

# **AMH3 - Interest Rate Modelling**

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## CHAPTER 1

# Preliminaries

### 1. Introduction to Interest Rate Modelling

There is a one-to-one correspondence between the class  $\mathcal{Q}$  of all probability measures equivalent to  $\mathbb{P}$  and the class  $\Lambda$  of all  $\mathbb{F}$ -adapted (or  $\mathbb{F}$ -predictable) process  $\lambda_t$  satisfying

$$\mathbb{P} \left( \int_0^{T^*} |\lambda_u|^2 du < \infty \right) = 1$$

and

$$\mathbb{E}_{\mathbb{P}} \left( \mathcal{E}_{T^*} \left( \int_0^{\cdot} \lambda_u dW_u \right) \right) = 1$$

Thus our correspondence is

$$\mathcal{Q} \ni \mathbb{P}^\lambda \iff \lambda \in \Lambda.$$

Consequently,

- (i)  $\frac{d\mathbb{Q}}{d\mathbb{P}} = \eta_{T^*}$
- (ii)

$$\begin{aligned} \frac{d\mathbb{Q}}{d\mathbb{P}} \big| \mathcal{F}_t &= \eta_t^{\mathbb{Q}} \\ &= \mathbb{E}_{\mathbb{P}}(\eta_{T^*} \mid \mathcal{F}_t) \\ &= \mathcal{E}_t \left( \int_0^{\cdot} \lambda_u dW_u \right) \end{aligned}$$

**Theorem 1.1** (Abstract Bayes formula). *Let  $\mathbb{Q} \sim \mathbb{P}$  with  $\frac{d\mathbb{Q}}{d\mathbb{P}} = \eta$ . Suppose that  $\mathcal{G} \subset \mathcal{F}$ . We then have*

$$\mathbb{E}_{\mathbb{Q}}(X \mid \mathcal{G}) = \frac{\mathbb{E}_{\mathbb{P}}(\eta X \mid \mathcal{G})}{\mathbb{E}_{\mathbb{P}}(\eta \mid \mathcal{G})}.$$

*Note that is  $\mathcal{G} = \{\emptyset, \Omega\}$  then the formula reduces to*

$$\mathbb{E}_{\mathbb{Q}}(X) = \mathbb{E}_{\mathbb{P}}(\eta X).$$

*If  $\mathbb{Q} \sim \mathbb{P}$  with  $\frac{d\mathbb{Q}}{d\mathbb{P}} \big| \mathcal{F}_t = \eta_t$ , for all  $t \in [0, T^*]$ , then*

$$\mathbb{E}_{\mathbb{Q}}(X \mid \mathcal{F}_t) = \frac{\mathbb{E}_{\mathbb{P}}(\eta_{T^*} X \mid \mathcal{F}_t)}{\mathbb{E}_{\mathbb{P}}(\eta_{T^*} \mid \mathcal{F}_t)}.$$

Hence if  $X$  is  $\mathcal{F}_t$  measurable for some  $T \in [0, T^*]$  then

$$\mathbb{E}_{\mathbb{Q}}(X | \mathcal{F}_t) = \frac{\mathbb{E}_{\mathbb{P}}(\eta_T X | \mathcal{F}_t)}{\eta_t} = \mathbb{E}_{\mathbb{P}}(\eta_t^{-1} \eta_T X | \mathcal{F}_t)$$

**Example 1.2.** If  $\eta_t = \mathcal{E}_t \left( \int_0^t \lambda_u dW_u \right)$ , then

$$\mathbb{E}_{\mathbb{Q}}(X | \mathcal{F}_t) = \mathbb{E}_{\mathbb{P}}(e^{\int_t^T \lambda_u dW_u - \frac{1}{2} \int_0^t |\lambda_u|^2 du} X | \mathcal{F}_t)$$

**Lemma 1.3.** A  $\mathbb{F}$ -adapted and  $\mathbb{Q}$ -integrable process  $M$  is a  $(\mathbb{Q}, \mathbb{F})$ -martingale if and only if the product  $M\eta$  is a  $(\mathbb{P}, \mathbb{F})$ -martingale.

PROOF.  $\mathbb{E}_{\mathbb{Q}}(M_t | \mathcal{F}_s) = M_s$ ,  $s \leq t$ , so

$$M_s = \mathbb{E}_{\mathbb{Q}}(M_t | \mathcal{F}_s) = \frac{\mathbb{E}_{\mathbb{P}}(\eta_t M_t | \mathcal{F}_s)}{\eta_s}$$

□

**Lemma 1.4.** If  $X$  and  $Y$  are two processes of the form

$$dX_t = \alpha_t dt + \beta_t dW_t$$

$$dY_t = \tilde{\alpha}_t dt + \tilde{\beta}_t dW_t$$

then the product satisfies the Itô product formula

$$d(X_t Y_t) = X_t dY_t + Y_t dX_t + d\langle X, Y \rangle_t$$

If  $X$  is of the form  $dX_t = \alpha_t dt + \beta_t dW_t$  and  $f$  is of class  $C^2(\mathbb{R})$ , then the continuous martingale part of  $Y_t = f(X_t)$  is given as

$$\int_0^t f'(X_u) \beta_u dW_u$$

**Proposition 1.5.**

PROOF OF PROPOSITION 1.1. Let  $\mathbb{P}^\lambda$  be equivalent to  $\mathbb{P}$ , so that

$$d\eta_t = \eta_t \lambda_t dW_t$$

and

$$\frac{d\mathbb{P}^\lambda}{d\mathbb{P}} = \eta_t$$

on  $(\Omega, \mathcal{F}_t)$ ,  $t \in [0, T^*]$ .

Define  $B(t, T)$  as follows, for all  $t \in [0, T]$ ,

$$\begin{aligned} B(t, T) &= B_t \mathbb{E}_{\mathbb{P}^\lambda} \left( \frac{1}{B_T} | \mathcal{F}_t \right) \\ &= \mathbb{E}_{\mathbb{P}^\lambda} \left( e^{-\int_t^T r_u du} | \mathcal{F}_t \right) \end{aligned}$$

For i), we simply apply Girsanov's theorem, replacing  $dW_t$  by  $dW_t = dW_t^\lambda - \lambda_t dt$  in the dynamics of  $r$  under  $\mathbb{P}$ .

For ii), we first recall that  $Z(t, T) = \frac{B(t, T)}{B_t}$  is given by

$$Z(t, T) = \mathbb{E}_{\mathbb{P}^\lambda} \left( \frac{1}{B_T} \mid \mathcal{F}_t \right)$$

is a  $(\mathbb{P}^\lambda, \mathbb{F})$ -martingale.

Note that  $\mathbb{F}^\lambda \neq \mathbb{F}$  in general. From Lemma 1.3, we know that  $\eta_t Z(t, T)$  is a  $(\mathbb{P}, \mathbb{F})$ -martingale. Thus applying the predictable representation property, there exists an  $\mathbb{F}$ -adapted process  $\gamma_t$  such that

$$M_t \equiv \eta_t Z(t, T) = Z(0, T) + \int_0^t \gamma_u dW_u$$

for all  $t \in [0, T]$ . Consequently,  $dM_t = \gamma_t dW_t$  and hence

$$dZ(t, T) = d(\eta_t^{-1} M_t) = M_t d\eta_t^{-1} + \eta_t^{-1} dM_t + d\langle \eta^{-1}, M \rangle_t$$

where

$$d\eta_t^{-1} = -\eta_t^{-1} \lambda_t dW_t^\lambda.$$

We obtain

$$\begin{aligned} dZ(t, T) &= \eta_t Z(t, T) (-\eta_t^{-1} \lambda_t dW_t^\lambda) + \eta_t^{-1} \gamma_t (dW_t^\lambda + \lambda_t dt) + (-\eta_t^{-1} \lambda_t \gamma_t) dt \\ &= \eta_t^{-1} (\gamma_t - M_t \lambda_t) dW_t^\lambda \end{aligned}$$

so that

$$dZ(t, T) = \tilde{b}^\lambda(t, T) dW_t^\lambda$$

Since  $B(t, T) = B_t Z(t, T)$ , using again the Itô formula we have

$$\begin{aligned} dB(t, T) &= B_t dZ(t, T) + Z(t, T) dB_t \\ &= \frac{B(t, T)}{B_t} r_t B_t dt + B_t \tilde{b}^\lambda(t, T) dW_t^\lambda \\ &= r_t B(t, T) dt + B(t, T) \underbrace{\frac{B_t \tilde{b}^\lambda(t, T)}{B(t, T)}}_{b^\lambda(t, T)} dW_t^\lambda. \end{aligned}$$

We conclude that for all  $T \in [0, T^*]$ , there exists an  $\mathbb{F}$ -adapted process  $b^\lambda(t, T)$ ,  $t \in [0, T]$  called the volatility of the bond, such that

$$dB(t, T) = B(t, T)(r_t dt + b^\lambda(t, T) dW_t^\lambda).$$

In fact, it does not depend on the choice of  $\lambda$ . For simplicity, we can write  $b(t, T) \equiv b^\lambda(t, T)$ .

The final formula is a special case of the well known result:

$$\begin{aligned} dX_t &= X_t(\alpha_t dt + \beta_t dW_t) \\ \Updownarrow \\ X_t &= X_0 e^{\int_0^t \alpha_u du} \mathcal{E}_t \left( \int_0^t \beta_u dW_u \right) \\ &= X_0 e^{\int_0^t \alpha_u du} e^{\int_0^t \beta_u dW_u - \frac{1}{2} \int_0^t |\beta_u|^2 du} \end{aligned}$$

This completes our proof of Proposition 1.1, under the assumption that  $\frac{1}{B_T}$  is  $\mathbb{P}^\lambda$ -integrable.  $\square$

There are still several issues given this pricing formula.

- (i) How to compute  $b(t, T)$  explicitly in terms of  $\mu$  and  $\sigma$  under the assumptions that

$$dr_t = \mu(r_t, t) dt + \sigma(r_t, t) dW_t$$

and  $\lambda_t = \lambda(r_t, t)$  is the risk premium.

