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4 **Manuscript Title:** Life on the edge: industrial footprint and edge effects variably influence the
5 spatial distribution of a boreal small mammal
6
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16
17 **Open research/data availability statement**
18 All data and scripts used to conduct this analysis are available in the public GitHub repository
19 https://github.com/aidanbrushett/OSM_red_squirrel_distribution.

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Abstract

Wildlife distributions are entrained by the integration of biophysical processes operating across spatial scales and levels of ecological organization. In the Boreal Plains of Alberta, Canada, development from the oil, gas, and timber industries is dramatically restructuring western boreal landscapes by altering both the composition and configuration of wildlife habitat. Using data from 430 camera-traps across 10 boreal landscapes, we examined how landscape structure at multiple spatial scales influences the distribution of the North American red squirrel. We used multi-model selection to compare the relative influence of disturbance composition and configuration metrics on red squirrel detections. Red squirrel abundance was best explained by natural habitat characteristics on local spatial scales but had a negative relationship with anthropogenic edge density at the population scale. Cumulative site disturbance positively affected red squirrels, possibly by providing resource complements at forest edges. Edge density interacted with cumulative site disturbance, such that edge effects were more strongly negative in heavily disturbed landscapes. We suggest that this pattern may be due to population-level shifts arising from increased encounters with mammalian and avian predators at forest edges. Anthropogenic disturbances that create a high amount of edge habitat, especially petroleum exploration “seismic” lines, may have a disproportionately negative impact on red squirrels. The ecological mechanisms through which landscape structure impacts wildlife extend beyond straightforward measures of habitat loss. Both habitat composition and spatial configuration must be considered on appropriate spatial scales when assessing wildlife responses to anthropogenic development.

41 **Introduction**

42 Human-caused alteration of landscapes is a dominant driver of global biodiversity loss
43 (Maxwell et al. 2016, Johnson et al. 2017). As natural resource extraction, agriculture, and urban
44 expansion encroach on wild spaces, wildlife are increasingly displaced or forced to compete for
45 space with anthropogenic activities (Shackelford et al. 2018). Uncovering the ecological
46 mechanisms that drive this process, however, is a long-standing and pressing challenge facing
47 ecologists. Landscapes are complex, multi-scale matrices of habitat, resources, energy, and
48 organisms (Forman and Godron 1981, Holling 1992) Wildlife distributions are entrained by the
49 integration of individual and population responses to biophysical processes across multiple
50 spatial scales (Levin 1992). Both the configuration and composition of landscapes, collectively
51 deemed landscape structure (Dunning et al. 1992), exert a strong influence on local wildlife
52 abundance—variations in either habitat amount or spatial patterning of habitat patches can lead
53 to markedly different distributions of organisms, ecological processes, and community structures
54 (Dunning et al. 1992, Tschamntke et al. 2012).

55 In terrestrial systems, anthropogenic disturbances change landscape structure through
56 changes to both composition and configuration. Natural resource development, for example,
57 reduces the total amount of habitat available to wildlife, while also reconfiguring remaining
58 habitat into smaller, more-isolated patches via fragmentation (Didham et al. 2012, Wilson et al.
59 2016). The resulting mosaic of dissimilar habitat patches can further affect wildlife by
60 influencing connectivity, edge habitat, resource availability, or predation risk (Haddad et al.
61 2015, Pfeifer et al. 2017). While the relationship between habitat loss and declines in terrestrial
62 biodiversity is well established (Brooks et al. 2002), the independent, additional influence of
63 landscape configuration remains more contentious in the literature (Fletcher et al. 2018). A

critical challenge in determining the relative importance of each process is that fragmentation is hierarchically connected to habitat loss in many real-world ecological systems—most habitat loss also results in net increase in fragmentation (Didham et al. 2012, Ruffell et al. 2016). The differential ecological mechanisms of habitat loss and fragmentation have distinct consequences for the assembly and functioning of wildlife communities (Swihart et al. 2006), and both must be investigated when seeking to understand how anthropogenic activity shapes species distributions.

Few regions embody the challenge of understanding landscape structure more than the Boreal Plains of North America. In the 20th and 21st centuries the Boreal Plains have undergone unprecedented structural changes at the collective hands of the timber, mining, and energy industries (Pickell et al. 2015). Superimposed on the naturally ‘patchy’ ecosystems of the Boreal Plains is a pervasive network of cut blocks, roads, seismic lines, well pads, and processing facilities dedicated to the extraction, refinement, and transportation of natural resources (Pasher et al. 2013, Pickell et al. 2015). For instance, over 1.8 million kilometers of seismic lines—linear features used to map underground oil and gas deposits—fragment the western boreal and can persist for decades if unrestored (Lee and Boutin 2006, Dabros et al. 2018). Persistent industrial features create a disproportionate amount of forest edge, increase growth of early-seral vegetation, and establish linear movement corridors through otherwise intact tracts of dense forest habitat. Structural and functional changes have introduced forage subsidies and movement subsidies to boreal landscapes, leading to behavioural- and population-level changes to wildlife communities that vary across species and across intensities of industrial development (Venier et al. 2014, Fisher and Ladle 2022, Barnas et al. 2024).

