Non-equilibrium evolution of volatility in origination and extinction explains fat-tailed fluctuations in Phanerozoic biodiversity

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Fluctuations in biodiversity, both large and small, are pervasive through the fossil record, yet we do not understand the processes generating them. Here we extend theory from non-equilibrium statistical physics to describe the previously unaccounted for fat-tailed form of fluctuations in marine invertebrate richness through the Phanerozoic. Using this theory, known as superstatistics, we show that the simple fact of heterogeneous rates or origination and extinction between clades and conserved rates within clades is sufficient to account for this fat-tailed form. We identifiy orders and the families they subsume as the taxonomic level at which clades experience inter-clade heterogeneity and within clade homogeneity of rates. Following superstatistics we would thus posit that orders and families are subsystems in local statistical equilibrium while the entire system is not in equilibrium. The separation of timescales between background origination and extinction within clades compared to the origin of major ecological and evolutionary innovations leading to new orders and families allows within-clade dynamics to reach equilibrium, while between-clade diversification is non-equilibrial. This between clade non-equilibrium accounts for the fat-tailed nature of the system as a whole. The distribution of shifts in diversification dynamics across orders and families is consistent with niche conservatism and pulsed exploration of adaptive landscapes by higher taxa. Compared to other approaches that have used simple birth-death processes, equilibrial dynamics, or non-linear theories from complexity science, super-statistics is superior in its ability to account for both small and extreme fluctuations in the richness of fossil taxa. Its success opens up new research directions to better understand the evolutionary processes leading to the stasis of order- and family-level occupancy in an adaptive land-

scape interrupted by innovations that lead to new orders and families.

1 Introduction

Biodiversity has not remained constant nor followed a simple trajectory through geologic time (*1–5*). Instead, it has been marked by fluctuations in the richness of taxa, both positive in the case of net origination, or negative in the case of net extinction. Major events, such as adaptive radiations and mass extinctions have received special attention (*6, 7*), but fluctuations of all sizes are ubiquitous (*2, 5, 8*) and follow a fat-tailed distribution where large events are more probable compared to, e.g. a Gaussian distribution. Understanding the fat-tailed nature of these fluctuations continues to elude paleobiologists and biodiversity theoreticians.

The fat-tailed distribution of fluctuations in taxon richness inspired earlier researchers to invoke ideas from complex systems with similar distributions. Such ideas include the hypotheses that biological systems self-organize to the brink of critical phase-transitions (*9, 10*), and that environmental perterbations are highly non-linear (*11*). New data and analyses have not, however, supported these hypotheses at the scale of the entire Phanerozoic marine invertebrate fauna (*5, 12*). Other studies have modeled the mean trend in taxon richness as tracking a potentially evolving equilibrium (*2, 13–15*) and yet ignore the role of stochasticity and non-equilibrium dynamics in producing observed patterns (*4, 8, 16–18*). Individual, population, and local ecosystem scale processes that could produce complex dynamics, such as escalatory co-evolutionary interactions (*19*), have not been documented to scale up to global patterns (*20*) and indeed should not be expected to do so (*?*). Thus, we still lack a new hypothesis to describe the striking fat-tailed nature of fluctuations throught the Panerozoic.

Despite the heterogeneity of explanations of Phanerozoic biodiversity, consensus has

emerged on one property of macroevolution: clades experience different rates of morphological evolution, origination and extinction (*2, 3, 21–24*). Here we show that the simple fact of conserved rates within clades and variable rates across clades is sufficient to describe pervasive, fat-tailed fluctuations in taxonomic richness throughout the marine Phanerozoic. This biological mechanism has a precise correspondence to the non-equilibrial theory from statistical physics known as "superstatistics" (*25*) which has been applied across the physical and social sciences (*26, 27*). We leverage this correspondence to explain the distribution of fluctuations in the standing richness of marine invertebrates preserved in the Phanerozoic fossil record. We further show that the specific mathematical form of this superstatistical distribution is consistent with niche conservatism (*28, 29*) and pulsed exploration on an adaptive landscape by higher taxa (*29–31*). We operationally define "adaptive landscape" to mean a clade's set characteristics that influences its macroevolution. Those characteristics could be ecological (e.g. substrate preference (*?, 29*), morphological (e.g. growth form (*?*)), or macroecological (e.g. range size (*32, 33*)).

1.1 Superstatistics of fossil biodiversity

Superstatistics (*25*) proposes that non-equilibrial systems can be decomposed into many local sub-systems, each of which attains a unique dynamic equilibrium. The evolution of these dynamic equilibria across sub-systems occurs more slowly. This separation in time scale allows local systems to reach equilibrium while the system as a whole is not in equilibrium (*25*). In the context of macroevolution we propose that a clade with conserved macroevolutionary rates corresponds to a sub-system in dynamic equilibrial.

