

TENTAMEN (EXAMINATION)

Tentamensdatum/Examination date: 2020-01-14

AID-nummer
AID number

Ifylles av student
Completed by student

Ifylles av vakt
Completed by supervisor

Utbildningskod/Education code: 732A97 Modul/Module: TENT

Kursnamn/Course title: Multivariate statistical method

Institution/Department: DDA

Jag intygar att varken mobil eller något annat otillåtet hjälpmedel finns tillgängligt under tentamen.
I confirm that no mobile or other non-permitted aids are available during the examination. ☒

Inlämnat: antal lösblad 6 tentamensformulär ☐
Enclosed: number of sheets exam booklet

Markera behandlade uppgifter med X/Mark tasks attempted with an X

X här/here	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15
Erhållna poäng Points obtained	X	X	X	6	X										
X här/here	16	17	18	19	20	21	22	23	24	25	26	27	28	29	30
Erhållna poäng Points obtained															

Anvisningar/Instructions

- Skriv AID-nummer, datum, utb.kod, modul på varje blad som lämnas in/Write AID number, date, edu.code and module on every sheet that is handed in
- På varje papper får högst en uppgift lösas om inget annat anges/
Maximum one task per sheet unless otherwise instructed
- Skriv endast på papprets ena sida om inget annat anges/
Use only one side of each sheet unless otherwise instructed
- Numrera de papper som lämnas in/Number every sheet that is handed in
- Använd inte röd penna/Do not use a red pen/pencil

Sen inlämning
Late hand in ☐

Klockslag
Time

Orsak
Reason

Σ Poäng/Points: 12 Betyg/Grade: D

Examinator/Examiner: K. B. B.

Problem 1.

Prove: $\Sigma = (1-\alpha)I + \alpha \vec{1}_p \vec{1}_p^T$

$$= A + B$$

let $(1-\alpha)I$ equals A ,

and $\alpha(\vec{1}_p \vec{1}_p^T)$ equals B .

So,

I prove symmetric.

① $A = (1-\alpha)I$

$\because I$ is the identity matrix ①

$$\alpha \geq 0 \quad \text{②}$$

$\Rightarrow A$ is symmetric.

② $B = \alpha \cdot \vec{1}_p \cdot \vec{1}_p^T$

$$\because \alpha \geq 0$$

$\vec{1}_p$ is a vector of p ones.

$\therefore \vec{1}_p \cdot \vec{1}_p^T$ is symmetric

$\therefore \alpha \cdot \vec{1}_p \cdot \vec{1}_p^T = B$ is symmetric.

From above proof above,

$$\Sigma = A + B$$

according to matrix calcu.

$\therefore \Sigma$ is symmetric.

II. prove positive-semi-definite.

according to definition, positive-semi-definite means all the eigenvalues of matrix should ≥ 0 . No

① $A = (1-\alpha)I$, Not in general

$$\text{so } A = P_A \cdot \Lambda_A \cdot P_A^T$$

and all the diag $(\Lambda) \geq 0$,

② $B = \alpha \cdot \vec{1}_p \cdot \vec{1}_p^T$

$$B = P_B \cdot \Lambda_B \cdot P_B^T$$

$\because \vec{1}_p$ is a vector of p ones,

\therefore all the diag $(B) \geq 0$

From proof above,

$$\Sigma = A + B = P_\Sigma \cdot \Lambda_\Sigma \cdot P_\Sigma^T$$

$$\Lambda_\Sigma = \Lambda_A + \Lambda_B$$

all ~~elements~~ diagonal elements of $\Lambda_\Sigma \geq 0$.

$\therefore \Sigma$ is positive-semi-definite. No

$\therefore \Sigma$ is symmetric-positive-semi-definite.

