

Data Analytics

EEE 4774 & 6777

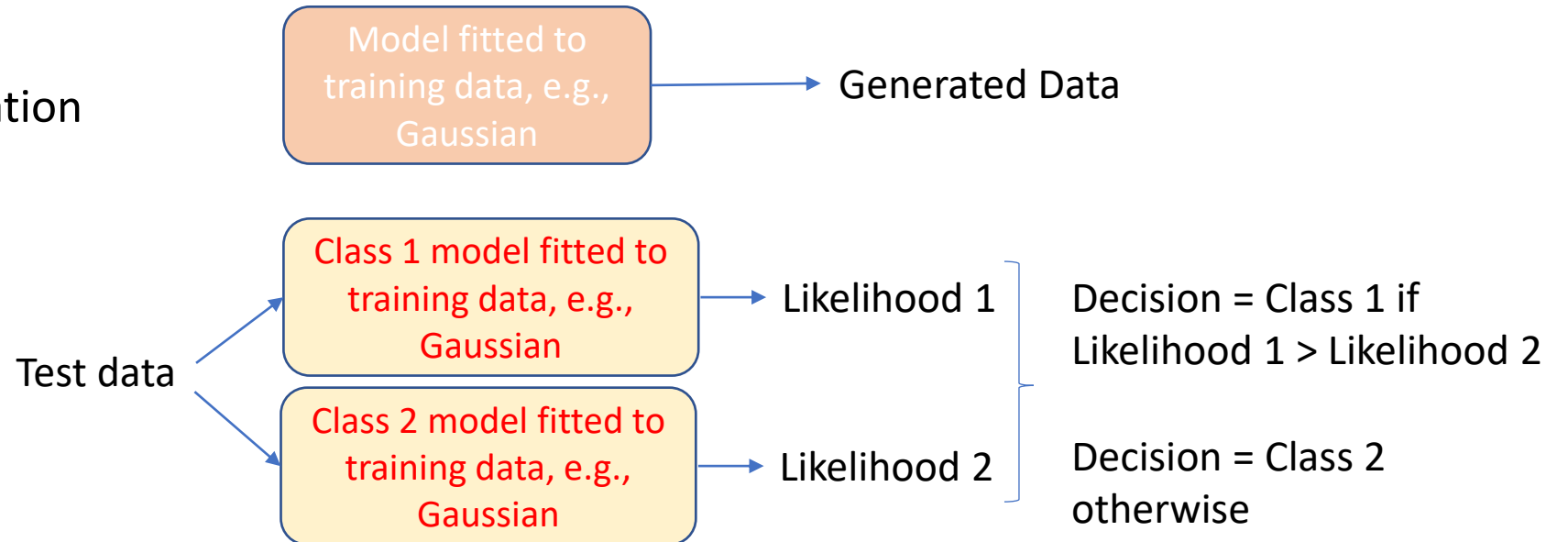
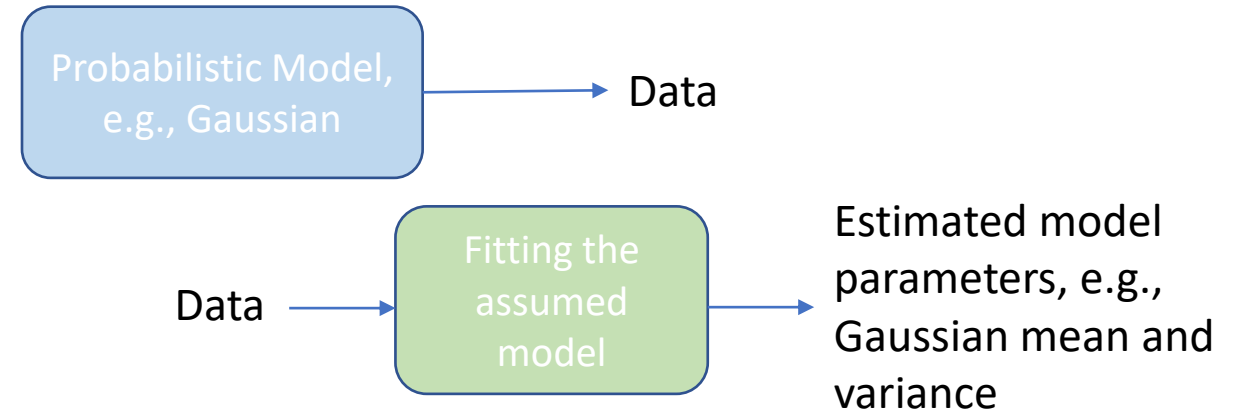
Module 2

