

# Statistics with Spa OWS

Lecture 11-b

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# Outline

- Linear models – going big
- Categorical and continuous predictors

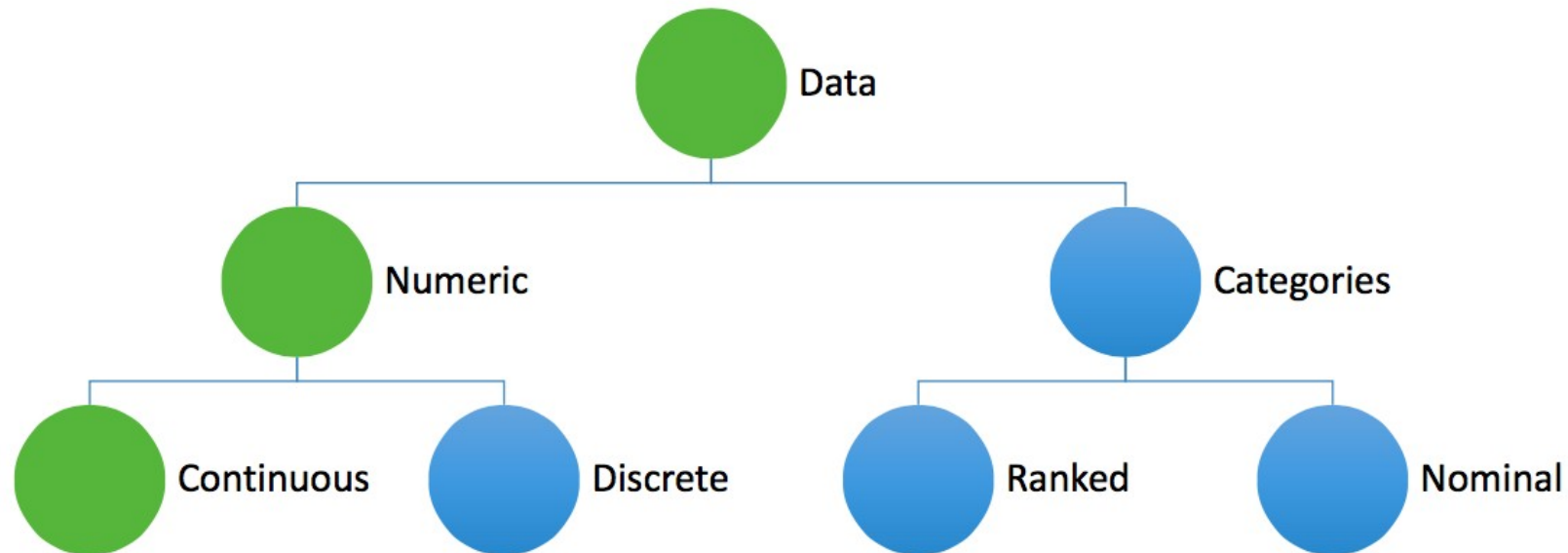
# Multiple linear models

```
lm(response~explanatory)
```

# Multiple linear models

`lm(response~explanatory)`

Data types



# Multiple linear models

`lm(response~explanatory)`

Response y:

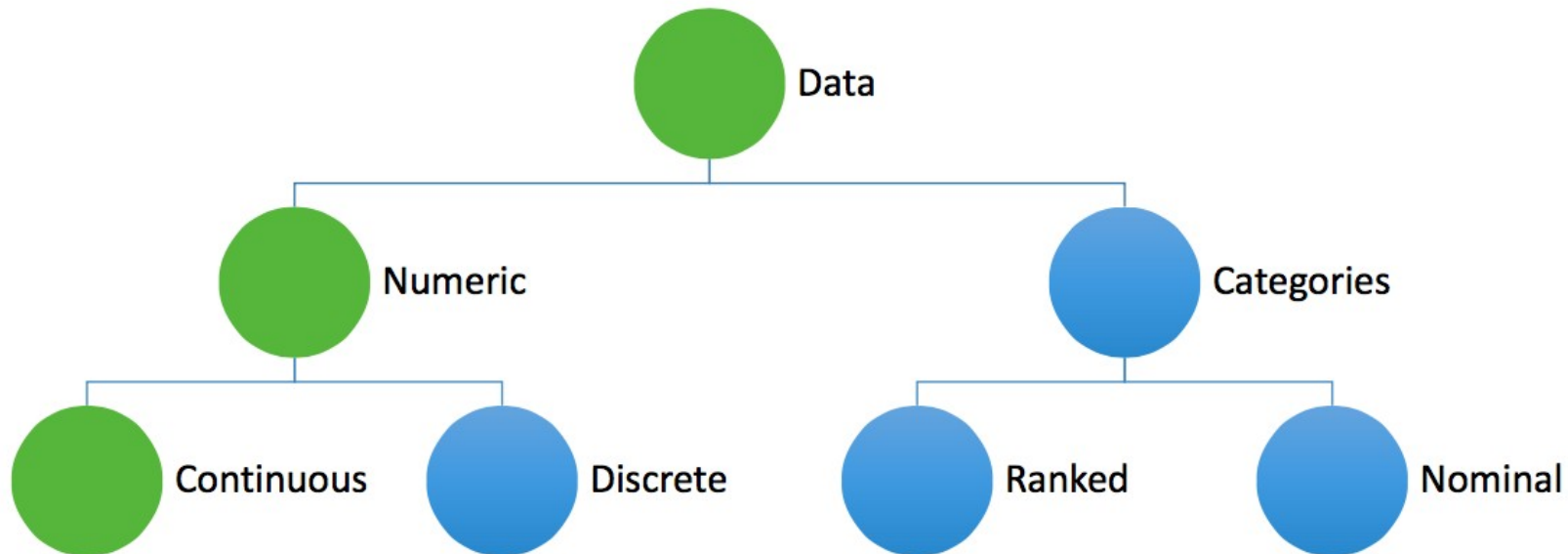
Continuous

Explanatory x:

Continuous (tarsus, wing, mass)

Categorical (Sex, Year, Observer, BirdID)

Data types



# Multiple linear models

Response y:

Continuous

Explanatory x:

Continuous (tarsus, wing, mass)

Categorical (Sex, Year, Observer, BirdID)

`lm(response~explanatory)`

$$y_i = b_0 + b_1 x_i + \varepsilon_i$$

# Multiple linear models

Response y:

Continuous

Explanatory x:

Continuous (tarsus, wing, mass)

Categorical (Sex, Year, Observer, BirdID)

```
lm(response~explanatory)
```

$$y_i = b_0 + b_1 x_i + \varepsilon_i$$

$$y_i = b_0 + b_1 x_{i0} + b_2 x_{i1} + b_3 x_{i2} + \varepsilon_i$$

# Multiple linear models

Response y:

Continuous

Explanatory x:

Continuous (tarsus, wing, mass)

Categorical (Sex, Year, Observer, BirdID)

```
lm(response~explanatory)
```

We can have more than one explanatory variable!

$$y_i = b_0 + b_1 x_{i0} + b_2 x_{i1} + b_3 x_{i2} + \varepsilon_i$$



# Multiple linear models

Response y:

Continuous

Explanatory x:

Continuous (tarsus, wing, mass)

Categorical (Sex, Year, Observer, BirdID)

```
lm(response~explanatory)
```

We can have more than one explanatory variable!

