

# Estimation of a TVP-VAR with MSV via Efficient Importance Sampling

## *Project Proposal*

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# Motivation

## Using VAR models to describe macroeconomic relations

“It should be feasible to estimate large-scale macromodels as unrestricted reduced forms, **treating all variables as endogenous.**”

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- ▶ RBC models x **VAR approach**
- ▶ Strong evidence of parameter instability in macroeconomic series:  
[Sims, 1999], [Bernanke and Mihov, 1998a], [Bernanke and Mihov, 1998b],  
[Mumtaz and Zanetti, 2015], [Galí and Gambetti, 2015],  
[Mumtaz and Theophilopoulou, 2015],  
[Mumtaz and Theophilopoulou, 2017], **[Cogley and Sargent, 2001]**,  
**[Cogley and Sargent, 2005]**, etc.

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Cogley and Sargent (2005)

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INFLATION-UNEMPLOYMENT  
OUTCOMES IN 1970S?

Bad Luck *vs* Bad Policy

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*Changes in volatility structure*

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It was not monetary strategy but  
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Changes in  
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*Drifting coefficients*

**Bad policy:**

The FED's view negates the natural  
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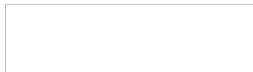


# Estimation of a TVP-VAR with MSV via EIS

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Economic Motivation

*You are here!*



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Economic Motivation

TVP-VAR + MSV

Cogley and Sargent (2001, 2005)

Prímícerí (2005)

Phillípov and Glickman (2006a, 2006b)

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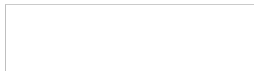
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IS, EIS, Sequential EIS

Sequential EIS + Wishart MSV Model



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# Estimation of a TVP-VAR with MSV via EIS

## Literature Review: TVP-VAR w/ MSV

Consider the following state-space representation of a TVP-VAR with MSV:

$$y_t = Z_t \alpha_t + \epsilon_t \quad \epsilon_t \sim \mathcal{N}_k(0_k, \Omega_t^{-1}) \quad (\text{measure eq.}) \quad (1)$$

$$\alpha_t = \alpha_{t-1} + v_t \quad v_t \sim \mathcal{N}_p(0_p, Q) \quad (\text{state transition eq.}) \quad (2)$$

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- ▶ If  $\Omega_t^{-1}$  were non-stochastic, the Kalman Filter could be used to obtain the ML function;
  - ▶ Since it is not the case, we have an **analytically intractable high-dimensional integral**.

# Estimation of a TVP-VAR with MSV via EIS

Literature Review: Cogley and Sargent (2001, 2005)

Their first work was a TVP-VAR of the form:

$$y_t = Z_t \alpha_t + \epsilon_t \quad \epsilon_t \sim \mathcal{N}_k(0_k, R) \quad (3)$$

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After being criticized in (2001) for the lack of MVS in their model, [Cogley and Sargent, 2005] proposed changing  $\epsilon_t$  in (3) for  $\epsilon_t = R_t^{1/2} \xi_t$ , where  $\xi_t$  follows a standard normal distribution and  $R_t = B^{-1} H_t B^{-1'}$  with:

$$B = \begin{bmatrix} 1 & 0 & \cdots & 0 \\ \beta_{21} & 1 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ \beta_{k1} & \beta_{k2} & \cdots & 1 \end{bmatrix}, \quad \text{and} \quad H_t = \begin{bmatrix} h_{1t} & 0 & \cdots & 0 \\ 0 & h_{2t} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & h_{kt} \end{bmatrix}$$

where  $\ln(h_{it}) = \ln(h_{it-1}) + \sigma_i \eta_{it}$ ,  $\eta_{it} \sim \mathcal{N}(0, 1)$ .

