients)



> summary(fit3)

Call:

 $lm(formula = lny \sim x)$

Residuals:

Min 1Q Median 3Q Max -0.35435 -0.09045 -0.01726 0.09740 0.26357

Coefficients:

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.1541 on 48 degrees of freedom Multiple R-Squared: 0.8635, Adjusted R-squared: 0.8607 F-statistic: 303.6 on 1 and 48 DF, p-value: < 2.2e-16

The fitted regression function is

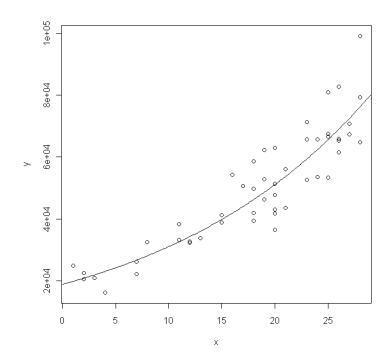
$$ln(y) = 9.841325 + 0.049978 * x.$$

That is,

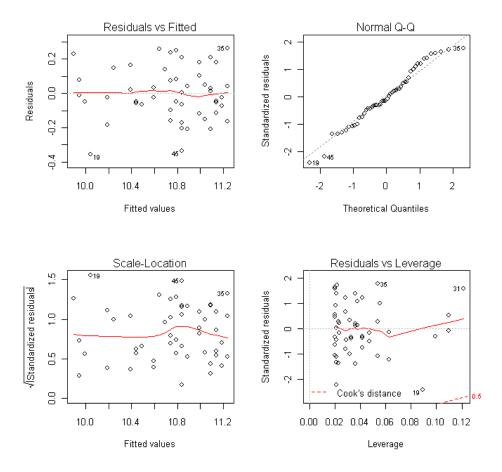
wage =
$$e^{9.841325} \cdot e^{0.049978 * x}$$
.

We would expect weekly wages to increase $e^{0.049978} = 1.051248$ times ("on average") [that is, by 5.125%] for every 1-year increase of years of experience.

```
> plot(x, y)
> xx = seq(0,30,by=0.1)
> yy2 = exp(fit3$coeff[1] + fit3$coeff[2]*xx)
> lines(xx,yy2)
```



> plot(fit3)



2. An experiment is conducted for the purpose of studying the effect of temperature on the life-length of an electrical insulation. Specimens of the insulation were tested under fixed temperatures, and their times to failure recorded.

FailureTime <i>y</i> (thousand hours)
7.3, 7.9, 8.5, 9.6, 10.3
1.7, 2.5, 2.6, 3.1
1.2, 1.4, 1.6, 1.9
0.6, 0.7, 1.0, 1.1, 1.2

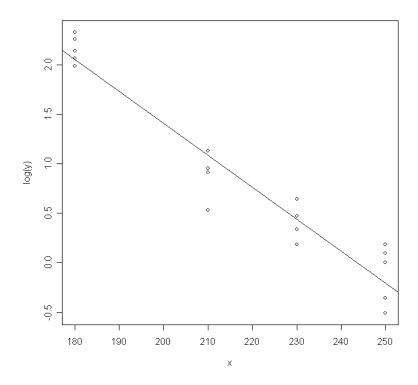
```
> x = c(rep(180,5), rep(210,4), rep(230,4), rep(250,5))

> y = c(7.3,7.9,8.5,9.6,10.3, 1.7,2.5,2.6,3.1, 1.2,1.4,1.6,1.9, 0.6,0.7,1.0,1.1,1.2)
```

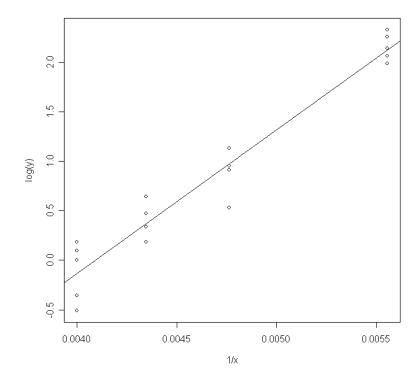
Relationship is non-linear, spread of y values is different for different x values.