In statistical mechanics, local sub-systems can be defined by a simple statistical parameter β often corresponding to inverse temperature. In macroevolutionary "mechanics" we define the β_k of clade k as the inverse variance of fluctuations x_k in the number of genera

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within that clade, i.e. fluctuations in the genus richness. The β_k thus represent the inverse variances, what we term volatilities, of the stationary distribution of a homogeneous origination-extinction processes of genera. Fluctuations from this stationary process will be approximately Gaussian if the clades' diversification dynamics are independent and in local equilibrium (see Supplemental Section S1; (*34, 35*)).

We make the hypothesis of dynamic equilibrium within a clade following MacArthur and Wilson (*36*) in recognition that while the identity and exact number of taxa will fluctuate stochastically from random origination and extinction (taking the place of local immigration and extinction in island biogeography), the overall process determining the number of taxa, and by extension, fluctuations in that number, is in equilibrium. Indeed, the different regions of adaptive space occupied by different clades can be conceptualized as islands with unique dynamic equilibria, albeit with macroevolutionary processes determining the "colonization" of adaptive peaks, as opposed to short timescale biogeographic processes.

The volatility of richness fluctuations will vary across these islands of adaptive space as an emergent trait of a clade. Ultimately volatility emerges from the life histories, ecologies, and evolutionary histories that characterize each clade's occupancy in different regions of an adaptive landscape. We do not attempt to disagnose which characteristics of different regions account for volatility differences, but others have found rates of origination and extinction to depend on larval type (*37*), body plan (*16*), body size (*32*), range size (*32,33*), and substrate preference (*29*) on rates of origination and extinction. Not all of these traits would be considered dimensions of an ecological niche or characteristics of a guild (*?*), but they all point to different strategies that influence a clade's macroevlutionarily success. These characteristics result from interactions between heritable traits and environments, which themselves may semi-heritable (*?*). Thus different regions of adaptive space, and

the clades occupying them, will experience different magnitudes of stochastic fluctuations in taxonomic richness. As clades occasionally split to fill new regions of adaptive space their pulsed diversification determines the non-equilibrium nature of the entire biota.

1.2   Real paleontological data to test superstatistics

To uncover the superstatistical nature of the marine invertebrate Phanerozoic fauna we analyzed the distribution of fluctuations in genus richness (the lowest reliably recorded taxonomic resolution) using the Paleobiology Database (PBDB; `paleobiodb.org`). We corrected these raw data for incomplete sampling and bias using a new approach described in the methods section. Occurrences from the PBDB were matched to 49 standard time-bins all of approximately 11MY duration following previous publicatoins (*5, 13*). Fluctuations in genus richness were calculated as the simple difference between bias-corrected richnesses in adjacent timebins.

To focus attention on the variance of fluctuations we zero-centered each clade's fluctuation distribution. In this way we focus on fluctuations about any possible trend toward net diversificaiton or extinction. Because "equilibrium" in the statistical mechanical sense means a system undergoes coherent, concerted responses to perturbation, the mean trend line (positive or negative) is of less interest than deviations from it. We also note that the distributions of fluctuations for most clades are already very close to a mean of 0 (mean of fluctuation distributions at the family level: 0.038 ± 0.176 SD), and so centering has little influence on clade-specific fluctuation distributions, consistent with the observation that origination is often roughly equal to extinction (*38*).

We define potentially equilibrial sub-systems based on taxonomic hierarchies as a full phylogenetic hypothesis for all marine invertebrates is lacking. Taxa ideally represent groups of organisms that descend from a common ancestor and share similar ecologically

and evolutionary relevant traits (*39–41*). Thus our model assumes that at a given higher taxonomic level, within-taxon fluctuations in richness are driven by equillibrial processes characterized by Gaussian distributions. We further assume that new higher taxa arrise due to the emergence of sufficiently novel traits (be they ecological, morphological, life history, or macroecological) so that those new taxa occupy a new region of an adaptive landscape. We lastly assume that different regions of adaptive space are characterized by different volatilities in origination and extinction.

To evaluate the optimal taxonomic level for sub-system designation, we test our super-statistical theory using taxonomic levels from family to phylum. Additionally, we compare our results to randomized taxonomies and confirm that the observed fit of superstatistical theory is not an artifact of arbitrary classification but instead represents real, biologically relevant diversification processes within and between clades. We find that families and orders conform to the assumptions of our superstatistical model while classes and orders do not.