Past boreal research has heavily emphasized the effects of habitat and disturbance composition on boreal mammals, wherein altered resource availability and movement has

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evoked a multitude of changes to population sizes, wildlife behaviours, and trophic and competitive interactions among species (McKenzie et al. 2012, Fisher and Burton 2018, Bugar et al. 2019, Tattersall et al. 2020a, Fisher and Ladle 2022). In comparison, the effects of landscape configuration in the western boreal have received little attention. Discriminating the interacting roles of habitat composition and configuration is critical for predicting wildlife outcomes (Côté et al. 2016) and informing conservation of wildlife in the Boreal Plains (Tattersall et al. 2020b, Beirne et al. 2021). Conceptual frameworks of fragmentation suggest that the importance of habitat configuration for wildlife populations varies non-linearly along a gradient of habitat availability, such that the effects of fragmentation are most pronounced in landscapes that have low or intermediate amounts of suitable habitat (Andrén 1994, Didham et al. 2012, Villard and Metzger 2014).

We present an empirical case study of the North American red squirrel (*Tamiasciurus hudsonicus*) to test species-specific relationships to landscape configuration and composition across spatial scales. Widespread throughout the Boreal Plains, red squirrels are omnivores that depend on stored conifer cones for overwinter survival (Rusch and Reeder 1978, Fisher et al. 2005). Natural and anthropogenic disturbances—including wildfire, timber harvesting, and energy infrastructure—may directly affect red squirrel habitat structure by removing cone-producing mature conifers (Fisher and Wilkinson 2005), changing vegetation structure along forest edges (Harper et al. 2015, Jackson et al. 2023) decreasing landscape permeability (Bakker and Van Vuren 2004) and influencing the abundance of co-occurring predators (Tattersall et al. 2020a). Remaining landscape structure also impacts red squirrels but has mostly been explored in the context of agricultural matrices (Bayne and Hobson 2000, Fisher et al. 2005, Patterson and Malcolm 2010).

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We investigate the degree to which habitat composition and configuration influence red squirrel distribution in the remaining habitat matrix of Boreal Plains landscapes, as well as the spatial scale of those effects. We hypothesized that: (i) both habitat loss and fragmentation from industrial development would negatively influence red squirrel distribution, but (ii) the effects of habitat fragmentation would be greatest in landscapes where the amount of anthropogenic disturbance was high—aligning with the idea that habitat loss and fragmentation represent separate ecological processes whose importance depends on landscape structure and amount of suitable habitat. We deployed 430 camera traps across the Boreal Plains to measure red squirrel relative abundance in ten landscapes with variable proportions and spatial patterns of forest habitat, seismic lines, cut blocks, and other industrial disturbances. We separately characterized landscape composition and configuration by determining the proportion of natural land cover or disturbance footprint in local landscapes and derived three metrics representing the spatial pattern of habitat and disturbances: edge density created by anthropogenic features, effective mesh size, cohesion (isolation), and evenness of remnant natural habitat. We predicted that the relative abundance of red squirrels would be highest in areas with a high proportion of mature conifer or mixed-wood forest relative to other natural habitat types. We also predicted that red squirrel relative abundance would be negatively related to the proportion of any cleared anthropogenic features (well pads, roads, and seismic lines). Finally, we expected that fragmentation effects (e.g., creation of anthropogenic edges or isolation of habitat patches) by industrial disturbances would have a negative, context-dependent relationship with red squirrel relative abundance, that was strongest in landscapes with high amounts of total disturbance.

Methods

Study area

The Boreal Plains ecozone spans 740 000 km² of the boreal forest in Treaty 8 territory of Alberta, Canada, and encompasses a wide variety of vegetative land cover, climate, productivity, and human activity (Fig. 1A). The region is characterized by a mosaic of aspen (*Populus tremuloides*) parkland, mixed broadleaf forests, white spruce (*Picea glauca*) and jack pine (*Pinus banksiana*) forest, black spruce (*Picea mariana*) lowland forest, as well as wetlands, muskeg, fens, and lakes. We studied ten landscapes (Fig. 1B) representing aggregations of 3–5 watersheds each 1000–3000 km² (Bayne 2021) across the Boreal Plains. Landscapes spanned a gradient of low–high industrial development intensity and a variety of anthropogenic land-use types that included timber harvesting, open pit bitumen mining, and *in situ* oil and gas extraction (Bayne et al., 2021). Common disturbance features included timber cut blocks, “seismic” petroleum exploration lines (Dabros et al. 2018), well pads, roads, pipelines, and processing facilities (Fig. 1C) with varying degrees of human use or active restoration. Other common small- and medium-sized mammals in our study area included food competitors snowshoe hare (*Lepus americanus*), northern flying squirrels (*Glaucomys sabrinus*) and diverse small mammals; and red squirrel predators including ermine (*Mustela richardsonii*) American marten (*Martes americana*), fisher (*Pekania pennanti*), Canada lynx (*Lynx canadensis*), red fox (*Vulpes vulpes*), and coyote (*Canis latrans*). Common avian predators of red squirrels included the northern goshawk (*Accipiter gentilis*) and great horned owl (*Bubo virginianus*).

