Non-equilibrium evolution of volatility in origination and extinction explains fat-tailed fluctuations in Phanerozoic biodiversity

One sentence summary: Phanerozoic marine invertebrate richness fluctuates in a non-equilibrium process due to pulsed adaptive evolution.

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Fluctuations in biodiversity, both large and small, are pervasive through the fossil record, yet we do not understand the processes generating them. Here we extend theory from non-equilibrium statistical physics to describe the previously unaccounted for fat-tailed form of fluctuations in marine invertebrate richness through the Phanerozoic. Using this theory, known as superstatistics, we show that the simple fact of heterogeneous rates of origination and extinction between clades and conserved rates within clades is sufficient to account for this fat-tailed form. We identify orders and the families they subsume as the taxonomic level at which clades experience inter-clade heterogeneity and within clade homogeneity of rates. Following superstatistics we would thus posit that orders and families are subsystems in local statistical equilibrium while the entire system is not in equilibrium. The separation of timescales between background origination and extinction within clades compared to the origin of major ecological and evolutionary innovations leading to new orders and families allows within-clade dynamics to reach equilibrium, while between-clade diversification is non-equilibrial. This between clade nonequilibrium accounts for the fat-tailed nature of the system as a whole. The distribution of shifts in diversification dynamics across orders and families is consistent with niche conservatism and pulsed exploration of adaptive landscapes by higher taxa. Compared to other approaches that have used simple birth-death processes, simple equilibrial dynamics, or non-linear theories from complexity science, superstatistics is superior in its ability to account for both small and extreme fluctuations in the richness of fossil taxa. Its success opens up new research directions to better understand the evolutionary processes leading to the stasis of order- and family-level occupancy in an adaptive landscape interrupted by innovations that lead to novel forms.

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$_{7}$ 1 Introduction

Biodiversity has not remained constant nor followed a simple trajectory through geologic time (1-5). Instead, it has been marked by fluctuations in the richness of taxa, both positive in the case of net origination, or negative in the case of net extinction. Major events, such as adaptive radiations and mass extinctions have received special attention (6,7), but fluctuations of all sizes are ubiquitous (2,5) and follow a fat-tailed distribution where large events are more probable compared to, e.g. a Gaussian distribution. Understanding the fat-tailed nature of these fluctuations continues to elude paleobiologists and biodiversity theoreticians.

The fat-tailed distribution of fluctuations in taxon richness inspired earlier researchers to invoke ideas from complex systems with similar distributions. Such ideas include the hypotheses that biological systems self-organize to the brink of critical phase-transitions (8,9) and that environmental perturbations are highly non-linear (10). New data and analyses have not, however, supported these hypotheses at the scale of the entire Phanerozoic marine invertebrate fauna (5,11). Other studies have modeled the mean trend in taxon richness as tracking a potentially evolving equilibrium (2,12,13) and yet ignore the role of stochasticity and non-equilibrium dynamics in producing observed patterns (4,14,15). Individual, population, and local ecosystem scale processes that could produce complex dynamics, such as escalatory co-evolutionary interactions (16), have not been documented to scale up to global patterns (17) and indeed should not be expected to do so (18). Thus, we still lack a theory to describe the striking fat-tailed nature of fluctuations throughout the Phanerozoic.

Despite the heterogeneity of explanations of Phanerozoic biodiversity, consensus has emerged on one property of macroevolution: clades experience different rates of morpho-

logical evolution, origination and extinction (2, 3, 19, 20). Here we show that the simple fact of conserved rates within clades and variable rates across clades is sufficient to describe pervasive, fat-tailed fluctuations in taxonomic richness throughout the marine Phanerozoic. This biological mechanism has a precise correspondence to the non-equilibrial theory from statistical physics known as "superstatistics" (21) which has been applied across the physical and social sciences (22, 23). We leverage this correspondence to explain the 56 distribution of fluctuations in the standing richness of marine invertebrates preserved in the Phanerozoic fossil record. We further show that the specific mathematical form of this superstatistical distribution is consistent with niche conservatism (24,25) and pulsed exploration on an adaptive landscape by higher taxa (19, 25-27). We operationally define "adaptive landscape" to mean a clade's set of characteristics, and the fitness they impart to the clade, that influences its macroevolution. Those characteristics could be 62 ecological (e.g. substrate preference (25, 28, 29)), morphological (e.g. body plan (14)), or 63 macroecological (e.g. range size (30, 31)).

