Regression analysis (in R)

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Outline for the course

- Introduction to statistics
- Hypothesis testing (in R)
- Regression analysis (in R)

In statistics, **regression analysis** is a tool to quantify relationships between 2 or more variables

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In regression, variables are clasified as

- ▶ **Response**: outcome variable of interest (dependent)
- **Explanatory**: covariate(s) to explain the response (independent)

Examples?

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- ► For a continuous response (e.g. blood sugar level), a typical choice is **linear regression**

Today, we will focus on linear regression

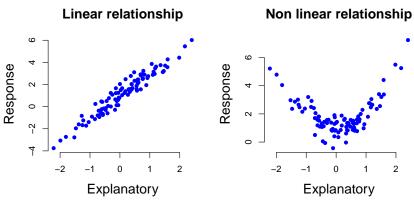
Linear regression

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In **simple linear regression**, the aim is to capture a linear relationship between a response (y) and a **single** covariate (x)

The simple linear regression model is writen as

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i, \qquad i = 1, \dots, n$$

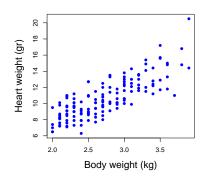
$$\beta_0 \text{ is the intercept}$$

$$\beta_1 \text{ is the slope (gradient)}$$

$$\epsilon_i \text{ is the error term}$$

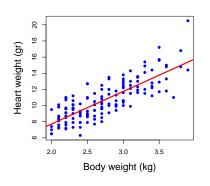
(*) Plot the response on the vertical axis and the covariate on the horizontal axis

The data in Fisher (1947) shows heart and body weights for 144 cats



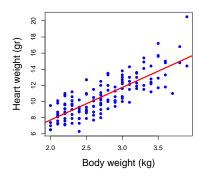
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Is there a linear relationship between heart and body weight?

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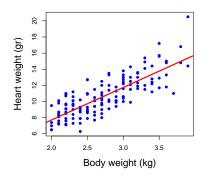
We can fit a simple linear regression

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$

where

- y_i = heart weight of cat i
- $> x_i = body weight of cat i$

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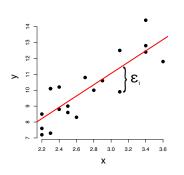
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How to estimate β_0 and β_1 ?

Idea: find β_0 and β_1 such that the line is a good fit for the data

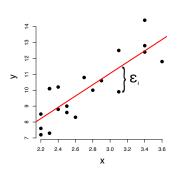
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For example, by minimizing the **residuals** of the regression

$$\epsilon_i = Y_i - (\beta_0 + \beta_1 x_i) \equiv Y_i - \hat{Y}_i$$

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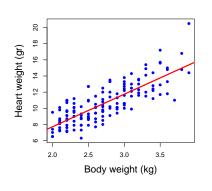
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Typically, this is done by minimizing the **Sum of Square Errors (SSE)**

$$SSE = \sum_{i=1}^{n} \epsilon_i^2$$

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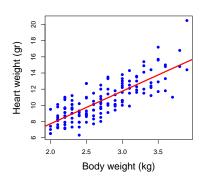


In this example we have

- $\beta_0 = -0.36$
- $\beta_1 = 4.03$

How to interpret this?

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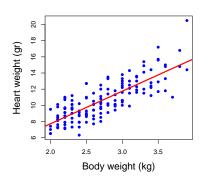
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For a 1kg increase in body weight, we expect heart weight to increase by 4.03gr

Is this increase statistically significant?

We can answer this using hypothesis testing:

$$H_0: \beta_1 = 0$$
 vs $H_1: \beta_1 \neq 0$

Recall:
$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$

 \rightarrow if $\beta_1 = 0$, y_i and x_i are not (linearly) dependent

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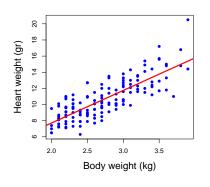
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Assumming $\epsilon_i \sim \text{Normal}(0, \sigma^2)$, we can derive a *t*-test for β_1

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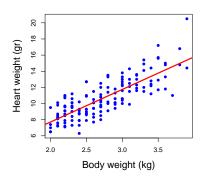


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We reject H_0 , i.e. we conclude that the linear relationship between heart and body weight is statistically significant ($\alpha=0.05$)