Camera trapping and wildlife occurrence data

We deployed motion-activated camera traps within each of our ten landscapes between 2021 and 2024 with the initial purpose of monitoring medium- to large-sized mammals ranging in size from red squirrel to moose (Fig. 1). Cameras were deployed using a constrained stratified sampling design. First, we divided landscapes into 2 km² hexagonal cells and randomly selected

~60 potential cells for camera deployment within strata corresponding to either upland forest (low moisture, >50% deciduous) and lowland forest (high moisture, >50% coniferous), or forest type (broadleaf, coniferous, and mixed). Camera placements in randomly selected hexagonal cells were constrained to locations >100-m from roads and >1 km from neighbouring camera placements in adjacent cells to maintain independent sampling. Infrared camera traps (Reconyx PC900 Hyperfire or Hyperfire II Professional; Holmen, WI, USA) were placed approximately 100 cm above the ground, facing a wildlife trail 3–7 meters away. To avoid confounding movement with habitat use (Stewart et al. 2018), cameras used in analyses were not directly deployed on human trails or seismic lines unless these features had clear wildlife signs including scat, digging, or abundant tracks. To maximize wildlife detectability, we applied a scent lure (O’Gorman’s Long Distance Call, MT, USA) once at each site at the time of deployment in either September or October (Stewart et al. 2019). Camera traps were programmed to take photos continuously at high sensitivity when movement was detected within the field of view. A total of 430 cameras were deployed across the ten landscapes (mean = 43 cameras per landscape, range = 36–50 cameras, SD = 5.7 cameras). Within the 2021–2024 sampling period, each camera was deployed for approximately one year (autumn to autumn; mean=347 days, SD= 67.2 days).

To quantify the occurrence of red squirrels in each landscape, all images of wildlife occurrences (e.g., Fig. 1D, Fig. 1E, Fig. 1F) were manually classified to species using Timelapse 2.0 (Greenberg et al. 2019), then aggregated to identify independent detections of red squirrels. We defined independent detection events for red squirrels as those occurring >30 minutes apart at a given camera site. Previous evidence suggests that independent detections of red squirrels correlate well with density estimates from other field methods (Villette et al. 2016). We binned red squirrel detections into monthly independent detections (Appendix S1: Fig. S1), an index

179 arising from both the local abundance and intensity of site use of red squirrels at that site that we
180 suggest indicates a site's quality (Burton et al. 2015, Parsons et al. 2017). We treated a non-
181 detection within a month as a true absence, given that an animal undetected on a lured trail site
182 within a full month is very likely not present. This gives a metric with more range and variability
183 than occupancy models (MacKenzie et al. 2002) which assume some zeros are error, but which
184 model presence/absence responses and rely on the assumption of closure which is surely violated
185 with mobile mammals (Rota et al. 2009).

186 *Measuring landscape structure*

187 We used the Alberta Biodiversity Monitoring Institute (ABMI) 2021 Human Footprint
188 Index dataset (ABMI, 2023) to determine the amount and type of anthropogenic disturbance
189 features across our study area, and the Satellite-Based Forest Inventory from Wulder et al. (2024)
190 to measure forest structure, amount of vegetative land cover, and fire disturbance across our
191 study area. To characterize landscape structure, we derived habitat variables describing the
192 anthropogenic and natural features surrounding each camera site which we hypothesized *a priori*
193 to be ecologically related to red squirrel relative abundance (Appendix S1: Table S1). Since
194 habitat relationships arise on spatial scales that are taxa- and context-specific (Wiens 1989,
195 Jackson and Fahrig 2015), we calculated all configuration and composition covariates at 22
196 spatial scales using concentric circular buffers of 50 m, 100 m, and 250–5000 m radii increasing
197 by 250 m increments. First, we measured landscape composition: this included the proportion of
198 area around each camera site covered by each dominant land cover type (e.g., forest type,
199 wetland, shrub) and recent (0–15 years) fire disturbance, as well as the proportion of area
200 covered by anthropogenic disturbances (e.g., well pads, roads, timber harvest blocks 0–15 years
201 old; Appendix S1: Table S1). As a representation of total disturbance amount, we calculated the

cumulative site disturbance at each camera location, which was the sum of the proportion of area covered by all anthropogenic disturbances (Bayne 2021). To measure landscape configuration, we superimposed and rasterized the Human Footprint Index from the Satellite Based Forest Inventory landcover types at a five-meter resolution and extracted four metrics which we hypothesized to correspond to potential mechanisms through which landscape structure affects the distribution of red squirrels (Appendix S1: Table S1). Our configuration metrics included *natural edge density*, representing the total length of patch edges and boundary complexity created by anthropogenic features, *effective mesh size*, representing the average size of habitat patches and degree of fragmentation, *cohesion*, representing the connectivity of similar habitat patches, and *Shannon evenness index*, representing dominance and how evenly different land cover types were distributed across the landscape. Cohesion was highly correlated with edge density (Spearman's $r = -0.95$) and had low variance among sites (Appendix S1: Table S1). All covariates were extracted from spatial layers using the *sf* (Pebesma 2018), *terra* (Hijmans et al. 2025), and *landscapemetrics* (Hesselbarth et al. 2019) packages in R v4.3.2 (R Core Team 2021).

Statistical analysis

Spatial responses of red squirrels to landscape structure were modeled using an information theoretic approach and model selection on a set of candidate models (Burnham and Anderson 2002). For all statistical analyses, we used generalized linear mixed models with a negative binomial distribution and log-link function (Zuur et al. 2009) to account for overdispersion of red squirrel detection data (Appendix S1: Equation S1). All models were fit using the *glmmTMB* package (Brooks et al. 2017) using maximum likelihood with the Laplace

224 approximation. All explanatory variables were scaled using standardized z-scaling to improve
225 model convergence and facilitate comparison of relative effect sizes.