2   Results

We first evaluate the local equilibria of clades from family level to phylum. We find that family level flucuation distributions are well approximated by Gaussians (Figs. 1 and **??**). Three exemplar family-level dynamics are highlighted in Figure 1 to illustrate how different volatility equilibra express themselves as actual richness timeseries. This Gaussian approximation also largely holds for orders, but classes and phyla increasingly show deiations from Gaussian with greater kurtosis corresponding to more requent outliers at these taxonomic levels.

To predict the super-statistical behavior of the entire marine invertebrate Phanerozoic fauna we must integrate over all possible local equilibria that each clade could experience.
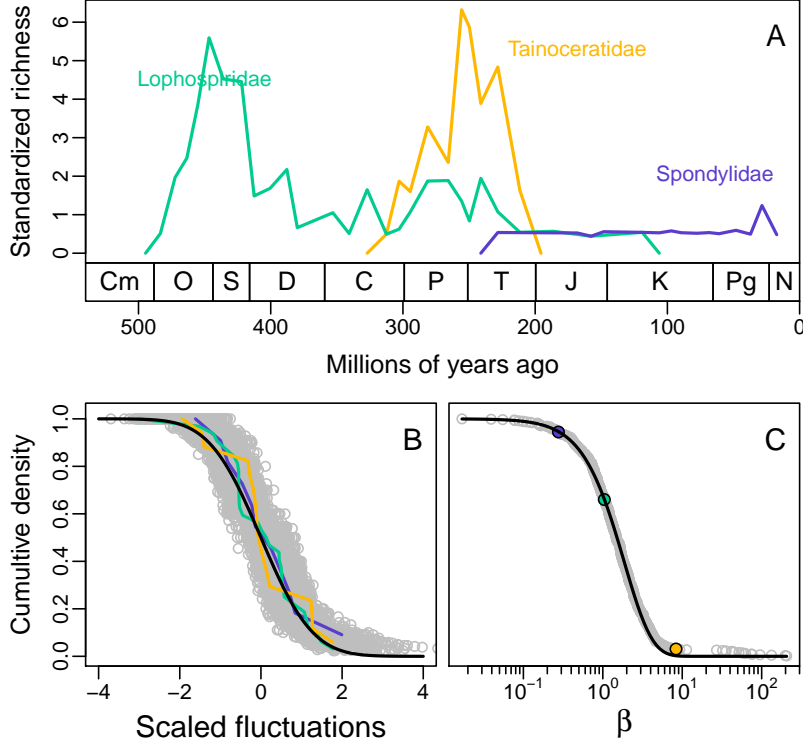
Figure 1: The distributions of within-order fluctuations in genus richness shown for the trajectories of three exemplar orders (A) and shown as an empirical cumulative density aggregated across all orders (B). To display all orders simultaneously we simply collapse their fluctuation distributions by dividing by their standard deviations. If orders conform to the Gaussian hypothesis their scaled fluctuations should fall along the cumulative density line of a normal N(0, 1) distribution, as shown in (B). In (C) the distribution of inverse variances β_k across all orders matches very closely to a Gamma distribution (black line); exemplar orders are again highlighted.

The stationary distribution of β_k values describes these possible equilibria, specifying the probability that a given clade, chosen at random, will occupy a region of adaptive space characterized by β_k.

We estimate the distribution of β_k's simply as the maximum likelihood distribution describing the set of volatilities for all families, orders, classes, or phyla. Phanerozoic marine invertebrate families clearly follow a Gamma distribution in their β_k values (Fig. 1). This Gamma shape also holds for orders and shows increasing deviations again for

classes and especially phyla (Fig. **??**).

Using the observation of within family statistical equilibrium and Gamma-distributed β_k parameters we can calculate, without further adjusting free parameters, the distributions of family-level fluctuations for the entire marine Phanerozoic, $P(x)$, as

$$P(x) = \int_0^\infty p_k(x \mid \beta) f(\beta) d\beta \tag{1}$$

where $p_k(x \mid \beta) = \sqrt{\frac{\beta}{2\pi}} e^{-\frac{\beta x^2}{2}}$ is the distribution of fluctuations within an order and $f(\beta) = \frac{1}{\Gamma(b_1/2)} \left(\frac{b_1}{2b_0}\right)^{b_1/2} \beta^{(b_1/2)-1} \exp\left(-\frac{b_1 \beta}{2b_0}\right)$ is the stationary distribution of volatilities in richness fluctuations. The integral in (1) leads to

$$P(x) = \frac{\Gamma\left(\frac{b_1+1}{2}\right)}{\Gamma\left(\frac{b_1}{2}\right)} \sqrt{\frac{b_0}{\pi b_1}} \left(1 + \frac{b_0 x^2}{b_1}\right)^{-\frac{b_1+1}{2}} \tag{2}$$

This corresponds to a non-Gaussian, fat-tailed prediction for $P(x)$ which closely matches aggregated family level fluctuations in the bias-corrected PBDB (Fig. 2).