- (ii) How can we calibrate our short-term rate model, meaning that

$$\mathbb{E}_{\mathbb{P}^\lambda} \left( \frac{1}{B_T} \right) = B(0, T) = P(0, T).$$

The issue of pricing bonds is related to solving a backward stochastic differential equation (BSDE). The general form is

$$X_t = X_0 + \int_0^t \mu(X_u, u) du + \int_0^t \xi_u dW_u \quad (\star)$$

where  $\mu : \mathbb{R} \times \mathbb{R}^+ \rightarrow \mathbb{R}$  is some function and  $\xi$  is some  $\mathbb{F}$ -adapted process. We also fix  $T > 0$  and postulate that  $X_T$  is a **known**  $\mathcal{F}_T$ -measurable random variable.

**Definition 1.6.** We say that  $(X, \xi)$  solves the BSDE with terminal condition with terminal condition  $Y$  ( $\mathcal{F}_T$ -measurable) if:

- (i)  $(X, \xi)$  satisfies  $(\star)$ ,
- (ii)  $X_T = Y$ .

This can be extended to cases where  $\mu : \mathbb{R} \times \mathbb{R}^+ \times \Omega \rightarrow \mathbb{R}$  is  $\mathbb{F}$ -adapted.

## CHAPTER 2

### Markovian Models of the Short Rate

Let  $\mathbb{P}^*$  be a martingale measure in the sense that

$$B(t, T) = \mathbb{E}_{\mathbb{P}^*} \left( e^{-\int_t^T r_u du} \mid \mathcal{F}_t \right).$$

In particular,

$$B(0, T) = \mathbb{E}_{\mathbb{P}^*} \left( e^{-\int_0^T r_u du} \right).$$

We postulate that

$$dr_t = \mu(r_t, t) dt + \sigma(r_t, t) dW_t^*, \quad (2.1)$$

where  $W^*$  is a Brownian motion under  $\mathbb{P}^*$ . The filtration  $\mathbb{F}$  is any filtration such that  $W^*$  is a BM with respect to  $\mathbb{F}$ . We assume that (2.1) has a unique (strong) solution.

Then it known that  $r_t$  has the Markov property with respect to  $\mathbb{F}$ , meaning that for any bounded continuous function  $h : \mathbb{R} \rightarrow \mathbb{R}$ ,

$$\mathbb{E}_{\mathbb{P}^*} (h(r_t) \mid \mathcal{F}_s) = \mathbb{E}_{\mathbb{P}^*} (h(r_t) \mid r_s)$$

for all  $s \leq t$ .

Hence

$$\mathbb{E}_{\mathbb{P}^*} \left( e^{-\int_t^T r_u du} \mid \mathcal{F}_t \right) = v(r_t, t, T) = \tilde{v}(r_t, t)$$

suppressing the dependence on  $T$ .

Goals:

- (i) Compute explicitly  $v(r_t, t, T)$  for some classical models
  - (a) Merton's model
  - (b) Vasicek's model
  - (c) CIR model (Bessel process)
 using either the probabilistic approach (martingale measure) or the analytic approach (PDEs).
- (ii) Represent the price of the bond as follows

$$B(t, T) = \exp (m(t, T) - n(t, T)r_t)$$

For a fixed maturity  $T$ ,

$$m(\cdot, T), n(\cdot, T) : [0, T] \rightarrow \mathbb{R}$$



can also be computed using the second method by separating variables in the PDE. Note that  $m(T, T), n(T, T) = 0$ .

- (iii) Compute explicitly the volatility  $b(t, T)$  of the bond by applying the Itô formula to the function  $v(r_t, t, T)$ .
- (iv) Extend the model to the time-inhomogeneous case in order to ensure that  $B(0, T) = P(0, T)$  for all  $T \in [0, T^*]$ .

### 1. Merton's model

Assume

$$r_t = r_0 + at + \sigma W_t^*$$

where  $W^* = W^\lambda$  for some  $\lambda$ . Hence

$$dr_t = a dt + \sigma dW_t^*, \quad r_0 > 0. \quad (2.2)$$

**Note.** The generator of the time homogeneous Markov diffusion can be represented as

$$A_r = a \frac{\partial}{\partial r} + \frac{1}{2} \sigma^2 \frac{\partial^2}{\partial r^2}.$$

**Proposition 2.1.** *The price  $B(t, T)$  is given by*

$$B(t, T) = e^{-r_t(T-t) - \frac{1}{2}a(T-t)^2 + \frac{1}{6}\sigma^2(T-t)^3}. \quad (2.3)$$

Hence

$$dB(t, T) = B(t, T) (r_t dt - \sigma(T-t)dW_t^*).$$

Thus we have the volatility of the bond  $b(t, T) = -\sigma(T-t)$ .

**PROOF.** It is enough to calculate  $B(0, T)$  and then establish the general formula for  $B(t, T)$  using the property that  $r_t$  is a time-homogeneous Markov process, thus

$$B(0, T) = v(r_0, T) \Rightarrow B(t, T) = v(r_t, T-t)$$

Computation of  $B(0, T)$  is done as follows:

$$B(0, T) = \mathbb{E}_{\mathbb{P}^*} \left( e^{-\int_0^T r_u du} \right) = \mathbb{E}_{\mathbb{P}^*} (e^{-\xi_T})$$

where the distribution of  $\xi_T$  can be found explicitly. We argue that

$$\xi_T \sim N \left( r_0 T + \frac{1}{2} a T^2, \frac{1}{3} \sigma^2 T^3 \right)$$

We have

$$\begin{aligned}\xi_T &= \int_0^T r_u du \\ &= \int_0^T (r_0 + au + \sigma W_u^*) du \\ &= \int_0^T (r_u + au) du + \sigma \int_0^T W_u^* du\end{aligned}$$

The rest proceeds quite simply.

We then derive the dynamics of  $B(t, T)$ . By the Itô formula, we have that since  $B(t, T) = v(r_t, t, T)$ , we must have

$$dB(t, T) = r_t B(t, T) dt + b(t, T) B(t, T) dW_t^*.$$

Note that the martingale component comes from

$$\frac{\partial v}{\partial r} dr_t$$

and

$$\frac{\partial}{\partial r} v(r_t, t, T) = -(T - t) v(r_t, t, T)$$

so that

$$\begin{aligned}\frac{\partial}{\partial r} v(r_t, t, T) dr_t &= -(T - t) v(r_t, t, T) (a dt + \sigma dW_t^*) \\ &\sim -\sigma (T - t) B(t, T) dW_t^*\end{aligned}$$

We then obtain the equality  $B(t, T) = -\sigma(T - t)$ . In particular,  $B(t, T) = 0$ .  $\square$

**Exercise 2.2.** Apply the PDE approach to obtain (2.3).

## 2. Vasicek's Model

Consider the dynamics

$$dr_t = (a - br_t) dt + \sigma dW_t^*. \quad (2.4)$$

**Lemma 2.3.** The unique solution to Vasicek's equation is

$$r_t = r_0 e^{-bt} + \frac{a}{b} (1 - e^{-bt}) + \sigma \int_0^t e^{-b(t-u)} dW_u^*. \quad (2.5)$$

**Proposition 2.4.** The bond price in the Vasicek model is given by

$$\begin{aligned}B(t, T) &= \exp(m(t, T) - n(t, T)r_t) \\ n(t, T) &= \frac{1}{b} (1 - e^{-b(T-t)})\end{aligned}$$

and  $m(t, T)$  is also known explicitly.

The volatility of the bond satisfies

$$b(t, T) = -\sigma n(t, T) = -\frac{\sigma}{b} \left(1 - e^{-b(T-t)}\right)$$

and

$$dB(t, T) = B(t, T) (r_t dt - \sigma n(t, T) dW_t^*).$$

**Theorem 2.5** (Stochastic Fubini's theorem). *In the computation above, we obtain the following double integral*

$$\int_0^T \int_0^t e^{-b(t-u)} dW_u^* dt = \frac{1}{b} \int_0^T \left(1 - e^{-b(T-u)}\right) dW_u^*.$$

To obtain this result, we must use the stochastic Fubini theorem

$$\int_0^T \int_0^t f(t, u) dW_u^* dt = \int_0^T \int_u^T f(t, u) dt dW_u^*$$

where  $f$  is a continuous function.

**2.1. PDE Approach to Vasicek's model.** We can either use some known results or provide some simple arguments.

We start by postulating that  $B(t, T) = v(r_t, t, T)$  where  $v \in C^{2,1}(\mathbb{R} \times [0, T^*], \mathbb{R})$ . On the other hand, we may apply the Itô formula and obtain

$$dv(r_t, t, T) = \left( \frac{\partial v}{\partial t} + \mu(r_t, t) \frac{\partial v}{\partial r} + \frac{1}{2} \sigma^2(r_t, t) \frac{\partial^2 v}{\partial r^2} \right) dt + \sigma(r_t, t) \frac{\partial v}{\partial r} dW_t^*.$$

On the other hand, from Proposition 1.5 we have

$$dB(t, T) = dv(r_t, t, T) = r_t v(r_t, t, T) dt + b(t, T) v(r_t, t, T) dW_t^*.$$

This means that

$$\underbrace{\left( \frac{\partial v}{\partial t} + \mu \frac{\partial v}{\partial r} + \frac{1}{2} \sigma^2 \frac{\partial^2 v}{\partial r^2} - r_t v \right)}_{A_t} dt = \underbrace{\left( b(t, T) v - \sigma \frac{\partial v}{\partial r} \right)}_{M_t} dW_t^*.$$

**Lemma 2.6.** *If  $(M_t)_{t \in [0, T^*]}$  is a continuous local martingale and a process of finite variation then  $M_t = M_0$  for  $t \in [0, T^*]$ .*

Since  $r_t$  is a Gaussian process, we note that the unknown function should necessarily satisfy the following pricing PDE for  $v = v(r_t, t, T)$ ,

$$\begin{cases} \frac{\partial v}{\partial t} + \mu \frac{\partial v}{\partial r} + \frac{1}{2} \sigma^2 \frac{\partial^2 v}{\partial r^2} - r_t v = 0 \\ v(r_t, T, T) = h(r_t). \end{cases}$$

For the bond maturing at  $T$ , we set  $h(r) = 1$ .