# Estimation of a TVP-VAR with MSV via EIS

## Literature Review: Primiceri (2005)

[Primiceri, 2005] proposed the same specification as [Cogley and Sargent, 2005] but allowed the covariances in  $B$  to evolve over time:

$$B_t = \begin{bmatrix} 1 & 0 & \cdots & 0 \\ \beta_{21,t} & 1 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ \beta_{k1,t} & \beta_{k2,t} & \cdots & 1 \end{bmatrix} \quad \text{with} \quad \beta_t = \beta_{t-1} + \nu_t \quad (5)$$

where  $\nu_t$  are mean-zero normal errors with constant covariance matrix.

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**Drawback:** (5) is not invariant to the observable variables  $y_t$  order!

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The problem lies in the **MSV structure**!

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Literature Review: Phillipov and Glickman (2006a, 2006b)

[Phillipov and Glickman, 2006b] and [Phillipov and Glickman, 2006a] proposes the following model:

$$y_t = \epsilon_t, \quad \epsilon \sim \mathcal{N}_k(0_k, \Omega_t^{-1}) \quad (6)$$

$$\Omega_t | \Omega_{t-1} \sim \mathcal{W}(\nu, S_{t-1}), \quad \text{with } S_t = \frac{1}{\nu} \mathbf{A}^{1/2} \Omega_t^{\textcolor{red}{d}} \mathbf{A}^{1/2'} \quad (7)$$

where

- ▶  $\mathbf{A}$  is a positive definite symmetric matrix containing parameters determining the intertemporal sensitivity of each element of the precision matrix and
- ▶  $\textcolor{red}{d} \in [0, 1)$  is a scalar that accounts for overall persistence of the process.

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This specification is **more flexible** than Primiceri's one,  
and it is **invariant to ordering**!



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- ▶ [Philipov and Glickman, 2006a] had trouble estimating simultaneously the elements of the matrix  $A$  and persistence parameter  $d$  and proposed to use a fixed  $d$ ;

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- ▶ [Asai and McAleer, 2009] reported convergence problems in the Gibbs sampler.

# Estimation of a TVP-VAR with MSV via EIS

Where are we?

## Primiceri (2005)

- The order of the variables  $y_t$  matters;
- Simple random walk structure for covariances;
- Missing steps in Gibbs sampler.

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- MVS specification
- Invariant to the order in  $y_t$ ;
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- Almost same number of parameters.

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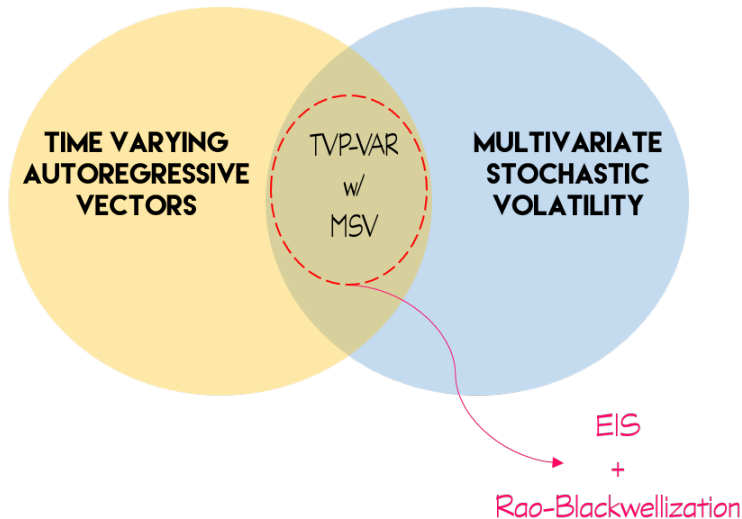
# Estimation of a TVP-VAR with MSV via EIS

## Objective

Develop and implement efficient statistical inference procedures in TVP-VAR with MSV models.

# Estimation of a TVP-VAR with MSV via EIS

## Proposal





# Estimation of a TVP-VAR with MSV via EIS

## Proposal

Build a **TVP-VAR** with **Wishart MSV** model based on [Philipov and Glickman, 2006b] and obtain **maximum likelihood** estimates using **EIS** [Richard and Zhang, 2007] combined with a **Rao-Blackellization** step [Moura and Turatti, 2014].