226 To test our ecological hypotheses about the relative importance of landscape composition
227 and configuration on red squirrels, we first divided our explanatory variables into three
228 categories: [1] composition of natural habitat, [2] composition of anthropogenic features, and [3]
229 landscape configuration (Appendix S1: Table S1). Since red squirrels may respond to different
230 aspects of landscape structure at separate spatial scales (Fisher et al. 2005), we identified a best-
231 fit scale within each category of predictors before testing our hypotheses. Specifically, within
232 each of our three categories of predictors, we fit a model using all predictors at each of the 22
233 spatial scales (i.e., 50–5000 m) and compared the Akaike Information Criterion corrected for
234 small sample size (AICc; Akaike 1998) scores across spatial scales to determine which one was
235 best supported for statistical modelling. This approach allowed us to empirically evaluate the
236 appropriate scales for modelling and flexibly incorporate the effect of natural habitat
237 composition, anthropogenic composition, and landscape configuration on red squirrel
238 distribution at separate spatial scales.

239 We combined covariates from each of the three categories at their respective best-fit
240 spatial scales (Appendix S1: Table S1) into a candidate model set describing statistical
241 hypotheses about the effects of natural habitat, anthropogenic disturbance, and landscape
242 configuration on red squirrel distribution (Table 1). Our first five candidate models included only
243 fire and natural land cover predictors. From these candidate models, we identified the best-
244 supported natural land cover predictors and deemed this a ‘core’ model, which we carried
245 forward into all other candidate models. The remainder of our final model set included
246 anthropogenic disturbance and configuration metrics in addition to these ‘core’ predictors (Table

1). We did not include roads in any candidate models due to their high correlation with most other anthropogenic disturbances (Appendix S1: Table S2) at our selected spatial scale. Configuration metrics were highly correlated with variables for anthropogenic disturbance type, so our candidate model set was not fully nested. Instead, we represented disturbance in any configuration models using interactions with the aggregate cumulative site disturbance metric, which was not highly correlated with any configuration variables. We performed model selection on the candidate set and ranked candidate models using AICc scores to compare the importance of various components of landscape structure for red squirrel distribution and identify a best-supported model (Burnham and Anderson 2002). We treated any candidate models with $\Delta AIC < 2$ as best supported and interpreted them accordingly, barring uninformative parameter problems (Arnold 2010). Model validation was performed using diagnostic plots of residuals vs. fitted values; collinearity among predictors was assessed using variance inflation factors (VIF < 3 ; Appendix S1: Fig. S2; Zuur and Ieno 2016). All model selection and top model evaluation were performed using the *MuMIn* (Barton 2015) and *performance* (Lüdecke et al. 2021) packages in R. Note that throughout, we follow the latest guidelines on interpreting p-values and do not rely on alpha values to weigh significance (Greenland et al. 2016). We further evaluated model performance with parametric simulations (Zuur and Ieno 2016) by generating 1000 datasets from the top model and comparing the empirical bias and precision of refitted coefficient and random effect estimates to those of the original model (Appendix S1: Fig. S3).

Results

We gathered 149,273 camera trap-days of data and 683,259 images, producing 7,399 images and 4,747 independent detections of red squirrels. Red squirrels were detected at least once at 305 of 430 camera trap sites (naïve occupancy = 0.709). Across all sites, our final

270 monthly dataset represented 4918 camera-months of data with a mean monthly detection rate of
271 0.93 squirrels.

272 ***Red squirrel modeling***

273 Red squirrel occurrence varied with natural land cover, anthropogenic composition, and
274 natural configuration at three separate spatial scales that had strong or moderate model support
275 compared to other scales (Appendix S1: Fig. S4). Specifically, red squirrel occurrence varied
276 with natural land cover around the camera trap at a 100-m spatial scale but varied with
277 composition of anthropogenic features and landscape structure around the camera trap at much
278 larger spatial scales: 4250-m and 2250-m, respectively (Appendix S1: Fig. S4).

279 The highest performing ‘core’ natural land cover model variables included the proportion
280 of conifer forest, mixedwood forest, broadleaf forest, and areas burned within 0–15 years (Table
281 1). However, a model containing only natural land cover predictors was not well-supported
282 compared to models that included effects for anthropogenic disturbance or configuration
283 variables (Table 1). Instead, red squirrel monthly occurrence was best explained by a model
284 including natural land cover at 100-m and an interaction between edge density and site
285 disturbance at 2250-m ($AIC_w = 0.559$; Table 1). Red squirrel monthly occurrence decreased with
286 proportion of broadleaf forest (Fig. 2, Table 2) and increased with proportion of mixed wood
287 forest (Fig. 2, Table 2) at the 100-m spatial scale. Paradoxically, red squirrel occurrence was
288 unrelated to the proportion of conifer forest at the 100-m spatial scale in this multivariate model
289 (Fig. 2, Table 2). The amount of recent fire had a weak positive effect on red squirrels (Fig. 2,
290 Table 2). Red squirrel monthly occurrence decreased nominally with anthropogenic edge density
291 at the 2250-m spatial scale in less developed landscapes (Fig. 2, Table 2) and decreased
292 markedly in more developed landscapes via an interaction term (Fig. 2, Table 2). Unexpectedly,

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Monthly detections instead of monthly occurrence.. Detection frequency??
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293 red squirrel occurrence increased with cumulative site disturbance at the 2250-m spatial scale
294 (Fig. 2, Table 2). Random intercepts for site and array had high variance ($\sigma^2_{\text{site(array)}} = 2.503$ and
295 $\sigma^2_{\text{array}} = 0.410$; Table 2), and improved model fit compared to a model with no random effects
296 based on a likelihood ratio test ($\chi^2 = 1349.1$, $p < 0.001$). Model validation simulations showed
297 that coefficient estimates were generally unbiased and precise across variables, with variance
298 consistent with model standard errors, except for recent fire, which exhibited higher deviation
299 and very low precision (Appendix S1: Fig. S3).