To quantitatively evaluate how well the super-statistical prediction matches the family-level data we constructed a 95% confidence envelope from bootstrapped maximum likelihood estimates for $P(x)$. Observed fluctuations fall within this 95% confidence envelope (Fig. 2), indicating that the data do not reject the super-statistical prediction. For further comparison, we fit a Gaussian distribution to the observed fluctuations, which corresponds to the equilibrium hypothesis that all orders conform to the same dynamic. Using Akaike Information Criterion (AIC) we find that observed fluctuations are considerably better explained by the super-statistical prediction than by the Gaussian hypothesis (ΔAIC = 11285.18). Thus, as expected under the superstatistical hypothesis, the fat-tailed distribution of fluctuations arise from the superposition of independent Gaussian statistics of fluctuations within families.

Computing the distribution of aggregated fluctuations using orders also closely matches the observec data (Fig. 2) but as we further coarsen the taxonomy to classes and phyla
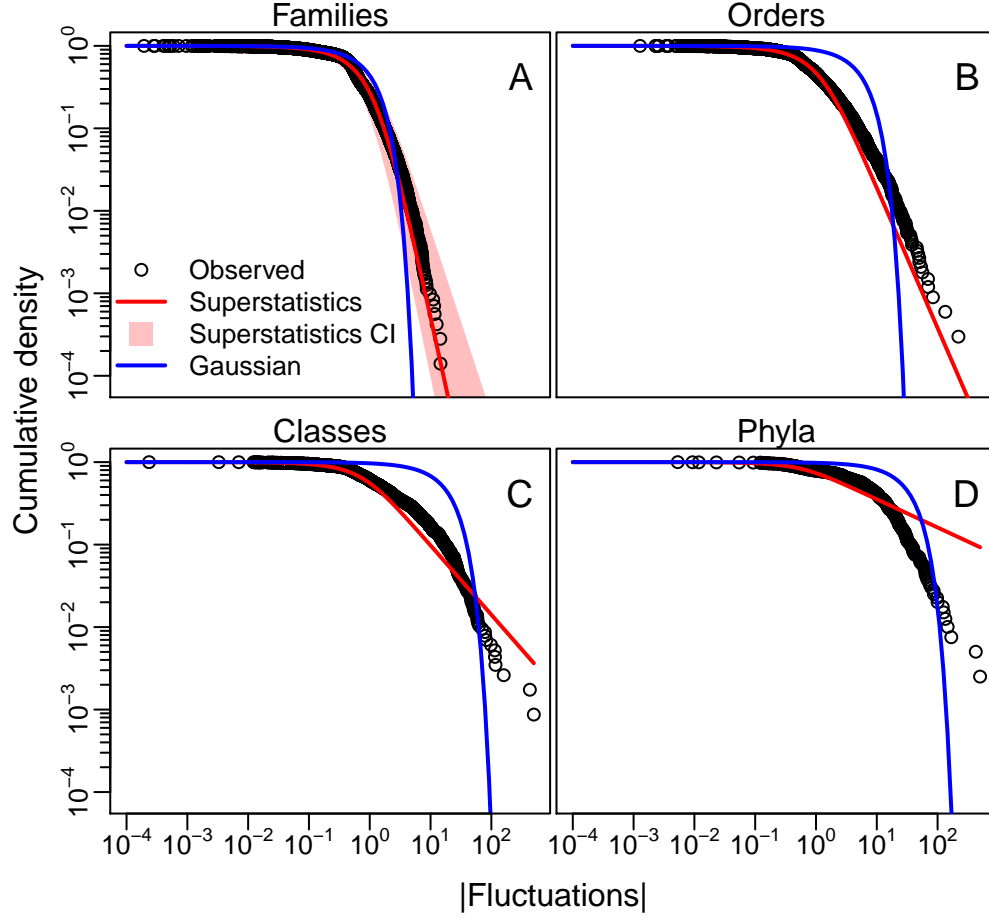
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Figure 2: Distribution of fluctuations in genus richness within orders of marine inverte-brates in the Paleobiology Database (5) after sampling correction. The distribution is fat-tailed as compared to the maximum likelihood estimate of the normal distribution (blue line). At the order level the empirical distribution of fluctuations are well described by our super-statistical approach, both when computed from integrating over the distribution of observed variances (red line) and when fit via maximum likelihood (95% confidence interval; red shading).

