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Problem 1. Let $\Omega = \{\omega_1, \omega_2, \omega_3\}$, $P(\omega_1) = P(\omega_2) = P(\omega_3) = 1/3$, and define X, Y and Z as follows:

$$X(\omega_1) = 1, X(\omega_2) = 2, X(\omega_3) = 3;$$

$$Y(\omega_1) = 2, Y(\omega_2) = 3, Y(\omega_3) = 1;$$

$$Z(\omega_1) = 3, Z(\omega_2) = 1, Z(\omega_3) = 2.$$

Show that these three random variables have the same probability distribution. Find the probability distributions of $X + Y$, $Y + Z$, and $Z + X$.

Solution:

$$\begin{cases} F_X(1) = P(X(\omega_1)) = 1/3 \\ F_X(2) = P(X(\omega_1)) + P(X(\omega_2)) = 2/3 \\ F_X(3) = P(X(\omega_1)) + P(X(\omega_2)) + P(X(\omega_3)) = 1 \end{cases}$$

$$\begin{cases} F_Y(1) = P(Y(\omega_3)) = 1/3 \\ F_Y(2) = P(Y(\omega_3)) + P(Y(\omega_1)) = 2/3 \\ F_Y(3) = P(Y(\omega_3)) + P(Y(\omega_1)) + P(Y(\omega_2)) = 1 \end{cases}$$

$$\begin{cases} F_Z(1) = P(Z(\omega_2)) = 1/3 \\ F_Z(2) = P(Z(\omega_2)) + P(Z(\omega_3)) = 2/3 \\ F_Z(3) = P(Z(\omega_2)) + P(Z(\omega_3)) + P(Z(\omega_1)) = 1 \end{cases}$$

So X, Y, Z have the same probability distribution.

$$\begin{cases} F_{X+Y}(3) = P((X+Y)(\omega_1)) = 1/3 \\ F_{X+Y}(4) = P((X+Y)(\omega_1)) + P((X+Y)(\omega_3)) = 2/3 \\ F_{X+Y}(5) = P((X+Y)(\omega_1)) + P((X+Y)(\omega_3)) + P((X+Y)(\omega_2)) = 1 \end{cases}$$

$$\begin{cases} F_{Y+Z}(3) = (Y+Z)(\omega_3) = 1/3 \\ F_{Y+Z}(4) = (Y+Z)(\omega_3) + (Y+Z)(\omega_2) = 2/3 \\ F_{Y+Z}(5) = (Y+Z)(\omega_3) + (Y+Z)(\omega_2) + (Y+Z)(\omega_1) = 1 \end{cases}$$

$$\begin{cases} F_{Z+X}(3) = (Z+X)(\omega_2) = 1/3 \\ F_{Z+X}(4) = (Z+X)(\omega_2) + (Z+X)(\omega_1) = 2/3 \\ F_{Z+X}(5) = (Z+X)(\omega_2) + (Z+X)(\omega_1) + (Z+X)(\omega_3) = 1 \end{cases}$$

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Problem 2. In No.1 find the probability distribution of

$$X + Y - Z, \sqrt{(X^2 + Y^2)Z}, \frac{Z}{|X - Y|}$$

Solution:

$$\begin{aligned} & \begin{cases} (X + Y - Z)(\omega_1) = 0 \\ (X + Y - Z)(\omega_2) = 4 \\ (X + Y - Z)(\omega_3) = 2 \end{cases} \\ \Rightarrow & \begin{cases} F_{X+Y-Z}(0) = P((X + Y - Z)(\omega_1)) = 1/3 \\ F_{X+Y-Z}(2) = P((X + Y - Z)(\omega_1)) + P((X + Y - Z)(\omega_3)) = 2/3 \\ F_{X+Y-Z}(4) = P((X + Y - Z)(\omega_1)) + P((X + Y - Z)(\omega_3)) + P((X + Y - Z)(\omega_2)) = 1 \end{cases} \\ & \begin{cases} (\sqrt{(X^2 + Y^2)Z})(\omega_1) = \sqrt{15} \\ (\sqrt{(X^2 + Y^2)Z})(\omega_2) = \sqrt{13} \\ (\sqrt{(X^2 + Y^2)Z})(\omega_3) = \sqrt{20} \end{cases} \\ \Rightarrow & \begin{cases} F_{\sqrt{(X^2 + Y^2)Z}}(\sqrt{13}) = P((\sqrt{(X^2 + Y^2)Z})(\omega_2)) = 1/3 \\ F_{\sqrt{(X^2 + Y^2)Z}}(\sqrt{15}) = P((\sqrt{(X^2 + Y^2)Z})(\omega_2)) + P((\sqrt{(X^2 + Y^2)Z})(\omega_1)) = 2/3 \\ F_{\sqrt{(X^2 + Y^2)Z}}(\sqrt{20}) = P((\sqrt{(X^2 + Y^2)Z})(\omega_2)) + P((\sqrt{(X^2 + Y^2)Z})(\omega_1)) \\ \quad + P((\sqrt{(X^2 + Y^2)Z})(\omega_3)) = 1 \end{cases} \\ & \begin{cases} (\frac{Z}{|X-Y|})(\omega_1) = 3 \\ (\frac{Z}{|X-Y|})(\omega_2) = 1 \\ (\frac{Z}{|X-Y|})(\omega_3) = 1 \end{cases} \\ \Rightarrow & \begin{cases} F_{\frac{Z}{|X-Y|}}(1) = P((\frac{Z}{|X-Y|})(\omega_2)) + P((\frac{Z}{|X-Y|})(\omega_3)) = 2/3 \\ F_{\frac{Z}{|X-Y|}}(3) = P((\frac{Z}{|X-Y|})(\omega_2)) + P((\frac{Z}{|X-Y|})(\omega_3)) + P((\frac{Z}{|X-Y|})(\omega_1)) = 1 \end{cases} \end{aligned}$$

□

Problem 3. Let X be integer-valued and let F be its distribution function. Show that for every x and $a < b$

$$\begin{aligned} P(X = x) &= \lim_{\varepsilon \downarrow 0} [F(x + \varepsilon) - F(x - \varepsilon)] \\ P(a < X < b) &= \lim_{\varepsilon \downarrow 0} [F(b - \varepsilon) - F(a + \varepsilon)] \end{aligned}$$

[The results are true for any random variable but require more advanced proofs even when Ω is countable.]

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Solution:

1.

$$F(x + \varepsilon) - F(x - \varepsilon) = \sum [P(X < x + \varepsilon) - P(X < x - \varepsilon)] = \sum P(x - \varepsilon \leq X < x + \varepsilon)$$

And because X is integer-valued, $P(X)$ is not equal to zero around x only when $X = x$.

$$\lim_{\varepsilon \downarrow 0} \sum P(x - \varepsilon \leq X < x + \varepsilon) = P(X = x)$$

So we have

$$P(X = x) = \lim_{\varepsilon \downarrow 0} [F(x + \varepsilon) - F(x - \varepsilon)]$$

2.

$$\begin{aligned} \lim_{\varepsilon \downarrow 0} [F(b - \varepsilon) - F(a + \varepsilon)] &= P(a + \varepsilon \leq X < b - \varepsilon) \\ &= P(a < X < b) - P(a < X < a + \varepsilon) - P(b - \varepsilon \leq X < b) \end{aligned}$$

When a is an integer, let $\varepsilon < 1$, $P(a < X < a + \varepsilon) = 0$.

When a is not an integer, and a_{int} is the integer next to a , let $\varepsilon < a_{int} - a$, $P(a < X < a + \varepsilon) \leq P(a < X < a_{int}) = 0$.

The same to b , Thus $P(a < X < a + \varepsilon) = P(b - \varepsilon \leq X < b) = 0$,

$$P(a \leq X < b) = \lim_{\varepsilon \downarrow 0} [F(b - \varepsilon) - F(a + \varepsilon)]$$

□

Problem 4. (a) Is there a discrete distribution with support $1, 2, 3, \dots$, such that the value of the PMF at n is proportional to $1/n$?

(b) Is there a discrete distribution with support $1, 2, 3, \dots$, such that the value of the PMF at n is proportional to $1/n^2$?

Solution:

(1) Suppose $P(x) = \frac{k}{n}$, Then $\sum P(x) = k \sum_1^{+\infty} \frac{1}{n} = +\infty$, so it's impossible.

