

But now let's control for the effect of our co-variate (which we first need to scale and centre)...

```
> cond$Gaming <- scale(cond$Gaming)
> model_ancova <- aov_4(Ability ~ Gaming + Condition + (1 | Participant),
data = cond, factorize = FALSE)
Contrasts set to contr.sum for the following variables: Condition
> anova(model_ancova)
Anova Table (Type 3 tests)
```

Response: Ability

	num	Df	den	Df	MSE	F	ges	Pr(>F)	
Gaming		1		41	0.55171	53.5636	0.56643	5.87e-09	***
Condition		2		41	0.55171	0.8771	0.04103	0.4236	

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

The factor Condition is now not significant with an $F < 1$. However, our covariate *Gaming Frequency* is significant. Adding it means a lot of the variance we previously attributed to our experimental factor is actually explained by our covariate. Note, the F values are calculated using Type III Sum of Squares by the `aov_4()` function - more on that in a bit...

Rather than calculating over the raw means which are:

Water Group = 4.82

Double Espresso Group = 9.02

Single Espresso Group = 6.69

```
> describeBy(cond$Ability, group = cond$Condition)
```

```
Descriptive statistics by group
group: Double Espresso
  vars  n mean   sd median trimmed  mad  min   max range  skew kurtosis   se
X1     1 15 9.02 1.19   9.01   9.07 0.73 6.24 11.07  4.83 -0.26    0.16 0.31
-----
group: Single Espresso
  vars  n mean   sd median trimmed  mad  min   max range  skew kurtosis   se
X1     1 15 6.69 0.98   6.37   6.63 0.78 5.25  8.88  3.63  0.69   -0.53 0.25
-----
group: Water
  vars  n mean   sd median trimmed  mad  min   max range  skew kurtosis   se
X1     1 15 4.82 1.16   4.75   4.82 0.8  2.53  7.09  4.56  0.03   -0.57 0.3
```