We can even mix continuous and factorial explanatory variables!

$$y_i = b_0 + b_1 x_{i0} + b_2 x_{i1} + b_3 x_{i2} + \varepsilon_i$$

# Let's try this

- $b_0$  = intercept

$$y_i = b_0 + b_1 x_{i0} + b_2 x_{i1} + \varepsilon$$

- $b_1$  = estimates effect of continuous variable  $x_0$
- $b_2$  = estimates effect of 2-level factor  $x_1$

# Let's try this

- $b_0$  = intercept

$$y_i = b_0 + b_1 x_{i0} + b_2 x_{i1} + \varepsilon$$

- $b_1$  = estimates effect of continuous variable  $x_0$  (tarsus)
- $b_2$  = estimates effect of 2-level factor  $x_1$  (sex)

# Let's try this

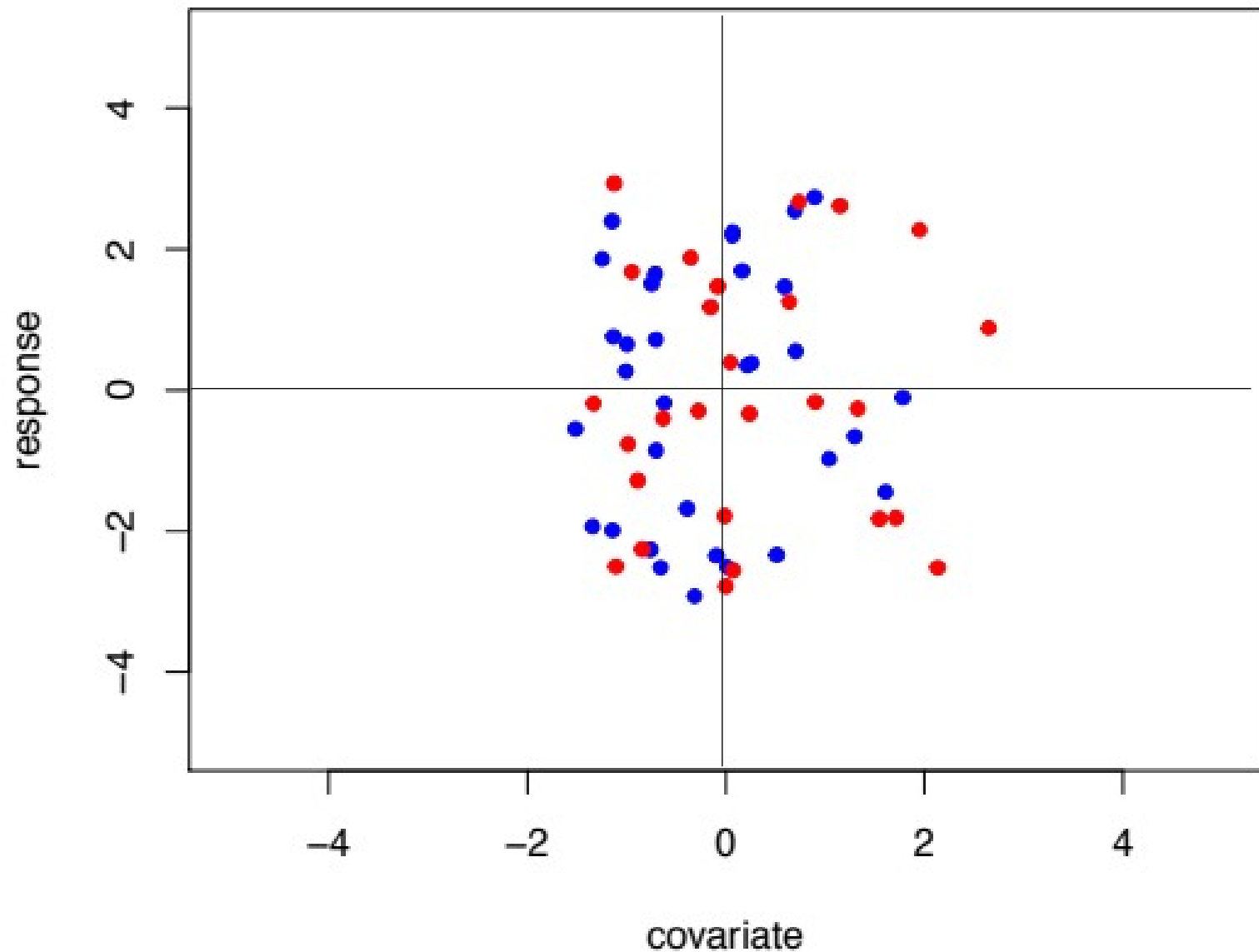
- $b_0$  = intercept

$$y_i = b_0 + b_1 x_{i0} + b_2 x_{i1} + \varepsilon$$

- $b_1$  = estimates effect of continuous variable  $x_0$  (tarsus)
- $b_2$  = estimates effect of 2-level factor  $x_1$  (sex)
- Sex will be re-coded internally. Females are 0 (and blue).

$$y_i = b_0 + b_1 x_{i0} + b_2 x_{i1} + \dots$$

Let's try this



Sex:  $y_i = b_0 + b_1 x_{i0} + b_2 x_{i1} + \varepsilon_i$

Female ( $x_{i2} = 0$ )  $y_i = b_0 + b_1 x_{i0} + b_2 0 + \varepsilon_i$

Male ( $x_{i2} = 1$ )  $y_i = b_0 + b_1 x_{i0} + b_2 1 + \varepsilon_i$

Female ( $x_{i2} = 0$ )  $0 + 0 * x_{i0} + \varepsilon_i$

Male ( $x_{i2} = 1$ )  $0 + 0 * x_{i0} + 0 + \varepsilon_i$

Call:  
lm(formula = y ~ x + sx)