5 1.1 Superstatistics of fossil biodiversity

Superstatistics (21) proposes that non-equilibrial systems can be decomposed into many local sub-systems, each of which attains a unique dynamic equilibrium. The evolution of these dynamic equilibria across sub-systems occurs more slowly. This separation in time scales allows local systems to reach equilibrium while the system as a whole is not (21). In the context of macroevolution we propose that a clade with conserved macroevolutionary rates corresponds to a sub-system in dynamic equilibrial.

In statistical mechanics, local sub-systems can be defined by a simple statistical parameter β often corresponding to inverse temperature. In macroevolutionary "mechanics" we define the β_k of clade k as the inverse variance of fluctuations x_k in the number of

genera within that clade, i.e. fluctuations in the genus richness. The β_k thus represent the inverse variances, what we term volatilities, of the origination-extinction processes of genera with clades. The details of this origination-extinction process, e.g. whether it is linear or subject to a carrying capacity, are not of central importance to our analysis; so long as fluctuations can reach a stationary distribution and are observed over time-averaged intervals in a temporally coarse-grained fossil record they will be approximately Gaussian (see Supplemental Section S1; (32)).

We make the hypothesis of dynamic equilibrium within a clade following MacArthur and Wilson (33) in recognition that while the identity and exact number of taxa will fluctuate stochastically from random origination and extinction (taking the place of local immigration and extinction in island biogeography), the overall process determining the number of taxa, and by extension, fluctuations in that number, is in equilibrium. Indeed, the different regions of adaptive space occupied by different clades can be conceptualized as islands with unique dynamic equilibria, albeit with macroevolutionary processes determining the "colonization" of adaptive peaks, as opposed to short timescale biogeographic processes.

The volatility of richness fluctuations will vary across these islands in adaptive space as an emergent trait of a clade resulting from the macroevolutionary fitness of the clade and the shape of the surrounding adaptive landscape. Ultimately, volatility emerges from the life histories, ecologies, and evolutionary histories that drive each clade's macroevolutionary fitness and characterize its occupancy of different regions of an adaptive landscape. We do not attempt to diagnose which characteristics of different regions account for volatility differences, but others have found rates of origination and extinction to depend on larval type (34), body plan (14), body size (30), range size (30,31), and substrate preference (25). Not all of these traits would be considered dimensions of an ecological niche or

characteristics of a guild (28, 29, 35), but they all point to different strategies that influence a clade's macroevolutionary success. These characteristics result from interactions between heritable traits and environments, which themselves may be semi-heritable (36). Thus different regions of adaptive space, and the clades occupying them, will experience different magnitudes of stochastic fluctuations in taxonomic richness. As clades occasionally split to fill new regions of adaptive space their pulsed diversification determines the non-equilibrium nature of the entire biota.

1.2 Real paleontological data to test superstatistics

To uncover the superstatistical nature of the marine invertebrate Phanerozoic fauna we 108 analyzed the distribution of fluctuations in genus richness (the lowest reliably recorded 109 taxonomic resolution) using the Paleobiology Database (PBDB; paleobiodb.org). We 110 corrected these raw data for incomplete sampling and bias using a new approach described 111 in the methods section. Occurrences from the PBDB were matched to 49 standard time 112 bins all of approximately 11MY duration following previous publications (5, 12). Fluctu-113 ations in genus richness were calculated as the simple difference between bias-corrected 114 richnesses in adjacent time bins. 115

To focus attention on the variance of fluctuations we zero-centered each clade's fluctu117 ation distribution. In this way we focus on fluctuations about any possible trend toward
118 net diversification or extinction. Because "equilibrium" in the statistical mechanical sense
119 means a system undergoes coherent, concerted responses to perturbation, the mean trend
120 line (positive or negative) is of less interest than deviations from it. We also note that the
121 distributions of fluctuations for most clades are already very close to a mean of 0 (mean at
122 the family level: 0.038 ± 0.176 SD), and so centering has little influence on clade-specific
123 fluctuation distributions, consistent with the observation that origination is often roughly

equal to extinction (37).