Parameter Estimation

Spring 2022

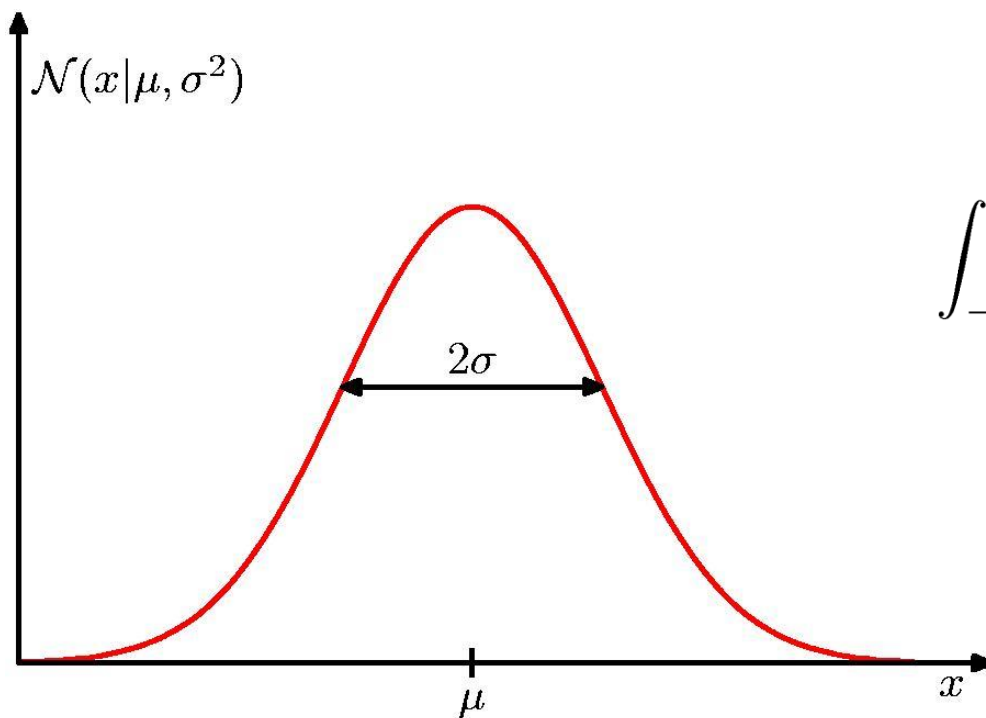
Parameter Estimation for Model Fitting

- Generative model assumption
- Model inference/fitting from data
- Used for
 - Data generation
 - Missing value estimation
 - Classification



Gaussian Distribution

$$\mathcal{N}(x|\mu, \sigma^2) = \frac{1}{(2\pi\sigma^2)^{1/2}} \exp \left\{ -\frac{1}{2\sigma^2} (x - \mu)^2 \right\}$$

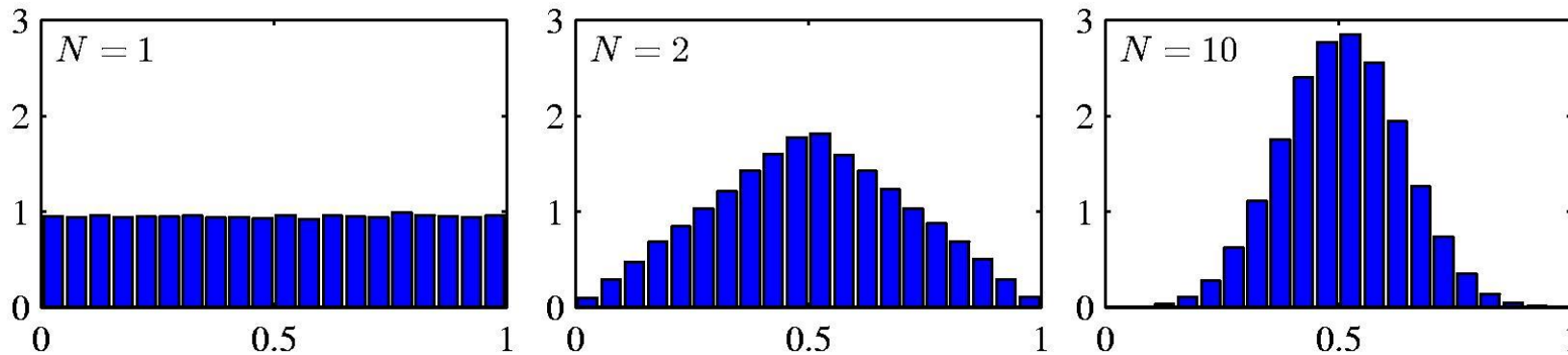


$$\mathcal{N}(x|\mu, \sigma^2) > 0$$

$$\int_{-\infty}^{\infty} \mathcal{N}(x|\mu, \sigma^2) \, dx = 1$$

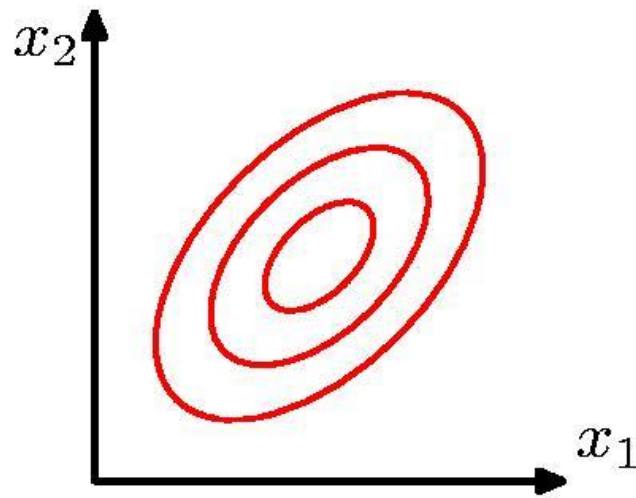
Central Limit Theorem

- The distribution of the sum of N i.i.d. random variables becomes increasingly Gaussian as N grows.
- Example: N uniform $[0,1]$ random variables.