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Problem 2:

$$\vec{X} \sim N(\vec{\mu}, \Sigma)$$

$$\therefore E(\vec{X}) = \vec{\mu}$$

$$V(\vec{X}) = \Sigma$$

$$\text{when } \vec{\mu} = \vec{0}$$

$$\Sigma = \sigma^2 I$$

$$\therefore E(\vec{X}) = \vec{0}$$

$$V(\vec{X}) = \sigma^2 I$$

which I is the identity matrix of appro size.

$$E(G\vec{X}) = 1 \cdot E(\vec{X}) = \vec{0}$$

$\therefore G$ is an orthogonal matrix,

\therefore All row and column vectors are orthogonal

$$\text{Var}(G\vec{X}) = G \cdot \text{Var}(\vec{X}) \cdot G'$$

$$= 1 \cdot \text{Var}(\vec{X}) \cdot 1$$

$$= \sigma^2 I$$

$$\therefore E(\vec{X}) = E(G\vec{X}) = \vec{0}$$

$$V(\vec{X}) = V(G\vec{X}) = \sigma^2 I$$

$\therefore \vec{X}$ and $G\vec{X}$ have the same distribution.

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Problem 3.

(a)

$$H_0: \bar{x} = \mu$$

$$H_1: \bar{x} \neq \mu$$

$$T^2 = (\bar{x} - \bar{\mu})^T \left(\frac{1}{n} S \right)^{-1} (\bar{x} - \bar{\mu})$$

$$\text{if } T^2 > \frac{(n-1)p}{n-p} F_{p, n-p}(\alpha)$$

reject H_0 .

$$\mu = (182, 182)^T \quad (1)$$

$$\bar{x} = (185.72, 183.84)^T \quad (2)$$

$$n = 25 \quad (3)$$

$$S^{-1} = \begin{bmatrix} 0.022 & -0.015 \\ -0.015 & 0.021 \end{bmatrix} \quad (4)$$

\Rightarrow put (1), (2), (3), (4) to T^2 .

$$\Rightarrow T^2 = \left[\begin{pmatrix} 185.72 \\ 183.84 \end{pmatrix} - \begin{pmatrix} 182 \\ 182 \end{pmatrix} \right]^T \times 25$$

$$\times \begin{bmatrix} 0.022 & -0.015 \\ -0.015 & 0.021 \end{bmatrix}$$

$$\times \left[\begin{pmatrix} 185.72 \\ 183.84 \end{pmatrix} - \begin{pmatrix} 182 \\ 182 \end{pmatrix} \right]$$

$$= (3.72, 1.84) \times 25 \times \begin{bmatrix} 0.022 & -0.015 \\ -0.015 & 0.021 \end{bmatrix} \times \begin{pmatrix} 3.72 \\ 1.84 \end{pmatrix}$$

$$= 25 \times 10.05424 - 0.01718 \times \begin{pmatrix} 3.72 \\ 1.84 \end{pmatrix}$$

$$= 25 \times 0.1701616$$

$$= 4.25404$$

$$n = 25 \quad \alpha = 0.05$$

$$p = 2$$

$$F_{2, 23}(0.05) = 3.42$$

$$\therefore \frac{(n-1)p}{n-p} \cdot F_{p, n-p}(\alpha)$$

$$= \frac{24 \times 2}{23} \times 3.42$$

$$\approx 7.13739$$

$$\therefore T^2 < \frac{(n-1)p}{n-p} \cdot F_{p, n-p}(\alpha)$$

~~accept~~ ~~fail to reject~~ H_0 conclude

the observed sample comes from the desired distribution.

(b)

$$C^2 = 7.13739 = \frac{(n-1)p}{n-p} \cdot F_{p, n-p}(\alpha)$$

$$\sqrt{\lambda_1} \cdot \sqrt{\frac{1}{n} \cdot C^2}$$

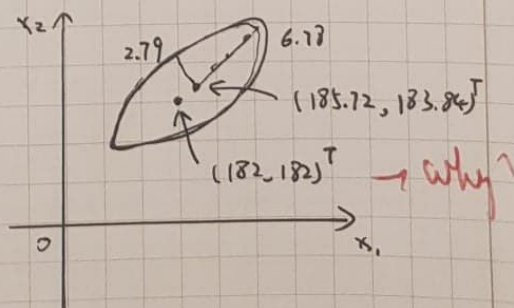
$$= \sqrt{161.055} \times \sqrt{\frac{1}{25} \times 7.13739}$$

$$\approx 6.78$$

$$\sqrt{\lambda_2} \cdot \sqrt{\frac{1}{n} \cdot C^2}$$

$$= \sqrt{27.201} \times \sqrt{\frac{1}{25} \times 7.13739}$$

$$\approx 2.79$$



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From the graph and (a).

we know $\mu = (182, 182)^T$

~~no~~ falls in 95% confidence interval, so

$(182, 182)^T$ must be in the ellipse.