To solve this PDE in the Vasicek case, we postulate that

$$v(r_t, t, T) = e^{m(t, T) - n(t, T) r_t}$$

and derive a system of two ODEs satisfied by the function  $m$  and  $n$ .

### 3. Valuation of Bond Options

Consider a European call option on a  $U$ -maturity zero-coupon bond with expiry  $T$  and strike  $K$  where  $t \leq T < U$  and  $K > 0$ . The payoff at time  $T$  equals

$$C_T = (B(T, U) - K)^+ = (B(T, U) - KB(T, T))^+$$

We postulate that

$$\begin{aligned} C_t &= B_t \mathbb{E}_{\mathbb{P}^*} (B_T^{-1} C_T | \mathcal{F}_t) \\ &= \mathbb{E}_{\mathbb{P}^*} \left( e^{-\int_t^T r_u du} (v(r_T, T, U) - K)^+ \right) \end{aligned}$$

The idea is to change the martingale measure  $\mathbb{P}^*$  to another probability measure  $\mathbb{Q}$  such that

$$\begin{aligned} C_t &= B(t, T) \mathbb{E}_{\mathbb{Q}} (C_T | \mathcal{F}_t) \\ &= B(t, T) \mathbb{E}_{\mathbb{Q}} \left( (F_t \xi - K)^+ | \mathcal{F}_t \right) \end{aligned}$$

where  $F_t = \frac{B(t, U)}{B(t, T)}$ . The measure  $\mathbb{Q}$  is equivalent to  $\mathbb{P}$  on  $(\Omega, \mathcal{F}_T)$  and it is chosen in such a way such that  $(F_t)_{t \in [0, T]}$  is a  $\mathbb{Q}$ -martingale.

Alternatively, consider a claim  $X = C_T$  maturing at time  $T$ . Then

$$\begin{aligned} C_t &= B_t \mathbb{E}_{\mathbb{P}^*} \left( \frac{C_T}{B_T} | \mathcal{F}_t \right) \\ \Phi_t(X) &= B_t \mathbb{E}_{\mathbb{P}^*} \left( \frac{X}{B_T} | \mathcal{F}_t \right) \end{aligned}$$

**Example 2.7.** In the context of equity options this approach yields the following representation of the price of a call option:

$$C_t = S_t \hat{P}(S_T > K | \mathcal{F}_t) - KB(t, T) \mathbb{P}^*(S_T > K | \mathcal{F}_t)$$

where

$$\begin{aligned} \frac{B_t}{S_t} &\text{ is a } \hat{P}\text{-martingale} \\ \frac{S_t}{B_t} &\text{ is a } P^*\text{-martingale} \end{aligned}$$

If  $B_t = e^{rt}$  (deterministic) then  $\hat{P} = P^*$ .

### 4. The CIR Model

We postulate that

$$dr_t = (a - br_t) dt + \sigma \sqrt{r_t} dW_t^*,$$

where  $a, b\sigma$  are positive constants. Using Yamada-Watanabe theorem, we obtain uniqueness and existence of solutions. A suitable comparison theorem tells us that if  $r_0 > 0$  then  $r_t \geq 0$  for  $t \in [0, T]$ . It is known that the solution  $r$  to the CIR equation is related to the Bessel process. It is known that

- (i)  $B(t, T) = e^{m(t, T) - n(t, T)r_t}$  where  $m$  and  $n$  can be computed explicitly using the PDE approach.
- (ii) The price of a call option can be computed explicitly using the probabilistic approach.

One can prove that

$$C_t = B(t, U)\Phi_1(B(t, U), B(t, T), t, T, U) - KB(t, T)\Phi_2(B(t, U), B(t, T), t, T, U)$$

where  $\Phi_1, \Phi_2$  are given explicitly in terms of the distribution of a Bessel process.

## 5. Calibration

We denote by  $\hat{B}(0, T)$  the market price of a zero coupon bond with maturity  $T$ . We assume that

$$\hat{B}(0, T) = e^{-\int_0^T \hat{f}(0, u) du}$$

where the instantaneous forward rate is a differentiable function such that

$$\hat{f}_T(0, t)$$

exists for  $t \in 0, T$ . In general, we can fit to market data a model of the form

$$dr_t = (a(t) - br_t) dt + \sigma r_t^\beta W_t^\star$$

for  $\beta \in [0, 1]$ .

**Proposition 2.8.** *Let  $\beta = 0$ . Then the model fits the market data if and only if  $a(t) = \hat{f}_T(0, t) + h'(t) + b(\hat{f}(0, t) + h(t))$  where*

$$h(t) = \frac{\sigma^2 (1 - e^{-bt})^2}{2b^2}.$$

It is essential here to assume that the function  $\hat{f}(0, T)$  is differentiable with respect to  $T$ . If we wish to produce a model such that  $f(0, T) = \hat{f}(0, T)$ .

## CHAPTER 3

# The HJM Approach to Modelling Bond Prices

### 1. Introduction

Take as inputs the following objects

- (i)  $(\Omega, \mathbb{F}, \mathbb{P})$ ,  $W$ , a  $d$ -dimensional Brownian motion.
- (ii) The dynamics of a family of processes

$$\{f(t, T), t \in [0, T]\}, T \in [0, T^*]$$

where  $f(\cdot, T)$  is an  $\mathbb{F}$ -adapted process such that

$$df(t, T) = \alpha(t, T) dt + \sigma(t, T) \cdot dW_t$$

with some initial condition  $f(0, \cdot) : [0, T^*] \rightarrow \mathbb{R}$ .

As an output, we obtain the family of bond prices

$$\{B(t, T), t \in [0, T]\}, T \in [0, T^*]$$

given by

$$B(t, T) = \exp \left( - \int_t^T f(t, u) du \right)$$

We must first derive the dynamics of  $B(\cdot, T)$  under  $\mathbb{P}$  for any maturity  $T$  in the following form

$$dB(t, T) = B(t, T) (a(t, T) dt + b(t, T) dW_t^*)$$

where  $a$  and  $b$  are given in terms of  $\alpha$  and  $\beta$ .

Next, we will find out under which assumptions on  $\alpha$  and  $\beta$  the HJM model admits a spot martingale measure  $\mathbb{P}^*$  or equivalently, a forward martingale measure  $\mathbb{P}_{T^*}$ .

By definition,  $\mathbb{P}^*$  is any probability measure on  $(\Omega, \mathcal{F}_{T^*})$  such that  $\mathbb{P}^* \sim \mathbb{P}$  and the processes

$$Z_t = \frac{B(t, T)}{B_t} = \frac{B(t, T)}{\exp \left( \int_0^t f(u, u) du \right)}$$

are  $\mathbb{P}^*$ -(local) martingales. Similarly,  $\mathbb{P}_{T^*} \sim \mathbb{P}$  and the processes

$$F(t, T, T^*) = \frac{B(t, T)}{B(t, T^*)}$$

are  $\mathbb{P}_{T^*}$ -(local) martingales.

**Note.** Let  $F(t, T, U) = \frac{B(t, T)}{B(t, U)}$ .

- (i) If  $U \leq T$ , then  $F(t, T, U)$  is the forward price of a  $T$ -maturity bond for the settlement date at time  $U$ .
- (ii) If  $U \geq T$  then  $F(t, T, U)$  represents the forward rate in the FRA initiated at time  $t$  for the future time interval  $[T, U]$ .

**Definition 3.1** (HJM approach). Assume that

$$df(t, T) = \alpha(t, T) dt + \sigma(t, T) \cdot dW_t$$

with  $W$  a  $d$ -dimensional Brownian motion and

$$\sigma(t, T) \cdot dW_t = \sum_{i=1}^d \sigma^i(t, T) dW_t^i.$$

All processes are specified under  $\mathbb{P}$ .

We define  $B(t, T) = e^{-\int_t^T f(t, u) du}$ .

**Lemma 3.2.** Let  $\alpha^*(t, T) = \int_t^T \alpha(t, u) du$ , and  $\sigma^*(t, T) = \int_t^T \sigma(t, u) du$ . These are  $\mathbb{F}$ -adapted processes.

Then we claim that

$$dB(t, T) = B(t, T) (a(t, T) dt + b(t, T) \cdot dW_t)$$

where

$$\begin{aligned} a(t, T) &= f(t, t) - \alpha^*(t, T) + \frac{1}{2} (\sigma^*(t, T))^2 \\ b(t, T) &= -\sigma^*(t, T). \end{aligned}$$

Let  $Z(t, T) = \frac{B(t, T)}{B_t}$ , with  $B_t = e^{\int_0^t f(u, u) du}$ , so that

$$dZ(t, T) = Z(t, T) \left( \left( \frac{1}{2} (\sigma(t, T))^2 - \alpha^*(t, T) \right) dt - \sigma^*(t, T) \cdot dW_t \right)$$

Under which assumptions on  $\alpha$  and  $\sigma$  does there exists a probability measure  $\mathbb{Q} \sim \mathbb{P}$  on  $(\Omega, \mathcal{F}_{T^*})$  such that  $Z(t, T), t \in [0, T]$  is a  $\mathbb{Q}$ -martingale for every  $T \in [0, T^*]$ .

We can also form process

$$F_B(t, T, T^*) = F(t, T, T^*) = \frac{B(t, T)}{B(t, T^*)}.$$

## 2. Trading Strategies

We first choose  $\tau = \{T_1 < T_2 < \dots < T_k \leq T^*\}$  and take some  $\mathbb{F}$ -adapted process  $\varphi = (\varphi^1, \dots, \varphi^k)$ .  $\tau$  represents the maturities of traded bonds.  $\varphi^i$  represents the number of shares of  $\tau_i$ -maturity bonds.

