Group

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```
In [1]: import numpy as np
   import pandas as pd
   import matplotlib.pyplot as plt
   import seaborn as sns
   import datetime as dt
   import statsmodels.api as sm
   import statsmodels.formula.api as smf
   from statsmodels.regression.linear_model import OLS
```

Question 1: Fixed effects and within transformations

You will find a modified version the imports-85.csv (imports85_modified.csv) file attached to this assignment. Again, make sure that all continuous variables of interest are numeric.

```
In [2]: data = pd.read_csv('data/imports85_modified-1.csv')
   data['city.mpg']=pd.to_numeric(data['city.mpg'], errors='coerce')
   data['horsepower'] = pd.to_numeric(data['horsepower'], errors='coerce')
   data.head()
```

Out[2]:		symboling	normalized- losses	make	fuel.type	aspiration	num.of.doors	body.style	dr
	0	3	NaN	alfa- romero	gas	std	two	convertible	
	1	3	NaN	alfa- romero	gas	std	two	convertible	
	2	1	NaN	alfa- romero	gas	std	two	hatchback	
	3	2	164.0	audi	gas	std	four	sedan	
	4	2	164.0	audi	gas	std	four	sedan	

5 rows × 26 columns

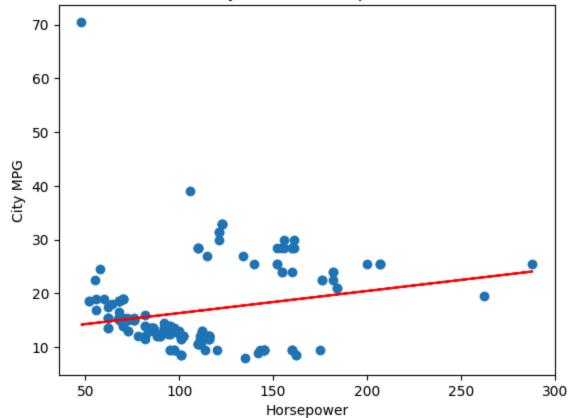
1. Regress fuel efficiency (city.mpg) on horsepower without fixed effects. What would

you conclude based on that regression?

```
In [3]: limited_df = data[['city.mpg', 'horsepower']].dropna()
    y = limited_df['city.mpg']
    X = limited_df['horsepower']
    X = sm.add_constant(X)
    model = OLS(y, X).fit()
    # print(model.summary())

plt.scatter(limited_df['horsepower'], limited_df['city.mpg'])
    plt.plot(limited_df['horsepower'], model.predict(X), color='red')
    plt.xlabel('Horsepower')
    plt.ylabel('City MPG')
    plt.title('City MPG vs Horsepower')
    plt.show()
```





Based on that visualization, it seems that there are two separate clusters, so a linear regression on this whole data without any pre treatment doesn't make sense statistically.

2. Repeat the same regression but this time, add a fixed effect for number of cylinders

being "two" or "four". What would you conclude based on this new regression? What do you think drives the results in part 1?

```
In [4]: limited_df['two_cylinders'] = (data['num.of.cylinders'] == 'two').astype(int
limited_df['four_cylinders'] = (data['num.of.cylinders'] == 'four').astype(int)
```

```
# limited_df['intersection'] = limited_df['two_cylinders'] * limited_df['fou
limited_df = limited_df.dropna()
# print(limited_df.head())
# print(limited_df['intersection'].sum())

X = limited_df[['horsepower', 'two_cylinders', 'four_cylinders']]
X = sm.add_constant(X)
y = limited_df['city.mpg']
model = OLS(y, X).fit()
print(model.summary())

