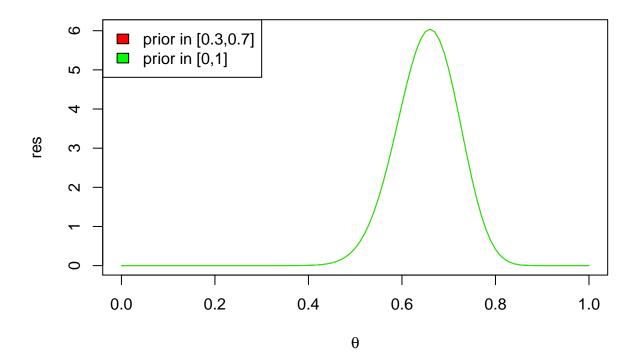
Bayesian Learning

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```
1
###b
posterior = function(theta, x){
 return(dbeta(theta, x+1, 51-x))
}
thetaGrid = seq(0,1,0.01)
# prior in [0.3, 0.7] p(theta) = 1/(0.7 - 0.3)
res = sapply(1:length(thetaGrid), function(i)posterior(thetaGrid[i], x=33))
plot(thetaGrid, res, type = "l", col="red", xlim =c(0,1), main = "NormalizedPosterior",
     xlab = expression(theta))
# prior in [0, 1] p(theta) = 1
res2 = sapply(1:length(thetaGrid), function(i)posterior(thetaGrid[i], x=33))
lines(thetaGrid, res2/(0.01*sum(res2)), col="green" )
legend("topleft",
       legend = c("prior in [0.3,0.7]", "prior in [0,1]"),
       fill = c("red", "green"))
```

NormalizedPosterior



```
c
print("\nPrior in [0.3, 0.7]: ")

## [1] "\nPrior in [0.3, 0.7]: "
cat("Pr(posterior < 0.5: ", sum(res < 0.5)/length(res), "\n\n\n")

## Pr(posterior < 0.5: 0.7128713

print("Prior in [0, 1]: ")

## [1] "Prior in [0, 1]: "
cat("Pr(posterior < 0.5: ", sum(res2 < 0.5)/length(res2), "\n")

## Pr(posterior < 0.5: 0.7128713</pre>
```

When the prior is uniform the normalized posterior for that will be the same regardless of the prior

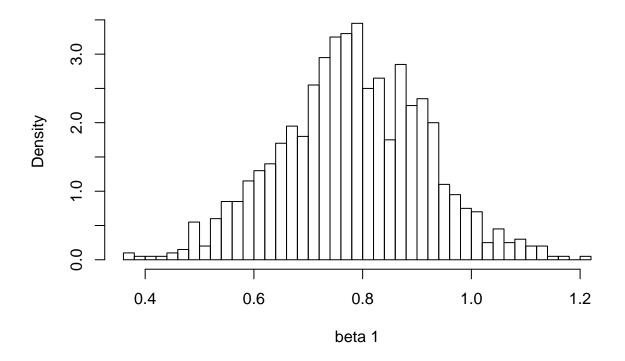
2

```
load(file = 'titanic.RData')

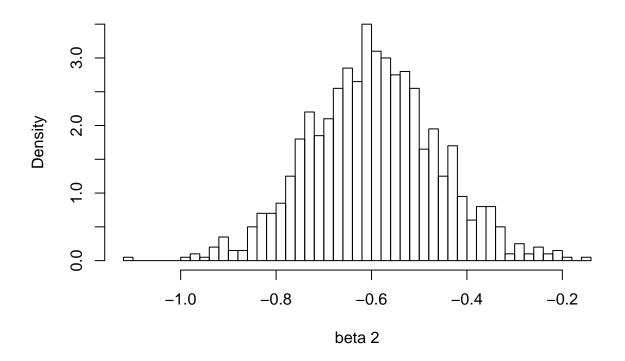
BayesProbReg <- function(y, X, mu_0, tau, nIter){
    # Gibbs sampling in probit regression using data augmentation:</pre>
```