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3. Develop an empirical application using the model developed in (2), in line with [Mumtaz and Theophilopoulou, 2017].

# Estimation of a TVP-VAR with MSV via EIS

## Methodology: Monte Carlo Methods and Importance Sampling

- ▶ Monte Carlo (MC) Methods are an alternative way to solve complex integrals;
  - ▶ Specially used for high-dimension problems, where non-stochastic algorithms are way too slow;
- ▶ The idea is to resample values from a probability density (thus, is a stochastic method), using a pseudo-random number generator and the inverting theorem.

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If  $X$  is a random vector with pfd  $f_X(\cdot)$ , then (8) is the expectation of  $g(X)$  with respect to  $f(\cdot)$ :

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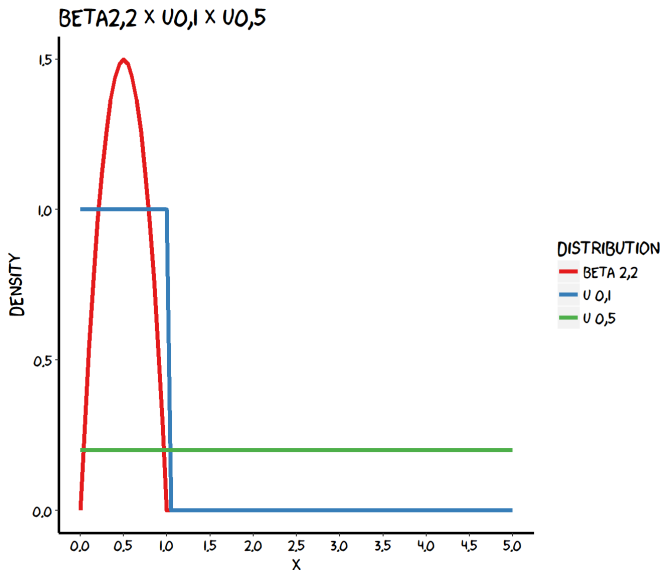
In cases where the function  $f$  is known, using the law of large numbers we can use the sample mean to approximate (8):

$$I \approx \frac{1}{N} \sum_{i=1}^N g(x_i) \quad (10)$$



# Estimation of a TVP-VAR with MSV via EIS

Methodology: Monte Carlo Methods and Importance Sampling



# Estimation of a TVP-VAR with MSV via EIS

## Methodology: Monte Carlo Methods and Importance Sampling

Consider the following integral:

$$I = \int_{\mathcal{X}} g(x) \frac{f(x)}{m(x)} m(x) dx \quad (11)$$

Now we have an expectation with respect to the density  $m(\cdot)$ :

$$I = \mathbb{E}_m \left[ \frac{g(x)f(x)}{m(x)} \right] \quad (12)$$

In cases where the function  $m$  is known, using the law of large numbers we can use the sample mean to approximate (11):

$$I \approx I_N^{IS} = \frac{1}{N} \sum_{i=1}^N \omega(x_i) g(x_i), \quad \text{with} \quad \omega_i = \frac{f(x_i)}{m(x_i)} \quad (13)$$

# Estimation of a TVP-VAR with MSV via EIS

## Methodology: Efficient Importance Sampling

Two main “problems” arise when dealing with IS:

- ▶ The choice of the class of samplers,  $M$ ;

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- ⇒ EIS resolves the last one!