300 Our top model only had moderate separation from the second-ranked model ($\Delta\text{AICc} =$
301 2.59 , $\text{AIC}_w = 0.153$; Table 1), which included effects for the composition of most persistent
302 energy-sector anthropogenic disturbances. Timber harvesting was not included in any well-
303 supported models. In this model, red squirrel monthly occurrence was negatively related to the
304 proportion cover of seismic lines, positively related to pipelines and transmission lines, and
305 neutrally related to inactive well sites (Appendix S1: Fig. S5).

306 Discussion

307 Disturbance-mediated landscape configuration best explained the spatial distribution of
308 red squirrels in the Boreal Plains. As ecologists and conservationists focus on amount of habitat
309 lost, the cost to wildlife of reconfiguring remaining habitat—even when most remains intact—
310 often goes unnoticed. Even for red squirrels, a conifer-dependent species for which simple
311 habitat loss would be assumed as particularly important, the cost of losing mature forest is
312 eclipsed by the larger-scale effects of reshaped landscape configuration from industrial
313 development.

314 Here, using a large network of camera trap arrays, we can highlight a complex, context-
315 dependent relationship between red squirrel distribution, landscape composition, and landscape

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made an attempt.

configuration, as has been observed for larger boreal mammals (Barnas et al. 2024). Overall, red squirrel monthly occurrence decreased with anthropogenic edges, a landscape feature that held more explanatory power than composition of any specific disturbance features. Supporting our hypotheses, an increase in the amount of industrial disturbance on the landscape was accompanied by a more pronounced effect of edge habitat—as cumulative site disturbance increased, edge effects caused a sharper decline in red squirrel abundance. However, contrary to our predictions, the effect of cumulative site disturbance itself was positive. Within the boreal landscapes that we sampled, the total footprint of industrial disturbances seldom exceeded 25% of the total habitat matrix. The negative relationship between habitat amounts and wildlife abundance predicted by theory (Didham et al. 2012) may not manifest until a higher threshold of cumulative disturbance that is less common in Boreal Plains landscapes. Even in highly fragmented landscapes, red squirrels are to some degree adaptable and resilient to habitat loss (Bayne and Hobson 2000), potentially complementing lost habitat with novel resource subsidies (Dunning et al. 1992, Fisher and Merriam 2000). The industrial disturbances currently observed in the western boreal do not bring about a complete loss of functional habitat, instead influencing red squirrels through other ecological mechanisms.

Natural habitat and anthropogenic disturbance affected red squirrels at markedly different spatial scales that varied by over a full order of magnitude in our models, suggesting that they influence red squirrels via different individual- and population-level processes as predicted by theory (Holling 1992, Levin 1992). Individual red squirrels respond to resource availability and heterogeneity at small spatial scales, matching their small body size and home ranges of approximately 0.5 to 2 Ha (Rusch and Reeder 1978, Munroe et al. 2009). The 100-m scale at which natural composition influenced squirrels indicates that differences in squirrel detections

across land cover types are most likely due to individual foraging behaviour and local habitat use. For instance, the relative abundance of squirrels decreased sharply in landscapes with high proportions of deciduous-leading forests, which contain inadequate cone resources compared to conifer and mixedwood forests. The unexpected neutral effect of coniferous forest suggests two main possibilities. First, although we took care to avoid collinearity among variables, the many non-conifer variables may have collectively subsumed variation in red squirrel occurrence that would have otherwise been attributed to conifer forests alone. However, collinearity often appears as model instability in subsequent validation, and we did not note that in our simulations. Secondly, in conifer-dominated systems like the Boreal Plains, even low densities of cone-producing trees situated in a non-conifer, partially developed matrix are not a limiting component of red squirrels' resource requirements (Larsen et al. 1997). Variation beyond a baseline amount of conifer may have a limited influence on the local abundance of despotic species like the red squirrel. Our results also suggest that low-intensity fire, including prescribed burning, may not have a negative effect on squirrels when key habitat features are retained (Russell et al. 2010). Overall, we suggest that red squirrels are distributed locally relative to resource quantity and quality on a spatial grain that is a function of individuals' resource requirements.

In contrast to natural land cover, the effects of anthropogenic disturbance composition and configuration were best explained at much larger spatial scales (4250-m and 2250-m, respectively) that exceed the spatial perception of individual animals, suggesting an influence on population-level processes such as dispersal, mortality, or recruitment rather than individual space use (Kotliar and Wiens 1990). This aligns with Fisher et al. (2011), who proposed that large spatial scales (between 2000-m and 4000-m) best explained numerical population

responses of squirrels to land-cover type. Thus, we propose that local resource availability and the configuration of landscape disturbances simultaneously impact red squirrel populations via mechanisms operating on nested scales. Ecological systems represent the integration of many spatial processes (Levin 1992), and our results emphasize that variability in landscape structure must be studied at multiple levels of ecological organization (Wiens 1989) to achieve a complete understanding of species distributions.