Multivariate Gaussian

$$\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\boldsymbol{\Sigma}|^{1/2}} \exp \left\{ -\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu}) \right\}$$



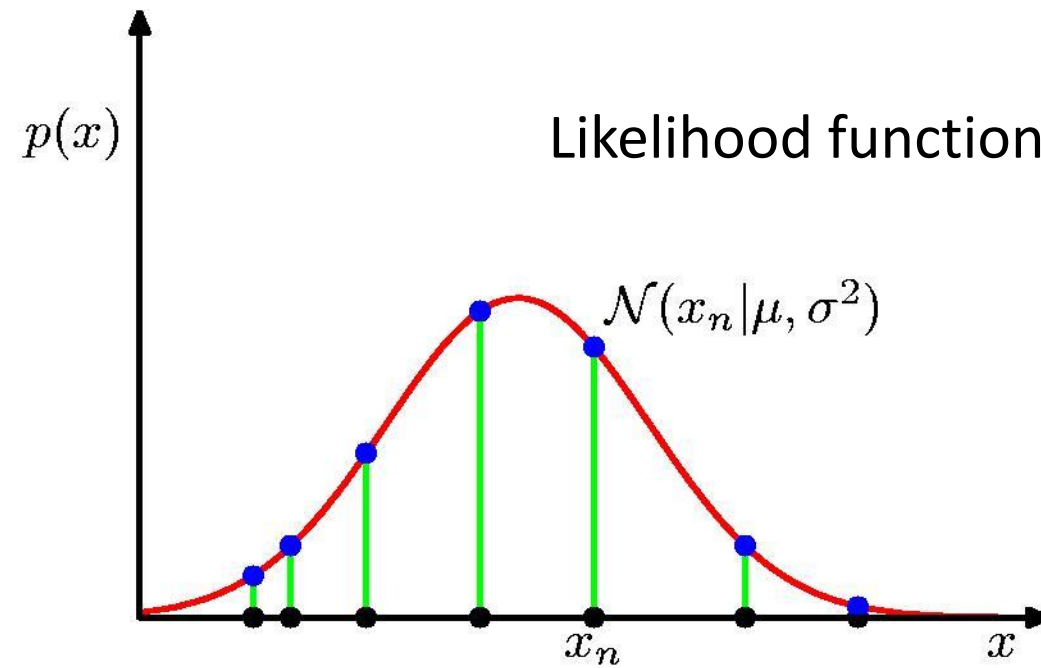
Gaussian Mean and Variance

$$\mathbb{E}[x] = \int_{-\infty}^{\infty} \mathcal{N}(x|\mu, \sigma^2) x \, dx = \mu$$

$$\mathbb{E}[x^2] = \int_{-\infty}^{\infty} \mathcal{N}(x|\mu, \sigma^2) x^2 \, dx = \mu^2 + \sigma^2$$

$$\text{var}[x] = \mathbb{E}[x^2] - \mathbb{E}[x]^2 = \sigma^2$$

Gaussian Parameter Estimation



$$p(\mathbf{x} | \mu, \sigma^2) = \prod_{n=1}^N \mathcal{N}(x_n | \mu, \sigma^2)$$

Maximum Likelihood (ML) Estimation

$$\ln p(\mathbf{x}|\mu, \sigma^2) = -\frac{1}{2\sigma^2} \sum_{n=1}^N (x_n - \mu)^2 - \frac{N}{2} \ln \sigma^2 - \frac{N}{2} \ln(2\pi)$$

$$\mu_{\text{ML}} = \frac{1}{N} \sum_{n=1}^N x_n \qquad \sigma_{\text{ML}}^2 = \frac{1}{N} \sum_{n=1}^N (x_n - \mu_{\text{ML}})^2$$

Properties of μ_{ML} and σ_{ML}^2

$$\mathbb{E}[\mu_{\text{ML}}] = \mu$$

$$\mathbb{E}[\sigma_{\text{ML}}^2] = \left(\frac{N-1}{N} \right) \sigma^2$$

$$\begin{aligned} \tilde{\sigma}^2 &= \frac{N}{N-1} \sigma_{\text{ML}}^2 \\ &= \frac{1}{N-1} \sum_{n=1}^N (x_n - \mu_{\text{ML}})^2 \end{aligned}$$

Parameter estimation: Bayesian for Gaussian

- Gaussian prior $p(\mu) = \mathcal{N}(\mu|\mu_0, \sigma_0^2)$.

- posterior $p(\mu|\mathbf{x}) \propto p(\mathbf{x}|\mu)p(\mu)$. $p(\mu|\mathbf{x}) = \mathcal{N}(\mu|\mu_N, \sigma_N^2)$