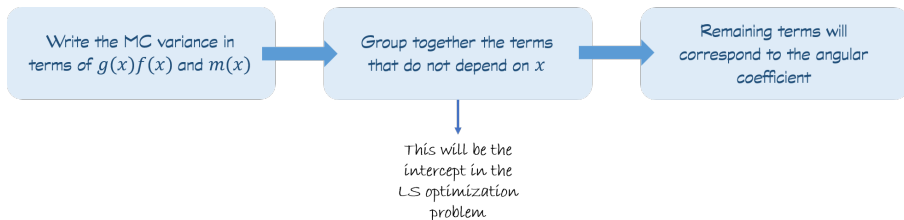
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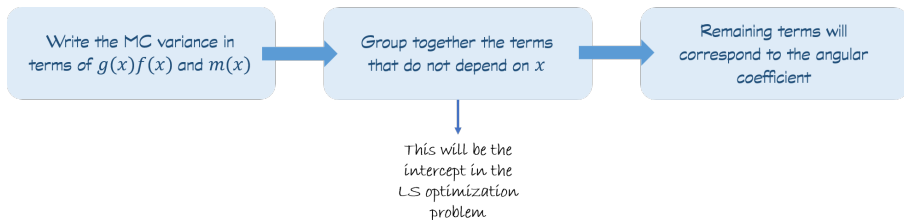
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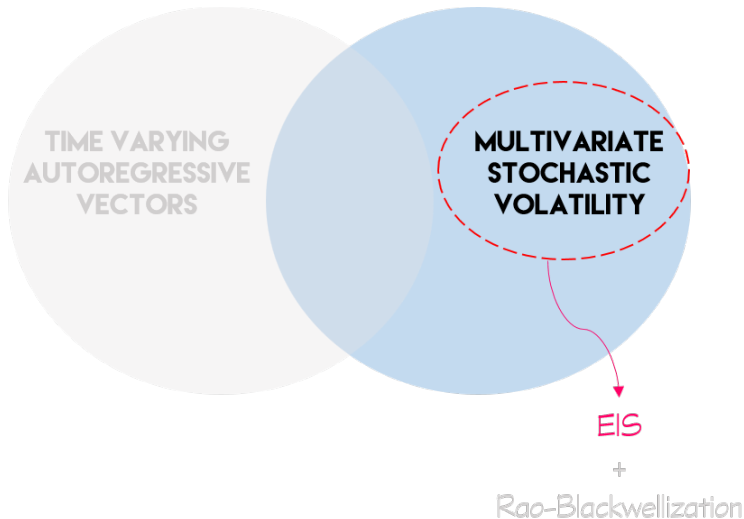
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- ▶ Sequential EIS is used in high-dimensional problems and consists in “breaking” the EIS algorithm in a sequence of low-dimension optimizations.

# Estimation of a TVP-VAR with MSV via EIS

Methodology: Wishart MSV model





# Estimation of a TVP-VAR with MSV via EIS

## Methodology: Wishart MSV model

Recall the model given in (6)-(7). Defining  $f(\underline{y}, \underline{\Omega}, \theta)$  as the joint density of  $\underline{y} = \{y_t\}_{t=1}^T$  and  $\underline{\Omega} = \{\Omega_t\}_{t=1}^T$ , where  $\theta$  is the vector of unknown parameters, the likelihood associated to  $\underline{y}_t$  is given by:

$$\begin{aligned} L(\theta|\underline{y}) &= \int f(\underline{y}, \underline{\Omega}, \theta) d\underline{\Omega} \\ &= \int \cdots \int \prod_{t=1}^T f(y_t, \Omega_t | \underline{\Omega}_{t-1}, \underline{y}_{t-1}; \theta) d\Omega_1 \dots d\Omega_T \end{aligned} \quad (14)$$

with  $\underline{y}_t$  e  $\underline{\Omega}_t$  given by  $\{y_t\}_{t=1}^T$  e  $\{\Omega_t\}_{t=1}^T$ , respectively.

# Estimation of a TVP-VAR with MSV via EIS

## Methodology: Wishart MSV model

The respective space-state representation is:

$$f(y_t, \Omega_t | \underline{\Omega}_{t-1}, \underline{y}_{t-1}; \theta) = g(y_t | \Omega_t, \underline{y}_{t-1}; \theta) \cdot p(\Omega_t | \underline{\Omega}_{t-1}, \underline{y}_{t-1}; \theta) \quad (15)$$

where

$$g(y_t | \Omega_t, \underline{y}_{t-1}; \theta) \propto |\Omega_t|^{1/2} \exp \left\{ -\frac{1}{2} y_t' \Omega_t y_t \right\} \quad (16)$$

$$p(\Omega_t | \underline{\Omega}_{t-1}, \underline{y}_{t-1}; \theta) \propto |\Omega_t|^{\frac{\nu-k-1}{2}} \exp \left\{ -\frac{1}{2} \text{tr} \left[ S_{t-1}^{-1} \cdot \Omega_t \right] \right\} \quad (17)$$