170 we see increasingly poorer correspondance between data and theory (Fig. 2). We quantify

171 this change in the goodness of fit with the Kolmogorov-Smirnov statistic (Fig. 3). We

172 can see that both families and orders have low Kolmogorov-Smirnov statistics, and in fact

173 order level designation of equilibrial subsystems performs slightly better than the family

level. Classes are substantially worse and phylo worse yet with the Kolmogorov-Smirnov statistic of phyla being no different than the null randomized taxonomies described below.

However, if super-statistical theory explains the data, this worsening fit with increasing taxonomic scale is expected as the different classes should not represent dynamically equilibrial sub-systems in their fluctuation dynamics. Instead, classes and phyla aggregate increasingly disparate groups of organisms, and thus effectively mix their associated Gaussian fluctuations, meaning that one statistic should no longer be sufficient to describe class-level dynamics. We see this confirmed by the increasing frequency of outlier fluctuations in within class and phylum level fluctuation distributions (Fig. **??**).

Our analysis indicates that orders and families are evolutionarily coherent units with all subsumed taxa sharing key ecological and evolutionary attributes allowing them to reach steady state diversification independently from other clades at global scale. The fact that both orders and families conform to theoretical predictions is consistent with superstatistics. If superstatistics operates at the order level, then the families subsumed by these orders should represent random realizations of their order's stationary $\beta_k^{(order)}$ volatility. The sum of Gamma random variables is still Gamma, but with new parameters, thus the family level distribution of $\beta_k^{(family)}$ is still Gamma.

To further test the evolutionary coherence of families we conducted a permutation experiment in which genera were randomly reassigned to families while maintaining the number of genera in each family. For each permutation, we calculated the super-statistical prediction and its Kolmogorov-Smirnov statistic. The permutation simulates a null model in which common evolutionary history is stripped away (genera are placed in random families) but the total number of observed genera per family is held constant, thus controlling for the distribution of genera per family which could result from arbitrary taxonomic desicions (**?**). Repeating this permutation 500 times yields a null distribution of Kolmogorov-
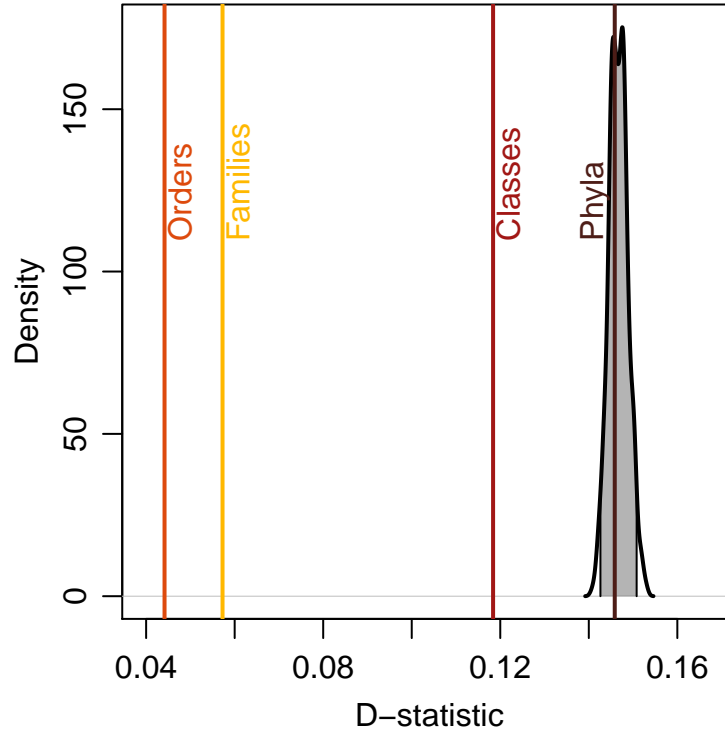
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Figure 3: Distribution of Kolmogorov-Smirnov (KS) statistics from randomly permuting genera within orders (gray shading represents 95% confidence interval). Solid black line is observed KS statistic at the order level, while the dashed black line shows the observed KS statistic at the class level.

Smirnov statistics that is far separated from the observed values at the family and order levels (Fig. 3) suggesting the good fit at these levels is not merely a statistical artifact of classification but carries important biological information. It should also be noted that the width of 95% confidence interval of this null distribution is not far from the distance between the Kolmogorov-Smirnov statistics of orders versus families, suggesting that differences of fit between these taxonomic levels is at least partially accounted for by the randomness of the sampling distribution of Kolmogorov-Smirnov statistics.

