

Self-adjoint kernels

Recap

Every H-S ^{kernel} K has a singular value expansion

$$K(x, y) = \sum_{j=1}^{\infty} \sigma_j u_j(x) \overline{v_j(y)}$$

which converges in $L^2(\mathbb{R}^d, \mathbb{R}^d)$.

\Rightarrow Best rank- n approx in $\|\cdot\|$ and $\|\cdot\|_{HS}$ is

$$K_n(x, y) = \sum_{j=1}^n \sigma_j u_j(x) \overline{v_j(y)}$$

$\Rightarrow T = \lambda I + K$ satisfies Fredholm alternative:

a) T^{-1} exists and $\|T^{-1}\| < \infty$, or

b) $N(T)$ is nontrivial and $Tu = f$

has solutions iff $\langle f, v \rangle = 0$ for all $v \in N(T^*)$.

$\Rightarrow K$ is a compact operator, meaning it can be approximated in norm (uniformly on all of $L^2(\mathbb{R}^d)$) by finite-rank operators.

Self-adjoint Kernels

If $K(x, y) = \overline{K(y, x)}$, then $K = K^*$ is self-adjoint.

The eigenvectors/values of K are solutions of

$$K u = \lambda u$$

↑ eigenvalue
↑ eigenvector

\Rightarrow Every self-adjoint compact operator $K: H \rightarrow H$ has an ONB of eigenvectors

\Rightarrow Eigenvalues are real and $\lambda_k \rightarrow 0$ as $k \rightarrow \infty$.

\Rightarrow Eigenvector expansion of Kernel is

$$K(x, y) = \sum_{j=1}^{\infty} \lambda_j u_j(x) \overline{u_j(y)}$$

which coincides with SVE when $\lambda_j \geq 0$, for $j = 1, 2, 3, \dots$, i.e., K is **semidefinite**.

Example: Principle Component Analysis (PCA)

Suppose that $x_1, x_2, x_3, \dots, x_m \in \mathbb{R}^n$ are a sequence of identically and independently distributed random vectors with mean

$$\mu = \mathbb{E}[X] = (\mathbb{E}[x^{(1)}], \dots, \mathbb{E}[x^{(n)}])^T$$

\uparrow expected value of each component \uparrow

and covariance matrix

$$C_{ij} = \mathbb{E}[(x^{(i)} - \mu^{(i)})(x^{(j)} - \mu^{(j)})] \Rightarrow C = \mathbb{E}[(x - \mu)(x - \mu)^T]$$

$\underbrace{\hspace{10em}}$
expected value of rank-1 matrix

Intuition: Think of vectors x_1, \dots, x_m as data collected from m identical, independent experiments. Each entry of x_i represents a particular "feature" measured during the i th experiment and the same "features" are measured across all m experiments.

In practice, one can use empirical estimates of the mean & covariance:

$$\tilde{\mu} = \frac{1}{m} \sum_{i=1}^m \mu_i \quad \left(\text{average features across all experiments} \right)$$

$$\tilde{C} = \frac{1}{m-1} \sum_{i=1}^m (\mu_i - \tilde{\mu})(\mu_i - \tilde{\mu})^T$$

Note that in general, the cov. matrix is not rank 1.

How should we analyze the data? **Principal Component Analysis** follows 2 guiding principles.

- (i) \Rightarrow Find a new coordinate system such that the leading coordinate directions captures as much variance in the data as possible.
- (ii) \Rightarrow Choose the coordinate system so that the features "decouple," i.e., the features "look like" independent variables (more precisely, diff. feat. have zero covariance).

The constraint is that the new coordinate system should be a rotation/reflection of the old coordinate system (features).

$$\text{new features} \rightarrow y_j = \underset{\substack{\uparrow \\ \text{orthogonal} \\ \text{transform}}}{U^T} (x_j - \mu) \leftarrow \text{old features}$$

Guided by principles (i)-(ii), how should we choose the orthogonal transformation U ?

(ii) Can we choose U so that the random variable $y = U^T(x - \mu)$ has diagonal covariance?

$$\begin{aligned} E[yy^T] &= E[U^T(x - \mu)(x - \mu)^T U] \\ \underbrace{\quad}_{\substack{\text{cov. of} \\ \text{r.v. } y}} &= U^T E[(x - \mu)(x - \mu)^T] U \\ &= U^T C U \end{aligned}$$

To diagonalize $\text{cov}(y) = E[yy^T]$, we choose U to be the eigenvector matrix of $C \Rightarrow C = U \Lambda U^T$

This is possible b/c covariance matrices are symmetric positive semidefinite matrices!

$$C_{ij} = \mathbb{E}[(x-u)(x-u)^T] = C_{ji} \quad (\text{symmetric})$$

$$\begin{aligned} v^T C v &= \mathbb{E}[v^T (x-u)(x-u)^T v] \\ &= \mathbb{E}[(v^T (x-u))^2] \geq 0 \quad (\text{Positive semi-def.}) \end{aligned}$$

$\Rightarrow U$ is an orthogonal matrix

$\Rightarrow \Lambda$ is diagonal w/ real non-negative entries

Therefore, y_1, \dots, y_m are distributed w/ mean zero and diagonal covariance matrix:

$$\mathbb{E}[y y^T] = U^T C U = \Lambda \quad \checkmark$$

$$\begin{aligned} \mathbb{E}[y] &= \mathbb{E}[U^T (x-u)] \\ &= U^T \mathbb{E}[x-u] \\ &= 0 \quad \checkmark \end{aligned}$$