We define potentially equilibrial sub-systems based on taxonomic hierarchies as a full 125 phylogenetic hypothesis for all marine invertebrates is lacking. Taxa ideally represent 126 groups of organisms that descend from a common ancestor and share similar ecologically 127 and evolutionary relevant traits (38,39). Thus our model assumes that at a given higher 128 taxonomic level, within-taxon fluctuations in richness are driven by equilibrial processes 129 characterized by Gaussian distributions. We further assume that new higher taxa arise 130 due to the emergence of sufficiently novel traits (be they ecological, morphological, life 131 history, or macroecological) so that those new taxa occupy a new region of an adaptive 132 landscape. We lastly assume that different regions of adaptive space are characterized by 133 different volatilities in origination and extinction. 134

To evaluate the optimal taxonomic level for sub-system designation, we test our superstatistical theory using taxonomic levels from family to phylum. Additionally, we compare
our results to randomized taxonomies and confirm that the observed fit of superstatistics
is not an artifact of arbitrary classification but instead represents real, biologically relevant diversification processes within and between clades. We find that families and orders
conform to the assumptions of our superstatistical model while classes and phyla do not.

141 2 Results

We first evaluate the local equilibria of clades from family level to phylum. We find that family level fluctuation distributions are well approximated by Gaussians (Figs. 1 and S3). Three exemplar family-level dynamics are highlighted in Figure 1 to illustrate how different volatility equilibria express themselves as actual richness timeseries. This Gaussian approximation also largely holds for orders, but classes and phyla increasingly show deviations from Gaussian with greater kurtosis corresponding to more frequent outliers

at these taxonomic levels (Fig. S3).

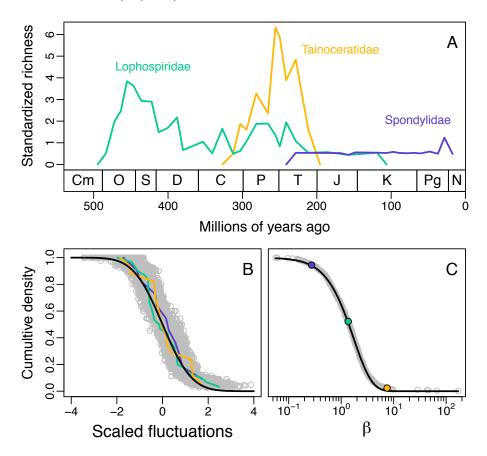


Figure 1: The distributions of within-family fluctuations in genus richness shown for the trajectories of three exemplar families (A) and shown as an empirical cumulative density functions aggregated across all families (B). To display all families simultaneously we simply collapse their fluctuation distributions by dividing by their standard deviations. If families conform to the Gaussian hypothesis their scaled fluctuations should fall along the cumulative density line of a normal N(0, 1) distribution, as shown in (B). We further confirm this normal distribution in the supplement (Fig. S3). In (C) the distribution of inverse variances β_k across all families matches very closely to a Gamma distribution (black line); exemplar families are again highlighted.

To predict the superstatistical behavior of the entire marine invertebrate Phanerozoic fauna we must integrate over all possible local equilibria that each clade could experience. The stationary distribution of β_k values describes these possible equilibria, specifying the probability that a given clade, chosen at random, will occupy a region of adaptive space characterized by β_k .

We estimate the distribution of β_k 's simply as the maximum likelihood distribution describing the set of volatilities for all families, orders, classes, or phyla. Phanerozoic marine invertebrate families clearly follow a Gamma distribution in their β_k values (Fig. 1). The Gamma distribution also holds for orders but shows increasing deviations again for classes and especially phyla (Fig. S4).

Using the observation of within family statistical equilibrium and Gamma-distributed β_k parameters we can calculate, without further adjusting free parameters, the distributions of family-level fluctuations for the entire marine Phanerozoic, P(x), as

$$P(x) = \int_0^\infty p_k(x \mid \beta) f(\beta) d\beta \tag{1}$$

where $p_k(x \mid \beta) = \sqrt{\frac{\beta}{2\pi}} e^{-\frac{\beta x^2}{2}}$ is the distribution of fluctuations within a family and $f(\beta) = \frac{1}{\Gamma(b_1/2)} \left(\frac{b_1}{2b_0}\right)^{b_1/2} \beta^{(b_1/2)-1} \exp\left(-\frac{b_1\beta}{2b_0}\right)$ is the stationary distribution of volatilities in richness fluctuations. The integral in (1) leads to

$$P(x) = \frac{\Gamma\left(\frac{b_1+1}{2}\right)}{\Gamma\left(\frac{b_1}{2}\right)} \sqrt{\frac{b_0}{\pi b_1}} \left(1 + \frac{b_0 x^2}{b_1}\right)^{-\frac{b_1+1}{2}} \tag{2}$$

This corresponds to a non-Gaussian, fat-tailed prediction for P(x) which closely matches aggregated family level fluctuations in the bias-corrected PBDB (Fig. 2).