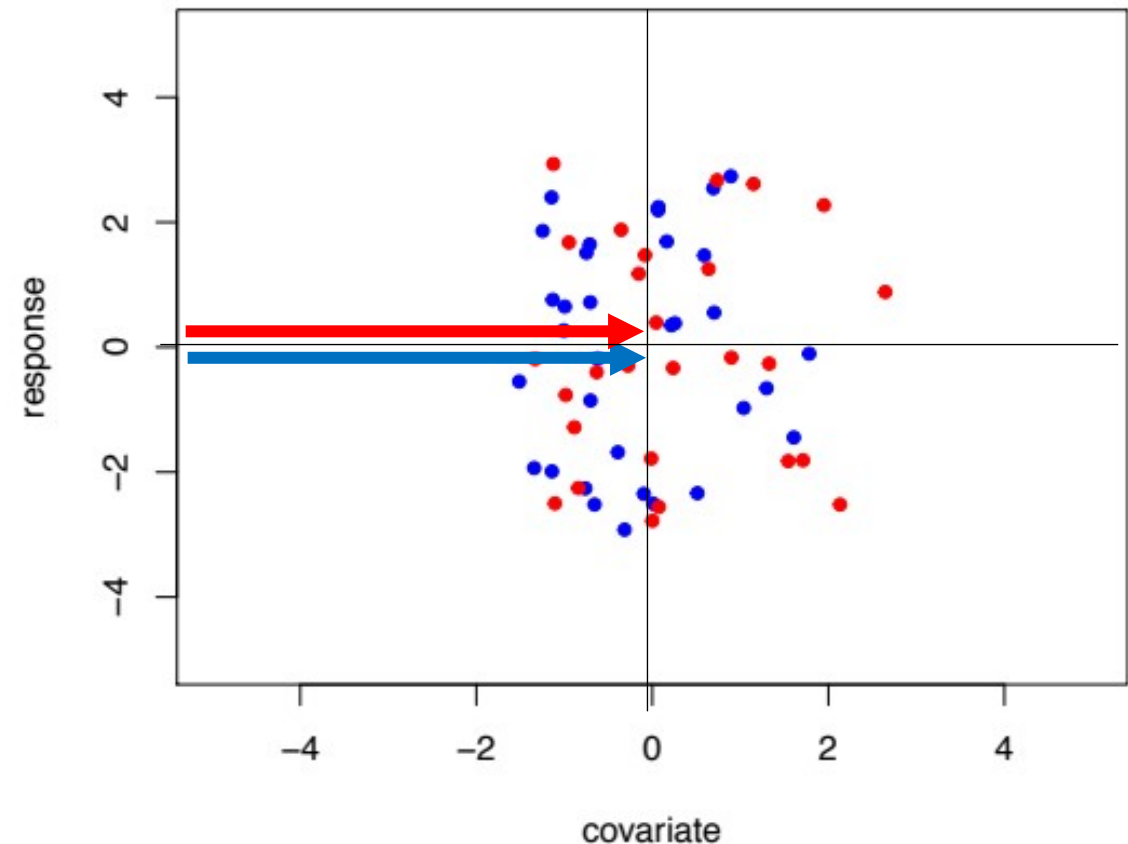
Residuals:

Min	1Q	Median	3Q	Max
-2.8702	-1.7023	-0.1178	1.5936	3.1370

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-0.02613	0.31286	-0.084	0.934
x	0.08222	0.23471	0.350	0.727
sx	-0.08598	0.47224	-0.182	0.856

Residual standard error: 1.784 on 57 degrees of freedom  
Multiple R-squared: 0.00238, Adjusted R-squared: -0.03262  
F-statistic: 0.06799 on 2 and 57 DF, p-value: 0.9343



Sex:

Female ( $x_{i2} = 0$ )

~~$b_0 + b_1 x_{i0} + \epsilon_i$~~

Male ( $x_{i2} = 1$ )

~~$b_0 + b_1 x_{i0} + b_2 + \epsilon_i$~~

Female ( $x_{i2} = 0$ )

$0 + 0 * x_{i0} + \epsilon_i$

Male ( $x_{i2} = 1$ )

$0 + 0 * x_{i0} + 3.91 + \epsilon_i$

Call:  
lm(formula = y ~ x + sx)

Residuals:

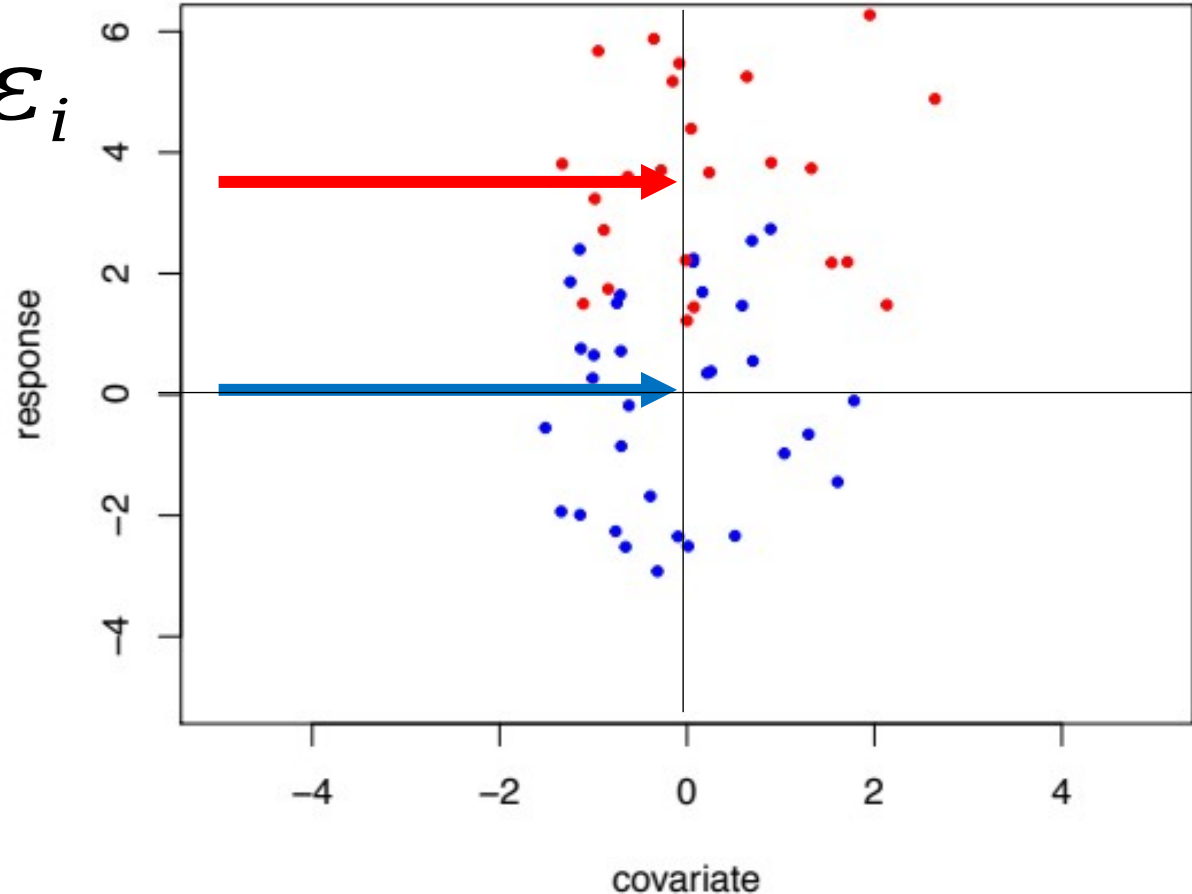
Min	1Q	Median	3Q	Max
-2.8702	-1.7023	-0.1178	1.5936	3.1370

