

- (5) The sample space is

$$\Omega = \{HH, THH, HTH, HTTH, THTH, TTHH, HTTTH, THTTH, \dots\}.$$

I.e. an element of Ω is a sequence in T and H of length k with the last element being H and exactly one of the first $k - 1$ elements being H .

There are exactly $k - 1$ sequences in Ω of length k , corresponding to the $k - 1$ choices for where to put an H in the initial sequence of length $k - 1$. Each length k sequence has probability $(1/2)^k$ of occurring. Hence the probability that exactly k tosses are required is $(k - 1)(1/2)^k$

- (8) We have $\mathbb{P}(\bigcap A_i) = 1 - \mathbb{P}(\bigcup A_i^c) \geq 1 - \sum \mathbb{P}(A_i^c)$ where \cdot^c denotes the complement. Since $\mathbb{P}(A_i^c) = 0$ for each i we have $\mathbb{P}(\bigcap A_i) = 0$, as claimed.
- (11) We have $\mathbb{P}(A^c B^c) = \mathbb{P}((A \cup B)^c) = 1 - \mathbb{P}(A \cup B) = 1 - \mathbb{P}(A) - \mathbb{P}(B) + \mathbb{P}(AB)$. Since A and B are independent this is

$$1 - \mathbb{P}(A) - \mathbb{P}(B) + \mathbb{P}(A)\mathbb{P}(B) = (1 - \mathbb{P}(A))(1 - \mathbb{P}(B)) = \mathbb{P}(A^c)\mathbb{P}(B^c).$$

- (12) The sample space of possible sequences of sides that we can see is $\Omega = \{RR, GG, RG, GR\}$ where R denotes a red side and G denotes a green side. The elements of Ω have probability $1/3$, $1/3$, $1/6$, and $1/6$. Hence the probability that the other side is green, given that the first side we see is green is

$$\frac{1/3}{1/3 + 1/6} = \frac{2}{3}.$$

- (13) (a) The sample space is

$$\Omega = \{HT, TH, HHT, TTH, HHHT, TTTH, HHHHT, TTTTH, \dots\}$$

I.e. an element of Ω is a sequence of length n where the first $n - 1$ elements are H and the last element is T , or vice versa. In particular, for any $n \geq 2$ there are exactly two sequences in Ω of length n .

- (b) The probability of three tosses being required is $2(1/2)^3 = 1/4$ since there are two elements of Ω of length 3, both with probability $(1/2)^3$.
- (15) (a) Order the children from youngest to oldest. Thus the sample space consists of sequences of length three in B and N , B indicating that a child has blue eyes and N indicating that they do not have blue eyes:

$$\Omega = \{NNN, BNN, NBN, NNB, BBN, BNB, NBB, BBB\}.$$

If at least one child has blue eyes then the element of the sample space Ω can be anything except for NNN . The probability of this occurring is

$$3(1/4)(3/4)^2 + 3(1/4)^2(3/4) + (1/4)^3.$$

Here the first term is the probability of one of the sequences BNN, NBN, NNB , the second term is the probability of one of the sequences BBN, BNB, NBB , and

the third term is the probability of BBB . So the probability that at least two children have blue eyes is

$$3(1/4)^2(3/4) + (1/4)^3.$$

The probability that at least two children have blue eyes given that one child has blue eyes is

$$\frac{3(1/4)^2(3/4) + (1/4)^3}{3(1/4)(3/4)^2 + 3(1/4)^2(3/4) + (1/4)^3} = \frac{10}{37}.$$

- (b) If the first child has blue eyes then the sequence is BNN , BBN , BNB , or BBB . The probability for this event is $(1/4)(3/4)^2 + 2(1/4)^2(3/4) + (1/4)^3$. The probability that the first child has blue eyes and at least two children have blue eyes is $2(1/4)^2(3/4) + (1/4)^3$. So the conditional probability is

$$\frac{2(1/4)^2(3/4) + (1/4)^3}{(1/4)(3/4)^2 + 2(1/4)^2(3/4) + (1/4)^3} = \frac{7}{16}.$$

- (19) Let M stand for Mac, W for Windows, L for Linux, and V for virus. By Bayes's Theorem,

$$\mathbb{P}(W|V) = \frac{\mathbb{P}(V|W)\mathbb{P}(W)}{\mathbb{P}(V|M)\mathbb{P}(M) + \mathbb{P}(V|W)\mathbb{P}(W) + \mathbb{P}(V|L)\mathbb{P}(L)}.$$

Substituting yields that $\mathbb{P}(W|V)$ is

$$\frac{(82/100)(50/100)}{(65/100)(30/100) + (82/100)(50/100) + (50/100)(20/100)} = \frac{82}{141} \approx 58.156\%$$

- (20) (a) By Bayes's Theorem:

$$\mathbb{P}(C_i|H) = \frac{\mathbb{P}(H|C_i)\mathbb{P}(C_i)}{\sum_j \mathbb{P}(H|C_j)\mathbb{P}(C_j)} = \frac{p_i(1/5)}{\sum_j p_j(1/5)} = \frac{p_i}{1/4 + 1/2 + 3/4 + 1} = \frac{2p_i}{5}.$$

I.e. the probabilities are

$$0, 1/10, 1/5, 3/10, 2/5.$$

- (b) We have $\mathbb{P}(H_2|H_1) = \mathbb{P}(H_1H_2)/\mathbb{P}(H_1)$ and $\mathbb{P}(H_1) = \sum p_i/5 = 1/2$. We also have

$$\mathbb{P}(H_1H_2) = \sum_{i=1}^5 \mathbb{P}(H_1H_2C_i) = \sum_i \mathbb{P}(H_1H_2|C_i)\mathbb{P}(C_i) = 1/5 \sum_i p_i^2 = 3/8.$$

Therefore

$$\mathbb{P}(H_2|H_1) = \frac{3/8}{1/2} = \frac{3}{4}.$$

- (c) By Bayes's Theorem:

$$\mathbb{P}(C_i|B_4) = \frac{\mathbb{P}(B_4|C_i)\mathbb{P}(C_i)}{\sum_j \mathbb{P}(B_4|C_j)\mathbb{P}(C_j)} = \frac{(1-p_i)^3 p_i(1/5)}{\sum_j (1-p_j)^3 p_j(1/5)} = \frac{128(1-p_i)^3 p_i}{23}.$$

The probabilities are

$$0, 27/46, 8/23, 3/46, 0.$$

- (22) See the Jupyter Notebook 22.ipynb.

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