

## Distributions: $T$ , $F$ , and Normal Approximation

The purpose of both of these distributions is to allow for inferences about  $\mu$  and  $\sigma$  in an unknown distribution. Both are quotients of known distributions.

### Preliminaries

**Sample Mean:** Let  $Y_1, \dots, Y_n$  be a random, independent sample from a distribution with mean  $\mu$  and variance  $\sigma^2$ . Then,

$$\bar{Y} = \frac{1}{n} \sum_{i=1}^n Y_i \quad \text{Sample Mean}$$

is a distribution with mean  $\bar{\mu} = \mu$  and variance  $\bar{\sigma}^2 = \frac{\sigma^2}{n}$ . If the underlying distribution is a normal distribution, then  $\frac{\bar{Y} - \mu}{\sigma/\sqrt{n}}$  is a *standard* normal distribution.

**Sample Variance:** The *sample variance* is defined as

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2. \quad \text{Sample Variance}$$

It is important to note that the sample variance is found for samples drawn from a distribution; for population standard deviation/variance, we use  $n$  instead of  $n-1$  in the denominator.

When  $Y_i$  is a normal distribution, then  $\frac{(n-1)S^2}{\sigma^2}$  is a  $\chi^2$  distribution with  $n-1$  df —  $S^2$  and  $\bar{Y}$  are independent.

### Definition of $T$ Distribution

Let  $Z$  be a standard normal distribution,  $W$  be  $\chi^2$  with  $\nu$  df, and  $Z$  and  $W$  be independent. Then,

$$T = \frac{Z}{\sqrt{W/\nu}}$$

has a  $T$  distribution with  $\nu$  df.

**Creating a  $T$  Distribution:** Let  $Y_i$  be sampled from a normal distribution with mean  $\mu$  and standard deviation  $\sigma$ .

Then,  $Z = \frac{\bar{Y} - \mu}{\sigma/\sqrt{n}}$  is a standard normal distribution, and  $W = \frac{(n-1)S^2}{\sigma^2}$  is  $\chi^2$  with  $n-1$  df.

So,

$$\begin{aligned} T &= \frac{Z}{\sqrt{W/(n-1)}} \\ &= \frac{(\bar{Y} - \mu)\sqrt{n}}{\sigma} \sqrt{\frac{(n-1)\sigma^2}{S^2}} \\ &= \frac{(\bar{Y} - \mu)\sqrt{n}}{S} \end{aligned}$$

has a  $T$  distribution with  $n-1$  df.

**$T$  Distribution:** Let  $Y_1, \dots, Y_6$  be samples from a normal distribution with unknown  $\mu, \sigma$ . Estimate  $P(|\bar{Y} - \mu| < (2S/\sqrt{n}))$ .

Thus, we have

$$\begin{aligned} P\left(|\bar{Y} - \mu| \leq \frac{2S}{\sqrt{n}}\right) &= P\left(-2 \leq \frac{\sqrt{n}(\bar{Y} - \mu)}{S} \leq 2\right) \\ &= P(-2 \leq T \leq 2) \end{aligned}$$

Thus, for  $n = 6$ , we have that our random variable  $T$  has 5 df. By looking at a  $T$  distribution table, we can find that  $P \approx 0.9$ . We can also use R.

## Definition of $F$ Distribution

Let  $W_1$  and  $W_2$  be independent  $\chi^2$  distributions with  $\nu_1$  and  $\nu_2$  df respectively. Then, the  $F$  distribution with  $\nu_1$  numerator df and  $\nu_2$  denominator df is found as follows:

$$F = \frac{W_1/\nu_1}{W_2/\nu_2}$$

**Simplifying an  $F$  Distribution:** Let  $n_1$  samples be drawn from normal distribution with mean  $\mu_1$  and variance  $\sigma_1^2$ , and  $n_2$  samples be drawn from normal distribution with mean  $\mu_2$  and variance  $\sigma_2^2$ . Both distributions are independent.

From each of these samples, we find the sample variance, and create  $\chi^2$  distributions with their respective df.

$$\begin{aligned} W_1 &= \frac{(n_1 - 1)S_1^2}{\sigma_1^2} \\ W_2 &= \frac{(n_2 - 1)S_2^2}{\sigma_2^2} \end{aligned}$$

Therefore, we have

$$\begin{aligned} F &= \frac{W_1/(n_1 - 1)}{W_2/(n_2 - 1)} \\ &= \frac{(n_1 - 1)S_1^2 \sigma_2^2 (n_2 - 1)}{\sigma_1^2 (n_1 - 1) (n_2 - 1) S_2^2} \\ &= \frac{\sigma_2^2 S_1^2}{\sigma_1^2 S_2^2} \end{aligned}$$

as an  $F$  distribution with  $n_1 - 1$  numerator df and  $n_2 - 1$  denominator df.