3 Discussion

Our analysis makes no assumption that orders and families should correspond to super-statistical subsystems, but identifies them as the appropriate level for marine invertebrates. As we show, orders and the families they subsume differ only in the volatilities of their richness fluctuations (Fig. 1).

Our study is the first to demonstrate that complex patterns in the fluctuation of taxon richness in the fossil record are the result of a simple underlying process analogous to the statistical mechanisms by which complexity emerges in large, non-equilibrium physical (26) and social systems (27). We do so by identifying the biological scale at which clades conform to locally independent dynamic equilibria in fluctuations. Equilibrium could result from many processes, including neutrality (36, 42, 43), diversity-dependence (14, 44, 45) and processes that dampen—rather than exacerbate—fluctuations in complex ecological networks (46). These candidate processes are directly opposed to the instability hypothesis underlying the self-organized criticality hypothesis for paleo biodiversity (9, 10).

We show the distribution describing the evolution to different equilibria between orders is Gamma (Fig. 1). A Gamma distribution, while consistent with multiple processes (e.g., (47)), could result from evolution of diversification rates across an adaptive landscape that promotes niche conservatism and pulsed exploration of niche space. Specifically, if β_k values are associated with a clade's macroevolutionarily-relevant traits, and those traits evolve via Ornstein-Uhlenbeck-like exploration of an adaptive landscape, the resulting stationary distribution of β_k will be Gamma (47, 48). For macroevolutionary rates to vary across an adaptive landscape, this landscape cannot be flat, and thus niche conservatism interrupted by adaptive exploration is likely (31, 49). The specifics of how this adaptive

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landscape is shaped and is traversed by evolving clades determine the exact form of the distribution of β_k volatilities, in the case of the marine Phanerozoic resulting in a Gamma. Our work thus motivates further study of the trait spaces and evolutionary shifts consistent with Gamma-distributed equilibria in richness fluctuation volatilities.

We then show that the pulsed shift to different equilibria between orders and the families they subsume is sufficient to explain the characteristically fat-tailed distribution of richness fluctuations when the marine Phanerozoic fauna is viewed as a whole macrosystem. Armed with an understanding of the statistical origin of this diversification pattern we can explore which models of niche conservatism and pulsed adaptive radiation in are consistent with the statistical behavior of the Phanerozoic. Our statistical theory provides new motivation for identifying the eco-evolutionary causes of innovations between lineages and how those innovations are eventually conserved within lineages. With a theoretical baseline, we can also go on to identify and robustly examine the mechanisms underlying patterns unexplained by our statistical theory. For example, clades wax and wane systematically, and possibly non-symetrically, through time ($4, 8, 33$), a pattern that we cannot explain with superstatistics alone.

Superstatistics could also be applied to other areas of evolution and macroecology. For example new phylogenetic models already consider heterogeneous rates of diversification (e.g., (24)). The superstatistics of clades in adaptive landscapes could motivated models that jointly predict changes in trait and diversification, a research area currently strugelling with model inadequecy (50). This framework could also provide a new paradigm in modeling the distributions of richness, abundance and resource use in non-neutral communities. Non-neutral models in ecology are criticized for their over-parameterization (51), yet a persistent counter argument to neutral theory (42) is the unrealistic assumption of ecological equivalency (52) and poor prediction of real dynam-

255 ics (*53*). If ecosystems are viewed as the super-position of many individualistically evolv-

256 ing clades, each exploiting the environment differently and thus obeying a different set of

257 non-equivalent statistics, then diversity dynamics could be parsimoniously predicted with

258 superstatistics while incorporating real biological information on ecological differences

259 between taxa.

260   Superstatistics is a powerful tool to derive macro-scale predictions from locally fluc-

261 tuating sub-systems whose evolution is driven by interesting, but complex and difficult

262 to model, biological mechanisms. As such, applications of superstatistics from islands to

263 populations to clades are ripe for exploration.

4   Methods and Materials

265 All data processing and analyses were preformed in R (*54*) and all code needed to repro-

266 duce our study are provided, with added explaination, in the supplement.