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-0.02613	0.31286	-0.084	0.934
x	0.08222	0.23471	0.350	0.727
sx	3.91402	0.47224	8.288	2.29e-11 ***

---  
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.784 on 57 degrees of freedom  
Multiple R-squared: 0.5608, Adjusted R-squared: 0.5454  
F-statistic: 36.4 on 2 and 57 DF, p-value: 6.524e-11



Sex:

Female ( $x_{i2} = 0$ )

$$b_0 + b_1 x_{i0} + \varepsilon_i$$

Male ( $x_{i2} = 1$ )

$$b_0 + b_1 x_{i0} + b_2 + \varepsilon_i$$

Female ( $x_{i2} = 0$ )

$$0 + 3.08 * x_{i0} + \varepsilon_i$$

Male ( $x_{i2} = 1$ )

$$0 + 3.08 * x_{i0} + 0 + \varepsilon_i$$

Call:

```
lm(formula = y ~ x + sx)
```

Residuals:

Min	1Q	Median	3Q	Max
-2.8702	-1.7023	-0.1178	1.5936	3.1370

Coefficients:

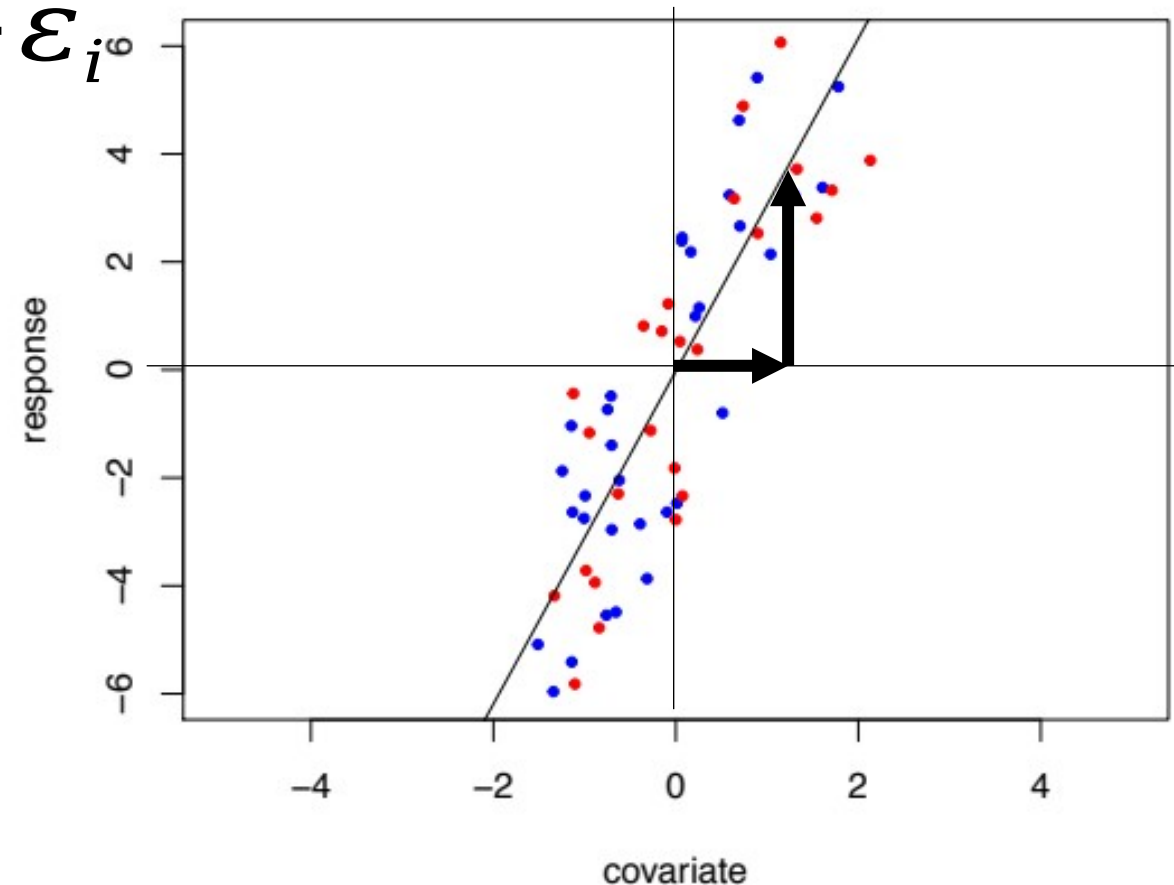
	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-0.02613	0.31286	-0.084	0.934
x	3.08222	0.23471	13.132	<2e-16 ***
sx	-0.08598	0.47224	-0.182	0.856

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.784 on 57 degrees of freedom

Multiple R-squared: 0.7579, Adjusted R-squared: 0.7494

F-statistic: 89.24 on 2 and 57 DF, p-value: < 2.2e-16





Sex:

Female ( $x_{i2} = 0$ )  ~~$b_0 + b_1 x_{i0} + \varepsilon_i$~~

Male ( $x_{i2} = 1$ )  ~~$b_0 + b_1 x_{i0} + b_2 + \varepsilon_i$~~

Female ( $x_{i2} = 0$ )  $0 + 3.08 x_{i0} + \varepsilon_i$

Male ( $x_{i2} = 1$ )  $0 + 3.08 x_{i0} - 2.09 + \varepsilon_i$   
 $- 2.09 + 3.08 x_{i0} + \varepsilon_i$