Then the wealth process of  $(\varphi, \tau)$  equals

$$V_t(\varphi) = \sum_{i=1}^k \varphi_t^i B(t, T_i).$$

**Definition 3.3** (Self-financing). We say that  $\varphi$  is self financing if

$$dV_t(\varphi) = \sum_{i=1}^k \varphi_t^i dB(t, T_i).$$

**Lemma 3.4.**

(i) Let  $V_t^*(\varphi) = \frac{V_t(\varphi)}{B_t}$ . Then  $\varphi$  is self-financing if and only if

$$dV_t^*(\varphi) = \sum_{i=1}^k \varphi_t^i dZ(t, T_i).$$

(ii) Let  $F_v(t, T) = \frac{V_t(\varphi)}{B(t, T)}$  for some  $0 < T \leq T^*$ . Then  $\varphi$  is self-financing if and only if

$$dF_v(t, T) = \sum_{i=1}^k \varphi_t^i d\left(\frac{B(t, T_i)}{B(t, T)}\right) = \sum_{i=1}^k \varphi_t^i dF(t, T_i, T)$$

where we assume  $T \geq T_k$ .

### 3. Martingale Measures

We will first address the issue of existence of the so-called *forward martingale measure*, that is, a martingale measure for processes  $\frac{V_t(\varphi)}{B(t, T^*)}$  or equivalently, a martingale measure for processes

$$F_B(t, T, T^*) = \frac{B(t, T)}{B(t, T^*)}, t \in [0, T], T \in [0, T^*].$$

**Lemma 3.5.** For any  $T \in [0, T^*]$ ,

$$dF_B(t, T, T^*) = F_B(t, T, T^*) (\tilde{a}(t, T) dt + (b(t, T) - b(t, T^*)) dW_t)$$

where

$$\tilde{a}(t, T) = a(t, T) - a(t, T^*) - b(t, T^*) (b(t, T) - b(t, T^*))$$

We denote by  $\hat{\mathbb{P}} = \mathbb{P}^*$  the martingale equivalent to  $\mathbb{P}$  on  $(\Omega, \mathcal{F}_{T^*})$  by

$$\frac{d\hat{\mathbb{P}}}{d\mathbb{P}} = \mathcal{E}_t \left( \int_0^\cdot h_u dW_u^* \right)$$

If  $h$  is such that

$$\mathbb{E} \left( \mathcal{E}_{T^*} \left( \int_0^\cdot h_u dW_u \right) \right) = 1$$



the  $\hat{\mathbb{P}}$  is well defined and we can compute the dynamics of  $F_B(t, T, T^*)$  under  $\hat{P}$  with respect to  $\hat{W}$ , where

$$\hat{W} = W_t - \int_0^t h_u du, t \in [0, T^*]$$

Assume that

$$a(t, T) - a(t, T^*) = (b(t, T^*) - h_t) \cdot (b(t, T) - b(t, T^*)) \quad (3.1)$$

Condition (3.1) in the lecture notes ensures that there is no drift term in the dynamics of  $F_B(t, T, T^*)$  under  $\hat{P}$  for all maturities  $T$ . After some computations, (3.1) can be represented as follows

$$\alpha(t, T) + \sigma(t, T) \left( h_t + \int_T^{T^*} \sigma(t, u) du \right) = 0.$$

Later on we will denote by  $\mathbb{P}_T$  the forward measure for the date  $T$ . Thus  $\hat{P} = \mathbb{P}_{T^*}$ .

**3.1. Spot Martingale Measure.** We know that

$$dZ(t, T) = -Z(t, T) \left( \left( \alpha^*(t, T) - \frac{1}{2} |\sigma^*(t, T)|^2 \right) dt + \sigma^*(t, T) dW_t \right)$$

Now, the conditions for the drift term in  $dZ(t, T)$  disappearing reads

$$\alpha^*(t, T) = \frac{1}{2} |\sigma^*(t, T)|^2 - \sigma^*(t, T) \lambda_t$$

$$\Updownarrow$$

$$\alpha(t, T) = \sigma(t, T) (\sigma^*(t, T) - \lambda_t)$$

The last formula can be seen as a tool for simple derivations of processes of interest interest under the measure  $\mathbb{P}^*$  (setting  $\lambda = 0$ ). We denote

$$W_t^* = W_t - \int_0^t \lambda_u du$$

**3.2. Forward Measure.** We are going to examine the relationship between  $\mathbb{P}^*$  and  $\mathbb{P}_T$  in a general term structure model.

**Note.** Define the following.

$$\begin{aligned} dB(t, T) &= B(t, T) (r_t dt + b(t, T) dW_t^*) \\ d\zeta_t^i &= \zeta_t^i (r_t dt + \sigma_t^i dW_t^*) \end{aligned}$$

By definition,

$$\pi_t(X) = B_t \mathbb{E}_{\mathbb{P}^*} \left( \frac{X}{B_T} \mid \mathcal{F}_t \right)$$

Can we find  $\mathbb{Q}$  such that  $\mathbb{Q} \sim \mathbb{P}^*$  and

$$B_t \mathbb{E}_{\mathbb{P}^*} \left( \frac{X}{B_T} \mid \mathcal{F}_t \right) = B(t, T) \mathbb{E}_{\mathbb{Q}} (X \mid \mathcal{F}_t)$$

for any claim  $X \in \mathcal{F}_T$  where  $B(t, T) = B_t \mathbb{E}_{\mathbb{P}^*} \left( \frac{1}{B_T} \mid \mathcal{F}_t \right)$ . Formally,

$$\mathbb{E}_{\mathbb{Q}} (X \mid \mathcal{F}_t) = \frac{\mathbb{E}_{\mathbb{P}^*} \left( \frac{X}{B_T} \mid \mathcal{F}_t \right)}{\mathbb{E}_{\mathbb{P}^*} \left( \frac{1}{B_T} \mid \mathcal{F}_t \right)}$$

We are guessing that  $\mathbb{Q} \sim \mathbb{P}^*$  with density on  $(\Omega, \mathcal{F}_t)$

$$\frac{d\mathbb{Q}}{d\mathbb{P}^*} = \frac{1}{B(0, T)B_T}, \mathbb{P}^* - a.s.$$

$$\mathbb{E}_{\mathbb{P}^*} \left( \frac{1}{B_T} \right) = B(0, T).$$

**Definition 3.6.** Suppose that  $\mathbb{P}^*$  is a spot martingale measure for our model. Then for any maturity  $T \in [0, T^*]$ , we define the forward martingale measure for the date  $T$  by setting on  $(\Omega, \mathcal{F}_{T^*})$

$$\frac{d\mathbb{P}_T}{d\mathbb{P}^*} = \frac{1}{B(0, T)B_T}, \mathbb{P}^* - a.s.$$

**Proposition 3.7.**

(i)

$$\begin{aligned} \frac{d\mathbb{P}_T}{d\mathbb{P}^*} \mid \mathcal{F}_t &= \mathbb{E}_{\mathbb{P}^*} \left( \frac{d\mathbb{P}_T}{d\mathbb{P}^*} \mid \mathcal{F}_t \right) \\ &= \mathbb{E}_{\mathbb{P}^*} \left( \frac{B_0 B(T, T)}{B(0, T)B_T} \mid \mathcal{F}_t \right) \\ &= \frac{B_0}{B(0, T)} \mathbb{E}_{\mathbb{P}^*} \left( \frac{B(T, T)}{B_T} \mid \mathcal{F}_t \right) \\ &= \frac{B_0}{B(0, T)} \frac{B(t, T)}{B_t}, \mathbb{P}^* - a.s. \end{aligned}$$

Recall that  $\frac{\pi_t(X)}{B_t}$  is a  $\mathbb{P}^*$ -martingale. Similarly,  $\frac{\pi_t(X)}{B(t, T)}$  is a  $\mathbb{P}_T$ -martingale. If  $\eta_t = \frac{d\mathbb{P}_T}{d\mathbb{P}^*} \mid \mathcal{F}_t$  then  $M$  is a  $\mathbb{P}_T$ -martingale if and only if  $M\eta$  is a  $\mathbb{P}^*$ -martingale.

**Exercise 3.8.** If we know that under  $\mathbb{P}$  processes  $\frac{X_t}{Z_t}$  are martingales where  $Z$  is a fixed, positive process and under  $\mathbb{Q}$  process  $\frac{X_t}{Y_t}$  are martingales for a fixed positive process  $Y$  then we can find a density of  $\mathbb{Q}$  with respect to  $\mathbb{P}$  in terms of  $Z$  and  $Y$ .

We consider an arbitrage free model of bond prices and stock prices in which the spot martingale measure  $\mathbb{P}^*$  exists, such that  $\frac{B(t, T)}{B_t}$  and  $\frac{S_t^i}{B_t}$  are  $\mathbb{P}^*$ -martingales.

We do not postulate that our model is complete.

Assume that  $X$  is an attainable claim in this model. We know that the arbitrage price  $\pi_t(X)$  is unique and it can be computed using the risk-neutral valuation

formula

$$\pi_t(X) = B_t \mathbb{E}_{\mathbb{P}^*} \left( \frac{X}{B_T} \mid \mathcal{F}_t \right).$$

**Remark.** How do we find the forward price of  $X$  at the time  $t$  in the forward contract with settlement date  $T$ .

**Definition 3.9** (Forward contract). The forward contract written at time  $t$  on a time  $T$  contingent claim is represented by the time  $T$  contingent claim

$$G_T = X - F_X(t, T)$$

such that

- (i)  $F_X(t, T)$  is an  $\mathcal{F}_t$ -measurable random variable,
- (ii) the arbitrage price at time  $t$  on a contingent claim  $G_T$  equals zero, that is,  
 $\pi_t(G_T) = 0$ .