# Estimation of a TVP-VAR with MSV via EIS

## Methodology: Wishart MSV model

Likelihood:

$$L(\theta|\underline{y}) = \int \prod_{t=1}^T \left[ \frac{f(y_t, \underline{\Omega}_t | \underline{\Omega}_{t-1}, \underline{y}_{t-1}; \theta)}{m(\underline{\Omega}_t | \underline{\Omega}_{t-1}; \gamma_t)} \right] \prod_{t=1}^T m(\underline{\Omega}_t | \underline{\Omega}_{t-1}; \gamma_t) d\underline{\Omega} \quad (18)$$

with  $m(\underline{\Omega}_t | \underline{\Omega}_{t-1}; \gamma_t) = \frac{k(\underline{\Omega}_t; \gamma_t)}{\chi(\underline{\Omega}_{t-1}; \gamma_t)}$  where  $\chi(\underline{\Omega}_{t-1}; \gamma_t) = \int k(\underline{\Omega}_{t-1}; \gamma_t) d\underline{\Omega}_t$ .

MC estimate:

$$\tilde{L}(\theta|\underline{y}) = \frac{1}{N} \sum_{i=1}^N \left[ \prod_{\ell=1}^L \frac{f(y_t, \tilde{\underline{\Omega}}_t^i | \tilde{\underline{\Omega}}_{t-1}^i, \underline{y}_{t-1}; \theta) \cdot \chi(\tilde{\underline{\Omega}}_{t-1}^i; \gamma_t)}{k(\tilde{\underline{\Omega}}_t^i; \gamma_t)} \right] \quad (19)$$

Minimization problem:

$$\hat{\gamma}_t(\theta) = \min_{\gamma_t} \sum_{i=1}^N \left\{ \ln [f(y_t, \tilde{\underline{\Omega}}_t^i | \tilde{\underline{\Omega}}_{t-1}^i, \underline{y}_{t-1}; \theta) \cdot \chi(\tilde{\underline{\Omega}}_{t-1}^i; \gamma_t)] - c_t - \ln k(\underline{\Omega}_t; \gamma_t) \right\} \quad (20)$$

# Estimation of a TVP-VAR with MSV via EIS

## Workplan - Completed tasks

	2017				2018												2019		
	S	O	N	D	J	F	M	A	M	J	J	A	S	O	N	D	J	F	M
1. Delimitation of the problem																			
2. LR: Efficient Importance Sampling																			
3. Implement EIS algorithm to replicate results from Richard and Zhang (2007)																			
														</					

Legend:

- Econometrics II article
- Macroeconomics II seminar and/or article
- Key dates

LR := literature revision

# Estimation of a TVP-VAR with MSV via EIS

## Workplan - Econometrics II related tasks

	2017				2018												2019		
	S	O	N	D	J	F	M	A	M	J	J	A	S	O	N	D	J	F	M
2. LR: Efficient Importance Sampling																			
3. Implement EIS algorithm to replicate results from Richard and Zhang (2007)																			
4. LR: MSV models																			
5. Pre-Project Deliver																			
7. Implement EIS algorithm in Phillipov and Glickman (2006) model																			