**Applying the  $F$  Distribution:** Let  $n_1 = 6$  and  $n_2 = 10$  be two samples from independent normal distributions with the same  $\sigma^2$ . Find  $b$  such that  $P\left(\frac{S_1^2}{S_2^2} \leq b\right) = 0.95$ .

$$\frac{S_1^2}{S_2^2} = \frac{S_1^2/\sigma^2}{S_2^2/\sigma^2}$$

The given  $F$  distribution has 5 numerator df and 9 denominator df. Therefore, we want to find  $0.95 = P(F_{5,9} < b)$ , or find the 0.95 quantile; in R, we find this with the `qt` function.

## Normal Approximation of Binomial

Recall that a binomial distribution  $Y$  with  $n$  trials and  $p$  probability of success has probabilities found below:

$$P(Y \leq \ell) = \sum_{k=0}^{\ell} \binom{n}{k} p^k (1-p)^{n-k}.$$

For very large  $n$ , this sum is hard to calculate. We could approximate with the Poisson distribution, but this still requires a lot of calculations and large factorial values. Instead, we will try the following:

$$\begin{aligned} X_i &= \begin{cases} 1 & i \text{ trial success} \\ 0 & i \text{ trial failure} \end{cases} \\ E(X_i) &= p \\ E(X_i^2) &= p \\ V(X_i) &= p(1-p) \\ \bar{X} &= \frac{1}{n} \sum_{i=1}^n X_i = \frac{Y}{n} \\ E(\bar{X}) &= p \\ V(\bar{X}) &= \frac{p(1-p)}{n} \end{aligned}$$

By the Central Limit Theorem, we approximate  $\bar{X}$  as a normal distribution with mean  $p$  and standard deviation  $\sqrt{\frac{p(1-p)}{n}}$ .

Alternatively, we can create, for large fixed  $n$ ,  $Y = n\bar{X}$  with mean  $np$  and standard deviation  $\sqrt{np(1-p)}$ .

For example, consider  $p = 0.5$ ,  $n = 100$ ,  $Y = \text{number of successes}$ . To find  $P(\frac{Y}{n} > 0.55)$ . By the Central Limit Theorem, this is approximately a normal distribution with mean 0.5 and standard deviation 0.05.

**Applying Central Limit Theorem:** Let  $Y$  be a binomial distribution with  $n = 25$  and  $p = 0.4$ . Then,  $\mu = np = 10$ , and standard deviation  $\sigma = \sqrt{\frac{p(1-p)}{n}} = 5\sqrt{0.24}$ .

To find  $P(Y \leq 8)$ , we can potentially approximate with  $P(X \leq 8.5)$  — the reason we use 8.5 instead of 8 is due to the fact that  $n$  may not be large enough, a process known as the continuity correction.

Using standardization (or R), we find that this probability is approximately 0.269.

The actual probability  $P(Y \leq 8)$  is found as below:

$$\begin{aligned} P(Y \leq 8) &= \sum_{k=0}^8 \binom{25}{k} (0.4)^k (0.6)^{1-k} \\ &= 0.274 \end{aligned}$$

The normal approximation for the binomial is adequate when  $p \pm 3\sqrt{\frac{p(1-p)}{n}} \in (0, 1)$ . Essentially, the binomial trial needs to have an adequate sample size such that the “spread” is small. This is equivalent to  $n \geq 9 \frac{\max(p, 1-p)}{\min(p, 1-p)}$ .

## Estimators

Let  $Y$  be a random variable with an *unknown* distribution.

**Parameter:** Feature of  $Y$ 's distribution that are not computable from samples.

**Examples of Parameters:**  $\mu$ ,  $\sigma$ ,  $m'_k$ , interval  $(a, b) \ni P(y \in I) = 0.95$ .

**Statistic:** Random variable that is computable from samples.

**Examples of Statistics:** sample mean,  $\bar{Y}$ , sample variance,  $S^2$ ,  $Y_{(i)}$ .

**Estimator:** a statistic intended to approximate a parameter. A point estimator estimates a single value.

**Examples of Estimators:**  $\bar{Y}$  as an estimator for  $\mu$ , and  $S^2$  as an estimator of  $\sigma^2$ .

## Two Measures of Goodness of Point Estimators

We want to find  $\theta$ , a constant parameter of the underlying distribution —  $\hat{\theta}$  is a random variable.

If  $E(\hat{\theta})$  is close to  $\theta$ , we can say that  $\hat{\theta}$  is a good estimator — more precisely, we define the bias  $B(\hat{\theta}) = E(\hat{\theta}) - \theta$ , and if  $B(\hat{\theta}) = 0$ , then  $\hat{\theta}$  is an unbiased estimator.

In addition to minimizing bias, to see whether or not an estimator is good requires minimizing the variance of the estimator — the mean squared estimator  $MSE(\hat{\theta}) = V(\hat{\theta}) + B(\hat{\theta})^2$