

Principle Component Analysis (PCA)

Suppose that $x_1, x_2, x_3, \dots, x_n \in \mathbb{R}^n$ are a sequence of identically and independently distributed random vectors with mean

$$\mu = \mathbb{E}[x] = (\mathbb{E}[x^{(1)}], \dots, \mathbb{E}[x^{(n)}])^T \in \mathbb{R}^n$$

and covariance matrix

$$C_{ij} = E[(x^{(i)} - \mu^{(i)})(x^{(j)} - \mu^{(j)})] \Rightarrow C = E[(x - \mu)(x - \mu)^T]$$

What does PCA aim to do?

View 1: Choose a new coordinate system for "feature space" \mathbb{R}^n in which the data or random variable has independent entries/words.

new features $y_i = u^T (x_i - \mu)$ old features μ mean correction

Find this to make y_i entries i.i.d.

$$\overset{\text{mean zero}}{\Rightarrow} E[y] = E[U^T(x_j - \mu)] = U^T E(x_j) - U^T \mu = 0$$

$$\begin{aligned} \Rightarrow E[yy^T] &= E[U^T(x - \mu)(x - \mu)^T U] \\ &= U^T E[(x - \mu)(x - \mu)^T] U \\ &= U^T C U \end{aligned}$$

Q: How should we choose the transform U ?

$$\begin{array}{ccc} i=1, \dots, n & [v_1 \dots v_n] & [\lambda_1 \dots \lambda_n] \\ C v_i = \lambda_i v_i & \Rightarrow & C V = V \Lambda \end{array}$$

$$U=V \Rightarrow E[yy^T] = V^T C V = V^T V \Lambda$$

Key Point: Covariance Matrix C is symmetric

$$C_{ij} = E[(x^{(i)} - \mu^{(i)})(x^{(j)} - \mu^{(j)})] = C_{ji} \Rightarrow C = C^T$$

$$\Rightarrow V \text{ is orthogonal so } V^T V = I.$$

$$\Rightarrow E[yy^T] = \Lambda = \begin{bmatrix} \lambda_1 & & \\ & \ddots & \\ & & \lambda_n \end{bmatrix}$$

Note that $\lambda_1 \geq 0, \dots, \lambda_n \geq 0$ and C is pos. def.,

$$U = V \begin{bmatrix} \frac{1}{\sqrt{\lambda_1}} & & \\ & \ddots & \\ & & \frac{1}{\sqrt{\lambda_n}} \end{bmatrix} \Rightarrow U^T C U = \sqrt{\Lambda}^{-1} V^T C V \sqrt{\Lambda}^{-1} \\ = \sqrt{\Lambda}^{-1} \Lambda \sqrt{\Lambda}^{-1} \\ = I$$

Principal Components

After transformation with $U = V =$ eigenvectors of C ,

new features $y_i = U^T (x_i - \mu)$ old features

\uparrow Find this to make y_i entries i.i.d. \nwarrow mean correction

$$\Rightarrow E[y] = 0 \quad \text{and} \quad E[yy^T] = I$$

Convention: $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_n \geq 0$

The k^{th} principal component of x is

$$y^{(k)} = u_k^T (x - \mu)$$

"Coordinate of $x - \mu$ along u_k ."

View 2: The leading principle component(s) capture the "most" variance in the distribution of random vector in the following sense.

$$\arg \max_{\|v\|=1} \mathbb{E}[v^T(x-\mu)(x-\mu)v^T] = \arg \max_{\|v\|=1} v^T \mathbb{E}[(x-\mu)(x-\mu)^T] v$$

$$\frac{v^T C v}{v^T v} = \text{"Rayleigh Quotient"}$$

$$= \arg \max_{\|v\|=1} v^T C v$$

The Rayleigh Quotient $\frac{v^T C v}{v^T v}$ is maximized when $v = u_1$ (the leading eigenvector of C).

pf sketch

$$v^T C v = v^T \left(\sum_{j=1}^n \lambda_j u_j u_j^T \right) v$$

$$= \sum_{j=1}^n \lambda_j (v^T u_j)(u_j^T v) = \sum_{j=1}^n \lambda_j |v^T u_j|^2$$

$$\leq \lambda_1 \sum_{j=1}^n |v^T u_j|^2 = \lambda_1 \frac{\|v\|^2}{v^T v}$$

$$u_1^T C u_1 = \sum_{j=1}^n \lambda_j |u_1^T u_j|^2 = \lambda_1 \quad \checkmark$$

So u_1 maximizes the variance of $y^{(1)}$, the first principle component of x . The variance of $y^{(1)}$ is λ_1 .

We can use similar principle to "find" the remaining rows of the transformation.

$$u_2 = \underset{\substack{\|v\|=1 \\ v^T u_1 = 0}}{\operatorname{argmax}} v^T C v$$

PS Sketch

$$v^T C v = \sum_{j=1}^n \lambda_j |v^T u_j|^2 = \sum_{j=2}^n \lambda_j |v^T u_j|^2$$

$$\leq \lambda_2 \sum_{j=2}^n |v^T u_j|^2 \leq \lambda_2 (v^T v) = \lambda_2$$

$$u_2^T C u_2 = \sum_{j=1}^n \lambda_j |u_2^T u_j|^2 = \lambda_2$$

$\Rightarrow u_2$ maximizes the variance of $y^{(2)}$, subject to $\|v\|=1$, and $v^T u_1 = 0$.

$$u_k = \underset{\substack{\|v\|=1, v \in V_{k-1}^\perp}}{\operatorname{argmax}} v^T C v \quad V_{k-1}^\perp = \operatorname{span}\{u_1, u_2, \dots, u_{k-1}\}^\perp$$

Courant-Fisher-Weyl Min-Max Principle

We can formulate the eigenvectors themselves as extrema of the Rayleigh-Quotient:

$$\lambda_k = \min_{\substack{v \in M \\ \|v\|=1}} \max_{\substack{v \in M \\ \|v\|=1}} v^T C v \quad (1 \leq k \leq n)$$