To quantitatively evaluate how well the superstatistical prediction matches the familyand order-level data we constructed a 95% confidence envelope from bootstrapped maximum
likelihood estimates of P(x). Observed fluctuations for both taxonomic levels fall within
these 95% confidence envelopes (Fig. 2), indicating that the data do not reject the superstatistical prediction. For further comparison, we fit a Gaussian distribution to the
observed fluctuations, which corresponds to the equilibrium hypothesis that all families
conform to the same dynamic. Using Akaike Information Criterion (AIC) we find that

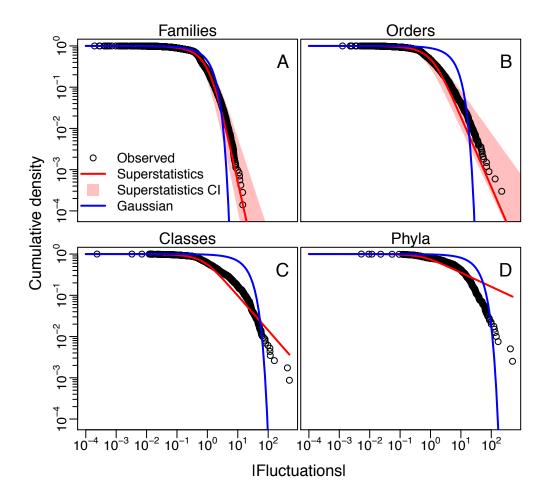


Figure 2: Distribution of fluctuations in genus richness within different taxonomic groupings of marine invertebrates in the Paleobiology Database (5) after sampling correction. The distribution is fat-tailed as compared to the maximum likelihood estimate of the normal distribution (blue line). At the family and order level the empirical distribution of fluctuations are well described by our superstatistical approach, both when computed from integrating over the distribution of observed variances (red line) and when fit via maximum likelihood (95% confidence interval; red shading in (A) and (B)).

observed fluctuations are considerably better explained by the superstatistical prediction than by the Gaussian hypothesis (Δ AIC = 1895.622). Thus, as expected under the superstatistical hypothesis, the fat-tailed distribution of fluctuations arise from the superposition of independent Gaussian statistics of fluctuations within families. Computing the

distribution of aggregated fluctuations using orders also closely matches the observed data (Fig. 2) but as we further coarsen the taxonomy to classes and phyla we see increasingly poorer correspondence between data and theory (Fig. 2).

We quantify this change in the goodness of fit with the Kolmogorov-Smirnov statistic (Fig. 3). We can see that both families and orders have low Kolmogorov-Smirnov
statistics, and in fact order level designation of equilibrial subsystems performs slightly
better than the family level. Classes are substantially worse and phyla worse yet with
the Kolmogorov-Smirnov statistic of phyla being no different than the null randomized
taxonomies described below.

However, if superstatistical theory explains the data, this worsening fit with increas-181 ing taxonomic scale is expected as the different classes and phyla should not represent 182 dynamically equilibrial sub-systems in their fluctuation dynamics. Instead, classes and 183 phyla aggregate increasingly disparate groups of organisms, and thus effectively mix their 184 associated Gaussian fluctuations, meaning that one statistic should no longer be sufficient 185 to describe class- and phylum-level dynamics. We see this confirmed by the increasing 186 frequency of outlier fluctuations in within class and phylum level fluctuation distributions 187 (Fig. S3). We can also see that families and orders represent, on average, 1 to 2 ecospace 188 hypercubes (defined by taxon environment, motility, life habit, vision, diet, reproduction, and ontogeny (28, 29, 35)), respectively. In contrast, classes and phyla represent, on average, 8 to 30 hypercubes, respectively (Fig. S5). 191

Our analysis indicates that orders and families are evolutionarily coherent units with all subsumed taxa sharing key ecological and evolutionary attributes allowing them to reach steady state diversification independently from other clades at global scale. The fact that both orders and families conform to theoretical predictions is consistent with superstatistics. If superstatistics operates at the order level, then the families subsumed by these orders should represent random realizations of their order's stationary $\beta_k^{(order)}$ volatility. The sum of Gamma random variables is still Gamma, but with new parameters, thus the family level distribution of $\beta_k^{(family)}$ is still Gamma.