ANOVA aims to decompose the total variance of y

Recall: Var(y) =
$$\sum_{i=1}^{n} (y_i - \bar{y})^2/(n-1)$$
, with $\bar{y} = (\sum_{i=1}^{n} y_i)/n$

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$$\sum_{i=1}^{n} (\hat{y}_i - \bar{y})^2 = \sum_{i=1}^{n} (y_i - \bar{y})^2 - \sum_{i=1}^{n} (y_i - \hat{y}_i)^2$$
SSR (Regression) SST (Total) SSE (Error)

- SST quantifies the total variability of y
- SSE quantifies residual variability of y (unexplained by x)
- SSR quantifies how much of the variability of y is explained by x

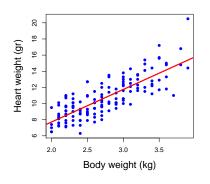
The **coefficient of determination** R^2 is defined as the proportion of total variability of y that is explained by x

$$R^2 = \frac{\text{SSR}}{\text{SST}} = 1 - \frac{\text{SSE}}{\text{SST}}$$

- ▶ $0 \le R^2 \le 1$
- ▶ If $R^2 = 0$, x explains none of the variability of y
- ▶ If $R^2 = 1$, x explains all of the variability of y (perfect fit!)

Warning: this assumes the relationship between x and y is linear

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$$R^2 = 0.65$$

Body weight explains 65% of the variability of a cat's heart weight

This variance decomposition is summarized by the **ANOVA table**

Source	Degrees of freedom	Sum of Squares	Mean Square	F ratio
Regression	1	SSR	SSR/1	
Error	n – 2	SSE	SSE/(n-2)	$\frac{SSR}{1} \ / \ \frac{SSE}{n-2}$
Total	<i>n</i> − 1	SST	SST/(n-1)	

▶ Assuming ϵ_i ~ Normal(0, σ^2), the F ratio can be used to test

 $H_0: \beta_1 = 0$ versus $H_1: \beta_1 \neq 0$

How does this compare to the *t*-test introduced earlier?

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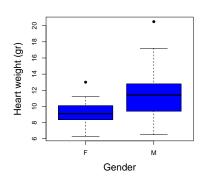
Examples?

Let's play: http://guessthecorrelation.com



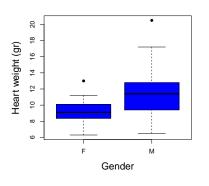
http://bioinformatics.cruk.cam.ac.uk/apps/winter-school/linear/

The data in Fisher (1947) also contains gender information



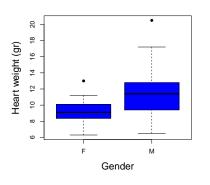
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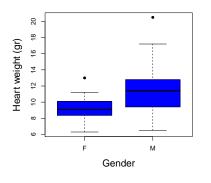


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In this example we have:

$$t = -5.35$$
 p-value = 3.38×10^{-7}

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Can we answer the same question using linear regression?

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However, categorical covariates require an special treatment . . .

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This can be done using dummy variables

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$$D_i = \begin{cases} 1, \text{if cat } i \text{ is male;} \\ 0, \text{if cat } i \text{ is female.} \end{cases}$$

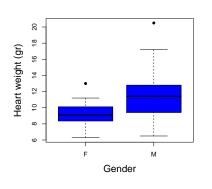
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Note: gender has **two** levels, but we only define **one** dummy variable (here we left *female* as a **reference category**)

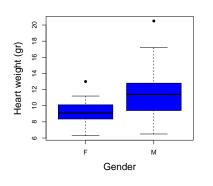
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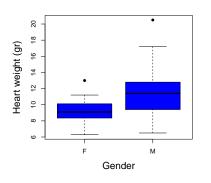


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- ▶ For a male cat: $\hat{y}_i = \beta_0 + \beta_1$
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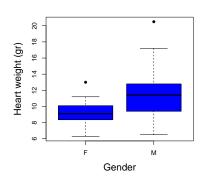
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- ▶ For a male cat: $\hat{y}_i = \beta_0 + \beta_1$
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 $\Rightarrow \beta_1$ quantifies the difference between female and male cats

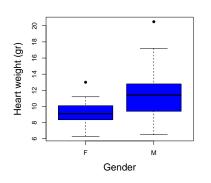
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$$y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \cdots + \beta_p x_{ip} + \epsilon_i$$