(2) Suppose $P(x) = \frac{k}{n^2}$, Then $\sum P(x) = k \sum_1^{+\infty} \frac{1}{n^2} = \frac{k\pi^2}{6} \implies k = \frac{6}{\pi^2}$, so $P(x) =$

$$\frac{6}{\pi^2 n^2}.$$

□

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Problem 5. Let X have PMF

$$P(X = k) = cp^k/k \text{ for } k = 1, 2, \dots$$

where p is a parameter with $0 < p < 1$ and c is a normalizing constant. We have $c = -1/\log(1-p)$, as seen from the Taylor series

$$-\log(1-p) = p + \frac{p^2}{2} + \frac{p^3}{3} + \dots$$

This distribution is called the *Logarithmic* distribution (because of the log in the above Taylor series), and has often been used in ecology. Find the mean of X .

Solution:

$$E(X) = \sum_1^{+\infty} kP(k) = c \sum_1^{+\infty} p^k = \frac{cp}{1-p} = -\frac{p}{(1-p)\log(1-p)}$$

□

Problem 6. Suppose F is some cumulative distribution function. Then for any real number y , the function F_y defined by $F_y(x) = F(x-y)$ is also a cumulative distribution function. In fact, F_y is just a “shifted” version of F

Solution:

(1)

$$\lim_{x \rightarrow +\infty} F_y(x) = \lim_{x-y \rightarrow +\infty} F(x-y) = 1$$

$$\lim_{x \rightarrow -\infty} F_y(x) = \lim_{x-y \rightarrow -\infty} F(x-y) = 0$$

(2) If $x_1 < x_2$, then $x_1 - y < x_2 - y$, so $F_y(x_1) < F_y(x_2)$.

(3)

$$\lim_{x \rightarrow x_0^+} F_y(x) = \lim_{x-y \rightarrow x_0^+-y} F(x-y) = F(x_0-y) = F_y(x_0)$$

□

Problem 7. Let X be a random variable, with cumulative distribution function F_X . Prove that $P(X = a) = 0$ if and only if the function F_X is continuous at a .

Solution:

$F(a^+) - F(a^-) = \sum_{a^-}^{a^+} P(x)$, if the function F_X is continuous at a , iff $F(a^+) - F(a^-) = 0$. Thus $F(a^+) - F(a^-) = \sum_{a^-}^{a^+} P(x) = 0 \iff P(a) = 0$. □

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Problem 8. Suppose that

$$p_n = cq^{n-1}p, 0 \leq n \leq m$$

where c is a constant and m is a positive integer; cf. (4.4.8). Determine c so that $\sum_{n=1}^m p_n = 1$. (This scheme corresponds to the waiting time for a success when it is supposed to occur within m trials.)

Solution:

$$\sum_{n=1}^m p_n = cp \sum_{n=1}^m q^{n-1} = cp \frac{1 - q^m}{1 - q} = 1$$

So that

$$c = \frac{1 - q}{p(1 - q^m)}.$$

□

Problem 9. A perfect coin is tossed n times. Let Y_n denote the number of heads obtained minus the number of tails. Find the probability distribution of Y_n and its mean. [Hint: there is a simple relation between Y_n and the S_n in Example 9 of 4.4]

Solution: Let H be times of heads, and T be times of tails.

$$Y_n = H - T = H - (n - H) = 2H - n.$$

$$H = \frac{n + Y_n}{2}.$$

$$P_Y(Y_n = x) = \binom{H}{n} / (2^n) = \binom{\frac{n+x}{2}}{n} / (2^n)$$

$$E_Y(x) = \sum_{x=-n}^n \frac{\binom{\frac{n+x}{2}}{n} x}{2^n} = 0$$

□

Problem 10. Let

$$P(X = n) - p_n = \frac{1}{n(n+1)}, n \geq 1$$

Show that it is a probability distribution for X ? Find $P(X \geq m)$ for any m and $E(X)$.

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Solution:

1. For any $n > 0$, $p_n > 0$.

$$2. \sum_1^{+\infty} \frac{1}{n(n+1)} = \sum_1^{+\infty} \frac{1}{n} - \frac{1}{n+1} = 1$$

For reason 1,2, p_n is a probability distribution for X.

$$P(X \geq m) = \sum_m^{+\infty} \frac{1}{n} - \frac{1}{n+1} = \frac{1}{m}$$

$$E(x) = \sum_1^{+\infty} n \left(\frac{1}{n(n+1)} \right) = \sum_1^{+\infty} \frac{1}{n+1} = +\infty$$

□