> plot(x,y)

- > plot(x, log(y))
- > abline(lm(log(y)~x)\$coefficients)



- > plot(1/x,log(y))
- > abline(lm(log(y)~I(1/x))\$coefficients)



In y vs. $\frac{1}{x}$ seems to be a better fit.

a) Fit a straight-line regression to the transformed data $x' = \frac{1}{x}$ and $y' = \ln y$

b) Is there strong evidence that an increase in temperature reduces the life of the insulation?

Yes. β_1 is very significant. $\hat{\beta}_1$ is positive. The evidence is strong that $\ln y$ depends on $\frac{1}{\chi}$. The expected value of $\ln(\text{Failure Time})$ decreases as $\frac{1}{\text{Temperature}}$ decreases. Thus increasing the temperature reduces the life of the insulation.

c) Predict the Failure Time y when x = 200 C.

$$x = 200$$
 \Rightarrow $x' = 0.005$
 \Rightarrow $\hat{y}' = -5.9489 + 1453.9049 \cdot 0.005 = 1.3206$
 \Rightarrow $\hat{y} = e^{1.3206} = 3.746$.
> predict (fit, data.frame (x=200))
[1] 1.320650
> exp(1.320650)
[1] 3.745855
> predict.lm(fit, data.frame (x=200), interval=c("prediction"))
fit lwr upr
[1,] 1.320650 0.8331375 1.808162
> exp(0.8331375)
[1] 2.300525
> exp(1.808162)
[1] 6.099227
95% prediction interval: (2.3,6.1)

3. To determine the maximum stopping ability of cars when their brakes are fully applied, 10 cars are driven each at a specified speed and the distance each requires to come to a complete stop is measured. The various initial speeds selected for each of the 10 cars and the stopping distances recorded are given in the table on the right. Use transformations to find a good model for predicting the stopping distance from the initial speed and use it to predict the stopping distance y if the initial speed is x = 55 mph.

```
Initial
               Stopping
Speed
               Distance
x \text{ (mph)}
                 y (ft)
   20
                  16.3
  20
                 26.7
                 39.2
   30
                 63.5
   30
   30
                 51.3
   40
                 98.4
   40
                 65.7
   50
                 104.1
   50
                 155.6
   60
                 217.2
```

```
> x = c(20,20,30,30,30,40,40,50,50,60)
> y = c(16.3,26.7,39.2,63.5,51.3,98.4,65.7,104.1,155.6,217.2)
> fit1 = lm(y ~ x)
> summary(fit1)
Call:
lm(formula = y ~ x)
```

Residuals:

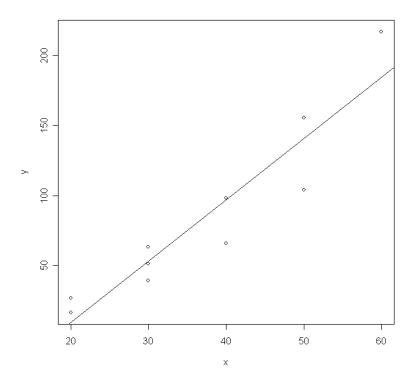
Min 1Q Median 3Q Max -36.666 -10.901 4.224 13.719 32.614

Coefficients:

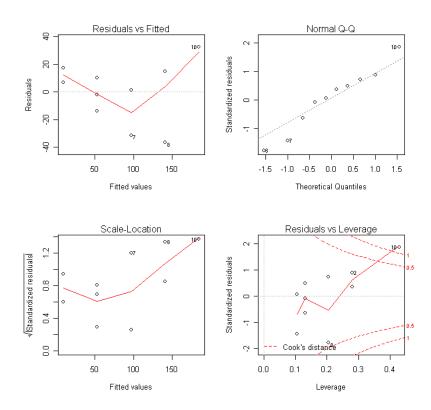
Residual standard error: 23.08 on 8 degrees of freedom Multiple R-Squared: 0.8788, Adjusted R-squared: 0.8637 F-statistic: 58.02 on 1 and 8 DF, p-value: 6.206e-05

```
> plot(x,y)
```