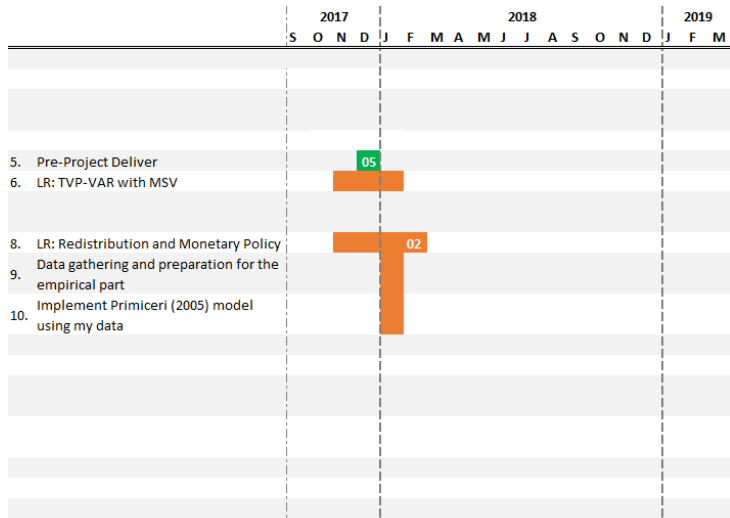
Legend:

- Econometrics II article
- Macroeconomics II seminar and/or article
- Key dates

LR := literature revision

# Estimation of a TVP-VAR with MSV via EIS

## Workplan - Macroeconomics II related tasks



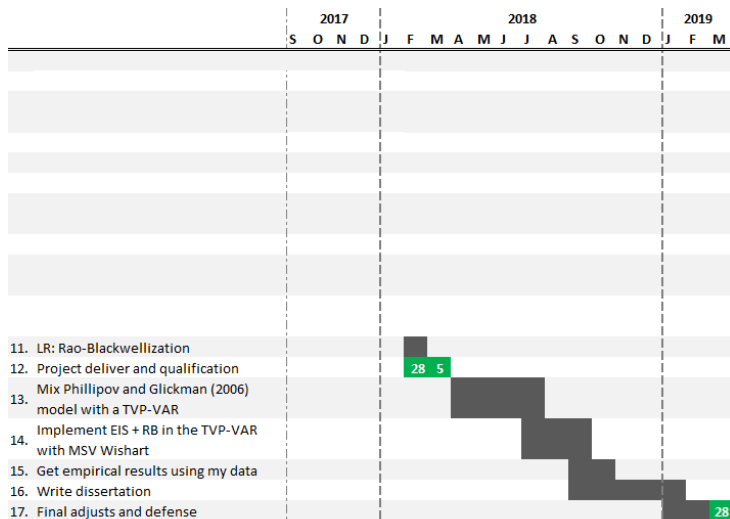
Legend:

- Econometrics II article
- Macroeconomics II seminar and/or article
- Key dates

LR := literature revision

# Estimation of a TVP-VAR with MSV via EIS

## Workplan - What future holds



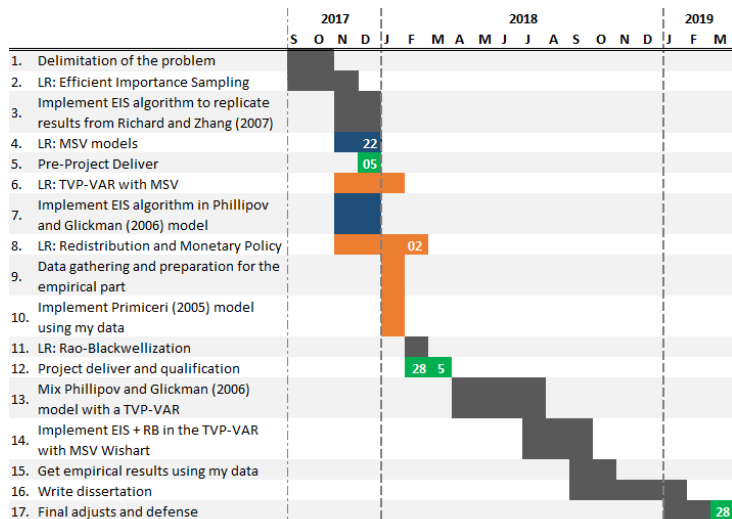
Legend:

- Econometrics II article
- Macroeconomics II seminar and/or article
- Key dates

LR := literature revision

# Estimation of a TVP-VAR with MSV via EIS

## Workplan - General Overview



Legend:

■ Econometrics II article

■ Macroeconomics II seminar and/or article

■ Key dates

LR := literature revision



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