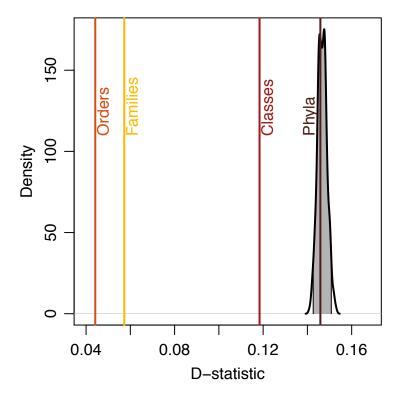


Figure 3: Distribution of Kolmogorov-Smirnov (KS) statistics from randomly permuting genera within families (gray shading represents 95% confidence interval). Solid colored lines are observed KS statistics at different taxonomic levels as indicated.

To further test the evolutionary coherence of families we conducted a permutation experiment in which genera were randomly reassigned to families while maintaining the number of genera in each family. For each permutation, we calculated the superstatistical prediction and its Kolmogorov-Smirnov statistic. The permutation simulates a null model in which common evolutionary history is stripped away (genera are placed in random families) but the total number of observed genera per family is held constant.

Controlling for the total number of genera per family is key because this could be purely

an artifact of an arbitrary taxonomic process (40, 41) and genus richness alone could be solely responsible for differences in the β_k across clades. Indeed, the number of general 208 in a family and that family's β_k value are <u>negatively</u> correlated (Fig. S6). A <u>negative</u> 209 correlation between clade richness and β_k is not unexpected because fluctuations are the 210 sums of the random variables representing genus origination and extinction events; the 211 more of these random variables in the summation (i.e. the more genus richness in a clade) 212 the higher the variance of the summation. Because $\beta_k \equiv 1/\sigma_k^2$ increasing richness should 213 lead to decreasing β_k values. Thus we want to know if this correlation alone accounts for 214 all downstream superstatistical results. The permutation test is specifically designed to 215 determine if the β_k imposed by this correlation with richness are sufficient to explain the 216 observed superstatistical fit. 217

Repeating the null permutation of genera in families 500 times yields a null distribution
of Kolmogorov-Smirnov statistics that is far separated from the observed values at the
family and order levels (Fig. 3) suggesting that the good fit at these levels is not merely a
statistical artifact of classification or the richness of clades, but carries important biological
information. Classes approach the null and phyla are no different. It should also be noted
that the width of 95% confidence interval of this null distribution is not far from the
distance between the Kolmogorov-Smirnov statistics of orders versus families, suggesting
that differences of fit between these taxonomic levels is at least partially accounted for by
the randomness of the sampling distribution of Kolmogorov-Smirnov statistics.

3 Discussion

Our analysis makes no assumption that orders and families should correspond to superstatistical subsystems, but identifies them as the appropriate level for marine invertebrates. Our study is the first to demonstrate that complex patterns in the fluctu-

ation of taxon richness in the fossil record are the result of a simple underlying process analogous to the statistical mechanisms by which complexity emerges in large, non-232 equilibrium physical (22) and social systems (23). We do so by identifying the biological 233 scale at which clades conform to locally independent dynamic equilibria in fluctuations. 234 Equilibrium could result from many processes, including neutrality (33, 42), diversity-235 dependence (43,44) and processes that dampen—rather than exacerbate—fluctuations in 236 complex ecological networks (45). These candidate processes are directly opposed to the 237 presumption of instability underlying the self-organized criticality hypothesis for paleo 238 biodiversity (8, 9). 239