plt.scatter(limited_df['horsepower'], limited_df['city.mpg'])
plt.scatter(limited_df['horsepower'], model.predict(X), color='red')
plt.xlabel('Horsepower')
plt.ylabel('City MPG')
plt.title('City MPG vs Horsepower')
plt.show()
```

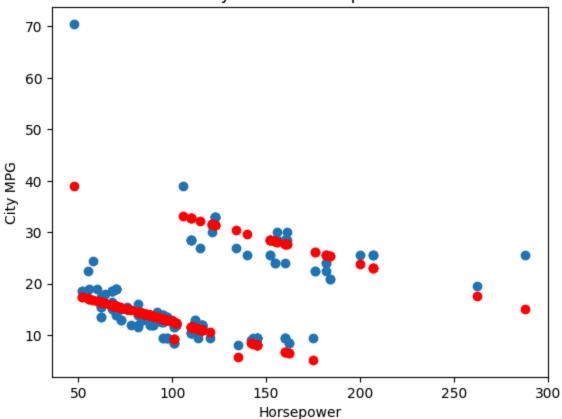
===========		========				====
== Dep. Variable:			R-squared:			0.8
27		0=1)1pg	oqua. ou.			
Model:		0LS	Adj. R−squ	ared:	0.8	
25		-+	F -4-4:-4:			21
Method: 7.7	Lea	st Squares	F-STatisti	.C:		31
Date:	Sun. 1	4 Apr 2024	Prob (F-st	atistic):	1.	34e-
75	,	Γ .				
Time:		18:01:42	Log-Likeli	hood:	-	-516.
05						
No. Observations	:	203	AIC:			104
<pre>0. Df Residuals:</pre>		199	BIC:			105
3.		199	DIC.			103
Df Model:		3				
Covariance Type:		nonrobust				
=============	=======	========	========	========	========	====
=====	coof	std err	+	D~ I + I	[0 025	
0.975]	COET	Stu en	Ĺ	F> C	[0.023	
const	43.6865	1.229	35.559	0.000	41.264	
46.109	0 0006	0 007	12 645	0 000	0 114	
horsepower -0.085	-0.0996	0.007	-13 . 645	0.000	-0.114	
two_cylinders	-24.4047	1.658	-14.715	0.000	-27.675	_
21.134						
four_cylinders	-21.0808	0.716	-29.457	0.000	-22.492	_
19.670						
=======================================	=======	========	========	:=======	========	====
Omnibus:		269.395	Durbin-Wat	son:		1.7
33						
<pre>Prob(Omnibus):</pre>		0.000	Jarque-Ber	a (JB):	250	82.7
81			>			_
Skew:		5.515	Prob(JB):			0.
00 Kurtosis:		56.327	Cond. No.			91
0.		501527	Cond No.			91
=======================================	=======	========	========	========	========	====

Notes:

==

 $\[1\]$ Standard Errors assume that the covariance matrix of the errors is correctly specified.

City MPG vs Horsepower



This new regression fits the data way better. We can see it visually and also thanks to the p-values of the t-statistic.

Based on this regression, one could conclude that the relationship between city mpg and horsepower is different wether the engine has two or four cylinders.

3. (Within transformation) Now obtain the mean city.mpg and horsepower for each

group. Use these group means to demean horsepower and city.mpg. Run the same regression you ran in part 1. Are the results different? Are the results obtained here different from the results in part 2? What does this tell you about the relation between fixed effect regressions and within transformations?

```
In [5]: two_cylinder_mean = limited_df.where(limited_df['two_cylinders'] == 1).dropr
four_cylinder_mean = limited_df.where(limited_df['four_cylinders'] == 1).dro
limited_df['city.mpg'] = limited_df['city.mpg'] - limited_df['two_cylinders'
hp_two_cylinder_mean = limited_df.where(limited_df['two_cylinders'] == 1).dr
hp_four_cylinder_mean = limited_df.where(limited_df['four_cylinders'] == 1).
limited_df['horsepower'] = limited_df['horsepower'] - limited_df['two_cylinders']
X = sm.add_constant(X)
y = limited_df['city.mpg']
```

```
model = OLS(y, X).fit()
model.summary()

# plt.scatter(limited_df['horsepower'], limited_df['city.mpg'])
# plt.scatter(limited_df['horsepower'], model.predict(X), color='red')
# plt.xlabel('Horsepower')
# plt.ylabel('City MPG')
# plt.title('City MPG vs Horsepower')
# plt.show()
```

Out[5]:

OLS Regression Results

Dep. Variable:	city.mpg	R-squared:	0.937
Model:	OLS	Adj. R-squared:	0.936
Method:	Least Squares	F-statistic:	981.0
Date:	Sun, 14 Apr 2024	Prob (F-statistic):	6.36e-119
Time:	18:01:43	Log-Likelihood:	-516.05
No. Observations:	203	AIC:	1040.
Df Residuals:	199	BIC:	1053.
Df Model:	3		

Covariance Type: nonrobust

		coef	std err	t	P> t	[0.025	0.975]
	const	43.6865	1.229	35.559	0.000	41.264	46.109
	horsepower	-0.0996	0.007	-13.645	0.000	-0.114	-0.085
	two_cylinders	-43.6865	1.980	-22.065	0.000	-47.591	-39.782
f	four_cylinders	-43.6865	1.253	-34.857	0.000	-46.158	-41.215

Omnibus:	269.395	Durbin-Watson:	1.733
Prob(Omnibus):	0.000	Jarque-Bera (JB):	25082.781
Skew:	5.515	Prob(JB):	0.00
Kurtosis:	56.327	Cond. No.	836.

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

We notice that the coefficients are equal and the R^2 is better (0.937 > 0.827). Fixed effect regressions and within transformations both achieve the same goal but they lead to different coefficients, which makes the interpretation of the regerssions different.

Question 2: On marginal significance and trading strategy improvements

You come up with a signal of stock outperformance: log total asset growth. You realize that your professor has conveniently already coded up this variable for you in the dataset StockRetAcct_DT.csv. The variable is called "InInv".

1. Using the Fama-MacBeth regression approach, what are the average return, standard

deviation and Sharpe ratio of the trading strategy implied by using only an intercept and InInv on the right hand side in the regressions?