The exact population-level mechanisms through which edge density negatively affected red squirrels cannot yet be discerned and pose new questions for future research. We propose several potential explanations for this relationship. First, the negative effect of edge density could reflect decreased dispersal. Long-range sciurid dispersal is impacted by landscape structure (Merrick and Koprowski 2016, Hämäläinen et al. 2019), and red squirrels are generally averse to crossing open habitat including timber harvest blocks (Bakker and Van Vuren 2004) and disperse shorter distances in younger forest stands (Larsen 2009). However, we note that in our second-supported model, the proportion of pipelines and transmission lines positively affected squirrel detections whereas the proportion of seismic lines negatively affected squirrel detections, despite both types of features representing a similar challenge to red squirrel dispersal. In addition, landscape cohesion—a strong proxy for structural connectivity and dispersal potential (Schumaker 1996)—performed poorly in our models, so we are cautious to accept this interpretation of edge density’s effect.

Alternatively, we hypothesize that the negative impact of edge density is caused by changes to predator-prey relationships. The movement of mammalian boreal predators, including coyotes, increases on industrial linear features like seismic lines (Tattersall et al. 2020a). These features may additionally change predator functional responses and increase predation events for

secondary prey like red squirrels near edges (Kremsater and Bunnell 1999), especially if red squirrels are frequently exposed by crossing cleared disturbance features. In particular, avian predators are likely influential for red squirrel populations since red squirrels are a main prey species of raptors like the northern goshawk (Salafsky et al. 2007). Preferential use of edge features by raptors may provide increased visibility, which has been shown to increase predation rates in multiple species raptors (Mirski and Väli 2021, Shonfield and Bayne 2023).

Contrary to our hypotheses, red squirrel detections were positively related to cumulative site disturbance itself. Cleared anthropogenic features may offer resource complements (Dunning et al. 1992) to red squirrels by altering edge vegetation characterized by denser berry-producing shrubs, fungi, and early successional herbaceous plants (Harper et al. 2015, Jackson et al. 2023). Red squirrels use a variety of food resources aside from conifer cones (Smith 1968, Benhamou 1996), and home range size may decrease with increasing resource heterogeneity (Fisher 1999). Specific industrial features may represent varying amounts of predation risk and resource complementation based on their width, orientation, and degree of human use, which is supported by our second-performing model (Appendix S1: Fig. S5). For instance, pipelines and transmission lines, which had a positive effect in our model, are generally much wider features than seismic lines. These wider features allow for stronger intensity of light, microclimate changes, and edge effects that extend further into surrounding habitat (Abib et al. 2019, Franklin et al. 2021), in turn increasing their potential value as a resource complement. In addition, there is widespread evidence for mesocarnivore use of seismic lines, especially coyotes (Tattersall et al. 2020a, Clare et al. 2023), but there is less evidence for predator use along more exposed features like transmission lines. As such, we suggest that pipelines and transmission lines provide complementary resources for red squirrels while seismic lines provide comparatively

408 fewer resources, while also representing increased predation risk. Previous research aligns with
409 this inference: Beirne et al. (2021) found a negative relationship between seismic line density
410 and red squirrel abundance accompanied by increased predator use, and Skatter et al. (2020)
411 found a positive relationship between red squirrels and pipelines or transmission lines.

412 **Caveats paragraph (collapsed into discussion)**

413 Small species such as red squirrels can be difficult to measure with precision using
414 camera traps. Our camera traps were designed to simultaneously capture a wide variety of
415 mammal species across a large geographic extent that would be unfeasible using other field-
416 based methods. If non-detections of red squirrels, which we assumed were true absences, arose
417 instead from poor detectability, our modelled estimates of red squirrel relative abundance may be
418 biased. We do not think this is likely given our use of lures, game trails, and long sample periods.
419 Our stratified design was intended to measure mammal distribution within the matrix of
420 remaining boreal habitat, and camera traps were not deployed directly on industrial features
421 themselves. As a result, this study may not have captured localized, individual responses to
422 disturbance features, such as the absence of squirrels within timber harvest blocks documented in
423 other studies (Fisher and Wilkinson 2005).

424 The proposed mechanism of changes to predation at forest edges warrants further
425 investigation of predation rates at anthropogenic edges. We suspect that a large proportion of this
426 shift in predation is driven by raptors like great-horned owls and northern goshawks, but raptor
427 dynamics in disturbed boreal landscapes are not yet understood (Shonfield and Bayne 2023).

428 **Conclusions (collapsed into discussion)**

429 The novel landscape configuration generated by industrial resource extraction can
430 produce important, often underappreciated, outcomes for wildlife species. The ecological

Commented [AB9]: Tie this into our age old question better? What happens if cumulative footprint increases even further? Will these effects be generalizable or become more pronounced through other mechanisms as predicted by theory?

431 mechanisms through which landscape structure impacts wildlife extend beyond straightforward
432 measures of habitat loss. Red squirrel distribution in the Boreal Plains is influenced by landscape
433 composition and configuration across multiple landscapes and multiple spatial scales. We
434 successfully generalized this relationship in a single model describing interacting positive effects
435 of disturbance amount and negative effects of edge density, which we propose are mediated
436 through resource complementation and increased predation, respectively. Although red squirrels
437 are not a conservation concern in the Boreal Plains, they are an important prey species for many
438 terrestrial and avian predators. There are few industrial development scenarios in which
439 cumulative site disturbance is high but impacts on configuration are low, and a focus on
440 landscape composition may underestimate the full ecological impacts of industrial development.
441 The multiscale modeling framework presented in this study encompassed an integration of
442 multiple individual- and population-level processes and could be applied to other species to
443 better understand scale-dependent responses of wildlife to disturbances. As human influence
444 over wild landscapes expands, the management of industrial activities will require cumulative
445 risk approaches (Tyack et al. 2022) that encompass interacting stressors on wildlife and an
446 increasing appreciation of the multiple mechanisms through which disturbance reshapes the
447 structure and function of wildlife communities.