To compute  $F_X(t, T)$ , we will use the risk-neutral formula

$$\begin{aligned} \pi_t(G_T) &= B_t \mathbb{E}_{\mathbb{P}^*} \left( \frac{G_T}{B_T} \mid \mathcal{F}_t \right) \\ &= B_t \mathbb{E}_{\mathbb{P}^*} \left( \frac{X}{B_T} \mid \mathcal{F}_t \right) - F_X(t, T) B_t \mathbb{E}_{\mathbb{P}^*} \left( \frac{1}{B_T} \mid \mathcal{F}_t \right) \\ &= \pi_t(X) - F_X(t, T) B(t, T) \\ &= 0 \end{aligned}$$

and so

$$F_X(t, T) = \frac{\pi_t(X)}{B(t, T)}.$$

Define

$$\begin{aligned} F_Z(t, T) &= \frac{Z_t}{B(t, T)} & Z_t = S_t \text{ or } B(t, T) \\ F_S(t, T) &= \frac{S_t}{B(t, T)} & \text{forward price of stock } S \\ F_B(t, U, T) &= \frac{B(t, U)}{B(t, T)} & \text{forward price of } U\text{-maturity bond.} \end{aligned}$$

**Definition 3.10** (Forward measure). We assume that  $\mathbb{P}^*$  is given. The corresponding forward measure for the date  $T, T \in [0, T^*]$  is defined by

$$\frac{d\mathbb{P}_T}{d\mathbb{P}^*} = \frac{1}{B(0, T) B_T}, \quad \mathbb{P}^* - a.s.$$

so that

$$\frac{d\mathbb{P}_T}{d\mathbb{P}^*} \mid \mathcal{F}_t = \frac{B_0}{B(0, T)} \frac{B(t, T)}{B_t}$$

for every  $t \in [0, T]$ .

**Lemma 3.11.** Assume that  $W_t^\star$  is a Brownian motion under  $\mathbb{P}^\star$  and

$$dB(t, T) = B(t, T) (r_t dt + b(t, T) dW_t^\star)$$

Then  $\eta_t \equiv \frac{d\mathbb{P}_T}{d\mathbb{P}^\star} |_{\mathcal{F}_t}$  equals

$$\eta_t = \exp \left( \int_0^t b(u, T) dW_u^\star - \frac{1}{2} \int_0^t |b(u, T)|^2 du \right).$$

That is,

$$\eta_t = \mathcal{E}_t \left( \int_0^t b(u, T) dW_u^\star \right). \quad (\star)$$

It then follows that

$$d\eta_t = \eta_t b(t, T) dW_t^\star, \quad \eta_0 = 1.$$

and

$$W_t^T = W_t^\star - \int_0^t b(u, T) du$$

is a Brownian motion under  $\mathbb{P}_T$ .

PROOF. Equation  $(\star)$  follows from

$$\frac{d\mathbb{P}_T}{d\mathbb{P}^\star} |_{\mathcal{F}_t} = \frac{B_0}{B(0, T)} \frac{B(t, T)}{B_t}$$

The corollaries follow from differentiation and Girsanov's theorem, respectively.  $\square$

**Exercise 3.12.** Let  $T \leq U$ . Find the dynamics of the forward price  $F_B(t, U, T)$  under  $\mathbb{P}_T$ . Apply the Itô formula under  $\mathbb{P}^\star$ , use Girsanov's theorem to express the dynamics of  $F_B(t, U, T)$  in terms of  $b(t, T)$ ,  $b(t, U)$  and  $W^T$ . Compute the volatility  $\gamma(t, U, T)$  of  $F_B(t, U, T)$ . Apply the above the the HJM model  $(\alpha(t, T), \sigma(t, T), W)$ .

### 3.3. Applications of forward measures.

- (i) Valuation of contingent claims.
- (ii) Construction of models for market rates.

Application (i) is based on the following equality

$$B_t \mathbb{E}_{\mathbb{P}^\star} \left( \frac{X}{B_T} | \mathcal{F}_t \right) = B(t, T) \mathbb{E}_{\mathbb{P}_T} (X | \mathcal{F}_t).$$

**Lemma 3.13.** If  $X$  is an attainable claim and settles at time  $T$ , then

$$\pi_t(X) = B(t, T) \mathbb{E}_{\mathbb{P}_T} (X | \mathcal{F}_t)$$

3.3.1. *Valuation of claims with maturity  $U \neq T$ .* Assume that  $U \leq T$ . Then the payoff  $X$  at  $U$  is equivalent to the payoff  $Y = \frac{X}{B(U, T)}$  at time  $T$ . Equivalence is understood in the sense that

$$X \text{ at } U \sim Y \text{ at } T \iff \pi_t(X) = \pi_t(Y), t \in [0, U].$$

So

$$\pi_t(X) = B(t, U) \mathbb{E}_{\mathbb{P}_U} (X | \mathcal{F}_t) = \pi_t(Y) = B(t, T) \mathbb{E}_{\mathbb{P}_T} \left( \frac{X}{B(U, T)} | \mathcal{F}_t \right).$$

To establish this equality, observe that for  $t \in [0, U]$ ,

$$\frac{d\mathbb{P}_U}{d\mathbb{P}_T} |_{\mathcal{F}_t} = \frac{\frac{d\mathbb{P}_U}{d\mathbb{P}^*} |_{\mathcal{F}_t}}{\frac{d\mathbb{P}_T}{d\mathbb{P}^*} |_{\mathcal{F}_t}} = \frac{\frac{B_0}{B(0, U)} \frac{B(t, U)}{B_t}}{\frac{B_0}{B(0, T)} \frac{B(t, T)}{B_t}} = \frac{B(0, T)}{B(0, U)} \frac{B(t, U)}{B(t, T)}.$$

We then need only apply the Bayes formula and apply the previous result.

Assume now that  $U \geq T$ . We postulate that  $X$  is  $\mathcal{F}_T$ -measurable. Then the claim  $Y = B(T, U)X$  is equivalent to  $X$ , in the sense that  $\pi_t(X) = \pi_t(Y)$ .

- (i)  $U \leq T$ . Then  $\pi_t(X) = B(t, T) = \mathbb{E}_{\mathbb{P}_T} \left( \frac{X}{B(U, T)} | \mathcal{F}_t \right)$ .
- (ii)  $U \geq T$  and  $X \in \mathcal{F}_T$ . Then  $\pi_t(X) = B(t, T) \mathbb{E}_{\mathbb{P}_T} (B(T, U)X | \mathcal{F}_t)$ .

#### 4. The Gaussian HJM Model

Under  $\mathbb{P}^*$ ,

$$dB(t, T) = B(t, T) (r_t dt - \sigma^*(t, T) dW_t^*) \quad (3.2)$$

where

$$-\sigma^*(t, T) = \int_t^T \sigma(t, u) du = b(t, T). \quad (3.3)$$

Moreover,

$$df(t, T) = \sigma(t, T) \sigma^*(t, T) dt + \sigma(t, T) dW_t^* \quad (3.4)$$

and

$$r_t = f(0, t) + \int_0^t \sigma(u, t) \sigma^*(u, t) du + \int_0^t \sigma(u, t) dW_u^* \quad (3.5)$$

**Remark.** From (3.2) and (3.4), we see that for any fixed  $T$ , processes  $B(t, T)$  and  $f(t, T)$  are continuous semimartingales. In (3.5), we integrate a different process for each  $t$ . Also, as an additional input we take some function  $f(0, t)$ .

Can we then compute  $dr_t$ ? The answer to this question is positive in some special cases.

We now always postulate that  $\sigma(t, T)$  is deterministic. Then we say that we deal with the *Gaussian HJM model* since  $r_t$  has a normal distribution for any  $t \in [0, T^*]$ .

Several examples of the Gaussian HJM model include:

- (i) The Ho-Lee model. We take  $d = 1$  and  $\sigma(t, T) = \sigma$ . Since  $b(t, T) = -\sigma(T - t)$ , it can also be seen as a counterpart to Merton's model.
- (ii) The bond price satisfies under  $\mathbb{P}^*$ ,

$$dB(t, T) = B(t, T) (r_t dt - \sigma(T - t) dW_t^*).$$

The short term rate equals

$$r_t = f(0, t) + \frac{1}{2} \sigma^2 t^2 + \sigma W_t^*,$$

so that

$$dr_t = \underbrace{(f_T(0, t) + \sigma^2 t)}_{a(t)} dt + \sigma dW_t^*.$$

where the function  $a : [0, T^*] \rightarrow \mathbb{R}$  can also be derived if we start from the extended merton model  $dr_t = a(t) dt + \sigma dW_t^*$  and we fit this model to the yield curve  $\mathbb{E}_{\mathbb{P}^*} \left( e^{-\int_0^T r_t dt} \right) = e^{-\int_0^T f(0, t) dt}$ . We also need to show that  $r_0 = f(0, 0)$ .

To solve this problem, we need to assume that  $f_T(0, t)$  exists.

- (iii) Vasicek's model. Take  $d = 1$  and  $\sigma(t, T) = \sigma e^{-b(T-t)}$  where  $\sigma, b$  are positive numbers. Then

$$b(t, T) = -\sigma^*(t, T) = -\frac{\sigma}{b} \left( e^{-b(T-t)} - 1 \right),$$

and other computations are given in the course notes.

## CHAPTER 4

# Valuation of Options in Gaussian Models

### 1. Options on Bonds

Consider any term structure in which at least some bonds are traded. If the short term rate process is given then under  $\mathbb{P}^*$ ,

$$dB(t, T_i) = B(t, T_i) (r_t dt + b(t, T_i) dW_t^*)$$

where  $b(t, T_i)$  is a deterministic function and  $0 < T_1 < \dots < T_m$ . If  $r$  is not explicitly specified then we should focus on the dynamics of the forward prices, for example

$$F_B(t, T_i, T_j) = \frac{B(t, T_i)}{B(t, T_j)}, \quad i = 1, \dots, m$$

under the forward measure  $\mathbb{P}_{T_j}$ .