```
# beta / tau ~ N(mu_0, tau^2*I)
  # INPUTS:
  # y - n-by-1 vector with response data observations
  {\it X-n-by-nCovs} matrix with covariates, first column should be ones if you want an intercept.
  # mu_0 - prior mean for beta
  # tau - prior standard deviation for beta
  # nIter - Number of samples from the posterior (iterations)
  # OUTPUTS:
  # betaSample - Posterior samples of beta. nIter-by-nCovs matrix
  # Prior
  priorCov <- tau^2*diag(nPara)</pre>
  priorPrec <- solve(priorCov)</pre>
  # Compute posterior hyperparameters
  n = length(y) # Number of observations
  n1 = sum(y)
  n0 = n - n1
  nCovs = dim(X)[2] # Number of covariates
  XX = t(X)%*%X
  # The actual sampling
  betaSample = matrix(NA, nIter, nCovs)
  u <- matrix(NA, n, 1)
  beta <- solve(XX,crossprod(X,y)) # OLS estimate as initial value
  for (i in 1:nIter){
    xBeta <- X%*%beta
    # Draw u | beta
    u[y == 0] \leftarrow rtnorm(n = n0, mean = xBeta[y==0], sd = 1, lower = -Inf, upper = 0)
    u[y == 1] \leftarrow rtnorm(n = n1, mean = xBeta[y==1], sd = 1, lower = 0, upper = Inf)
    # Draw beta / u
    betaHat <- solve(XX,t(X)%*%u)</pre>
    postPrec <- XX + priorPrec</pre>
    postCov <- solve(postPrec)</pre>
    betaMean <- solve(postPrec,XX%*%betaHat + priorPrec%*%mu_0)</pre>
    beta <- t(rmvnorm(n = 1, mean = betaMean, sigma = postCov))
    betaSample[i,] <- t(beta)</pre>
  }
  return(betaSample=betaSample)
a
y = as.vector(titanic\survived)
x = as.matrix(titanic[,-1])
```

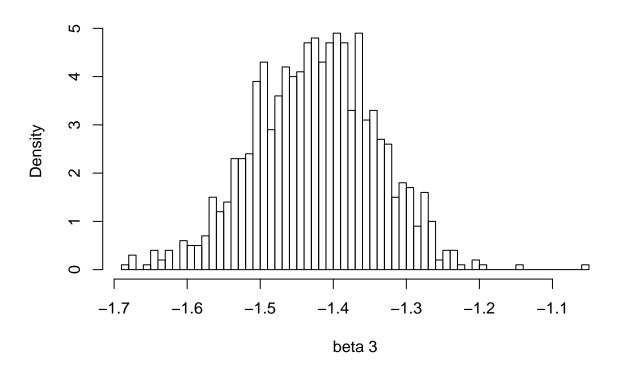
Posterior distribution of beta for intercept



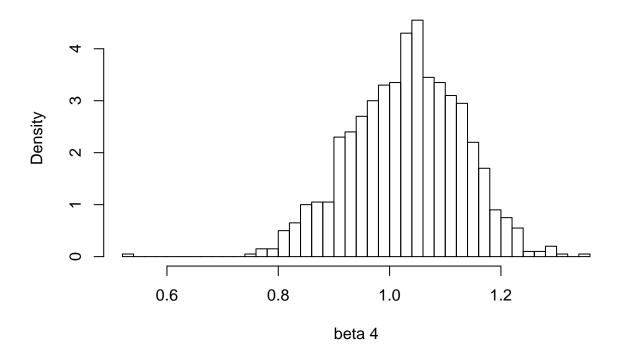
Posterior distribution of beta for adult



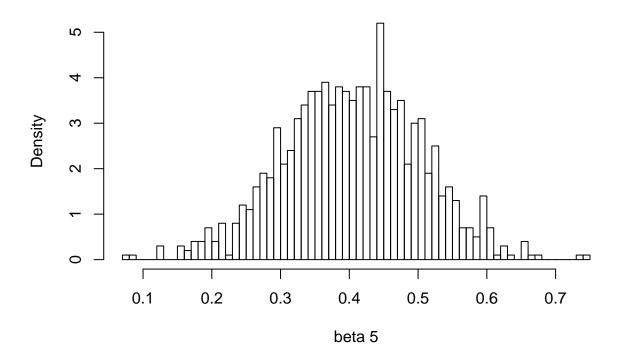
Posterior distribution of beta for man



Posterior distribution of beta for class1



Posterior distribution of beta for class2



###b

Point estimate for a linear loss function would be the Median