Call:

`lm(formula = y ~ x + sx)`

Residuals:

Min	1Q	Median	3Q	Max
-2.8702	-1.7023	-0.1178	1.5936	3.1370

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-0.02613	0.31286	-0.084	0.934
x	3.08222	0.23471	13.132	< 2e-16 ***
sx	-2.08598	0.47224	-4.417	4.53e-05 ***

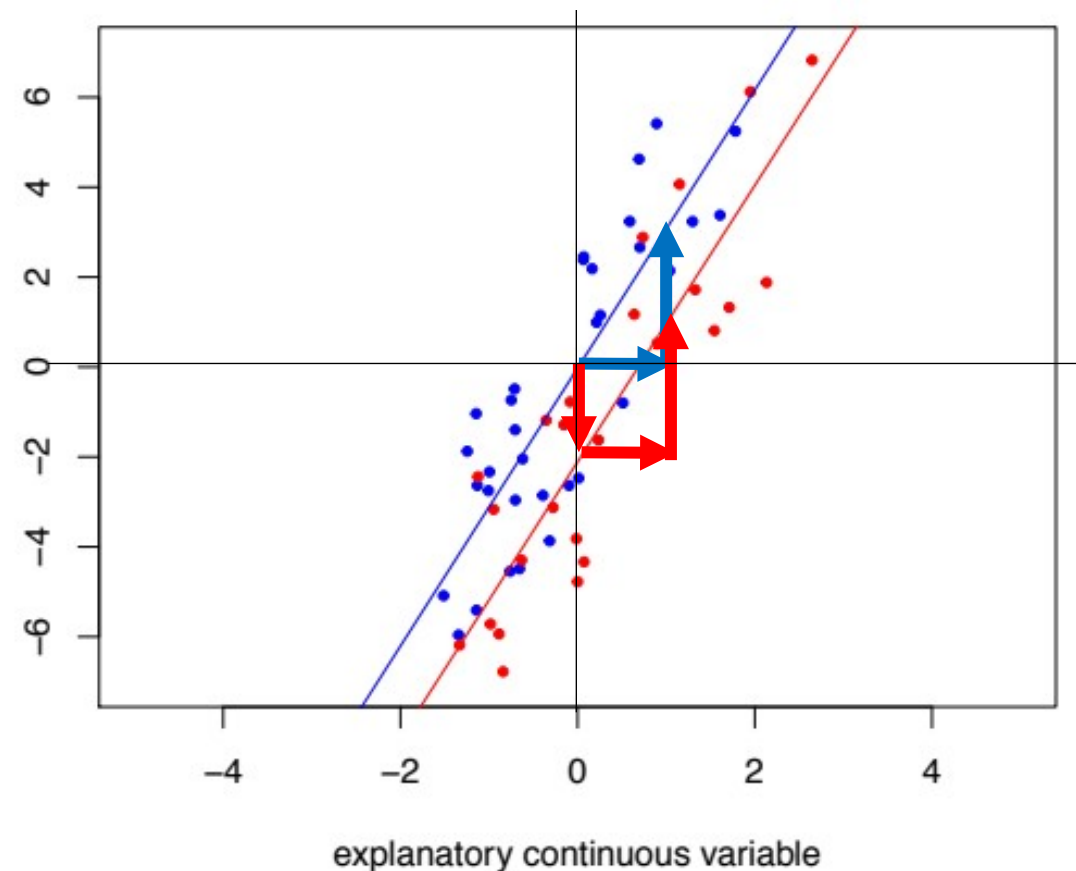
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Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.784 on 57 degrees of freedom

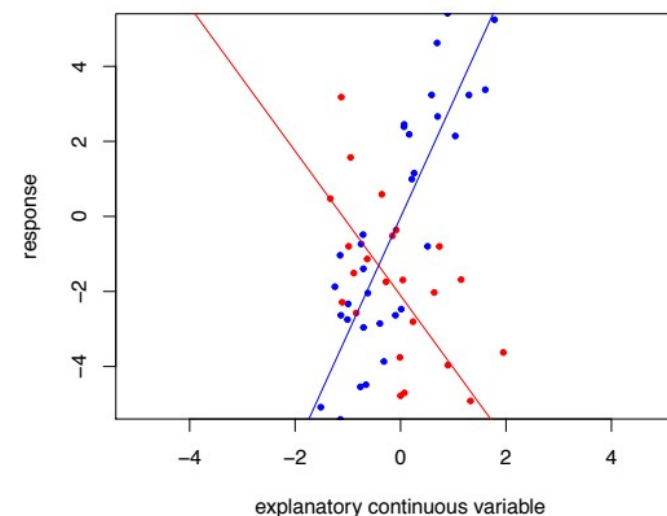
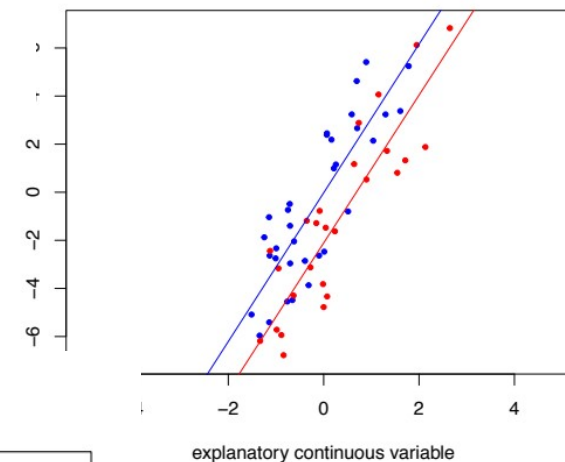
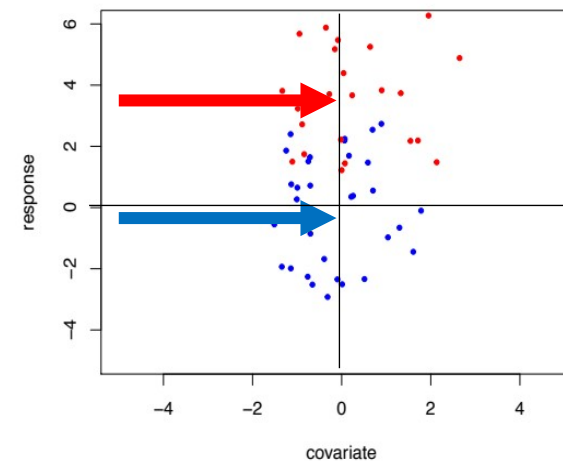
Multiple R-squared: 0.7553, Adjusted R-squared: 0.7467