How do we value and hedge European bond options with maturity  $T$  and the underlying zero coupon bond maturing at  $U > T$ . The payoff at  $T$  equals

$$\begin{aligned} C_T &= (B(T, U) - K)^+ \\ P_T &= (K - B(T, U))^+ \end{aligned}$$

so that

$$C_T - P_T = B(T, U) - K$$

and thus for  $t \in [0, T]$ ,

$$C_t - P_t = B(t, U) - KB(t, T).$$

Instead of computing the expectation under  $\mathbb{P}^*$ ,

$$C_t = B_t \mathbb{E}_{\mathbb{P}^*} \left( \frac{C_T}{B_T} \mid \mathcal{F}_t \right),$$

we will compute the equivalent probability measure  $P_T$

$$C_t = B(t, T) \mathbb{E}_{\mathbb{P}_T} (C_T \mid \mathcal{F}_t).$$

Let  $D = \{B(T, U) > K \in \mathcal{F}_T\}$ . Then

$$C_T = B(T, U) \mathbf{1}_D - K \mathbf{1}_D = X_1 - X_2$$

So that

$$C_t = \pi_t(X_1) - \pi_t(X_2) = I_1 - I_2.$$

For  $I_2$ , we compute

$$I_2 = \pi_t(K\mathbf{1}_D) = KB(t, T)\mathbb{P}_T(D | \mathcal{F}_t).$$

We observe that

$$B(T, U) = \frac{B(T, U)}{B(T, T)} = F_B(T, U, T)$$

where under  $\mathbb{P}_T$  the forward price  $F_B(t, U, T)$ ,  $[t \in [0, T]]$  satisfies

$$dF_B(t, U, T) = F_B(t, U, T) (b(t, U) - b(t, T)) dW_t^T$$

so that  $F_t = F_B(t, U, T)$  satisfies

$$F_T = F_t \exp\left(\zeta(t, T) - \frac{1}{2}v^2(t, T)\right)$$

where

$$\zeta(t, T) = \int_0^T \gamma(u, U, T) dW_u^T, \quad v^2(t, T) = \int_t^T |\gamma(u, U, T)|^2 du$$

where  $\gamma(u, U, T) = b(u, U) - b(u, T)$ .

We need to compute

$$\begin{aligned} \mathbb{P}_T(D | \mathcal{F}_t) &= P_T(B(T, U) > K | \mathcal{F}_t) \\ &= \mathbb{P}_T(F_B(T, U, T) > K | \mathcal{F}_t) \\ &= \mathbb{P}_T\left(F_t e^{\zeta(t, T) - \frac{1}{2}v^2(t, T)} | \mathcal{F}_t\right), \end{aligned}$$

where  $\zeta(t, T)$  is independent of  $\mathcal{F}_t$  and  $\zeta(t, T) \sim N(0, v^2(t, T))$ . Hence

$$\begin{aligned} \mathbb{P}_T(D | \mathcal{F}_t) &= \mathbb{P}_T\left(F e^{\zeta(t, T) - \frac{1}{2}v^2(t, T)} | F = F_t\right) \\ &= \mathbb{P}_T\left(\frac{\zeta(t, T)}{v(t, T)} > \ln \frac{K}{F} + \frac{1}{2}v^2(t, T) | F = F_t\right) \\ &= N(\tilde{d}_-(F_t, t, T)) \end{aligned}$$

where  $\tilde{d}_2(F_t, t, T) = \frac{\ln \frac{F}{K} \pm \frac{1}{2}v^2(t, T)}{v(t, T)}$ .

For  $I_1$ , we need to compute the conditional expectation

$$I_1 = B(t, T)\mathbb{E}_{\mathbb{P}_T}(B(T, U)\mathbf{1}_D | \mathcal{F}_t)$$

where

$$\frac{B(T, U)}{C} = \frac{F_B(T, U, T)}{C} = \frac{d\tilde{\mathbb{P}}_T}{d\mathbb{P}_T}.$$



so that

$$\begin{aligned} \frac{d\tilde{\mathbb{P}}_T}{d\mathbb{P}_T} \big| \mathcal{F}_t &= \frac{F_B(t, U, T)}{C} \\ &= \exp \left( \int_0^t \gamma(u, U, T) dW_u^T - \frac{1}{2} \int_0^t |\gamma(u, U, T)|^2 du \right) \\ &= \tilde{\eta}_t \end{aligned}$$

for  $t \in [0, T]$ . Note also that

$$\begin{aligned} \mathbb{E}_{\tilde{\mathbb{P}}_T}(X \mid \mathcal{F}_t) &= \frac{\mathbb{E}_{\mathbb{P}_T}(X \tilde{\eta}_t \mid \mathcal{F}_t)}{\tilde{\eta}_t} \\ \frac{F_B(t, U, T)}{c} \mathbb{E}_{\tilde{\mathbb{P}}_T}(\mathbf{1}_D \mid \mathcal{F}_t) &= \mathbb{E}_{\mathbb{P}_T} \left( \mathbf{1}_D \frac{B(T, U)}{C} \mid \mathcal{F}_t \right) \end{aligned}$$

and

$$\mathbb{E}_{\mathbb{P}_T}(B(T, U) \mathbf{1}_D \mid \mathcal{F}_t) = \frac{B(t, U)}{B(t, T)} \tilde{P}_T(D \mid \mathcal{F}_t)$$

and thus

$$I_1 = B(t, U) \tilde{P}_T(D \mid \mathcal{F}_t)$$

and since  $dF_t = F_t \gamma(t, U, T) dW_t^T$  and

$$\tilde{W}_t^T - \int_0^t \gamma(u, U, T) du$$

is a  $\tilde{P}_T$ -Brownian motion, we obtain

$$dF_t = F_t \left( |\gamma(t, U, T)|^2 dt + \gamma(t, U, T) d\tilde{W}_t^T \right)$$

under  $\tilde{P}_T$ . Solving this equation, we obtain

$$F_T = F_t \exp \left( \int_t^T \gamma(u, U, T) d\tilde{W}_u^T + \frac{1}{2} \int_t^T |\gamma(u, U, T)|^2 du \right).$$

and so

$$\tilde{\mathbb{P}}_T(D \mid \mathcal{F}_t) = N(\tilde{d}_+(F_t, t, T)).$$

We conclude that

$$\begin{aligned} I_1 &= B(t, U) N(\tilde{d}_+(F_t, t, T)), \\ I_2 &= KB(t, T) N(d_-(F_t, t, T)). \end{aligned}$$

so that the price of the call bond option is now known explicitly. It remains to find out whether the call option can be replicated, for instance, by a trading strategy

$\varphi = (\varphi^1, \varphi^2)$  with the wealth process  $V(\varphi)$ ,

$$\begin{aligned} V_t(\varphi) &= \varphi_t^1 B(t, U) + \varphi_t^2 B(t, T) \\ dV_t(\varphi) &= \varphi_t^1 dB(t, U) + \varphi_t^2 dB(t, T) \\ V_T(\varphi) &= C_T = (B(T, U) - K)^+. \end{aligned}$$

Equivalently,

$$\begin{aligned} \frac{V_t(\varphi)}{B(t, T)} &= \varphi_t^1 F_B(t, U, T) + \varphi_t^2 \\ d\left(\frac{V_t(\varphi)}{B(t, T)}\right) &= \varphi_t^1 dF_B(t, U, T) \\ \frac{V_T(\varphi)}{B(T, T)} &= (F_B(T, U, T) - K)^+. \end{aligned}$$

Let  $F_V(t, T) = \frac{V_t(\varphi)}{B(t, T)}$ . Then we need to solve the following problem

$$\begin{aligned} dF_V(t, T) &= \varphi_t^1 dF_B(t, U, T) \\ F_V(T, T) &= (F_B(T, U, T) - K)^+ \end{aligned}$$

where

$$dF_B(t, U, T) = F_B(t, U, T) \gamma(t, U, T) dW_t^T.$$

To solve this equation, observe that

$$\frac{C_t}{B(t, T)} = \frac{B(t, U)}{B(t, T)} \left( N(\tilde{d}_+(F_t, t, T)) - KN(\tilde{d}_-(F_t, t, T)) \right),$$

and

$$F_C(t, T) = F_t \left( N(\tilde{d}_+(F_t, t, T)) - KN(\tilde{d}_-(F_t, t, T)) \right).$$

**Lemma 4.1.** *Let  $(Y_t)$  be given by*

$$\begin{aligned} Y_t &= X_t \left( N(\tilde{d}_+(X_t, t, T)) - KN(\tilde{d}_-(X_t, t, T)) \right) \\ dX_t &= X_t \sigma(t) dW_t \\ \tilde{d}_\pm(x, t, T) &= \frac{\ln \frac{x}{K} \pm 2}{v(t, T)}. \end{aligned}$$

Then

$$dY_t = N(\tilde{d}_-(X_t, t, T)) dX_t.$$

PROOF. Apply the Itô formula. Assume here that  $\sigma$  is deterministic.  $\square$

If we apply the lemma to  $F_c(t, T)$ , we obtain

$$\begin{aligned} dF_c(t, T) &= N(\tilde{d}_+(F_t, t, T)) dF_t \\ &= \varphi_t^1 dF_t. \end{aligned}$$

so that

$$\varphi_t^1 = N(\tilde{d}_1(F_t, t, T))$$

and

$$\varphi_t^2 = \frac{C_t - \varphi_t^1 B(t, U)}{B(t, T)}$$

Then

$$\begin{aligned} V_t(\varphi) &= C_t = \varphi_t^1 B(t, U) + \varphi_t^2 B(t, T). \\ dV_t(\varphi) &= dC_t = \varphi_t^1 dB(t, U) + \varphi^2 - tdB(t, T). \end{aligned}$$

In the future, we will deal with more general options of the form

$$C_T = (Z_T^1 - K Z_T^2)^+$$

where  $Z^i$  is some portfolio of bonds. Then the choice of a natural hedging strategy depends on the choice of traded assets.