4.1   Paleobiology Database data download and filtering

268 Data were downloaded from the Paleobiology Database (PBDB; `https://paleobiodb.org`)

269 on 16 November 2018 via the database's API (data retrival and processing script in the

270 supplement). Collections were filtered using the same approach as Alroy (*5*) to insure

271 that only well preserved marine invertebrate occurrences were used in subsequent analy-

272 ses. This filtering resulted in 815,222 unique genus-level occurrences. These were further

273 filtered to exclude those occurrences without family-level taxonomy and those collections

274 with age estimate resolutions outside the 11MY timebins proposed by Alroy (*5*) resulting

275 in 454,033 occurrences. These timebins were compiled from `http://fossilworks.org`

276 with a custum script reproduced in the supplement. The first and last of these time-

277 bins, corresponding to the earliest Cambrian and the latest Cenozoic, were excluded from

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analysis because their sampling completeness (see below) could not be assessed.

4.2   Correcting for imperfect and potentially biased sampling

We use a new and flexible method to correct for known sampling incompleteness and biases in publication-based specimen databases (*5, 13*). Incompleteness is inherent in all biodiversity samples, the fossil record being no exception (*55–58*). This incompleteness is only an issue if it varies across the fossil record, as indeed it does (*55–58*). In addition to variable incompleteness, bias may result from preferential publication of novel taxa (*13*) which exacerbates the difference between poorly-sampled and well-sampled time periods. We therefore develop a simple two-step method: we first correct genus richness for incomplete sampling using the "three-timer" correction (*5*) and then further correct this three-timer richenss estimate by accounting for any correlation between the number of genera and the number of publications in a time period.

The three-timer correction estimates the probability of failure to observe a genus in a given time period p_t as the number of times any genus is recorded before and after that period but not during, divided by the number of genera whoes occurrence histories span the period t. To calculate the sampling-corrected richness $\hat{D}_{kt}$ of a clade k in the time period in question, the observed genera within that clade and time period are divided by $1 - p_t$ and their occurrences summed:

$$\hat{D}_{kt} = \sum_{j \in k} \frac{I_{jt}}{1 - p_t} \tag{3}$$

where $j \in k$ designates genera in clade k and I_{jt} is an indicator equal to 1 if genus j occurs in time period t.

$\hat{D}_{kt}$ is the maximum likelihood estimator of richness in a simple occupancy through time type model assuming binomial sampling (*59*), and in that way mimics other proposed methods for the fossil record (*56, 57*). We avoid parametrically modeling the sampling

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process through time by instead taking a sliding window of timebins from the Cambrian to the Cenozoic. It should be noted that the three-timer correction compares favorably to other similar methods to account for imperfect detection (*60*)

To eliminate further bias due to preferential publication of novel taxa (*13*) we divide the three-timer-corrected number of genera per family per time period by the expected number of genera given publications in that time period. The expected number is calculated by regressing the log-transformed three-timer-corrected number of genera on log-transformed number of publications. There is only a weak trend toward higher richness with more publications (Fig. S1) meaning that the most important correction comes from the three timer correction.

Our new method re-scales each genus occurrence from 0 or 1 (absent or present) to a weighted number continuously ranging between 0 and 1. Because these weighted numbers represent sampling and bias-corrected *occurrences* we can add them arbitrarily, corresponding to the membership of any given genus in any given higher taxonomic group. We must, however, choose a taxonomic level at which to evaluate the relationship between richness and publications; we choose the level of family because this is the most finely resolved option.

We opt not to use subsampling methods (*13, 55, 61*) because these approaches would not be advisable for clades with few genera. However, our new method achieves similar results at the global scale across all clades to subsampling procedures. We directly compare our predicted time series of global fluctuations in genus richness with results derived from rarifaction and shareholder quarum subsampling (SQS) in Figure S2. Our method shows very minor differences with these subsampling-based predictions and any discrepancies do not impact the statistical distribution of fluctuations (Fig. S2).

4.3 Super-statistical methods

We first derive the super-statistical distribution $P(x)$ by fitting Gaussian distributions to clade-level distributions of fluctuations $p_k(x)$, extracting the inverse variances β_k of those $p_k(x)$, testing the best function to describe the distribution of β_k, and then integrating $P(x) = \int_\beta p_k(x|\beta)f(\beta)$. This process allows no free parameters to hone the fit of $P(x)$ to the data. However, each inverse variance must of course be estimated for each clade, making its good fit to data all the more surprising. To do so we use least squares instead of maximum likelihood because the asymmetric fluctuation distributions of small clades were more reliably fit with curve fitting than with maximum likelihood.

We then estimated $P(x)$ directly from the raw data using maximum likelihood to compare the fit of our super-statistical prediction and that of a simple Gaussian distribution using AIC. To calculate a likelihood-based confidence interval on our prediction we bootstrapped the data, subsampling fluctuations with replacement from all families and fit superstatistics using maximum likelihood to the aggregated fluction distribution of each bootstrap replicate.

