$$| (0) \hat{\theta} = \frac{1}{2} \frac{\partial^{2} e^{-\theta}}{\partial x_{1}} \Rightarrow (\hat{\theta}) = \frac{1}{2} (0) \left(\frac{\partial^{2} e^{-\theta}}{\partial x_{1}} \right)$$

$$= \frac{1}{2} (\log(\theta^{2}) - \log(x_{1}) - \theta_{1}) \cdot ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{1} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(x_{1})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(\theta^{2})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(\theta^{2})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(\theta^{2})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(\theta^{2})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(\theta^{2})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(\theta^{2})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(\theta^{2})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) - \log(\theta^{2})) - \theta_{2} \Rightarrow ((\theta)) = \frac{1}{2} (\log(\theta^{2}) -$$

b) According to CLT, $M(\hat{\theta} - \theta) \sim N(0, \theta)$ since the variance of a Poisson distribution is ? 2. OBD, then when n is large, $\overline{m}(\widehat{\theta}-\theta) \sim N(0, \widehat{\theta})$. (We don't know what I is, but we all know I, so we are use I to find a C.I. for I),

3.
$$f(X_i; \theta) = \prod_{i=1}^{n} \frac{\theta^{X_i} e^{-\theta}}{X_i!} = (\prod_{i=1}^{n} \theta^{X_i} e^{-\theta}) (\prod_{i=1}^{n} x_{i!}) = (e^{-n\theta} \theta^{\sum X_i}) (\prod_{i=1}^{n} x_{i!})$$

So $Y = \sum X_i$ is sufficient by factorization theorem.

$$f(x) = E(f(x))^2 = f(x)^2 =$$

b) $V_{ar}(Y) = V_{ar}(\frac{1}{2}X_{1}^{2}) = \frac{1}{n^{2}} \sum_{i=1}^{n} V_{ar}(X_{i}^{2}) = \frac{1}{n^{2}} \sum_{i=1}^{n} E(X_{i}^{4}) - E(X_{i}^{2})^{2}$ E(Xi2) = Var(Xi) + E(Xi)2 = B $E(X_i^4) = E[(X_i - \mu)^4] = 30^2$ since $\mu = 0$

So $Var(Y) = \frac{1}{R^2} \stackrel{\triangle}{=} E(X_1^4) - E(X_1^2)^2 = \frac{1}{R^2} \stackrel{\triangle}{=} (3\theta^2 - \theta^2) = \frac{2\theta^2}{R^2}$

5. a)
$$I(\theta) = -E[\frac{\partial^2 \log(f(x,\theta))}{\partial \theta^2}] = -E[\frac{\partial^2}{\partial \theta^2}\log(\frac{1}{\partial \theta}e^{-\frac{x^2}{\partial \theta}})]$$

$$= -E[\frac{\partial^2}{\partial \theta^2}I(\log(\theta) - (\log(\frac{1}{\partial \theta}) - \frac{x^2}{\partial \theta})] = -E[\frac{\partial^2}{\partial \theta^2}[-\frac{1}{2}\log(2\pi\theta) - \frac{x^2}{2\theta}]]$$

$$= -E[\frac{\partial}{\partial \theta}(-\frac{1}{2}(\frac{2\pi}{2\pi\theta}) + \frac{x^2}{2\theta^2})] = -E[\frac{\partial}{\partial \theta}(-\frac{1}{2\theta} + \frac{x^2}{2\theta^2})]$$

$$= -E[\frac{1}{2\theta^2} - \frac{x^2}{\theta^2}] = -\frac{1}{2\theta^2} + \frac{1}{\theta^2}E(x^2) = -\frac{1}{2\theta^2} + \frac{1}{\theta^2} = \frac{1}{2\theta^2}$$
b) Rao-Gramer Lover Bound
$$\frac{1}{nI(\theta)} = 2\theta^2 \qquad \text{MLE} \cdot L(\theta | x) = \left(\frac{1}{2\pi H}\right)^2 e^{-\frac{1}{2}(x^2/2\theta)}$$

$$L(\theta, x) = -\frac{1}{2}\ln(2\pi\theta) - \frac{1}{2}x^2/2\theta \implies \frac{\partial L(\theta, x)}{\partial \theta} = -\frac{1}{2\theta} + \frac{2}{2\theta^2} = \frac{1}{2\theta^2}$$

$$D(\theta, x) = -\frac{1}{2}\ln(2\pi\theta) - \frac{1}{2}x^2/2\theta \implies \frac{\partial L(\theta, x)}{\partial \theta} = -\frac{1}{2\theta} + \frac{2}{2\theta^2} = \frac{1}{2\theta^2}$$

$$D(\theta, x) = -\frac{1}{2}\ln(2\pi\theta) - \frac{1}{2}x^2/2\theta \implies \frac{\partial L(\theta, x)}{\partial \theta} = -\frac{1}{2\theta} + \frac{2}{2\theta^2} = \frac{1}{2\theta^2}$$

$$D(\theta, x) = -\frac{1}{2}\ln(2\pi\theta) - \frac{1}{2}\ln(2\pi\theta) - \frac{1}{2}\ln(2\pi\theta) + \frac{1}{2}$$

6.9)
$$f_{X}(X_{i};\theta) = \frac{1}{i!}(\theta+1)(1-X_{i})^{\theta} = (\theta+1)^{n} [\frac{1}{i!}(1-X_{i})^{\theta}]$$

$$= (\theta+1)^{n} [\frac{1}{i!}(1-X_{i})]^{\theta}$$

$$= (1-X_{i}) \text{ is sufficient.}$$

b)
$$\underline{T}(\theta) = - E\left(-\frac{1}{(\theta+1)^2}\right) = \frac{1}{(\theta+1)^2}$$