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461 **Author contributions**

462 Aidan Brushett and Emerald Arthurs conceived the study, designed the methodology, analyzed
463 the data, and wrote the first draft of the manuscript; All authors collected the data; Jason Fisher
464 oversaw the experimental design; All authors contributed meaningfully to the final draft of the
465 manuscript and gave final approval for publication.

466 **Conflict of interest**

467 The authors have no conflict of interest to declare.

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719 **Tables**

720 **Table 1.** Model selection results from candidate hypotheses describing impact of the composition
 721 of natural features, composition of anthropogenic features, and landscape configuration on red
 722 squirrel abundance in the Boreal Plains. The top performing model is shown in bold. The ‘core’
 723 land cover model, which contained only the best-supported natural land cover predictors, is
 724 shown in italics. The ‘core’ metrics we Degrees of freedom (df), small sample corrected Akaike
 725 Information Criterion (AICc), differences in AICc from the top model ($\Delta AICc$), and AICc model
 726 weights (AICw) are reported for each model. Models with $\Delta AICc \leq 2$ were considered as well-
 727 supported.

Main Hypothesis	Covariates	df	AICc	$\Delta AICc$	AICw
Context-dependent configuration	CORE + Edge Density \times CFI	11	9394.44	0.00	0.559
Composition	CORE + Pipelines & Transmission Lines + Seismic Lines + Inactive Well Sites	11	9397.03	2.59	0.153
Composition	CORE + Harvest <15 + Pipelines & Transmission Lines + Seismic Lines + Inactive Well Sites	12	9398.56	4.12	0.071
Configuration	CORE + Edge Density + Site	10	9399.35	4.91	0.048
Configuration	CORE + Mesh Index + Edge Density	10	9399.79	5.35	0.039
Composition	CORE + Industrial Facilities + Active Well Sites + Pipelines & Transmission Lines + Seismic Lines + Inactive Well Sites	13	9399.81	5.36	0.038
Configuration	CORE + Mesh Index + Shannon's Evenness + Edge Density	11	9400.87	6.43	0.022
Composition	CORE + Industrial Facilities + Active Well Sites + Harvest <15 + Pipelines & Transmission Lines + Seismic Lines + Inactive Well Sites	14	9401.30	6.86	0.018
Configuration	CORE + Shannon's Evenness + Edge Density	10	9402.07	7.63	0.012
Configuration	CORE + Edge Density	9	9402.24	7.80	0.011
Configuration	CORE + Cohesion + Mesh Index	10	9402.88	8.44	0.008
Configuration	CORE + Cohesion	9	9404.17	9.73	0.004
Configuration	CORE + Cohesion + Shannon's Evenness	10	9404.37	9.92	0.004
Configuration	CORE + Mesh Index	9	9405.30	10.86	0.002

Commented [MOU10]: See note on fig 5 about consistency between plots.

Also I think this is useful for a reviewer and for complete transparency but you have a lot of figures I might suggest moving to a supplemental info depending on the journal you submit to

Commented [AB11R10]: I agree, we added this as per advice from Barnas but I'm not sure it adds a ton of information that's critical for the main manuscript. Will move to SI

Commented [MOU12]: May seem intuitive to us, but you need to write out what all the abbreviations in the table mean specifically (e.g. df, AICc, etc.)

Commented [AB13R12]: Added

Commented [JF14]: Explicitly state spatial scale here in a way easily referable. Also state clearly in the text.

<i>Natural habitat (CORE)</i>	<i>Fire <15 + Broadleaf + Coniferous + Mixedwood</i>	8	9405.71	11.27	0.002
Configuration	CORE + Shannon's Evenness	9	9406.58	12.14	0.001
Composition	CORE + Harvest <15	9	9406.80	12.36	0.001
Configuration	CORE + Mesh Index + Shannon's Evenness	10	9406.89	12.45	0.001
Composition	CORE + Industrial Facilities + Active Well Sites	10	9407.02	12.58	0.001
Composition	CORE + Industrial Facilities + Active Well Sites + Harvest <15	11	9407.78	13.33	0.001
Natural habitat	Broadleaf + Coniferous + Mixedwood	7	9408.66	14.22	0.000
Natural habitat	Fire <15 + Broadleaf + Coniferous + Mixedwood + Shrubland + Wetland + Treed Wetland	11	9409.88	15.44	0.000
Natural habitat	Broadleaf + Coniferous + Mixedwood + Shrubland + Wetland + Treed Wetland	10	9411.74	17.30	0.000
Natural habitat	Fire <15	5	9418.76	24.32	0.000
Null	1	4	9423.79	29.35	0.000

729 Table 2. Coefficient estimates, standard errors (SE), p-values (p), and standard deviations of
 730 random intercepts for the top-performing model explaining red squirrel detections as a function
 731 of landscape structure. Models were fit using mixed-effect negative binomial models. Predictors
 732 were included in models at separate spatial scales, which were identified *a priori* by fitting
 733 models of natural land cover, disturbance, and configuration represented at 22 spatial scales (50
 734 m–5000 m) and comparing the small sample corrected Akaike Information Criterion (AICc). All
 735 models included nested random intercepts of site and array, which were assumed to have a
 736 normally distributed variance.