```
for (i in 1:5) {
  cat("Point estimate for beta of ", covNames[i], ": ",
      median(beta[,i]),"\n\n")
}
## Point estimate for beta of intercept : 0.7806523
## Point estimate for beta of
                               adult: -0.5968705
##
## Point estimate for beta of
                              man : -1.422684
##
## Point estimate for beta of class1: 1.036072
##
## Point estimate for beta of class2 : 0.4056117
\mathbf{c}
pr = numeric(nrow(beta))
for (i in 1:nrow(beta)) {
  if(beta[i,2]+beta[i,5] > 0){
    pr[i] = 1
  }else{
```

```
pr[i] = 0
  }
}
cat("Pr((Beta2 + Beta5) > 0) = ", sum(pr)/length(pr))
## Pr((Beta2 + Beta5) > 0) = 0.116
This probabilit says that if a person is adult and this person belong to the class1, it is roughly
11% probable to be survived
3
\mathbf{a}
y = c(5, 3, 17, 8)
x = c(\log(20), \log(20), \log(50), \log(40))
logPost = function(beta){
  logLik = sum(dpois(y, exp(x*beta), log = TRUE))
  logPrior = dnorm(beta, 1, 0.1, log = TRUE)
  return(logLik + logPrior)
}
initVal = 1
OptimResults<-optim(initVal,logPost,</pre>
                     method=c("BFGS"),
                     control=list(fnscale=-1),hessian=TRUE)
betaMean=OptimResults$par
Var = -solve(OptimResults$hessian)
cat("Posterior Mean :", betaMean, "\n\n")
## Posterior Mean : 0.6875749
cat("standard deviation: ", sqrt(Var))
## standard deviation: 0.03958642
\mathbf{b}
loss = function(x,y){
  4 + (\exp(x)/50) - \operatorname{sqrt}(y)
}
cost = c(20, 40)
x = log(cost)
Loss = matrix(0, 10000, length(x))
for (i in 1:10000) {
```

for (j in 1:length(x)) {

```
beta = rnorm(1,betaMean, sqrt(Var))
    pred = rpois(1,exp(beta*x[j]))
    Loss[i,j] = loss(x[j], pred)
}
}
cat("Expected value for loss when cost = 20: ", mean(Loss[,1]), "\n")
## Expected value for loss when cost = 20: 1.644435
cat("Expected value for loss when cost = 40: ", mean(Loss[,2]), "\n")
## Expected value for loss when cost = 40: 1.26471
According to the expected values, we select the one with the lowest loss, so the decision would
be keeping cost at 40 milion
4
y1 = 0.75 * dnorm(10, 14, sqrt(4 + 4/20)) * dnorm(250, 300, sqrt(2500 + 2500/20))
y0 = 0.25 * dnorm(10, 12, sqrt(4 + 4/20)) * dnorm(250, 280, sqrt(2500 + 2500/20))
unNorm = c(y1, y0)
print("UnNormalized Prob:")
## [1] "UnNormalized Prob:"
print(unNorm)
## [1] 0.0001051133 0.0001982958
print("********")
## [1] "********
Normalized = unNorm/sum(unNorm)
names(Normalized) = c("Feamale", "Male")
print("Normalized:")
## [1] "Normalized:"
print(Normalized)
```

Given the new measurements for length and weight and the previous data, the probability that the new fish is female is 0.35

##

Feamale

0.3464408 0.6535592

Male