(2)

(c)

$$\Sigma_1 = \begin{bmatrix} 100 & 0 \\ 0 & 100 \end{bmatrix}$$

From the Σ_1 , we see that

$\text{cov}(X_1, X_2) = 0$, and the

measurements assumed come

from a Normal distribution.

$\therefore X_1$ and X_2 are independent.

(1)

$$S = \begin{bmatrix} 91.481 & 66.875 \\ 66.875 & 96.775 \end{bmatrix}$$

the $\text{cov}(X_1, X_2) = 66.875$, which

means X_1 and X_2 are highly correlated to each other.

So it is not reasonable that the

sample comes from a distribution

with var-cov matrix, Σ_1 .

④ Problem 4.

(a) $H_0: \bar{X} = \mu$

$H_1: \bar{X} \neq \mu$

$T^2 = (\bar{X} - \mu)^T \left(\frac{1}{n} S \right)^{-1} (\bar{X} - \mu)$

if $T^2 > \frac{(n-1)p}{n-p} F_{p, n-p}(\alpha)$,

reject H_0 .

$\mu = (182, 182)^T$ ①

$\bar{X} = (185.72, 183.84)^T$ ②

$n = 25$ ③

$S^{-1} = \begin{bmatrix} 0.022 & -0.015 \\ -0.015 & 0.021 \end{bmatrix}$ ④

\Rightarrow put ① ② ③ ④ to T^2 ,

$\Rightarrow T^2 = (3.72, 1.84) \times 25 \times \begin{bmatrix} 0.022 & -0.015 \\ -0.015 & 0.021 \end{bmatrix}$

$\times \begin{pmatrix} 3.72 \\ 1.84 \end{pmatrix}$

$= 4.25404$

$n = 25$ $p = 2$ $\alpha = 0.05$

$F_{2, 23}(0.05) = 3.42$

$\therefore \frac{(n-1)p}{n-p} = F_{p, n-p}(\alpha)$

$\approx \frac{24 \times 2}{23} \times 3.42$

≈ 7.13739

$\therefore T^2 < \frac{(n-1)p}{n-p} \cdot F_{p, n-p}(\alpha)$

\therefore ~~accept~~ ^{Fail to reject} H_0 , conclude the observed sample comes from the desired distribution.

(b)

$C^2 = 7.13739 = \frac{(n-1)p}{n-p} \cdot F_{p, n-p}(\alpha)$

$\sqrt{\lambda_1} \cdot \sqrt{\frac{1}{n} C^2}$

~~$\approx \sqrt{6.78} \approx 2.6$~~

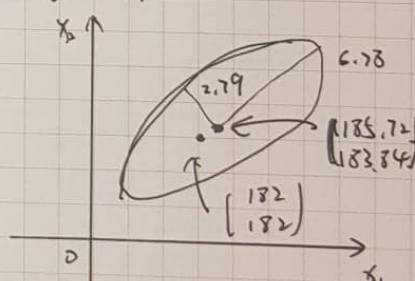
$= \sqrt{161.055 \times \frac{1}{25} \times 7.13739}$

≈ 6.78

$\sqrt{\lambda_2} \cdot \sqrt{\frac{1}{n} C^2}$

$= \sqrt{27.201 \times \frac{1}{25} \times 7.13739}$

≈ 2.79



From the graph and (a), we know $\mu = (182, 182)^T$ falls in 95% confidence interval, so $(182, 182)^T$ must be in the ellipse.



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(c)

$$\Sigma_2 = \begin{bmatrix} 100 & 50 \\ 50 & 100 \end{bmatrix}$$

Compared to Σ_1 , we can

see the $\text{cov}(X_1, X_2) = 50$,

which means these two variables
are correlated.

$$S = \begin{bmatrix} 91.481 & 66.875 \\ 66.875 & 96.775 \end{bmatrix}$$

the $\text{cov}(X_1, X_2) = \cancel{66.875} 66.875$,

So it seems more reasonable that
the sample shows similar variable
relationship with the population.