```
In [6]: # Using the code from the class slides we can compute the statistics of this
        StockRetAcct_DF = pd.read_csv('data/StockRetAcct_DT.csv')
        # We compute the excess annual return to use for the regression after
        StockRetAcct_DF['ExRet'] = np.exp(StockRetAcct_DF['lnAnnRet']) - np.exp(StockRetAcct_DF['lnAnnRet'])
        # define function that returns OLS coefficients from fitted regression
        def ols_coef(x,formula): return smf.ols(formula,data=x).fit().params
        # the below runs regression ExRet ~ lnInv by year, including intercept
        res = (StockRetAcct DF.groupby('year').apply(ols coef, 'ExRet ~ lnInv'))
       res['lnInv'].std())+'\n',
                      ', str(35**.5*(res['lnInv'].mean())/
           res['lnInv'].std()), sep="\n")
      Mean Return:
      -0.08679146319781365
      Std Dev:
      0.14864410759416122
      Sharpe Ratio
      -0.5838876804641185
      t-stat:
      -3.4543261019947002
```

for

2. What is the analytical expression for the portfolio weights in this case? (I'm looking

a formula)

 $\label{eq:likelihood} $$ w_{i, t-1} = \frac{1}{N} \frac{(x_{i, t-1}-\mathbb{E}_N[x_{i, t-1}])}{\mathcal E}_N[x_{i, t-1}]} $$$

- With:
 - \$x_{i, t-1}\$: Risk premium of the company \$i\$ at time \$t\$.
 - \$N\$: number of companies.
 - 3. You worry that there is industry-related noise associated with the characteristic lnlny

and want to clean up your trading strategy with the goal of reducing exposure to unpriced industry risks. What regressions to you run? Report mean, standard deviation, and Sharpe ratio of the 'cleaned-up' trading strategy.

We run a similar regression but with industry dummies.

Mean Returns: Industry[T.2.0] Industry[T.3.0] Industry[T.4.0] Industry[T.5.0] Industry[T.6.0] Industry[T.7.0] Industry[T.8.0] Industry[T.9.0] Industry[T.10.0] Industry[T.11.0] Industry[T.11.0] Industry[T.12.0] Influstry[T.12.0] Influstry[T.10.0]	-0.012202 -0.016387 -0.029597 -0.001101 0.006305 0.023252 -0.018057 0.008121 0.029135 0.012417 -0.023147 -0.082578
Std Dev: Industry[T.2.0] Industry[T.3.0] Industry[T.4.0] Industry[T.5.0] Industry[T.6.0] Industry[T.7.0] Industry[T.7.0] Industry[T.8.0] Industry[T.9.0] Industry[T.10.0] Industry[T.11.0] Industry[T.11.0] Industry[T.12.0] lnInv dtype: float64	0.130250 0.129407 0.297996 0.098524 0.280174 0.217234 0.143258 0.093976 0.184509 0.118847 0.090814 0.101964
Sharpe Ratio Industry[T.2.0] Industry[T.3.0] Industry[T.4.0] Industry[T.5.0] Industry[T.6.0] Industry[T.7.0] Industry[T.8.0] Industry[T.9.0] Industry[T.10.0] Industry[T.11.0] Industry[T.11.0] Industry[T.12.0] lnInv dtype: float64	-0.093681 -0.126628 -0.099320 -0.011177 0.022504 0.107039 -0.126045 0.086421 0.157905 0.104476 -0.254884 -0.809868
t-stat: Industry[T.2.0] Industry[T.3.0] Industry[T.4.0] Industry[T.5.0] Industry[T.6.0] Industry[T.7.0] Industry[T.7.0] Industry[T.8.0] Industry[T.9.0] Industry[T.10.0] Industry[T.11.0]	-0.554223 -0.749142 -0.587584 -0.066126 0.133133 0.633249 -0.745693 0.511272 0.934180 0.618091