> abline(fit1\$coefficients)



- > par(mfrow=c(2,2))
 > plot(fit1)



If the relationship between x and y is exponential, $y = A e^{b x}$, then $\ln y = \ln A + bx$, or z = a + bx, where $z = \ln y$, $a = \ln A$.

```
> fit2 = lm(log(y) \sim x)
> summary(fit2)
```

Call:

 $lm(formula = log(y) \sim x)$

Residuals:

Min 1Q Median 3Q Max -0.421385 -0.140194 -0.003966 0.159422 0.376411

Coefficients:

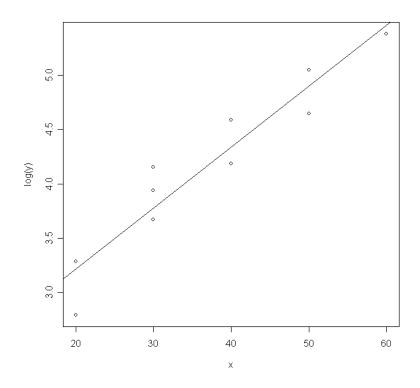
Estimate Std. Error t value Pr(>|t|) (Intercept) 2.088393 0.253648 8.233 3.55e-05 *** 0.006485 8.668 2.44e-05 *** 0.056208 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 Signif. codes:

Residual standard error: 0.2602 on 8 degrees of freedom Multiple R-squared: 0.9038, Adjusted R-squared: 0.8917

F-statistic: 75.13 on 1 and 8 DF, p-value: 2.441e-05

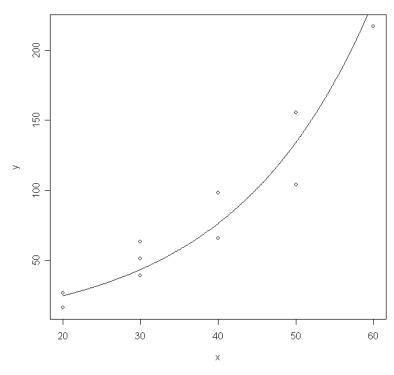
> plot(x,log(y))

> abline(fit2\$coefficients)

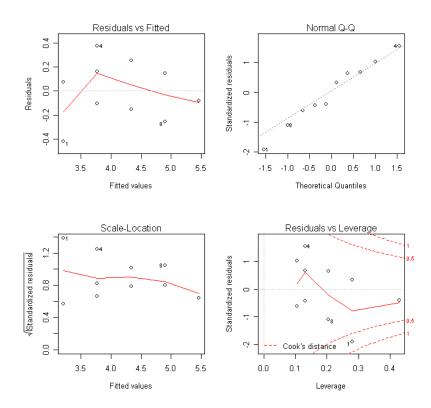


```
> plot(x,y)
```

- > xx = seq(20, 60, by=0.1)
- > yy2 = exp(fit2\$coeff[1]+fit2\$coeff[2]*xx)
 > lines(xx,yy2)



> plot(fit2)



If the relationship between x and y is power function, $y = A x^b$, then $\ln y = \ln A + b \ln x$, or z = a + b w, where $z = \ln y$, $w = \ln x$, $a = \ln A$.

> fit3 = lm(log(y) ~ log(x)) > summary(fit3)

Call:

 $lm(formula = log(y) \sim log(x))$

Residuals:

Min 1Q Median 3Q Max -0.26472 -0.24103 0.09071 0.14553 0.28115

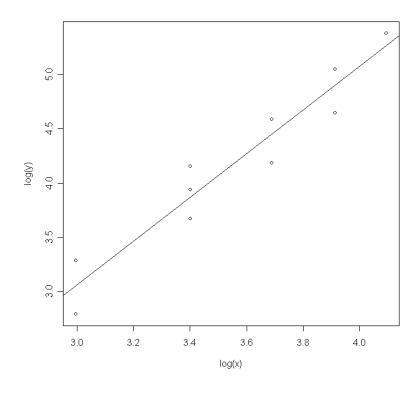
Coefficients:

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

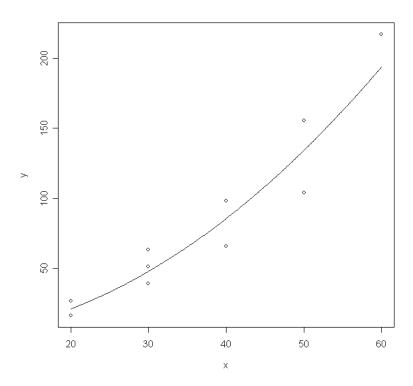
Residual standard error: 0.2332 on 8 degrees of freedom Multiple R-squared: 0.9227, Adjusted R-squared: 0.913 F-statistic: 95.47 on 1 and 8 DF, p-value: 1.009e-05

> plot(log(x),log(y))

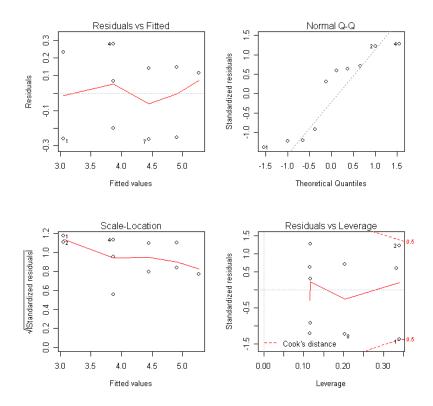
> abline(fit3\$coefficients)



- > plot(x,y)
- > yy3 = exp(fit3\$coeff[1]+fit3\$coeff[2]*log(xx))
 > lines(xx,yy3)



> plot(fit3)



> fit4 = $lm(y \sim x + I(x^2))$ > summary(fit4)

Call:

 $lm(formula = y \sim x + I(x^2))$

Residuals:

Min 1Q Median 3Q Max -33.326 -8.632 3.367 13.634 18.174

Coefficients:

Estimate Std. Error t value Pr(>|t|)
(Intercept) 37.90064 56.01200 0.677 0.5204

x -2.34538 3.09621 -0.757 0.4735

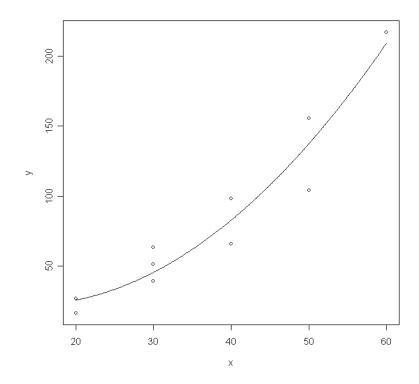
I(x^2) 0.08672 0.03944 2.199 0.0639 .

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

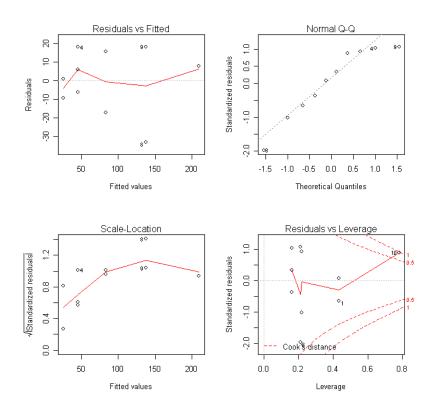
Residual standard error: 18.98 on 7 degrees of freedom Multiple R-squared: 0.9283, Adjusted R-squared: 0.9078 F-statistic: 45.33 on 2 and 7 DF, p-value: 9.861e-05

> plot(x,y)

- > yy4 = fit4\$coeff[1]+fit4\$coeff[2]*xx+fit4\$coeff[3]*xx^2
- > lines(xx,yy4)



> plot(fit4)



> fit5 = lm(y ~ I(x^2)) > summary(fit5)

Call:

 $lm(formula = y \sim I(x^2))$

Residuals:

Min 1Q Median 3Q Max -35.175 -7.146 5.528 13.910 16.325

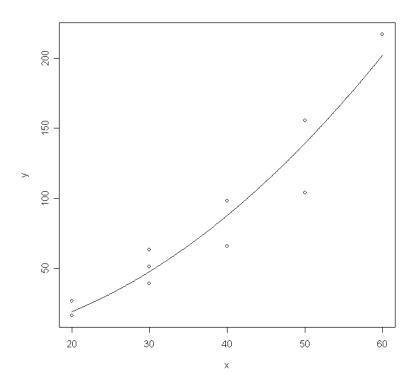
Coefficients:

Estimate Std. Error t value Pr(>|t|) (Intercept) -3.701935 10.703929 -0.346 0.738 $I(x^2)$ 0.057191 0.005863 9.754 1.02e-05 ***

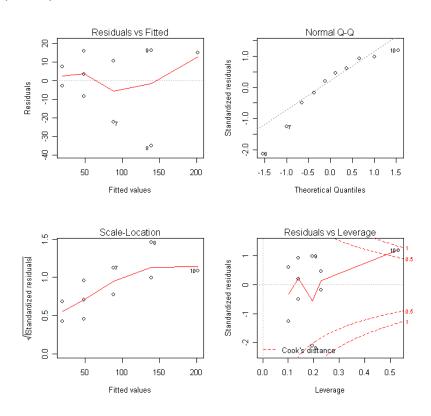
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 18.47 on 8 degrees of freedom Multiple R-squared: 0.9224, Adjusted R-squared: 0.9127 F-statistic: 95.15 on 1 and 8 DF, p-value: 1.022e-05

- > plot(x,y)
 > yy5 = fit5\$coeff[1]+fit5\$coeff[2]*xx^2
 > lines(xx,yy5)



> plot(fit5)



> fit6 = lm(sqrt(y) ~ x) > summary(fit6)

Call:

 $lm(formula = sqrt(y) \sim x)$

Residuals:

Min 1Q Median 3Q Max -1.4821 -0.6468 0.4055 0.6630 1.0243

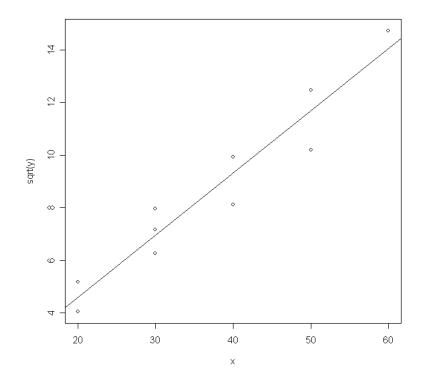
Coefficients:

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

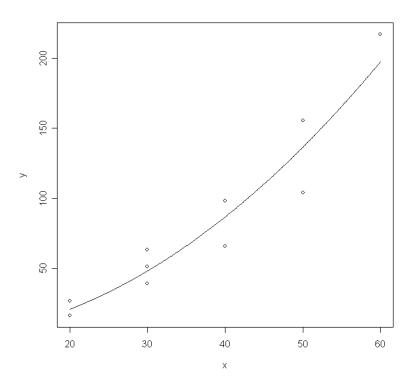
Residual standard error: 0.9563 on 8 degrees of freedom Multiple R-squared: 0.9252, Adjusted R-squared: 0.9158 F-statistic: 98.91 on 1 and 8 DF, p-value: 8.843e-06

> plot(x,sqrt(y))

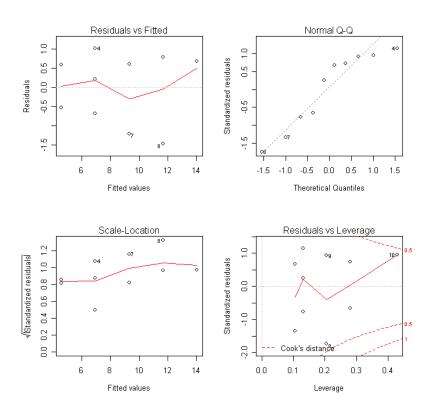
> abline(fit6\$coefficients)