For example, with 2 continuous covariates:

$$y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \epsilon_i,$$

where

- ▶ y_i: heart weight for cat i
- ▶ x_{i1}: body weight for cat i
- ► x_{i2}: age for cat i

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$$H_0: \beta_j = 0$$
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Global: to assess the joint effect of all covariates

$$H_0: \beta_1 = \cdots = \beta_p = 0$$
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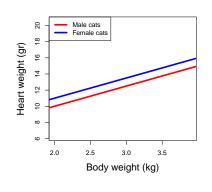
Assuming $\epsilon_i \sim N(0, \sigma^2)$, we can use an *F*-test (from ANOVA table)

... or mixing a continuous and a categorical covariate

$$y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \epsilon_i,$$

where

- ▶ y_i: heart weight for cat i
- ▶ x_{i1}: body weight for cat i

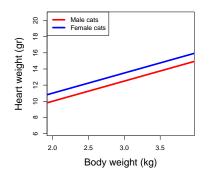


For a male cat:

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For a female cat:

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For a male cat:

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For a female cat:

$$\hat{y}_i = \beta_0 + \beta_1 x_{i1}$$

 \Rightarrow parallel lines, effect of body weight is independent of gender

 $\Rightarrow \beta_2$ quantifies a global difference between female and male cats

What if we think the effect of body weight depends on gender?

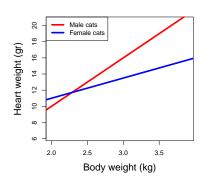
What if we think the effect of body weight depends on gender?

We can define an interaction effect

$$y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \beta_3 x_{i1} x_{i12} + \epsilon_i,$$

Recall:

- ▶ y_i: heart weight for cat i
- ► x_{i1}: body weight for cat i

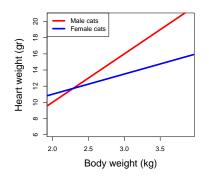


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⇒ crossing lines, effect of body weight depends on gender

 $\Rightarrow \beta_2$ quantifies a global difference between female and male cats

 $\Rightarrow \beta_3$ quantifies the differential effect of body weight

${\sf Questions} + {\sf practical}$



Assumptions in linear regression

Recall: the multiple linear regression model is written as

$$y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \cdots + \beta_p x_{ip} + \epsilon_i$$

What assumptions underlie this model?

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▶ Secondly, we assume that residuals ϵ_i are **independent** and **identically distributed** with

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▶ Secondly, we assume that residuals ϵ_i are **independent** and **identically distributed** with

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How can we assess this?

Idea: after estimating the regression coefficients, define residuals as

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Use these estimated residuals to diagnose model quality

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For example, we can prepare a _____ to see if residuals are normally distributed

More about this in the practical ...

What if things go wrong?

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If the residuals are not good enough to pass the diagnostic criteria, we need to revisit the model. For example

- We can transform some of the covariates or
- We can transform the response variable

Suppose a simple linear regression model

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$

Covariate transformation can be useful in situations where the relationship between y and x is **not linear**

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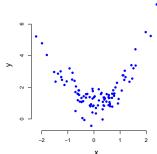
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Covariate transformation can be useful in situations where the relationship between y and x is **not linear**

Idea: replace x_i by a transformed version of x_i (e.g. $x_i^* = x_i^2$)

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$

Non linear relationship



$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i \qquad \qquad y_i = \beta_0 + \beta_1 x_i^2 + \epsilon_i$$
 Non linear relationship Linear relationship

Suppose a simple linear regression model

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Response transformation can be useful in situations where the variance of y is not constant as a function of x

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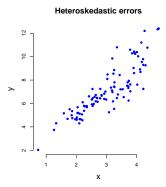
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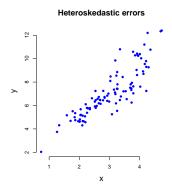
Idea: replace y_i by a transformed version of y_i (e.g. $y_i^* = \log(y_i)$)

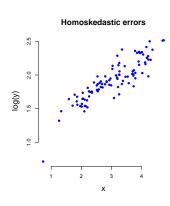
$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$



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$$\log(y)_i = \beta_0 + \beta_1 x_i + \epsilon_i$$





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We can use Generalized Linear Models ...

${\sf Questions} + {\sf practical}$