We show that the distribution describing the evolution to different equilibria between 240 orders and families is Gamma (Fig. 1). A Gamma distribution, while consistent with mul-241 tiple processes, could result from evolution of diversification rates across an adaptive land-242 scape that promotes niche conservatism and pulsed exploration of niche space (46). Specif-243 ically, if β_k values are associated with a clade's macroevolutionarily-relevant traits, and those traits evolve via Ornstein-Uhlenbeck-like exploration of an adaptive landscape, the 245 resulting stationary distribution of β_k will be Gamma (46). For macroevolutionary rates to vary in a way consistent with the observed superstatistical description of fluctuations this landscape cannot be flat (i.e. equal fitness everywhere), but instead must be rugged. Thus niche conservatism around local fitness optima in adaptive space interrupted by adaptive exploration is likely (27, 47). The specifics of how this adaptive landscape is 250 shaped and is traversed by evolving clades determine the exact form of the distribution 251 of β_k volatilities, in the case of the marine Phanerozoic resulting in a Gamma distribu-252 tion. Our work thus motivates further study of the trait spaces and evolutionary shifts 253 consistent with Gamma-distributed equilibria in richness fluctuation volatilities. 254

We show that the pulsed shift to different equilibria between orders and the families

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they subsume is sufficient to explain the characteristically fat-tailed distribution of richness fluctuations when the marine Phanerozoic invertebrate fauna is viewed as a whole macrosystem. Armed with an understanding of the statistical origin of this diversification 258 pattern we can explore which models of niche conservatism and pulsed adaptive radiation 259 are consistent with the statistical behavior of the Phanerozoic. Our statistical theory 260 provides new motivation for identifying the eco-evolutionary causes of innovations be-261 tween lineages and how those innovations are eventually conserved within lineages. Using 262 the superstatistical prediction as a theoretical baseline, we can also go on to identify and 263 robustly examine the mechanisms underlying deviations from statistical theory. For exam-264 ple, some clades wax and wane systematically, and possibly non-symmetrically, through 265 time (4, 13, 31), a pattern that we cannot explain with superstatistics alone. 266

Superstatistics could also be applied to other areas of evolution and macroecology. 267 For example new phylogenetic models already consider heterogeneous rates of diversifica-268 tion (e.g., (20)) as expected between different subsystems. The superstatistics of clades 269 in adaptive landscapes could motivate models that jointly predict changes in traits and 270 diversification, a research area currently struggling with model inadequacy (48). This 271 framework could also provide a new paradigm in modeling the distributions of richness, abundance, and resource use in non-neutral communities which can be viewed as emerging from the combination of locally equilibrium subsystems. Non-neutral models in ecology are criticized for their over-parameterization (49), yet a persistent counter argument to 275 neutral theory (42) is the unrealistic assumption of ecological equivalency and poor pre-276 diction of real dynamics (49). If ecosystems are viewed as the superposition of many 277 individualistically evolving clades, each exploiting the environment differently and thus 278 obeying a different set of statistics, then diversity dynamics could be parsimoniously pre-279 dicted with superstatistics while incorporating real biological information on ecological 281 differences between taxa.

Superstatistics is a powerful tool to derive macro-scale predictions from locally fluctuating sub-systems whose evolution is driven by interesting, but complex and difficult to model, biological mechanisms. As such, applications of superstatistics to a wide variety of patterns in ecological and evolutionary systems are ripe for exploration.

²⁸⁶ 4 Methods and Materials

All data processing and analyses were preformed in R (50) and all code needed to reproduce our study are provided, with added explanation, in supplemental Appendix A.

²⁸⁹ 4.1 Paleobiology Database data download and filtering

Data on individual fossil occurrences and the ecospace characteristics of Phanerozoic marine invertebrates were downloaded from the Paleobiology Database (PBDB; 291 https://paleobiodb.org) on 16 November 2018 via the database's API (data retrieval 292 and processing script available in the supplement). Collections were filtered using the same 293 approach as Alroy (5) to insure that only well preserved marine invertebrate occurrences 294 were used in subsequent analyses. This filtering resulted in 815,222 unique genus-level 295 occurrences. These were further filtered to exclude those occurrences without family-level 296 taxonomy and those collections with age estimate resolutions outside the 11MY time 297 bins proposed by Alroy (5) resulting in 454,033 occurrences. Time bins were compiled 298 from http://fossilworks.org with a custom script reproduced in the supplement. The 299 first and last of these time bins, corresponding to the earliest Cambrian and the latest 300 Cenozoic, were excluded from analysis because their sampling completeness (see below) 301 could not be assessed.