**Lemma 4.2.** *The price  $C_t$  of a call option equals*

$$C_t = B(t, U)\mathbb{P}_U(D | \mathcal{F}_t) - KB(t, T)\mathbb{P}_T(D | \mathcal{F}_t)$$

PROOF.

$$\begin{aligned} C_T &= B(T, U)\mathbf{1}_D - K\mathbf{1}_D = X_1 - X_2 \\ \pi_t(X_2) &= B(t, T)\mathbb{E}_{\mathbb{P}_T} K\mathbf{1}_D | \mathcal{F}_t = KB(t, T)\mathbb{P}_T(D | \mathcal{F}_t) \end{aligned}$$

and  $X_1 = B(T, U)\mathbf{1}_D$  is equivalent to  $Y_1 = \mathbf{1}_D$  at time  $U$ , so that

$$\pi_t(X_1) = \pi_t(Y_1) = B(t, U)\mathbb{P}_U(D | \mathcal{F}_t)$$

for  $t \in [0, T]$ . □

## 2. Options on Coupon Bonds

Let  $T_1 < T_2 < \dots < T_n \leq T^*$  be coupon dates and  $c_1, \dots, c_n$  the corresponding deterministic coupons. Then the price  $Z_t = B_c(t, T)$  of the coupon bond equals

$$Z_t = \sum_{j=1}^n c_j B(t, T_j).$$

We consider the call option with maturity  $T < T_1$  and the payoff

$$C_T = (Z_T - K)^+ = \sum_{j=1}^m c_j B(t, T_j)\mathbf{1}_D - K\mathbf{1}_D$$

where

$$D = \{Z_T > K\}.$$

One possible way of pricing this is to represent  $C_t$  as follows:

$$C_t = \sum_{j=1}^n c_j B(t, T_j) \mathbb{P}_{T_j}(D | \mathcal{F}_t) - KB(t, T) P_T(D | \mathcal{F}_t).$$

**Remark** (On the proof of Proposition 4.3). We know that if we set  $D = \{Z_T > K\}$ , then

$$C_t = \sum_{j=1}^m c_j B(t, T_j) \mathbb{P}_{T_j}(D | \mathcal{F}_t) - KB(t, T) \mathbb{P}_T(D | \mathcal{F}_t).$$

For simplicity, we may set  $t = 0$  - then we need to compute  $\mathbb{P}_{T_j}(D)$  and  $\mathbb{P}_T(D)$ . Recall that  $T = T_0 < T_1 \cdots < T_m$ . Then

$$D = \left\{ \sum_{j=1}^m c_j \underbrace{F_B(T, T, T_j)}_{F_B^j(T)} > K \right\}$$

where  $dF_B^j(t) = F_B^j(t)(b(t, T_j) - b(t, T))dW_t^T$ , and hence

$$\mathbb{P}_T(D) = \mathbb{P}_T \left( \sum_{j=1}^m c_j F_B^j(0) e^{\int_0^T \gamma(t, T_j, T) dW_t^T - \frac{1}{2} \int_0^T |\gamma(t, T_j, T)|^2 dt} > K \right)$$

If we denote  $\zeta_j = \int_0^T \gamma(t, T_j, T) dW_t^T$ , then the vector  $\zeta = (\zeta_1, \dots, \zeta_m)$  has a normal distribution under  $\mathbb{P}_T$ , with mean  $(0, 0, \dots, 0)$  and covariance  $(\nu_{kl})$  where

$$\nu_{kl} = \int_0^T \gamma(t, T_k, T) \cdot \gamma(t, T_l, T) dt.$$

To compute  $\mathbb{P}_{T_j}(D)$ , we need to know the distribution of  $\zeta$  under  $\mathbb{P}_{T_j}$ . Since  $W_t^{T_j} = W_t^T - \int_0^t \gamma(u, T_j, T) du$  it is clear that under  $\mathbb{P}_{T_j}$ , the forward price  $F_B^l(t) = F_B(t, T_l, T)$

$$dF_B^l(t) = F_B^l(t) \gamma(t, T_l, T) dW_t^T + F_B(t) \gamma(t, T_l, T) \gamma(t, T_j, T) dt$$

so that the joint distribution of  $\zeta_1, \dots, \zeta_m$  under each forward measure  $\mathbb{P}_{T_j}$  can also be computed. The joint distribution is Gaussian with the same covariance matrix but with means  $v_{lj}$

### 3. Pricing of General Contingent Claims

Let  $\zeta_i(t, T) = \int_t^T \gamma_i(u, T) dW_u^T$ . Then under  $\mathbb{P}_T$  the random variables  $\zeta_i(t, T), \dots, \zeta_n(t, T)$  are normally distributed with mean  $(0, \dots, 0)$  and covariance matrix  $(\gamma_{ij})$  given by

$$\gamma_{ij} = \int_t^T \gamma_i(u, T) \gamma_j(u, T) du.$$

**Proposition 4.3.** *Let  $X = g(Z_T^1, \dots, Z_T^n)$  at time  $T$ . Then the price of  $X$  at time  $t \in [0, T)$  is given by*

$$\pi_t(X) = B(t, T) \int_{\mathbb{R}^k} g\left(\frac{Z_t^1}{B(t, T)} \frac{n_k(x + \theta_1)}{n_k(x)}, \dots, \frac{Z_t^n}{B(t, T)} \frac{n_k(x + \theta_n)}{n_k(x)}\right) n_k(x) dx$$

where  $n_k$  is the standard  $n$ -dimensional Gaussian density on  $\mathbb{R}^k$  and  $(\theta_i)$  are elements of  $\mathbb{R}^k$  such that

$$\theta_i \theta_j = \gamma_{ij}$$

for all  $i, j$ . This follows from the Cholesky decomposition of the covariance matrix  $(\gamma_{ij})$

PROOF.

$$\pi_t(X) = B(t, T) \mathbb{E}_{\mathbb{P}_T} (g(F_{Z^1}(T, T)), \dots, F_{Z^n}(T, T) \mid \mathcal{F}_t)$$

$$F_{Z^i}(T, T) = F_{Z^i}(t, T) e^{\zeta_i(t, T) - \frac{1}{2} \gamma_{ii}}$$

$$\pi_t(X) = B(t, T) \mathbb{E}_{\mathbb{Q}} \left( g \left( F_{Z_t^i} e^{\theta_i \eta - \frac{1}{2} \gamma_{ii}} \right) \mid \mathcal{F}_t \right)$$

$$= B(t, T) \int_{\mathbb{R}^k} g \left( \frac{Z_t^i}{B(t, T)} e^{\theta_i \cdot x - \frac{1}{2} \underbrace{\gamma_{ii}}_{|\theta_i|^2}} \right) n_k(x) dx.$$

Since  $\frac{n_k(x + \theta_i)}{n_k(x)} = e^{\theta_i \cdot x - \frac{1}{2} |\theta_i|^2}$ , we obtain our result.  $\square$

## CHAPTER 5

# Modelling of Forward LIBORs

### 1. Introduction to LIBOR

Let  $\delta$  equal 3 months. If  $L(0) = 10\%$  then if we borrow  $N$  at time 0, we will pay back after three months the amount  $N(1 + \delta L(0))$  where the unit is one year so that  $\delta = \frac{1}{4}$ .

- (i) Spot LIBOR is (or was) the most commonly used rate for interbank funding and as an underlying for interest rate derivatives such as caps and floors.
- (ii) By convention, the pricing formula for caplets and floorlets was a version of the Black formula which reads

$$C_T = F_t N(d_+) - KN(d_-)$$

where  $F_t$  is the forward price of the underlying asset.

Let us consider a caplet with maturity  $T$  and settlement date  $T + \delta$ . Here, a caplet is a call option on LIBOR, in the sense that it pays the amount  $C_P = (L(T) - K)^+ \delta N$  at time  $T + \delta$  where  $T$  is the maturity date,  $N$  is the nominal value, and  $x$  the strike level.

**Definition 5.1** (Cap). A cap is a portfolio of caplets over non-overlapping periods

$$0 < T_0 < T_1 < \dots < T_n$$

so we have  $n$  caplets, struck at  $T_i$  for the period  $[T_i, T_{i+1}]$  and paying  $(L(T_i) - K)^+ N \delta_{i+1}$  at  $T_{i+1}$ , where  $\delta_{i+1} = T_{i+1} - T_i$ .

By convention, the price of a caplet over  $[T, T + \delta]$  equals

$$\text{CPL}_t = B(t, T + \delta) (L(t) N(d_+) - KN(d_-))$$

where

$$d_{\pm} = \frac{\ln \frac{L(t)}{K} \pm \frac{1}{2} \sigma^2 (T - t)}{\sigma \sqrt{T - t}}.$$

## 2. Caps and Floors in the LIBOR Market Model

A caplet (floorlet) is a protection against the rise (fall) in the LIBOR rate. The caplet (floorlet) pays off:

$$\begin{aligned} \mathbf{Cpl}_{T_j}^j N (L(T_{j-1}) - \kappa)^+ \delta_j \\ \mathbf{Frl}_{T_j}^j N (\kappa - L(T_{j-1}))^+ \delta_j \end{aligned}$$

paid at time  $T_j$ .

We clearly have the *cap-floor put call parity*,

$$\mathbf{Cpl}_{T_j}^j - \mathbf{Frl}_{T_j}^j = N_p (L(T_{j-1}) - \kappa) \delta_j.$$

**Exercise 5.2.** Using this relationship, find the difference  $\mathbf{Cpl}_t^j - \mathbf{Frl}_t^j$  for any  $t \in [0, T_{j-1}]$ .

Recall that

$$1 + \delta_j L(T_{j-1}) = \frac{1}{B(T_{j-1}, T_j)}$$

Hence

$$\mathbf{Cpl}_{T_j}^j = N \left( \frac{1}{B(T_{j-1}, T_j)} - \underbrace{(1 + \delta_j \kappa)}_{\tilde{\delta}_j} \right)^+ \delta_j$$

An equivalent payoff at time  $T_{j-1}$  equals

$$\begin{aligned} \tilde{\mathbf{Cpl}}_{T_{j-1}}^j &= B(T_{j-1}, T_j) \mathbf{Cpl}_{T_j}^j \\ &= \tilde{\delta}_j N \left( \frac{1}{\tilde{\delta}_j} - B(T_{j-1}, T_j) \right)^+. \end{aligned}$$

**Definition 5.3.** The forward swap rate  $\kappa(t, T_0, T_1, \dots, T_n) = \kappa(t, T, n)$  where  $T_0 = T$  is the  $\mathcal{F}_t$ -measurable random variable such that  $\mathbf{FS}_t(\kappa(t, T, n)) = 0$ .