F-statistic: 87.96 on 2 and 57 DF, p-value: < 2.2e-16



# Ok

- We can estimate two different intercepts
- We can estimate one slope and two intercepts
- How can we estimate a separate slope for each sex?



$$y_i = b_0 + b_1 x_{i0} + b_2 x_{i1} +$$

Interactions of terms:

$$y_i = b_0 + b_1 x_{i0} + b_2 x_{i1} + b_3 x_{i0} x_{i1} + \varepsilon_i$$

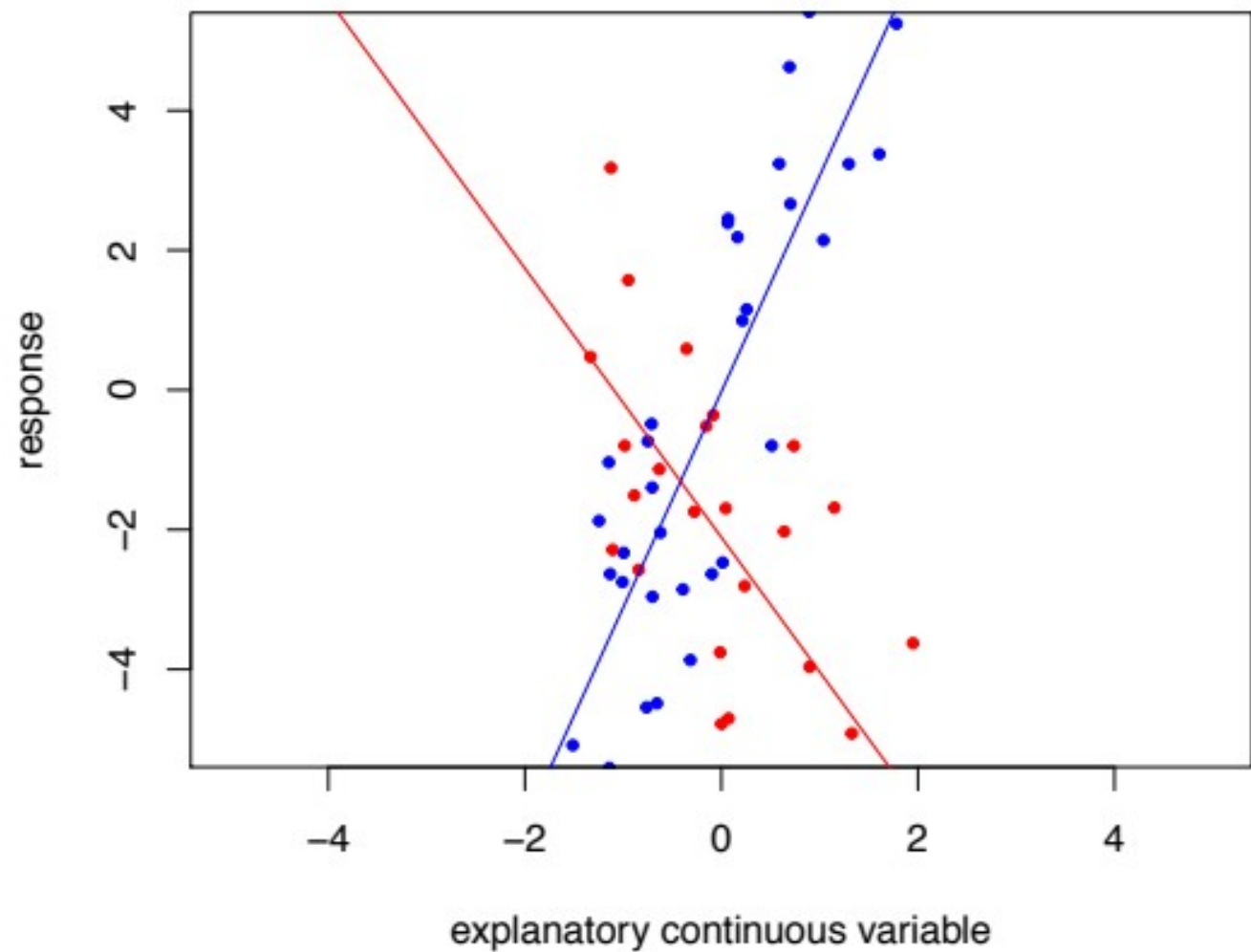
- interaction between sex and tarsus
- one more parameter estimate
- one more degree of freedom
- but not more variables

Sex:

$$y_i = b_0 + b_1 x_{i0} + b_2 x_{i1} + b_3 x_{i0} x_{i1} + \varepsilon_i$$

Female ( $x_{i2} = 0$ )  $b_0 + b_1 x_{i0} + b_2 \cdot 0 + b_3 x_{i0} \cdot 0 + \varepsilon_i$   $\hat{=} b_0 + b_1 x_{i0} + \varepsilon_i$

Male ( $x_{i2} = 1$ )  $b_0 + b_1 x_{i0} + b_2 + b_3 x_{i0} + \varepsilon_i$   $\hat{=} b_0 + b_2 + (b_1 + b_3) x_{i0} + \varepsilon_i$