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Acknowledgments

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⁴²² **Author contributions:** AJR, MAF and PAM designed the study; AJR and MAF pre-

⁴²³   formed the analyses; AJR, MAF and PAM interpreted the results and wrote the

⁴²⁴   manuscript.

⁴²⁵ **Competing interests:** none.

⁴²⁶ **Data and materials availability:** Data are available through the Paleobiology Database

⁴²⁷   (`paleobiodb.org`) and all code needed to interface with the `paleobiodb.org` API,

⁴²⁸   process, clean, and ultimately analyze the data are available online at `github.com/ajrominger/pale`

⁴²⁹   This github repository also hosts the exact download from `paleobiodb.org` used in

⁴³⁰   this analysis. All required scripts are also availible and explained in the Supplement.

Supplementary materials

S1 Limit distribution of a time-averaged homogeneous origination-extinction process

Fossil taxa gain and lose genera according to an origination-extinction process. We assume that most fossil occurrences of a taxon come from the period of its history when it is dominant and in steady state. In a time slice of duration τ during such a period of steady state the latent per capita rates of origination and extinction would be equal (i.e. $\lambda = \mu \equiv \rho$) and the number of origination or extinctions events (call such events Y) each follow an inhomogeneous Poisson process with rate ρN_t where N_t is the number of species or genera in the taxon of interest at time t. Allowing N_t to vary smoothly with time, and recognizing that the sum of Poisson random variables remains Poisson, we arrive at the number Y of extinction *or* origination events in τ being distributed

$$Y \sim \text{Pois}(\rho \int_{t=0}^{\tau} N(t)dt). \tag{4}$$

Under the steady state assumption we can approximate $N(t)$ by $\bar{N}$, the steady state richness, leading to

$$Y \sim \text{Pois}(\rho \bar{N} \tau). \tag{5}$$

This Poisson distribution is asymptotically Gaussian, which is a more appropriate distribution for our sampling and bias-corrected richness estimates because these estimates are not integer-valued but rather continuous random variables. Furthermore, because we ue standard time periods of average duration $\tau = 11\text{MY}$ the the distribution of fluctuations within taxa will be independent of the specific time periods considered. The

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⁴³⁹ Gaussian asymptotics of time-averaged birth-death processes have been proven and explored elsewhere as well (*34, 35*).

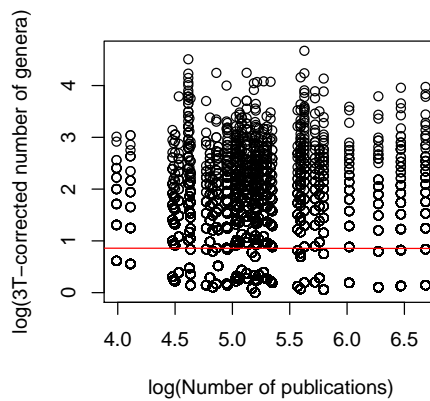## S2   Evaluation of sampling bias correction methods

(*61*)

Figure S1: Relationship between number of publications and genus richness as recorded by the PBDB.

S3   Understanding deviations from superstatistics at higher taxonomic levels
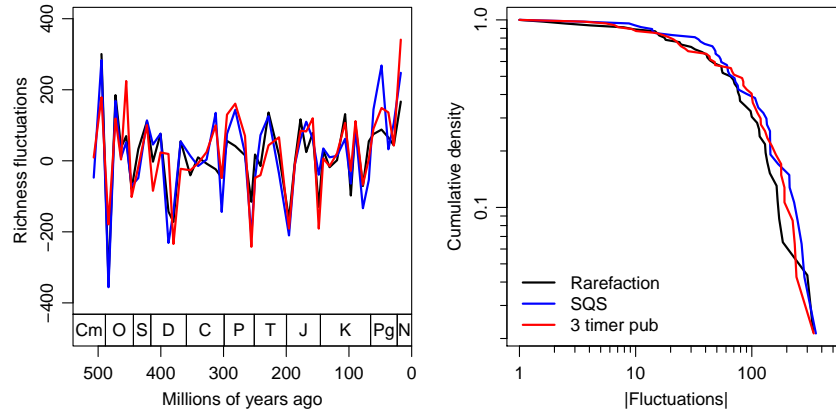
S4   Guild composition of higher taxa

Figure S2: Comparison of SQS method (*13*) (solid black line) with the raw data (dashed black) and our three-timer and publication bias correction method (red). The time-series of all marine invertebrate genera shows general agreement with the only major deviations toward the modern (A). Despite these differences the distribution of fluctuations in genus richness across all marine invertebrates show good agreement (B).