4.2 Correcting for imperfect and potentially biased sampling

We use a new and flexible method to correct for known sampling incompleteness and biases in publication-based specimen databases (5, 12). Incompleteness is inherent in all biodiversity samples, the fossil record being no exception (51-54). In addition to 306 incompleteness, bias may result from preferential publication of novel taxa (12) which 307 exacerbates the difference between poorly-sampled and well-sampled time periods. We 308 therefore develop a simple two-step method: we first correct genus richness for incomplete 309 sampling using the "three-timer" correction (5) and then further correct this three-timer 310 richness estimate by accounting for any correlation between the number of genera and the 311 number of publications in a time period. 312

The three-timer correction estimates the probability of failure to observe a genus in a given time period p_t as the number of times any genus is recorded before and after that period but not during, divided by the number of genera whose occurrence histories span the period t. To calculate the sampling-corrected richness \hat{D}_{kt} of a clade k in the time period in question, the observed genera within that clade and time period are divided by $1 - p_t$ and their occurrences summed:

$$\hat{D}_{kt} = \sum_{j \in k} \frac{I_{jt}}{1 - p_t} \tag{3}$$

where $j \in k$ designates genera in clade k and I_{jt} is an indicator equal to 1 if genus j occurs in time period t.

 \hat{D}_{kt} is the maximum likelihood estimator of richness in a simple occupancy through time type model assuming binomial sampling (55), and in that way mimics other proposed methods for the fossil record (52, 53). We avoid parametrically modeling the sampling process through time by instead taking a sliding window of time bins from the Cambrian to the Cenozoic. It should be noted that the three-timer correction compares favorably to other similar methods to account for imperfect detection (56)

To eliminate further bias due to preferential publication of novel taxa (12) we divide the
three-timer-corrected number of genera per family per time period by the expected number
of genera given publications in that time period. The expected number is calculated by
regressing the log-transformed three-timer-corrected number of genera on log-transformed
number of publications. There is only a weak trend toward higher richness with more
publications (Fig. S1) meaning that the most important correction comes from the three
timer correction.

Our new method re-scales each genus occurrence from 0 or 1 (absent or present)
to a weighted number continuously ranging between 0 and 1. Because these weighted
numbers represent sampling and bias-corrected occurrences we can add them arbitrarily,
corresponding to the membership of any given genus in any given higher taxonomic group.
We must, however, choose a taxonomic level at which to evaluate the relationship between
richness and publications; we choose the level of family because this is the most finely
resolved option.

We opt not to use subsampling methods (12, 51, 57) because these approaches would not be advisable for clades with few genera. However, our new method achieves similar results to subsampling procedures at the global scale across all clades. We directly compare our predicted time series of global fluctuations in genus richness with results derived from rarefaction and shareholder quorum subsampling (SQS) in Figure S2. Our method shows very minor differences with these subsampling-based predictions and any discrepancies do not impact the statistical distribution of fluctuations (Fig. S2).

4.3 Superstatistical methods

We first derive the superstatistical distribution P(x) by fitting Gaussian distributions to clade-level distributions of fluctuations $p_k(x)$, extracting the inverse variances β_k of those $p_k(x)$, testing the best function to describe the distribution of β_k , and then integrating $P(x) = \int_{\beta} p_k(x|\beta) f(\beta)$. This process allows no free parameters to hone the fit of P(x) to the data. However, each inverse variance must of course be estimated for each clade, making its good fit to data all the more surprising. To do so we use least squares instead of maximum likelihood because the asymmetric fluctuation distributions of small clades were more reliably fit with curve fitting than with maximum likelihood.

We also estimated P(x) directly from the family-level data using maximum likelihood to compare the fit of our superstatistical prediction and that of a simple Gaussian distribution using AIC. To calculate a likelihood-based confidence interval on our prediction we bootstrapped the data, subsampling fluctuations with replacement from all families and fit superstatistics using maximum likelihood to the aggregated fluctuation distribution of each bootstrap replicate.

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- Data and materials availability: Data are available through the Paleobiology Database

 (paleobiodb.org) and all code needed to interface with the paleobiodb.org API,

 process, clean, and ultimately analyze the data are available online at

 github.com/ajrominger/paleo_supStat. This github repository also hosts the ex
 act download from paleobiodb.org used in this analysis. All required scripts are

 also available and explained in supplemental Appendix A.