**Lemma 5.4.** The forward swap rate equals

$$\kappa(t, T, n) = \frac{B(t, T_0) - B(t, T_n)}{\sum_{j=1}^n \delta_j B(t, T_j)}$$

## CHAPTER 6

### Modelling of Forward Swap Rates

- (i) Definition and payoffs of an  $n$ -period forward swap.
- (ii) Valuation formula for a forward swap (6.4)
- (iii) Definition and formula for forward swap rates (6.5)
- (iv) Definition and equivalent representations for a swaption (Lemma 6.5)
- (v) Postulates of Jamshidian's model of co-terminal forward swap rates
- (vi) Valuation of a swaption (Proposition 6.3)
- (vii) Choice of a numeraire portfolio

Consider the family of co-terminal swap rates

$$\begin{aligned}
 \kappa(t, T_0; n) &= \frac{B(t, T_0) - B(t, T_n)}{\sum_{k=1}^n \delta_k B(t, T_k)} \\
 \kappa(t, T_1; n-1) &= \frac{B(t, T_1) - B(t, T_n)}{\sum_{k=2}^n \delta_k B(t, T_k)} \\
 &\downarrow \\
 \kappa(t, T_{n-1}; 1) &= \frac{B(t, T_{n-1}) - B(t, T_n)}{\delta_n B(t, T_n)} = L(t, T_{n-1})
 \end{aligned}$$

For ease of notation, we let  $\kappa(t, T; n-j) = \tilde{\kappa}(t, T_j)$ .

#### 1. Payer Swaptions

Let us take  $j = 0$  so that the underlying forward swap has  $n$  periods. Let  $\mathbf{FS}_t(\kappa)$  denote the value of the forward swap. We know that

$$\mathbf{FS}_t(\kappa) = B(t, T_0) - \sum_{j=1}^n c_j B(t, T_j)$$

where  $c_j = \kappa \delta_j$ ,  $j = 1, \dots, n-1$ ,  $c_n = (1 + \kappa \delta_n)$ .

**Lemma 6.1.** *The price  $\mathbf{FS}_t(\kappa)$  can be represented as follows:*

$$\begin{aligned}
 \mathbf{FS}_t(\kappa) &= \mathbf{FS}_t(\kappa) - \mathbf{FS}_t(\kappa(t, T_0; n)) \\
 &= \sum_{j=1}^n (\kappa(t, T_0; n) - \kappa) \delta_j B(t, T_j) \\
 &= G_t(n)
 \end{aligned}$$



where

$$G_t(n) = \sum_{\delta_k B(t, T_k)} , \quad G_t(n-j) = \sum_{k=j+1}^n \delta_k B(t, T_k)$$

A payer swaption with a fixed rate  $\kappa$ , maturing date  $T = T_0$  and the underlying  $n$ -period fixed-for-floating forward swap can be identified with the payoff  $(\mathbf{FS}_T(\kappa))^+$  at time  $T$ . A receiver swaption pays  $(-\mathbf{FS}_T(\kappa))^+$  at time  $T$ . Of course, we have a put call parity relationship

$$\mathbf{PS}_t(\kappa) - \mathbf{RS}_t(\kappa) = \mathbf{FS}_t(\kappa)$$

The inequality  $\mathbf{FS}_t(\kappa) > 0$  holds if and only if  $\kappa(T, T; n) > \kappa$  where  $\kappa(T, T; n)$  is the spot swap rate at time  $T_0$ . Hence if  $\kappa(T, T; n) \leq \kappa$  the swaption expires worthless, but it is still possible to enter at  $T$  a forward swap with fixed rate  $\kappa(T, T; n) \leq \kappa$ .

If we define

$$Y_k = \delta_k (\kappa(T, T; n) - \kappa)^+,$$

we know that

$$\begin{aligned} (\mathbf{FS}_t(\kappa))^+ &= \sum_{k=1}^n \delta_k B(t, T_k) (\kappa(T, T; n) - \kappa)^+ \\ &= \sum_{k=1}^n B(t, T_k) Y_k \end{aligned}$$

which is equivalent to a sequence of payoffs  $Y_1, \dots, Y_n$  at times  $T_1, \dots, T_n$ . Also for  $j = 0, 1, \dots, n-1$ ,

$$\begin{aligned} (\mathbf{FS}_{T_0}^0(\kappa))^+ &= G_{T_0}(n) (\kappa(T_0, T_0; n) - \kappa)^+ \\ (\mathbf{FS}_{T_j}^j(\kappa))^+ &= G_{T_j}(n-j) (\kappa(T_j, T_j; n) - \kappa)^+ \end{aligned}$$

We now seek to construct a model for the joint dynamics of a co-terminal family of forward swap rates

$$\kappa(t, T_j; n-j) = \tilde{\kappa}(t, T_j), t \in [0, T_j]$$

such that the volatility  $\nu(t, T_j)$  is given in advance by a deterministic function and the model is driven by a  $d$ -dimensional Brownian motion.<sup>1</sup>

We expect that each process  $\tilde{\kappa}(t, T_j)$  will be a martingale under some probability measure  $\tilde{P}_{T_{j+1}}$  so that

$$d\tilde{\kappa}(t, T_j) = \tilde{\kappa}(t, T_j) \nu(t, T_j) d\tilde{W}_t^{T_{j+1}}$$

---

<sup>1</sup>Any process that we can apply Girsanov's theorem to will be sufficient.

where  $\tilde{W}_t^{T_{j+1}}$  is a Brownian motion under  $\tilde{P}_{T_{j+1}}$  and the Radon-Nikodym densities for  $j = 0, \dots, n-1$  should be given by

$$\frac{d\tilde{P}_{T_{j+1}}}{d\tilde{\mathbb{P}}_{T_n}} = ?$$

which should be expressed in terms of  $\tilde{W}^{T_n}, \tilde{\kappa}(t, T_k), \nu(t, T_k)$  for  $k = n-j+1, \dots, n$ .

## 2. Valuation of Swaptions in Jamshidian's Model

Let us assume that the model is well defined. We will value the  $j$ -th swaption for  $j = 0, \dots, n-1$ . Suppose that it is attainable, so that the price can be computed using the martingale method, meaning here that

$$\pi_t(X) = G_t(n-j) \mathbb{E}_{\tilde{\mathbb{P}}_{T_{j+1}}} \left( \frac{X}{G_{T_j}(n-j)} \mid \mathcal{F}_t \right)$$

where  $X$  is any attainable claim in Jamshidian's model with maturity  $T$ . Observe that only a finite family of forward swaps are traded in this model. In our case,  $X = G_{T_j}(n-j) (\tilde{\kappa}(T_j, T_j) - \kappa)^+$ , and thus

$$\mathbf{PS}_t^j(\kappa) = G_t(n-j) \mathbb{E}_{\tilde{\mathbb{P}}_{T_{j+1}}} \left( (\tilde{\kappa}(T_j, T_j) - \kappa)^+ \mid \mathcal{F}_t \right).$$

Since  $\eta(t, T_j) : [0, T_j] \rightarrow \mathbb{R}^d$  is deterministic, we can evaluate this expression using the Black formula, and obtain

$$\tilde{\kappa}(t, T_j) \Phi \left( \tilde{d}_+^j(\tilde{\kappa}(t, T_j), t, T_j) \right) - \kappa \Phi \left( \tilde{d}_-^j(\tilde{\kappa}(t, T_j), t, T_j) \right)$$

where

$$\begin{aligned} \tilde{d}_\pm(x, t, T_j) &= \frac{\ln \frac{x}{\kappa} \pm \frac{1}{2} v_j^2(t, T_j)}{v_j(t, T_j)} \\ v_j(t, T_j) &= \int_t^{T_j} |v(u, T_j)|^2 du. \end{aligned}$$

For replication of a swaption, we formally define the relative price

$$\mathbf{F}_{S_j, G}(t, T_j) = \frac{\mathbf{PS}_t^j}{G_t(n-j)} = \tilde{\kappa}(t, T_j) \Phi \left( \tilde{d}_+^j(t) \right) - \kappa \Phi \left( \tilde{d}_-^j(t) \right).$$

In this case,

$$dF_{S_j, G}(t, T_j) = \Phi(\tilde{d}_+^j(t)) d\tilde{\kappa}(t, T_j).$$

It is possible to then hedge this option using forward swaps in discrete time.

Let  $\psi^j$  be any trading strategy in the  $j$ -th forward swap. At time 0 the value of our strategy is zero. Then the trading strategy:

$$\begin{aligned} t = 0 & \quad \psi_0^j \text{ positions in market forward swap with rate } \tilde{\kappa}(0, T_j) \\ t = t_1 & \quad \varphi_{t_1}^j \text{ positions in market forward swap with rate } \tilde{\kappa}(t_1, T_j) \\ \downarrow & \quad t = t_n = T_j \end{aligned}$$

Then gains and losses can be conveniently expressed in units of  $G_t(n-j)$ . For instance, the value of our  $\psi_0^j$  positions at time  $t_1$  equals

$$\begin{aligned} \mathbf{PL}_{t_1} &= G_{t_1}(n-j)\psi_0^j(\tilde{\kappa}(t_1, T_j) - \tilde{\kappa}(0, T_j)) \\ \tilde{\mathbf{PL}}_{t_1} &= \underbrace{\psi_0^j(\tilde{\kappa}(t_1, T_j) - \tilde{\kappa}(0, T_j))}_{\text{paid in installments at times } T_{j+1}, \dots, T_n}. \end{aligned}$$

After  $n$  steps,

$$\begin{aligned} \tilde{\mathbf{PL}}_{T_j} &= \sum_{k=0}^{n-1} \psi_{t_k}^j(\kappa(t_{k+1}, T_j) - \tilde{\kappa}(t_k, T_j)) \\ &\rightarrow_{\substack{n \rightarrow \infty \\ t_k = \frac{k}{n}T_j}} \int_0^{T_j} \psi_u^j d\kappa(u, T_j) \end{aligned}$$

The premium  $\mathbf{PS}_0^j$  is totally invested in the level portfolio  $G(n-j)$  so that the totla value of the profit and loss at time  $T_j$  equals

$$\frac{\mathbf{PS}_0^j}{G_0(n-j)} + \int_0^{T_j} \psi_u^j d\kappa(u, T_j)$$

Taking derivatives, we can show that by setting  $\psi_t^j = \Phi(\tilde{d}_+^j(t))$  we obtain the replicating strategy for the  $j$ -th swaption.