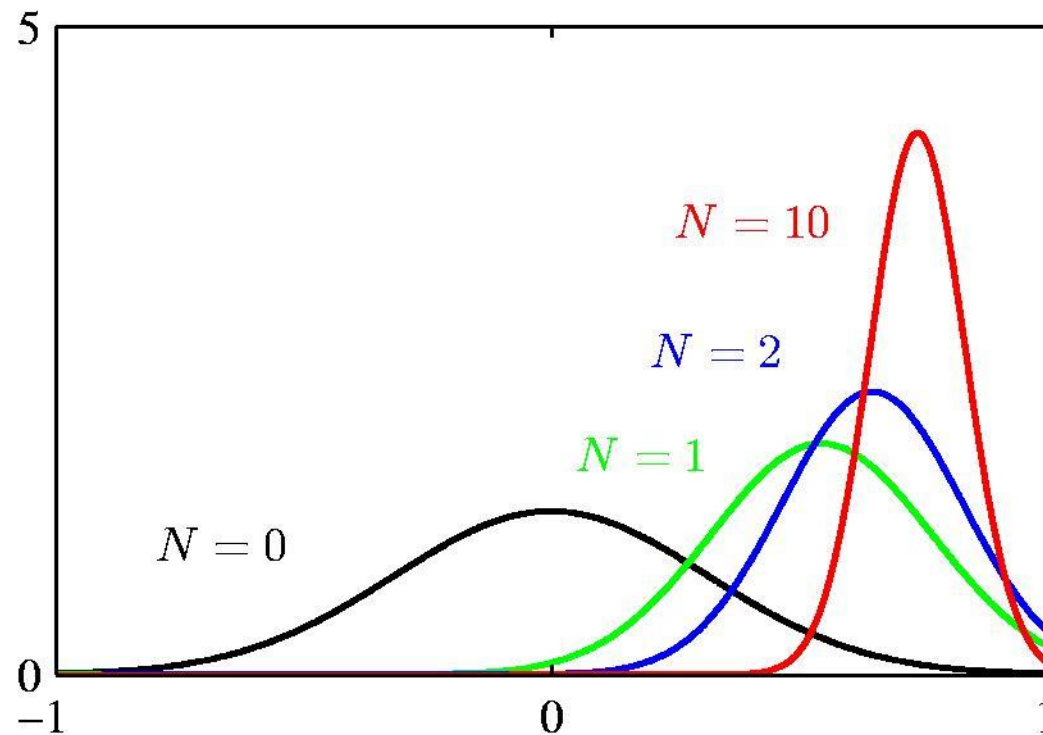
$$\mu_N = \frac{\sigma^2}{N\sigma_0^2 + \sigma^2}\mu_0 + \frac{N\sigma_0^2}{N\sigma_0^2 + \sigma^2}\mu_{\text{ML}}, \quad \mu_{\text{ML}} = \frac{1}{N} \sum_{n=1}^N x_n$$

$$\frac{1}{\sigma_N^2} = \frac{1}{\sigma_0^2} + \frac{N}{\sigma^2}.$$

	$N = 0$	$N \rightarrow \infty$
μ_N	μ_0	μ_{ML}
σ_N^2	σ_0^2	0

Parameter estimation: Bayesian for Gaussian

- Example: $p(\mu|\mathbf{x}) = \mathcal{N}(\mu|\mu_N, \sigma_N^2)$ for $N = 0, 1, 2$ and 10.
- True mean = 0.8



Python Example

- Dataset “adult.csv” includes information (attributes/features) of a number of people such as age, gender, occupation, and whether income less than or greater than 50K per year.
- Compute the mean and standard deviation of age for two groups of people:
 - Income \leq 50K
 - Income $>$ 50K
- Use the mean and standard deviation statistics to fit a Gaussian model to each group
- Compute the histogram to check if the Gaussian model assumptions are suitable.

Python Example

- Dataset “iris.csv” includes 4 measurements (attributes/features) from a number of iris plants.
- Each plant is from one of 3 classes.
- Fit a 4-dimensional multivariate Gaussian to data from each class.
- Compute the likelihood of each test instance under the 3 Gaussian models to make a classification decision.
- Choose the class (i.e., model) with the highest likelihood value as the class prediction for each test instance.
- Compute the accuracy by comparing the predicted class labels with the ground truth.

Bernoulli Distribution

- Coin flipping: heads=1, tails=0

$$p(x = 1|\mu) = \mu$$

- Bernoulli Distribution

$$\text{Bern}(x|\mu) = \mu^x(1 - \mu)^{1-x}$$

$$\mathbb{E}[x] = \mu$$

$$\text{var}[x] = \mu(1 - \mu)$$

Binomial Distribution

- N coin flips:

$$p(m \text{ heads} | N, \mu)$$

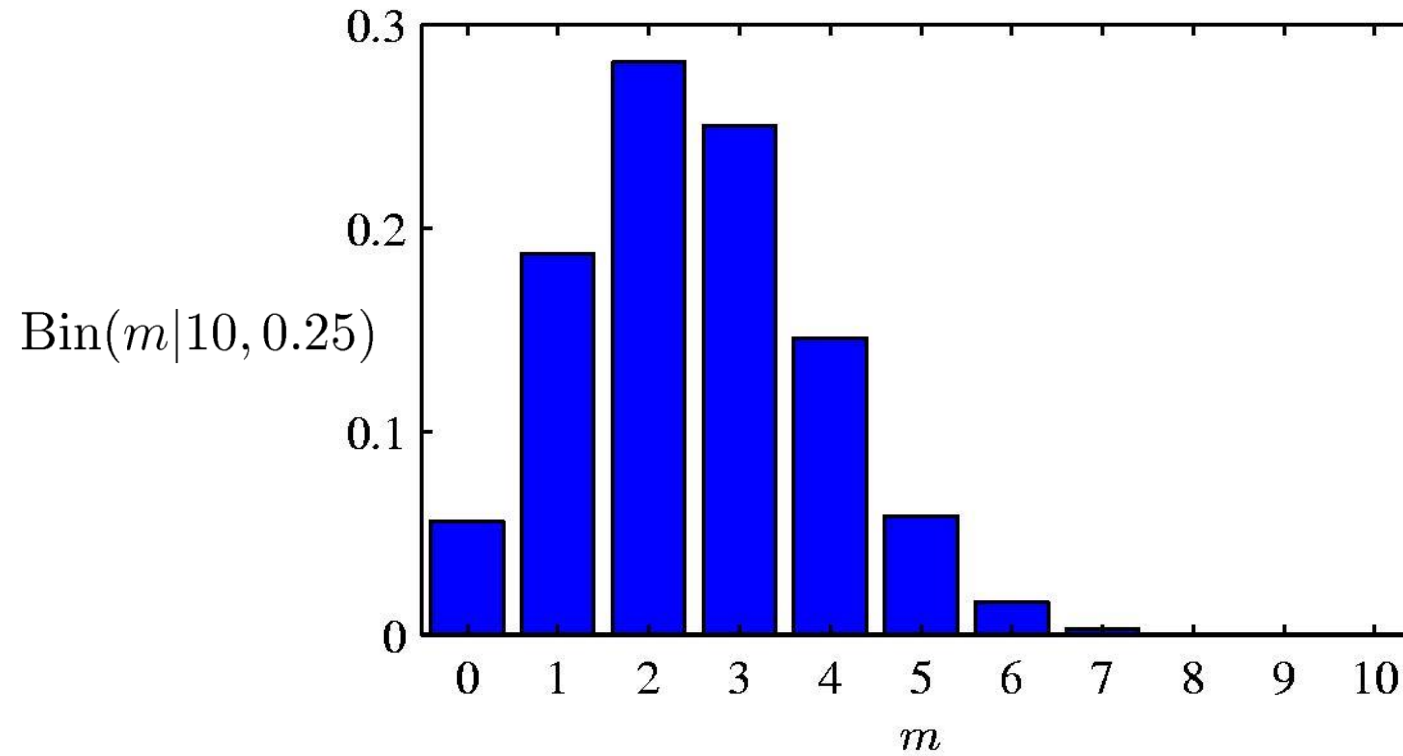
- Binomial Distribution

$$\text{Bin}(m | N, \mu) = \binom{N}{m} \mu^m (1 - \mu)^{N-m}$$

$$\mathbb{E}[m] \equiv \sum_{m=0}^N m \text{Bin}(m | N, \mu) = N\mu$$

$$\text{var}[m] \equiv \sum_{m=0}^N (m - \mathbb{E}[m])^2 \text{Bin}(m | N, \mu) = N\mu(1 - \mu)$$

Binomial Distribution



Parameter Estimation: ML for Bernoulli

Given: $\mathcal{D} = \{x_1, \dots, x_N\}$, m heads (1), $N - m$ tails (0)

$$p(\mathcal{D}|\mu) = \prod_{n=1}^N p(x_n|\mu) = \prod_{n=1}^N \mu^{x_n} (1 - \mu)^{1-x_n}$$

$$\ln p(\mathcal{D}|\mu) = \sum_{n=1}^N \ln p(x_n|\mu) = \sum_{n=1}^N \{x_n \ln \mu + (1 - x_n) \ln(1 - \mu)\}$$

$$\mu_{\text{ML}} = \frac{1}{N} \sum_{n=1}^N x_n = \frac{m}{N}$$

Parameter Estimation: ML for Bernoulli

Example: $\mathcal{D} = \{1, 1, 1\} \rightarrow \mu_{\text{ML}} = \frac{3}{3} = 1$

- Prediction: all future tosses will land heads up
- Overfitting to \mathcal{D}

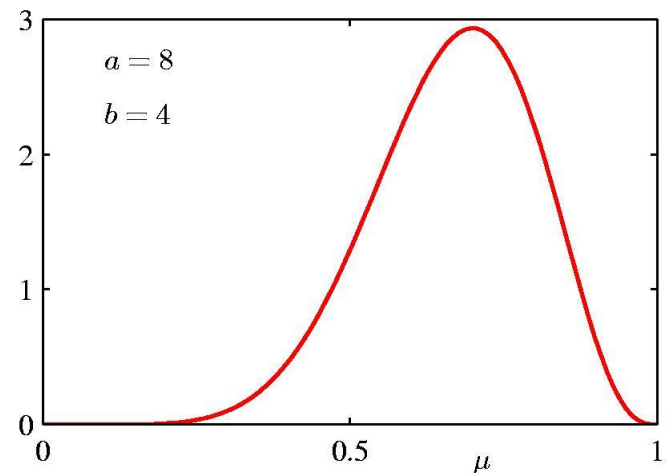
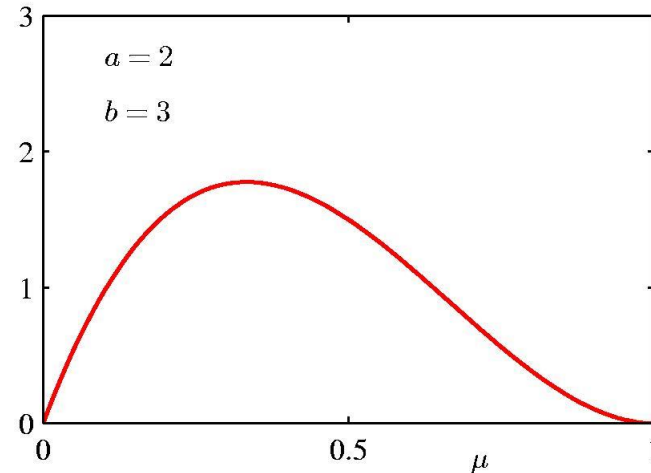
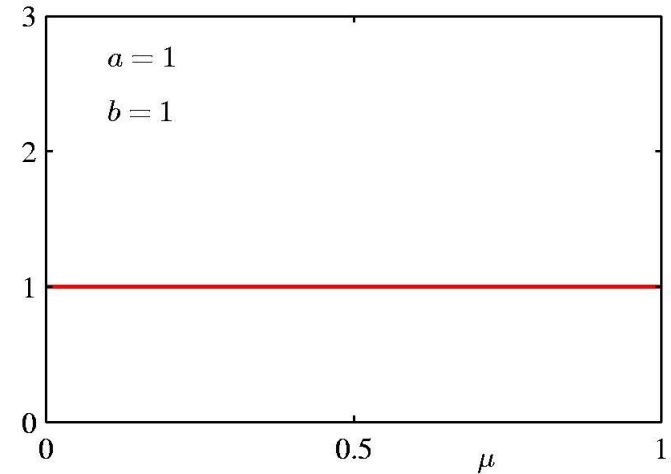
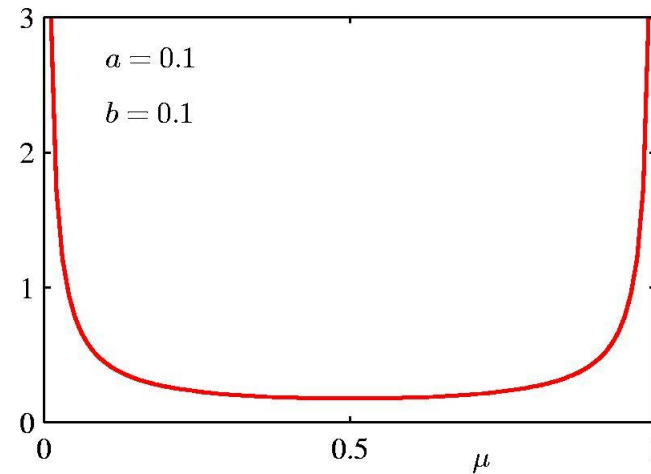
Beta Distribution

