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Reading: Chapter 5 in J.F.. Chapter. 13.1-13.2 in R.C.

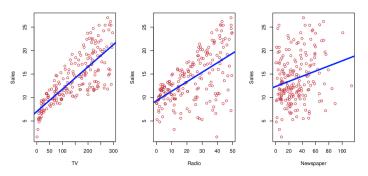
• Recall multiple linear regression:

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_p X_p + \varepsilon$$
, where  $\varepsilon \sim N(0, \sigma^2)$ .

- We interpret  $\beta_j$  as the mean change in Y per unit change in  $X_j$ , holding all other predictors fixed.
- E.g., consider the relationship between sales and advertising budget on various media:

Sale = 
$$\beta_0 + \beta_1 \text{TV} + \beta_2 \text{radio} + \beta_3 \text{newspaper} + \varepsilon$$
,  $\varepsilon \sim N(0, \sigma^2)$ .

## Advertising data



- Is at least one of the predictors  $X_1, X_2, ..., X_p$  useful in predicting the response?
- Do all the predictors help to explain Y, or is only a subset of the predictors useful?
- Which media contribute most to sales?
- Is there synergy among the advertising media?



## Results from advertising data

	Coefficient	Std. Error	t-statistic	p-value
Intercept	2.939	0.3119	9.42	< 0.0001
TV	0.046	0.0014	32.81	< 0.0001
radio	0.189	0.0086	21.89	< 0.0001
newspaper	-0.001	0.0059	-0.18	0.8599

#### Correlations:

	TV	radio	newspaper	sales	
TV	1.0000	0.0548	0.0567	0.7822	
radio		1.0000	0.3541	0.5762	
newspaper			1.0000	0.2283	
sales				1.0000	

How to we interpret  $\hat{\beta}_3 < 0$ , but Cov(newspage, sales) > 0?

- The ideal scenario is when the predictors are uncorrelated
  - ▶ Each coefficient can be estimated and tested separately.
  - ▶ Interpretation such as "a unit change in  $X_j$  is associated with a  $\beta_j$  average change in Y, while holding all other predictors fixed".
- Correlation amongst predictors cause problem:
  - ► The variance of all coefficient estimates tends to increase, sometimes dramatically.
  - Interpretations become hazardous.

## Two quotes by famous statisticians

 "Essentially, all models are wrong, but some are useful" George Box!

 "The only way to find out what will happen when a complex system is disturbed is to disturb the system, not merely to observe it passively" Fred Mosteller and John Tukey, paraphrasing George Box

- When the explanatory variables are correlated among themselves, **multicollinearity** among them is said to exist.
- Consider two extreme cases.
  - Case 1: Uncorrelated explanatory variables.
  - Case 2: Perfectly correlated explanatory variables.

## Case 1: Uncorrelated Explanatory Variables

- Suppose  $Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \varepsilon$ .
- Suppose  $X_1$  and  $X_2$  are orthogonal such that the sample correlation between  $X_1$  and  $X_2$  is 0.

$$\sum_{i=1}^{n} (X_{i1} - \bar{X}_1)(X_{i2} - \bar{X}_2) = 0$$

• We can show (why?)

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (Y_i - \bar{Y})(X_{i1} - \bar{X}_1)}{\sum_{i=1}^n (X_{i1} - \bar{X}_1)^2}, \quad \hat{\beta}_2 = \frac{\sum_{i=1}^n (Y_i - \bar{Y})(X_{i2} - \bar{X}_2)}{\sum_{i=1}^n (X_{i2} - \bar{X}_2)^2}.$$

- That is, the LS estimate of  $\beta_1$  is not affected by  $X_2$  and the LS estimate of  $\beta_2$  is not affected by  $X_1$ .
- Interpretation of regression coefficients is clear:  $\beta_1$  (or  $\beta_2$ ) is the expected change in Y for one unit increase in  $X_1$  (or  $X_2$ ) with  $X_2$  (or  $X_1$ ) held constant.

## Case 2: Perfectly Correlated Explanatory Variables

- Again, suppose  $Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \varepsilon$ .
- But  $X_2 = 2X_1 + 1$ .
- Suppose  $\beta_0 = 3, \beta_1 = 2, \beta_2 = 5$ .
- Then all the following models give the same fit for Y:
  - $Y = 3 + 2X_1 + 5X_2 + \varepsilon$ .
  - ►  $Y = 8 + 12X_1 + ε$ .
  - ►  $Y = 2 + 6X_2 + ε$ .

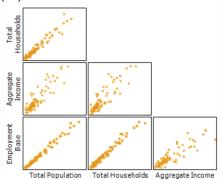
For example, with 1 unit increase in  $X_1$ , there are 2 units increase in  $X_2$  and  $\beta_1 + 2\beta_2$  change in Y.

## Consequences of Multicollinearity

- In practice, most cases are in between the two extreme cases.
- Effect of multicollinearity on the inference of regression coefficients.
  - ▶ Larger changes in the fitted  $\hat{\beta}_k$  when another X is added or deleted.
  - ▶ More difficult to interpret  $\hat{\beta}_k$  as the effect of  $X_k$  on Y, because the other X's cannot be held constant.
  - ▶ **X**<sup>t</sup>**X** ill-conditioned or rank-deficient
  - Estimates become sensitive to minor changes of data. (why?)

# Diagnostics for Multicollinearity

- Large changes in  $\hat{\beta}$ 's when an explanatory variable (or an observation) is added or deleted.
- Significant joint effects for the affected variables, but wide confidence intervals for  $\beta$ 's corresponding to important explanatory variables.
- The sign of  $\hat{\beta}$  is counter-intuitive.
- Explanatory variables are highly correlated. e.g. scatter plot matrix R command: pairs(...)



## Variance Inflation Factor (VIF)

• Variance inflation factor (VIF) for  $\hat{\beta}_k$ :

$$VIF_k = \frac{1}{1 - R_k^2}, \quad k = 1, \dots, p - 1$$

where  $R_k^2$  is the coefficient of multiple determination when  $X_k$  is regressed on the p-2 other X explanatory variables.

• That is,  $R_k^2$  is the coefficient of multiple determination  $R^2$  of the model

$$X_k = \beta_0 + \beta_1 X_1 + \dots + \beta_{k-1} X_{k-1} + \beta_{k+1} X_{k+1} + \dots + \beta_{p-1} X_{p-1} + \varepsilon.$$

- If the mean VIF values of VIF<sub>k</sub> (k = 1, ..., p 1) is considerably greater than 1, there may be serious multicollinearity problems.
- If the largest VIF value among VIF<sub>k</sub> (k = 1, ..., p 1) is larger than 10, multicollinearity may have a large impact on the inference.