- > plot(x,y)
- > yy6 = (fit6\$coeff[1]+fit6\$coeff[2]*xx)^2
 > lines(xx,yy6)



> plot(fit6)



fit3 and fit6 seem to be the best.

```
> new = data.frame(x=55)
```

> predict(fit3,new)

- [1] 5.09179
- $> \exp(5.09179)$
- [1] <mark>162.6808</mark>

> predict(fit6,new)

- [1] 12.87022
- > 12.87022^2
- [1] <mark>165.6426</mark>

STAT 420

Box-Cox method:

Transformation
$$g_{\lambda}(y) = egin{array}{cccc} \mathbb{X} & \frac{y^{\lambda}-1}{\lambda} & \lambda \neq 0 \\ \mathbb{X} & \text{In } y & \lambda = 0 \end{array}$$

Recall Example 3 from Examples for 10/11/12 (part 2):

3. Initial Stopping To determine the maximum stopping ability of cars Speed Distance when their brakes are fully applied, 10 cars are driven X(mph)y (ft) each at a specified speed and the distance each requires 20 16.3 to come to a complete stop is measured. The various 20 26.7 initial speeds selected for each of the 10 cars and the 30 39.2 stopping distances recorded are given in the table on 30 63.5 the right. Use transformations to find a good model 51.3 30 for predicting the stopping distance from the initial 40 98.4 speed and use it to predict the stopping distance y if 40 65.7 the initial speed is X = 55 mph. 50 104.1 155.6 50

>
$$x = c(20,20,30,30,30,40,40,50,50,60)$$

> $y = c(16.3,26.7,39.2,63.5,51.3,98.4,65.7,104.1,155.6,217.2)$
> fit = $lm(y \sim x)$

217.2

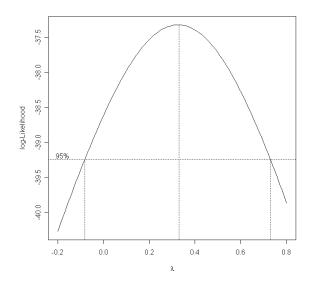
60

- > library(MASS)
- > boxcox(fit,plotit=T)

> boxcox(fit,plotit=T,lambda=seq(-0.2,0.8,by=0.01))

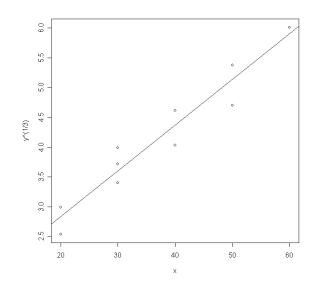
ln y ($\lambda = 0$) and \sqrt{y} ($\lambda = 0.5$) are acceptable.

 $\lambda \approx \frac{1}{3}$ seems to give the best transformation of the response variable.

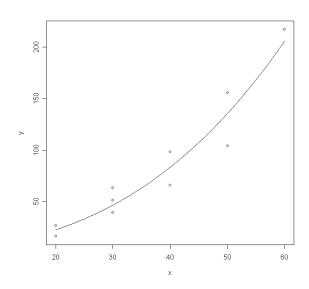


$$> fit2 = lm(y^{(1/3)} \sim x)$$

- > plot(x,y^(1/3))
- > abline(fit2\$coefficients)



- > plot(x,y)
- > xx = seq(20,60,by=0.1)
- > yy = (fit2\$coeff[1]+fit2\$coeff[2]*xx)^3
- > lines(xx,yy)



> summary(fit2)

Call:

 $lm(formula = y^{(1/3)} \sim x)$

Residuals:

Min 1Q Median 3Q Max -0.4315 -0.2727 0.1117 0.2215 0.3900

Coefficients:

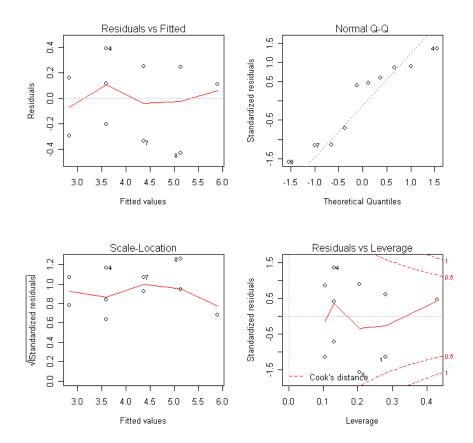
Estimate Std. Error t value Pr(>|t|)
(Intercept) 1.295415 0.298150 4.345 0.00246 **
x 0.076806 0.007622 10.076 8.02e-06 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.3058 on 8 degrees of freedom Multiple R-squared: 0.927, Adjusted R-squared: 0.9178 F-statistic: 101.5 on 1 and 8 DF, p-value: 8.019e-06

> par(mfrow=c(2,2))

> plot(fit2)



4. The data in the file were collected in a study of the effect of dissolved sulfur on the surface tension of liquid copper (Baes, C. and Kellogg, H. (1953). Effect of dissolved sulphur on the surface tension of liquid copper. *J. Metals*, 5, 643-648).

http://www.stat.umn.edu/alr/data/baeskel.txt

This data frame contains the following columns:

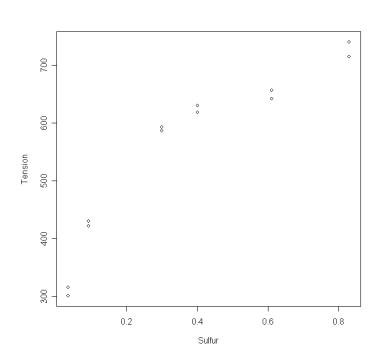
Sulfur Weight percent sulfur

Tension Decrease in surface tension, dynes/cm

- > tension = read.table("http://www.stat.umn.edu/alr/data/baeskel.txt",
 header=T)
- > attach(tension)
- > tension

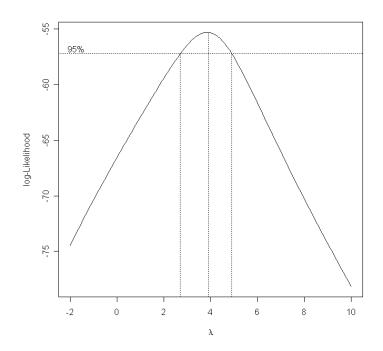
```
Sulfur Tension
1
    0.034
                301
2
    0.034
                316
3
    0.093
                430
4
    0.093
                422
    0.300
5
                593
6
    0.300
                586
7
    0.400
                630
    0.400
8
                618
9
    0.610
                656
    0.610
10
                642
11
    0.830
                740
12
    0.830
                714
```

> plot(Sulfur, Tension)



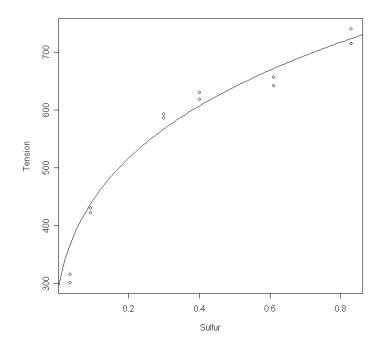
- > fit = lm(Tension ~ Sulfur)
- > boxcox(fit,plotit=T,lambda=seq(-2,10,by=0.1))

 $\lambda \approx 4$ seems to give the best transformation of the response variable.



```
> fit2 = lm(Tension^4 ~ Sulfur)
> plot(Sulfur, Tension)
> x = seq(0,0.9,by=0.01)
> y = (fit2$coeff[1]+fit2$coeff[2]*x)^0.25
```

> lines(x,y)



A better fit:

```
> fit3 = lm(Tension ~ log(Sulfur))
> y2 = fit3$coeff[1]+fit3$coeff[2]*log(x)
> lines(x,y2,col=2)
```