$$y_i = b_0 + b_1 x_{i0} + b_2 x_{i1} + b_3$$

Female ( $x_{i2} = 0$ )  $b_0 + b_1 x_{i0}$

$$0 + 3.09 x_{i0}$$

Male ( $x_{i2} = 1$ )  $b_0 + b_2 + (b_1 + b_3) x_{i0}$

$$(0 - 2.09) + (3.09 - 5.02) x_{i0}$$

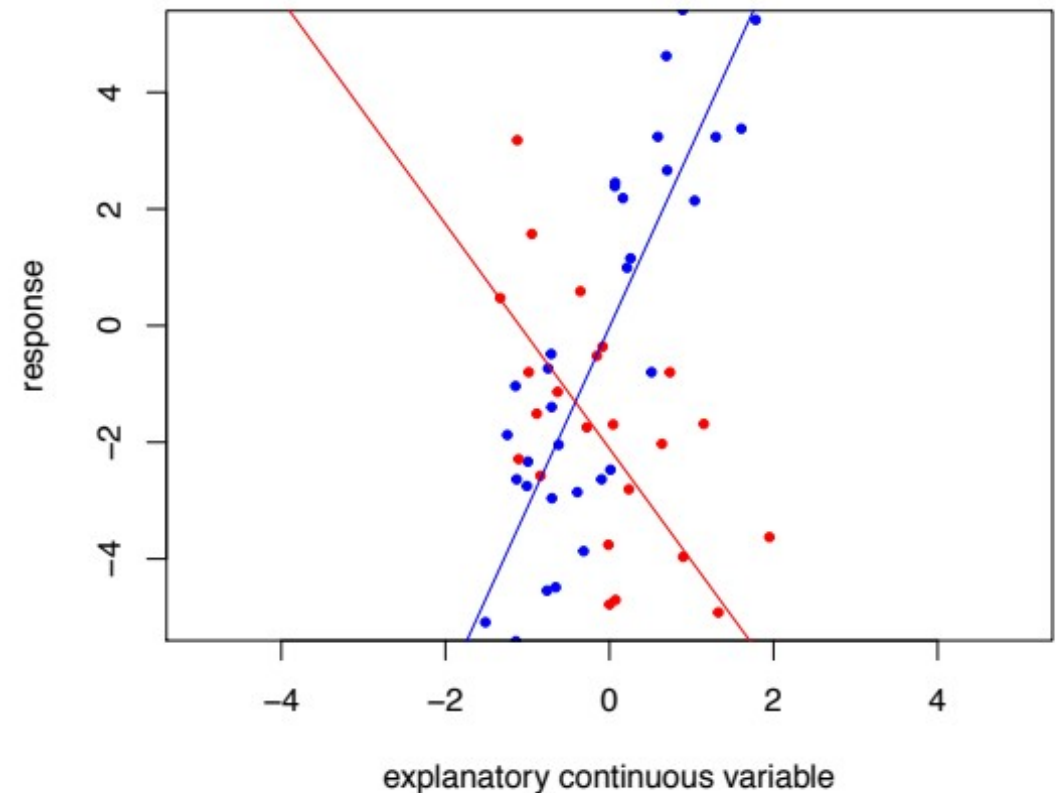
$$-2.09 + 1.93 x_{i0}$$

```
Call:
lm(formula = y ~ x * sx)

Residuals:
    Min       1Q   Median       3Q      Max
-2.8687 -1.7008 -0.1129  1.5931  3.1264

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.02453    0.31856  -0.077   0.939
x             3.09216    0.35685   8.665 6.30e-12 ***
sx           -2.08574    0.47648  -4.377 5.30e-05 ***
x:sx         -5.01776    0.47700 -10.519 7.07e-15 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
Residual standard error: 1.8 on 56 degrees of freedom
Multiple R-squared:  0.7008,    Adjusted R-squared:  0.6848
F-statistic: 43.73 on 3 and 56 DF,  p-value: 1.079e-14
```



$$y_i = b_0 + b_1 x_{i0} + b_2 x_{i1} + b_3$$

Female ( $x_{i2} = 0$ )  $b_0 + b_1 x_{i0}$

$0 + 3.09 x_{i0}$

Male ( $x_{i2} = 1$ )  $b_0 + b_2 + (b_1 + b_3) x_{i0}$

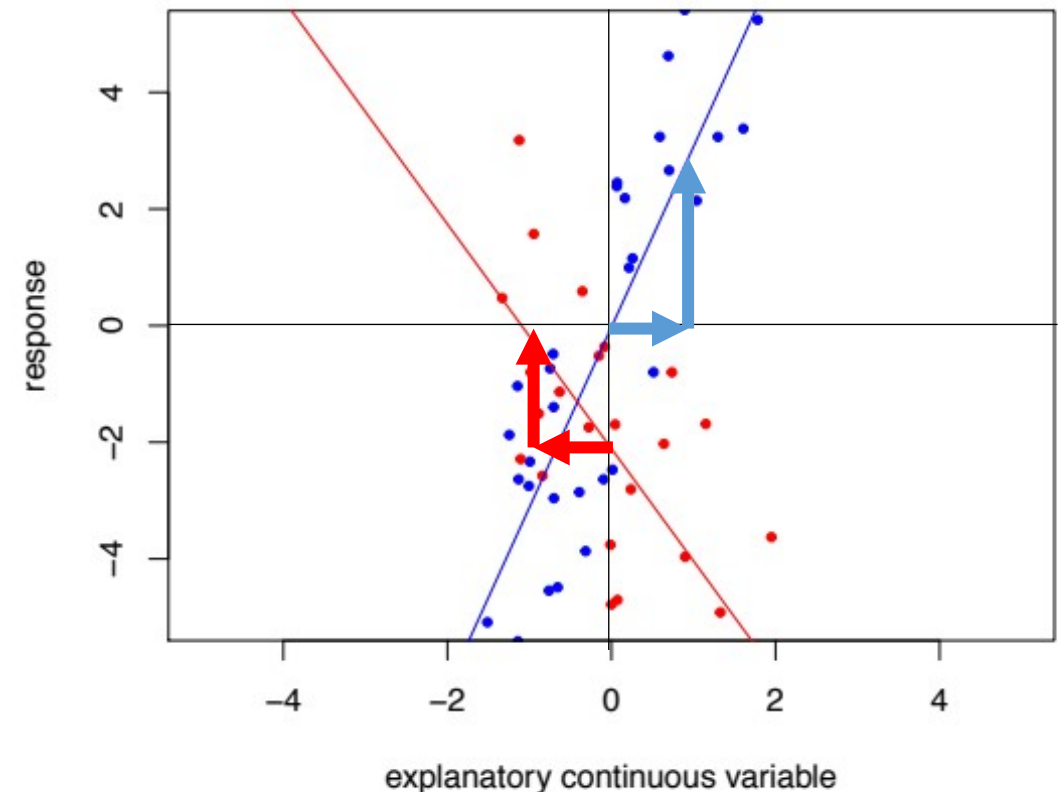
$-2.09 + 1.93 x_{i0}$

```
Call:
lm(formula = y ~ x * sx)

Residuals:
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x:sx         -5.01776    0.47700 -10.519 7.07e-15 ***
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Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.8 on 56 degrees of freedom
Multiple R-squared:  0.7008,    Adjusted R-squared:  0.6848
F-statistic: 43.73 on 3 and 56 DF,  p-value: 1.079e-14
```