7.
$$EY = 2EX_{1} = 2\int_{0}^{\infty} x \frac{3\theta^{3}}{(x+\theta)^{2}} dx$$

$$= 2\theta^{3} \left\{ x \left[-(x+\theta)^{\frac{3}{2}} \right] \left[-(x+\theta)^{\frac{3}{2}} \right] x \right\}$$

$$= 0$$

$$\frac{3^{2}}{3^{2}} \left[\log \frac{1}{3} (x;\theta) \right] = -\frac{1}{9^{2}} + \frac{1}{(x+\theta)^{2}} \right]$$

$$= \frac{1}{9^{3}} - E \left(\frac{1}{(x+\theta)^{3}} \right)$$

$$= \frac{1}{9^{3}} - E \left(\frac{1}{(x+\theta)^$$

10.
$$L(\theta) = \theta^{n} \left(\frac{1}{1!} \times i \right)^{\theta-1}$$

$$L(\theta; x) = n \log \theta + (\theta-1) \int_{i=1}^{n} \log x_{i}$$

$$\frac{\partial}{\partial \theta} \left(L(\theta; x) \right) = \frac{n}{\theta} + \int_{i=1}^{n} \log x_{i} = 0 \quad \Rightarrow \quad \hat{\theta} = -\frac{n}{\sum_{i=1}^{n} \log x_{i}}$$

$$A = \frac{L(1)}{L(\hat{\theta})} = \frac{1}{\hat{\theta}^{n} \left(\frac{1}{1!} \times i \right)^{\hat{\theta}-1}} = \left(-\frac{\sum_{i=1}^{n} \log x_{i}}{n} \right)^{n} \left(\frac{1}{1!} \times i \right)^{\frac{n}{2} \log x_{i}} + 1$$

$$-2 \ln \Lambda = -2 \left[n \log \left(-\frac{\sum_{i=1}^{n} \log x_{i}}{n} \right) + n + \sum_{i=1}^{n} \log x_{i} \right]$$

$$E(\tilde{X}_{n}^{2}) = Var(\tilde{X}_{n}) + \left(E(\tilde{X}_{n}) \right)^{2} = \frac{\sigma^{2}}{n} + \theta^{2}$$

$$E(\tilde{X}_{n}^{2}) = Var(\tilde{X}_{n}) + \left(E(\tilde{X}_{n}) \right)^{2} = \frac{\sigma^{2}}{n} + \theta^{2}$$

$$E(\tilde{X}_{n}^{2} - \sigma^{2} n) = \theta^{2} \Rightarrow \tilde{X}_{n}^{2} - \sigma^{2} n \quad \text{is unbiased for } \theta^{2}$$

$$I(0) = \frac{1}{1 \log x_{i}} \left(\log \frac{1}{2} \left(x_{i} \cdot \theta \right) \right) = -E\left(-\frac{1}{\sigma^{2}} \right) = \frac{1}{\sigma^{2}}$$

$$|ower bound| = \frac{\left(\frac{1}{2} (1 \log x_{i}) \right)^{2}}{n \log x_{i}} = \frac{1}{1 \log x_{i}}$$

$$Var(\tilde{X}_{n}^{2} - \frac{\sigma^{2}}{n}) = Var(\tilde{X}_{n}^{2}) = \frac{1}{2 \log x_{i}} = \frac{1}{1 \log x_{i}}$$

$$Var(\tilde{X}_{n}^{2} - \frac{\sigma^{2}}{n}) = Var(\tilde{X}_{n}^{2}) = \frac{1}{2 \log x_{i}} = \frac{1}{1 \log x_{i}}$$

$$\therefore eff(c)ency = \frac{4\theta^{2}\sigma^{2}}{n} + 2\sigma^{2} n^{2} + \frac{1}{1 \log x_{i}} = \frac{1}{1 \log x_{i}}$$

$$Var(\tilde{X}_{n}^{2} - \frac{\sigma^{2}}{n}) = Var(\tilde{X}_{n}^{2}) = \frac{1}{2 \log x_{i}} = \frac{1}{1 \log x_{i}}$$

$$\therefore eff(c)ency = \frac{4\theta^{2}\sigma^{2}}{n} + 2\sigma^{2} n^{2} + \frac{1}{1 \log x_{i}} = \frac{1}{1 \log x_{i}}$$

$$vhen \quad 0 < x < \frac{\pi}{4}, \quad \Lambda = \frac{\sin x}{\cos x} = \tan x \le C$$

$$P(\tan x \le c \mid H_{0}) = 0.1 \quad \Leftrightarrow \quad P(x \le \arctan c \mid H_{0}) = 0.1$$

$$\int_{0}^{2} \arctan c \sin x \, dx = 0.1 \quad \Leftrightarrow \quad C = \tan (\arcsin \theta) \quad \approx 0.4$$

$$P(\tan X \leq c \mid Ho) = 0.1 \iff P(X \leq \arctan c \mid Ho) = 0.1$$

$$\int_{0}^{\arctan c} \sin X \, dX = 0.1 \implies C = \tan(\arctan c \log 0.9) \approx 0.484$$

$$\therefore \text{ reject Ho when } \Lambda \leq 0.484$$

$$Power = P(\tan X \leq 0.484 \mid Ha) = \int_{0}^{\arctan 0.484} \cos X \, dx \approx 0.436$$