- Distribution over $\mu \in [0, 1]$

$$\text{Beta}(\mu|a, b) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \mu^{a-1} (1-\mu)^{b-1}$$

$$\mathbb{E}[\mu] = \frac{a}{a+b}$$

$$\text{var}[\mu] = \frac{ab}{(a+b)^2(a+b+1)}$$



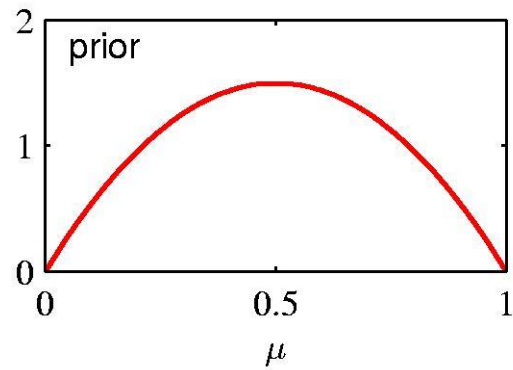
Parameter Estimation: Bayesian for Bernoulli

$$\begin{aligned} p(\mu|a_0, b_0, \mathcal{D}) &\propto p(\mathcal{D}|\mu)p(\mu|a_0, b_0) \\ &= \left(\prod_{n=1}^N \mu^{x_n} (1 - \mu)^{1-x_n} \right) \text{Beta}(\mu|a_0, b_0) \\ &\propto \mu^{m+a_0-1} (1 - \mu)^{(N-m)+b_0-1} \\ &\propto \text{Beta}(\mu|a_N, b_N) \end{aligned}$$

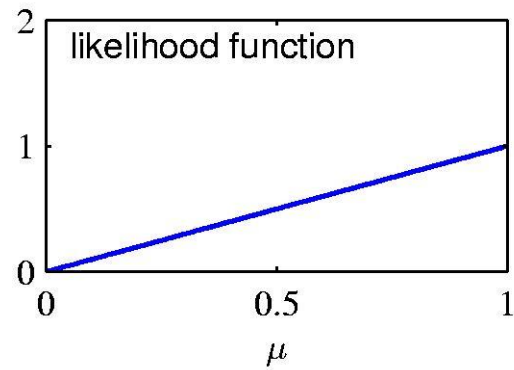
$$a_N = a_0 + m \quad b_N = b_0 + (N - m)$$

The **Beta** distribution provides the *conjugate prior* for the **Bernoulli** distribution.

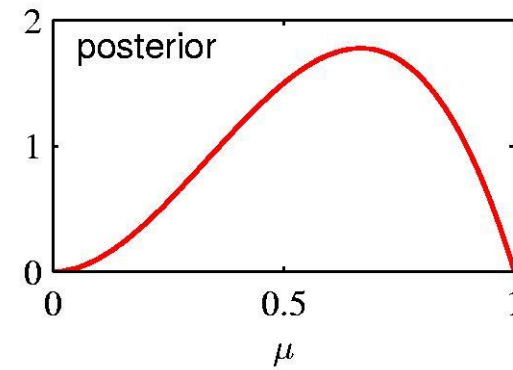
Prior · Likelihood = Posterior



$a=2, b=2$
Beta



$N=m=1$



$a=3, b=2$
Beta

Properties of the Beta Posterior

As the size of the data set, N , increases

$$a_N \rightarrow m$$

$$b_N \rightarrow N - m$$

$$\mathbb{E}[\mu] = \frac{a_N}{a_N + b_N} \rightarrow \frac{m}{N} = \mu_{\text{ML}}$$

$$\text{var}[\mu] = \frac{a_N b_N}{(a_N + b_N)^2 (a_N + b_N + 1)} \rightarrow 0$$

Empirical Bayes

- In full Bayes a probability distribution is assumed/known for hyperparameters (a_0, b_0)
- In empirical Bayes, hyperparameters (a_0, b_0) are estimated from data

- Example:

Consider a baseball dataset with the total number of batting attempts A_i and the total number of hits H_i given for each player i .