$$y_i = b_0 + b_1 x_{i0} + b_2 x_{i1} + b_3 x_{i0} x_{i1} + \varepsilon_i$$

# Let's try this

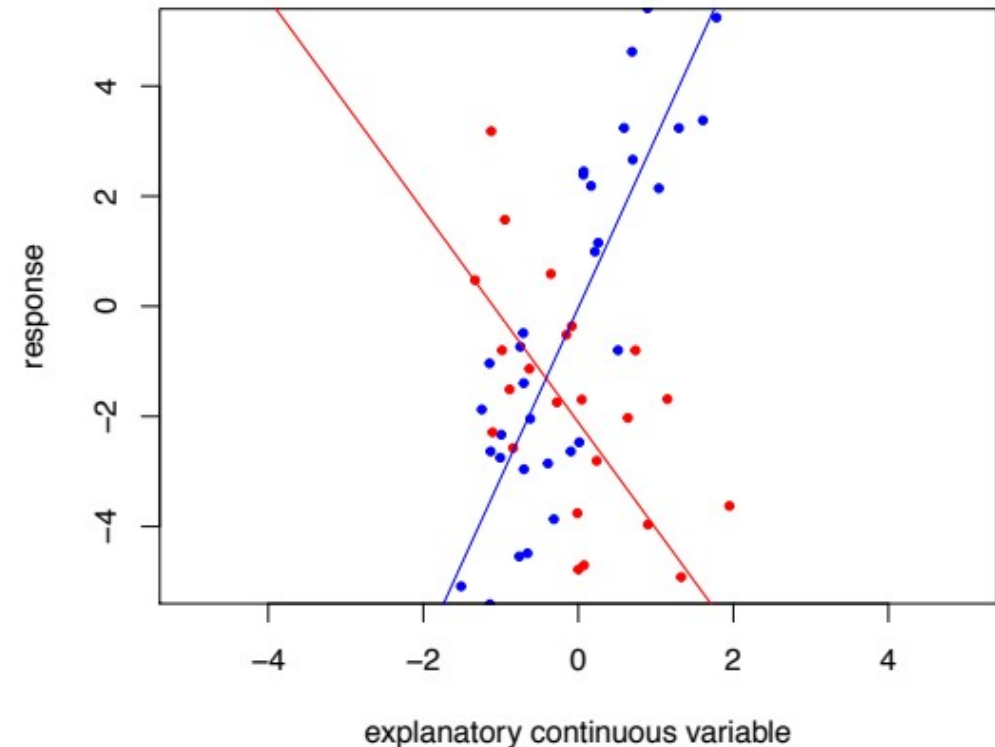
- Intercept:  $b_0$  reference group,  $b_2$ : difference of group 1 to reference
- Slope:  $b_1$  slope of reference group,  $b_3$  difference of slope of group 1 to reference
- Do not interpret estimates by themselves!

```
Call:
lm(formula = y ~ x * sx)

Residuals:
    Min       1Q   Median       3Q      Max
-2.8687 -1.7008 -0.1129  1.5931  3.1264

Coefficients:
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(Intercept) -0.02453    0.31856  -0.077    0.939
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x:sx         -5.01776    0.47700 -10.519 7.07e-15 ***
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Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

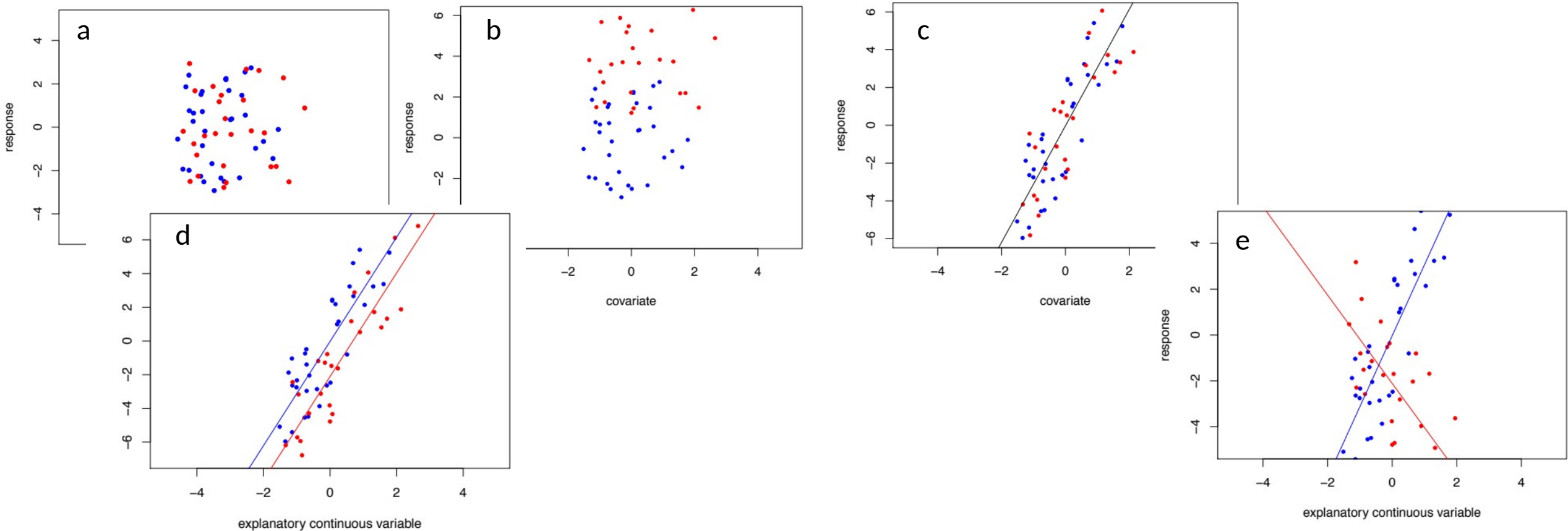
```
Residual standard error: 1.8 on 56 degrees of freedom
Multiple R-squared:  0.7008,    Adjusted R-squared:  0.6848
F-statistic: 43.73 on 3 and 56 DF,  p-value: 1.079e-14
```





$$y_i = b_0 + b_1 x_{i0} + b_2 x_{i1} + b_3 x_{i0} x_{i1} + \varepsilon$$

plot	a	b	c	d	e
Intercept					
$b_1$ (sex)	0	+	0	+	+
$b_2$ (tarsus)	0	0	+	+	+
$b_3$ (tarsus x sex)	0	0	0	0	-





# What else?

- Interactions between continuous variables and categorical predictors with more than 2 levels
  - Multiple continuous predictors
  - Interactions between categorical predictors
  - → hand outs!
- 
- Avoid (at all costs):
  - Interactions between 2 continuous predictors
  - 3- or more-way interactions

# Take home: categorical predictors:

- Interpreting categorical x continuous predictor interactions
- When interaction terms are present – never interpret the estimates in isolation
- You will need to do some math's to interpret your results