```
Industry[T.12.0] -1.507913
lnInv -4.791247
```

dtype: float64

4. As in the class notes, plot the cumulative returns to the simple and the 'cleaned-up'

trading strategies based on your new signal, InInv. Make sure both trading strategies result in portfolios with a 15% return standard deviation.

The function FamaMacBeth does not work for me, so I could not go further.

It always gives me the error: "ValueError: dependent and exog must have the same number of observations. The number of observations in dependent is 66788, and the number of observations in exog is 133576.".

Which is not true as print(y.shape, X.shape) gives (66788, 1) (66788, 2). I have verified everything and the documentation does not give any information that could help. It should not be the function dmatrices as splitting the dataset by hand gives the same error message.

The following code normally displays the cumulative returns to the simple and the 'cleaned-up' trading strategies based on the new signal lnlnv.

Both trading strategies result in portfolios with a 15% return standard deviation.

```
In [8]: # from linearmodels.panel.model import FamaMacBeth
        # from patsy import dmatrices
        # y, X = dmatrices('ExRet~lnInv', StockRetAcct_DF, return_type = 'dataframe
        # res1 = FamaMacBeth(y, X).fit()
        # LastDate = StockRetAcct DF[StockRetAcct DF['year']==2014].dropna()
        # LastDate = LastDate[['lnInv']].assign(c=1).sort_index(axis=1)
        # # Create dummy variables for each industry. We drop the last column becaus
        # # include a constant to avoid multicollinearity
        # StockRetAcct_DF['Industry']=StockRetAcct_DF['ff_ind'].astype(object)
        # LastDate = LastDate.join(pd.get dummies(StockRetAcct DF.Industry).iloc[:,:
        # # use (X'X)^{(-1)} X' formula to compute weights
        # portweights lnInv = np.matmul(np.linalg.inv(np.matmul(LastDate.transpose()
        # values,LastDate.values)),LastDate.transpose().values)
        # # lnInv is third row
        \# portweights_lnInv = np.matmul(np.array([1,0,0,0,0,0,0,0,0,0,0,0,0]).T, por
        # lnInvstdev = res.lnInv.std()
        # lnInvret = res.lnInv
        # # scale portfolio weights to get 15% standard deviation of returns
        # portweights lnInv = portweights lnInv*0.15/lnInvstdev
        # y, X = dmatrices('ExRet~lnInv+Industry', StockRetAcct_DF, return_type = 'd
        \# res = FamaMacBeth(y,X).fit()
        # # for plotting, get the scaled excess portfolio returns
```

```
# lnInv ret = lnInvret*0.15/lnInvstdev
# # create cumulative log return series
# cum ret lnInv = pd.DataFrame.cumsum(np.log(1+lnInv ret))
# # get "old" simple value strategy returns
# lnInvstdev = res1.lnInv.std()
# lnInvret = res1.lnInv
# lnInv old ret = pd.DataFrame.cumsum(np.log(1+lnInvret*0.15/lnInvstdev))
# summary = pd.DataFrame()
# summary['NewValue'] = np.exp(cum ret lnInv)
# summary['OldValue'] = np.exp(lnInv old ret)
# # Plot Old Value vs New Value
# sns.set style('darkgrid')
# plt.figure()
# ax=sns.lineplot(data=summary,dashes = False,linewidth = 3)
# ax.set(xlabel = 'Year',
     ylabel = 'Cumulative Return',
      title = "Old Value vs New Value")
```

Question 3: Predicting medium to long-run firm-level return variance

There are many return volatility models, such as GARCH. These work best at shorter horizons. As an alternative, we will explore a panel regression approach to predicting firm-level return variance. The data set StockRetAcct_DT.csv has annual realized variance (rv), calculated as the sum of squared daily returns to each firm, each year. Run panel forecasting regressions to forecast firm-level one-year ahead rv along the lines of what we did with InROE in class.

1. Try with and without industry and year fixed effects, with and without clustering of

standard errors. Discuss which specification makes most sense to you. In particular, discuss the effect of a year fixed effect. What is the intuition for the impact of this fixed effect?