Fixed effects	Modeled scale (m)	Estimate	SE	p
Intercept	100	-1.428	0.230	<0.001
Fire <15	100	0.158	0.112	0.16
Broadleaf	100	-0.518	0.128	<0.001
Coniferous	100	-0.032	0.111	0.771
Mixedwood	100	0.164	0.092	0.076
Edge Density	2250	-0.089	0.166	0.593
Site Disturbance	2250	0.295	0.121	0.014
	2250	-0.256	0.100	0.010
Edge Density × Site Disturbance				
Random effects	St. Dev. (intercept)			
Site [nested]	1.582			
Array	0.640			

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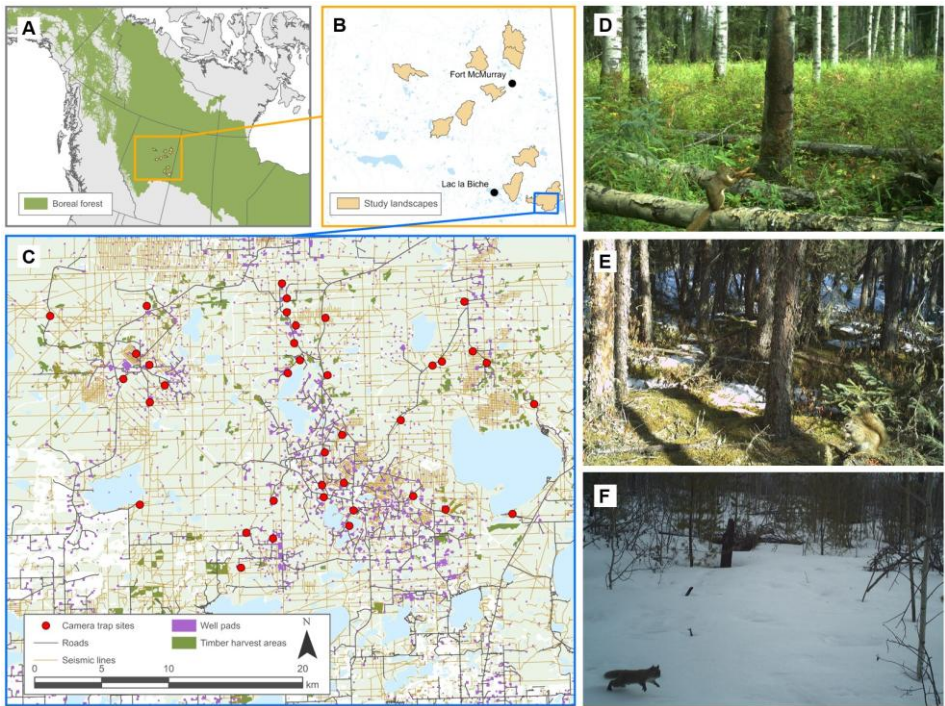
Figure Captions

Figure 1. [A] Boreal Plains ecozone within the western boreal forest. [B] Ten landscapes within the Boreal Plains selected for sampling, representing a gradient of anthropogenic land-use and environmental traits. [C] Camera trap placements within a given landscape followed a constrained stratified design based on vegetation characteristics. Red squirrels were detected year-round in a variety of habitat types, including [D] broadleaf forest, [E] conifer forest, and [F] mixed-wood forest.

Figure 2. Conditional effects of landscape covariate on the expected monthly count of red squirrels from the top-performing negative binomial model. The interaction plot represents the effect of edge density at high (90% quantile) and low (10% quantile) cumulative site disturbance. Density of lines along the x-axis indicates the frequency of the data value. All predictors represent the proportion of area covered within a set radius of camera trap locations (i.e. landscape composition), except edge density, which represents the density of anthropogenic edges in m/Ha. Note different y-axis scales for each landscape covariate.

755 **Figures**

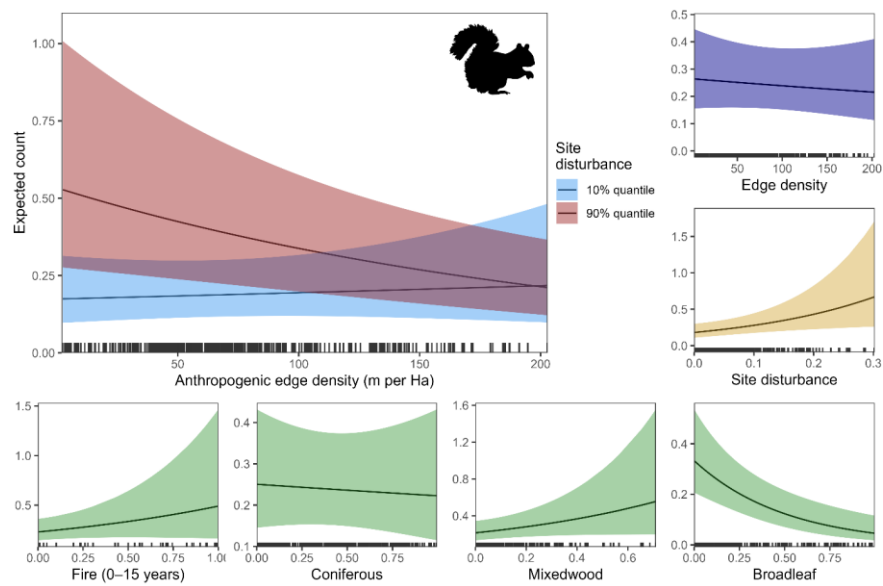
756 Figure 1.



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759 Figure 2.



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