The ML estimate for the probability of hit is given by the batting average H_i/A_i

To obtain a MAP estimate if you assume a Beta prior on the probability of hit, then you will either need to assume values for hyperparameters (a_0, b_0) such as $(100, 300)$ or estimate these values from the dataset.

Categorical Distribution

1-of-K coding scheme: $\mathbf{x} = (0, 0, 1, 0, 0, 0)^T$

$$p(\mathbf{x}|\boldsymbol{\mu}) = \prod_{k=1}^K \mu_k^{x_k}$$

$$\forall k : \mu_k \geq 0 \quad \text{and} \quad \sum_{k=1}^K \mu_k = 1$$

$$\mathbb{E}[\mathbf{x}|\boldsymbol{\mu}] = \sum_{\mathbf{x}} p(\mathbf{x}|\boldsymbol{\mu}) \mathbf{x} = (\mu_1, \dots, \mu_K)^T = \boldsymbol{\mu}$$

$$\sum_{\mathbf{x}} p(\mathbf{x}|\boldsymbol{\mu}) = \sum_{k=1}^K \mu_k = 1$$

Multinomial Distribution

$$\begin{aligned}\text{Mult}(m_1, m_2, \dots, m_K | \boldsymbol{\mu}, N) &= \binom{N}{m_1 m_2 \dots m_K} \prod_{k=1}^K \mu_k^{m_k} \\ \mathbb{E}[m_k] &= N\mu_k \\ \text{var}[m_k] &= N\mu_k(1 - \mu_k) \\ \text{cov}[m_j m_k] &= -N\mu_j \mu_k\end{aligned}$$

Parameter estimation: ML for Multinomial

- Given: $\mathcal{D} = \{\mathbf{x}_1, \dots, \mathbf{x}_N\}$

$$p(\mathcal{D}|\boldsymbol{\mu}) = \prod_{n=1}^N \prod_{k=1}^K \mu_k^{x_{nk}} = \prod_{k=1}^K \mu_k^{(\sum_n x_{nk})} = \prod_{k=1}^K \mu_k^{m_k}$$

- Ensure $\sum_k \mu_k = 1$, use a Lagrange multiplier,

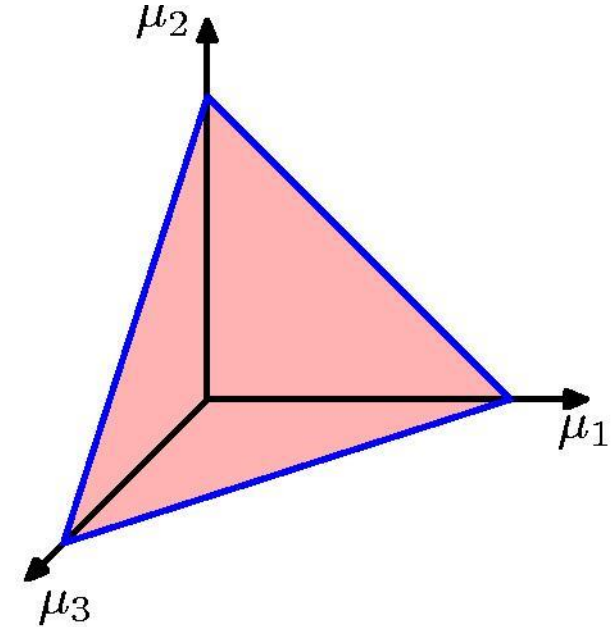
$$\sum_{k=1}^K m_k \ln \mu_k + \lambda \left(\sum_{k=1}^K \mu_k - 1 \right)$$

$$\mu_k = -m_k / \lambda \quad \mu_k^{\text{ML}} = \frac{m_k}{N}$$

Dirichlet Distribution

$$\text{Dir}(\boldsymbol{\mu}|\boldsymbol{\alpha}) = \frac{\Gamma(\alpha_0)}{\Gamma(\alpha_1) \cdots \Gamma(\alpha_K)} \prod_{k=1}^K \mu_k^{\alpha_k - 1}$$

$$\alpha_0 = \sum_{k=1}^K \alpha_k$$



Conjugate prior for the **multinomial** distribution.

Parameter estimation: Bayesian for Multinomial

$$p(\boldsymbol{\mu}|\mathcal{D}, \boldsymbol{\alpha}) \propto p(\mathcal{D}|\boldsymbol{\mu})p(\boldsymbol{\mu}|\boldsymbol{\alpha}) \propto \prod_{k=1}^K \mu_k^{\alpha_k + m_k - 1}$$

$$\begin{aligned} p(\boldsymbol{\mu}|\mathcal{D}, \boldsymbol{\alpha}) &= \text{Dir}(\boldsymbol{\mu}|\boldsymbol{\alpha} + \mathbf{m}) \\ &= \frac{\Gamma(\alpha_0 + N)}{\Gamma(\alpha_1 + m_1) \cdots \Gamma(\alpha_K + m_K)} \prod_{k=1}^K \mu_k^{\alpha_k + m_k - 1} \end{aligned}$$

Exponential Family

$$p(\mathbf{x}|\boldsymbol{\eta}) = h(\mathbf{x})g(\boldsymbol{\eta}) \exp \{ \boldsymbol{\eta}^T \mathbf{u}(\mathbf{x}) \}$$

