732A90 Computational Statistics Lecture 6 Derivation of EM algorithm for Slide 15

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E-step: derive

$$Q(\theta, \theta^k) = \mathrm{E}\left[\mathrm{loglik}(\theta|Y, Z) | \theta^k, Y\right] = \mathrm{E}\left[-\frac{1}{2}n\log(2\pi\sigma^2) - \frac{1}{2\sigma^2}\left(\sum_{i=1}^n (y_i - \mu)^2\right) | \theta^k, Y\right] \stackrel{*}{=} .$$

Our observed data is $Y = \{y_1, \dots, y_r\}$ and the latent variables (unobserved data) is $Z = \{y_{r+1}, \dots, y_n\}$. Hence, continuing we have

$$\begin{split} &\stackrel{*}{=} - \frac{1}{2} n \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \operatorname{E} \left[\sum_{i=1}^n y_i^2 - 2\mu \sum_{i=1}^n y_i + n\mu^2 | \theta^k, Y \right] \\ &= - \frac{1}{2} n \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \operatorname{E} \left[\sum_{i=1}^r y_i^2 + \sum_{i=r+1}^n y_i^2 - 2\mu \sum_{i=1}^r y_i - 2\mu \sum_{i=r+1}^n y_i + n\mu^2 | \theta^k, Y \right] \\ &= - \frac{1}{2} n \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^r y_i^2 + \sum_{i=r+1}^n \operatorname{E} \left[y_i^2 | \theta^k \right] - 2\mu \sum_{i=1}^r y_i - 2\mu \sum_{i=r+1}^n \operatorname{E} \left[y_i | \theta^k \right] + n\mu^2 \\ &= - \frac{1}{2} n \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^r y_i^2 + \sum_{i=r+1}^n (\mu_k + \sigma_k^2) - 2\mu \sum_{i=1}^r y_i - 2\mu \sum_{i=r+1}^n \mu_k + n\mu^2 \\ &= - \frac{1}{2} n \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^r (y_i - \mu)^2 + (n-r)(\mu - \mu_k)^2 + (n-r)\sigma_k^2 \\ &= \left\{ \operatorname{setting the unobserved } y_i \equiv \mu_k \right\} \\ &= - \frac{1}{2} n \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2 + \frac{n-r}{2\sigma^2} \sigma_k^2 = Q(\theta, \theta^k). \end{split}$$

M-step: taking the derivatives with respect to the parameters μ and σ^2

$$Q(\theta, \theta^k)'_{\mu} = \frac{2}{2\sigma^2} \sum_{i=1}^n (y_i - \mu),$$

$$Q(\theta, \theta^k)'_{\sigma^2} = -\frac{n}{2} \frac{1}{2\pi\sigma^2} \cdot 2\pi - \left(\sum_{i=1}^n (y_i - \mu)^2\right) \frac{-1}{2} \frac{1}{(\sigma^2)^2} - \frac{1}{2} (n - r)\sigma_k^2 (-1) \frac{1}{(\sigma^2)^2}.$$

Setting the first equation to 0 we obtain $0 = \sum_{i=1}^{n} (y_i - \mu)$, giving $n\mu = \sum y_i$ and in turn

$$\mu_{k+1} = \frac{1}{n} \sum_{i=1}^{n} y_i.$$

Turning to the second equation

$$0 = -\frac{n}{2} \frac{1}{2\pi\sigma^2} \cdot 2\pi - \left(\sum_{i=1}^n (y_i - \mu)^2\right) \frac{1}{2} \frac{1}{(\sigma^2)^2} - \frac{1}{2}(n-r)\sigma_k^2(-1) \frac{1}{(\sigma^2)^2}$$

$$0 = -\frac{n}{2} + \frac{1}{(\sigma^2)^2} \left(\sum_{i=1}^n (y_i - \mu)^2\right) + \frac{\sigma_k^2}{(\sigma^2)^2}(n-r)$$

$$n\sigma^2 = \left(\sum_{i=1}^n (y_i - \mu)^2\right) + (n-r)\sigma_k^2$$

$$\sigma^2 = \frac{1}{n} \left(\sum_{i=1}^n (y_i - \mu)^2\right) + \frac{1}{n}(n-r)\sigma_k^2$$

$$= \frac{1}{n} \left(\sum_{i=1}^n y_i^2 - 2\mu \sum_{i=1}^n y_i + n\mu^2 + (n-r)\sigma_k^2\right)$$

$$= \frac{1}{n} \left(\sum_{i=1}^r y_i^2 + \sum_{i=r+1}^n \mu_k^2 - 2\mu \sum_{i=1}^r y_i - 2\mu \sum_{i=r+1}^n \mu_k + n\mu^2 + (n-r)\sigma_k^2\right)$$

$$= \frac{1}{n} \left(\sum_{i=1}^r y_i^2 + (n-r)\mu_k^2 - 2\mu \sum_{i=1}^r y_i - 2(n-r)\mu\mu_k + n\mu^2 + (n-r)\sigma_k^2\right)$$

$$= \frac{1}{n} \left(\sum_{i=1}^r y_i^2 + (n-r)(\mu_k^2 + \sigma_k^2)\right) + \mu^2 - \frac{2}{n}\mu \sum_{i=1}^n y_i.$$

As we now have the estimate at the (k+1)st step of μ as μ_{k+1} , then

$$\mu_{k+1}^2 - \frac{2}{n}\mu_{k+1} \sum_{i=1}^n y_i = \mu_{k+1}^2 - 2\mu_{k+1} \frac{1}{n} \sum_{i=1}^n y_i = \mu_{k+1}^2 - 2\mu_{k+1}^2 = -\mu_{k+1}^2$$

and we obtain

$$\sigma_{k+1}^2 = \frac{1}{n} \left(\sum_{i=1}^r y_i^2 + (n-r)(\mu_k^2 \sigma_k^2) \right) - \mu_{k+1}^2.$$

The next step could then be to e.g. consider the second derivative and show that this is actually a maximum.