Assignment 2

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Here we consider the probability space (Ω, \mathcal{F}, P) , where \mathcal{F} is the σ -field, Ω is the sample space and P is the probability measure.

We will prove the strong law of large numbers. We state the theorem first

Theorem 1. If $\{X_i\}_{i\in\mathbb{N}}$ is an IID sequence and $\mathbb{E}[|X_1|] < \infty$. Then

$$\lim_{n \to \infty} \frac{S_n}{n} = \mathbb{E}\left[X_1\right] \ a.s.,$$

where $S_n = \sum_{i=1}^n X_i$.

We will use the following fact in this proof.

Lemma 1. Given p > 0 and X a non-negative r.v., i.e., $X \ge 0$, we have

$$\mathbb{E}\left[X^{p}\right] = \int_{0}^{\infty} px^{p-1}P(X > x)dx. \tag{1}$$

Proof of Theorem 1:

Proof. We may assume $\mathbb{E}[X_i] = 0$. Otherwise, one can replace X_i with $X_i - \mathbb{E}[X_i]$. This implies that we have to prove that $\frac{1}{n} \sum_{i=1}^n X_i$ converges to 0.

Let

$$Y_n = X_n I_{\{|X_n| \le n\}} \text{ and } Z_n = Y_n - \mathbb{E}[Y_n].$$
(2)

We will break up the proof into parts.

Part 1: We first prove $\lim_{n\to\infty} \frac{1}{n} \sum_{i=1}^n Z_i = 0$. Define

$$M_n = \sum_{i=1}^n \frac{Z_i}{i} \tag{3}$$

with $M_0 = 0$. Define the filtration $\{\mathcal{F}_n\}_{n \in \mathbb{N}}$ where $\mathcal{F}_n := \sigma(X_1, X_2 \dots X_n)$.

Q1: Prove that $\{M_n\}$ is a martingale w.r.t. the filtration \mathcal{F}_n . Now we will show that $\mathbb{E}\left[M_n^2\right]$ is bounded by a constant not dependent on n. For that, consider

$$\mathbb{E}\left[M_n^2\right] = Var(M_n) = \sum_{i=1}^n \frac{Var(Z_i)}{i^2} = \sum_{i=1}^n \frac{1}{i^2} Var(Y_i)$$

$$\leq \sum_{i=1}^n \frac{1}{i^2} \mathbb{E}\left[Y_i^2\right] \leq \sum_{i=1}^\infty \frac{1}{i^2} \int_0^i 2y P(|X_i| \geq y) dy \text{ (from Lemma 1)}$$

$$= 2 \sum_{i=1}^\infty \frac{1}{i^2} \int_0^\infty I_{\{y \leq i\}} y P(|X_i| \geq y) dy$$

$$\cdots$$

$$\leq c \mathbb{E}\left[|X_1|\right], \text{ where } c > 0. \tag{4}$$

Q2: Fill in the missing details (Use the fact that $\{X_i\}$ is IID).

Q3: Claim that $\mathbb{E}\left[M_n^2\right]$ is bounded by a constant not dependent on n using Eq. (4).

Q4: Further claim that the martingale $\{M_n\}$ converges.

Now note that from Eqs. (2) and (3), we have

$$Z_i = i(M_i - M_{i-1}). (5)$$

This implies that

$$\frac{1}{n} \sum_{i=1}^{n} Z_i = \frac{1}{n} \sum_{i=1}^{n} i(M_i - M_{i-1}) = \frac{1}{n} \left(\sum_{i=1}^{n} iM_i - \sum_{i=1}^{n-1} (i+1)M_i \right)
= M_n - \frac{n-1}{n} \left(\frac{1}{n-1} \sum_{i=1}^{n-1} M_i \right)$$
(6)

Q5: Now finally show that $\lim_{n\to\infty} \frac{1}{n} \sum_{i=1}^n Z_i = 0$.

Part 2: We next show that

$$\lim_{n \to \infty} \frac{1}{n} \sum_{i=1}^{n} Y_i = 0. \tag{7}$$

Q6: Using dominated convergence theorem and the fact that X_n are identically distributed show that

$$\lim_{n \to \infty} \mathbb{E}\left[Y_n\right] = \mathbb{E}\left[X_1\right] = 0. \tag{8}$$

This implies that $\lim_{n\to\infty} \frac{1}{n} \sum_{i=1}^n \mathbb{E}[Y_i] = 0$. Now observe that $Y_i = Z_i + \mathbb{E}[Y_i]$.

Q7: Finally, show (7).

Part 3: We next show that

$$\lim_{n \to \infty} \left| \frac{1}{n} \sum_{i=1}^{n} X_i - \frac{1}{n} \sum_{i=1}^{n} Y_i \right| = 0 \text{ a.s.}$$
 (9)

First show that

$$\sum_{i=1}^{\infty} P(|X_i| \ge i) = \sum_{i=1}^{\infty} \int_{i-1}^{i} P(|X_i| \ge i) dx$$

$$\dots$$

$$= \mathbb{E}[|X_1|] < \infty. \tag{10}$$

Q8: Fill the missing part in the above proof. Hint: Use Lemma 1.

(It is important to note that in Eq. (10) we are bounding the summation $\sum_{i=1}^{\infty} P(|X_i| \ge i)$ by the expectation of the modulus of X_1 , the first r.v. of the sequence.)

Q9: Now using Eq. (10) show that

$$\sum_{i=1}^{\infty} P(X_i \neq Y_i) < \infty. \tag{11}$$

Hint: Rewrite the above summation using the summation in Eq. (10).

Therefore by the Borel–Cantelli lemma, except for a set of probability zero, $X_i = Y_i$, for all i greater than some positive integer N (N depends on ω). Henceforth,

Q10: Using Eq. (7) and the above claim (Borel-Cantelli claim), show (9).

Q11: Finally, conclude that the proof is complete. \Box