- where $\boldsymbol{\eta}$ is the *natural parameter* and

$$g(\boldsymbol{\eta}) \int h(\mathbf{x}) \exp \{ \boldsymbol{\eta}^T \mathbf{u}(\mathbf{x}) \} d\mathbf{x} = 1$$

- so $g(\boldsymbol{\eta})$ can be interpreted as a normalization coefficient.
- For any member of the exponential family, there exists a *conjugate prior*, which makes the posterior the same distribution as itself

Nonparametric Methods

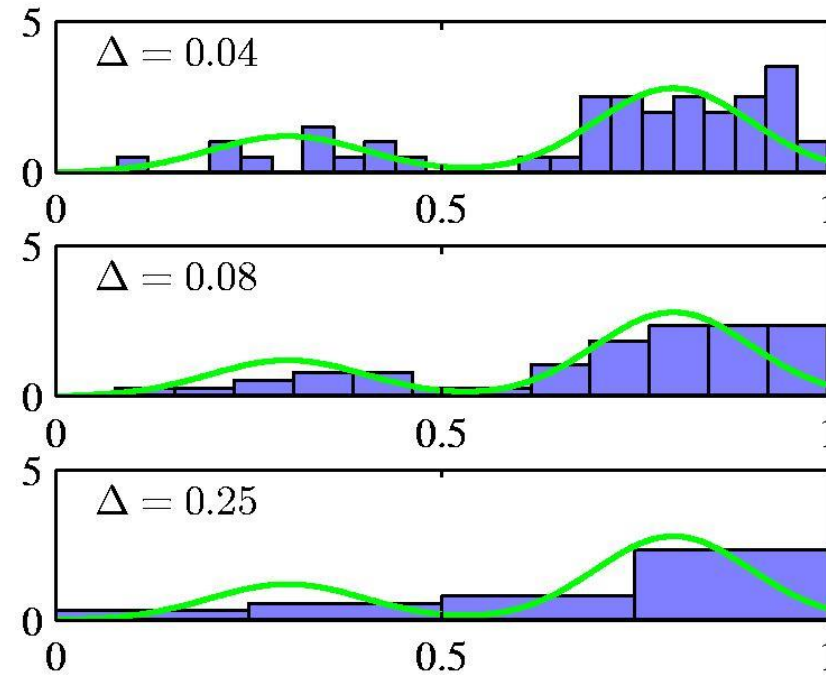
- Parametric distribution models are restricted to specific forms,
 - which may not always be suitable;
 - for example, consider modelling a multimodal distribution with a single, unimodal model.
- Nonparametric approaches
 - make few assumptions about the overall shape of the distribution being modelled.

Histogram

Histogram methods partition the data space into distinct bins with widths Δ_i and count the number of observations, n_i , in each bin.

$$p_i = \frac{n_i}{N\Delta_i}$$

- Often, the same width is used for all bins, $\Delta_i = \Delta$.
- Δ acts as a smoothing parameter.



- In a D-dimensional space, using M bins in each dimension will require M^D bins!

Curse of dimensionality!

Kernel Methods

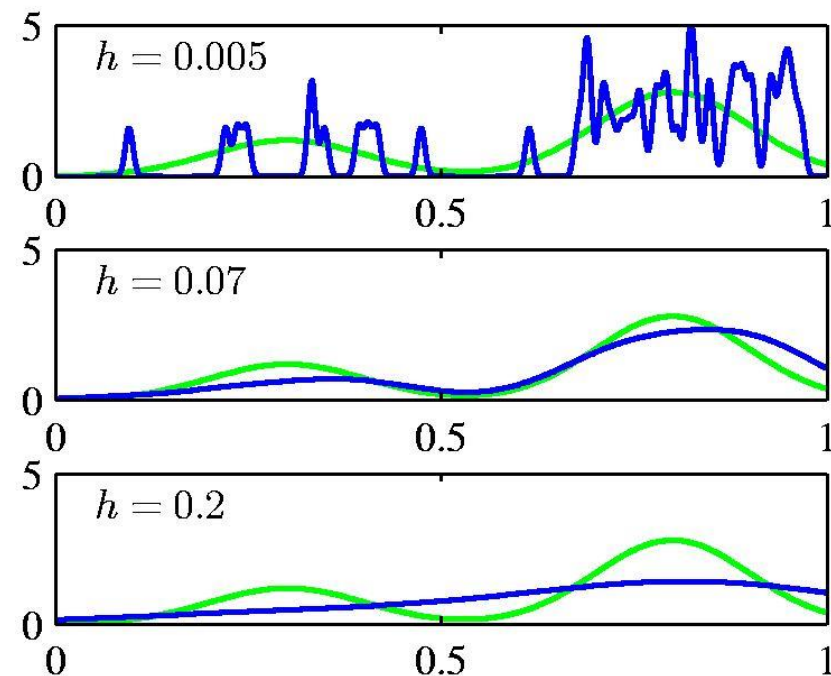
To avoid discontinuities in $p(x)$, use a smooth kernel, e.g. a Gaussian

$$p(\mathbf{x}) = \frac{1}{N} \sum_{n=1}^N \frac{1}{(2\pi h^2)^{D/2}} \exp \left\{ -\frac{\|\mathbf{x} - \mathbf{x}_n\|^2}{2h^2} \right\}$$

Any kernel such that

$$\begin{aligned} k(\mathbf{u}) &\geq 0, \\ \int k(\mathbf{u}) d\mathbf{u} &= 1 \end{aligned}$$

will work.



h acts as a smoother.