```
In [9]: retData = pd.read_csv('data/StockRetAcct_DT.csv')
#Vanilla, fixed-effect free model
X = retData['rv'].shift(1)[1:].fillna(0)
X = sm.add_constant(X)
y = retData['rv'][1:]
vanilla_model = OLS(y, X).fit()
print(vanilla_model.summary())
```

```
Dep. Variable:
                                      R-squared:
                                                                        n
Model:
                                0LS
                                      Adj. R-squared:
                                                                        n
an
                      Least Squares
                                      F-statistic:
Method:
                                                                        n
an
Date:
                    Sun, 14 Apr 2024
                                      Prob (F-statistic):
                                                                        n
an
Time:
                            18:01:43
                                      Log-Likelihood:
                                                                        n
No. Observations:
                              70755
                                      AIC:
                                                                        n
Df Residuals:
                              70753
                                      BTC:
                                                                        n
an
Df Model:
Covariance Type:
                          nonrobust
                coef std err t P>|t| [0.025
                                                                     0.97
51
const
                 nan
                           nan
                                      nan
                                                 nan
                                                            nan
                                                                        n
an
                           nan
                                                            nan
rv
                 nan
                                      nan
                                                 nan
                                                                        n
an
==
Omnibus:
                                nan
                                      Durbin-Watson:
                                                                        n
an
Prob(Omnibus):
                                      Jarque-Bera (JB):
                                nan
                                                                        n
Skew:
                                      Prob(JB):
                                nan
                                                                        n
an
                                      Cond. No.
                                                                       5.
Kurtosis:
                                nan
==
```

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

```
In [10]: inds = retData['ff_ind'].dropna().unique()
    regression_df = retData[['year', 'rv']]
    regression_df['lag_rv'] = regression_df['rv'].shift(1).fillna(0)

for year in regression_df['year'].unique():
    regression_df['is_'+str(year)] = (regression_df['year'] == year).astype(
    regression_df['ff_ind'] = retData['ff_ind'].fillna(0)
    for ind in inds:
        regression_df['is_ind_'+str(ind)] = (regression_df['ff_ind'] == ind).ast
```

```
regression_df.head()
# print(regression_df.head())

year_fe_features = regression_df[['lag_rv'] + [('is_' + str(year)) for year
year_fe_features = sm.add_constant(year_fe_features)
y = regression_df['rv'].fillna(0)
year_fe_model = OLS(y, year_fe_features).fit()
print(year_fe_model.summary())
```

```
/var/folders/r0/zk7h1dpx693qc15c5y61js8w0000qn/T/ipykernel 89533/2185439882.
py:3: SettingWithCopyWarning:
A value is trying to be set on a copy of a slice from a DataFrame.
Try using .loc[row indexer,col indexer] = value instead
See the caveats in the documentation: https://pandas.pydata.org/pandas-docs/
stable/user guide/indexing.html#returning-a-view-versus-a-copy
  regression_df['lag_rv'] = regression_df['rv'].shift(1).fillna(0)
/var/folders/r0/zk7h1dpx693qc15c5y61js8w0000qn/T/ipykernel 89533/2185439882.
py:6: SettingWithCopyWarning:
A value is trying to be set on a copy of a slice from a DataFrame.
Try using .loc[row_indexer,col_indexer] = value instead
See the caveats in the documentation: https://pandas.pydata.org/pandas-docs/
stable/user_guide/indexing.html#returning-a-view-versus-a-copy
  regression df['is '+str(year)] = (regression df['year'] == year).astype(in
/var/folders/r0/zk7h1dpx693gc15c5y61js8w0000gn/T/ipykernel_89533/2185439882.
py:6: SettingWithCopyWarning:
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stable/user_guide/indexing.html#returning-a-view-versus-a-copy
  regression_df['is_'+str(year)] = (regression_df['year'] == year).astype(in
t)
/var/folders/r0/zk7h1dpx693gc15c5y61js8w0000gn/T/ipykernel_89533/2185439882.
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t)
/var/folders/r0/zk7h1dpx693qc15c5y61js8w0000qn/T/ipykernel 89533/2185439882.
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  regression_df['is_'+str(year)] = (regression_df['year'] == year).astype(in
t)
/var/folders/r0/zk7h1dpx693qc15c5y61js8w0000qn/T/ipykernel 89533/2185439882.
py:6: SettingWithCopyWarning:
A value is trying to be set on a copy of a slice from a DataFrame.
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stable/user guide/indexing.html#returning-a-view-versus-a-copy
  regression_df['is_'+str(year)] = (regression_df['year'] == year).astype(in
/var/folders/r0/zk7h1dpx693qc15c5y61js8w0000qn/T/ipykernel 89533/2185439882.
py:6: SettingWithCopyWarning:
A value is trying to be set on a copy of a slice from a DataFrame.
```

```
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stable/user guide/indexing.html#returning-a-view-versus-a-copy
  regression_df['is_'+str(year)] = (regression_df['year'] == year).astype(in
t)
/var/folders/r0/zk7h1dpx693qc15c5y61js8w0000qn/T/ipykernel 89533/2185439882.
py:6: SettingWithCopyWarning:
A value is trying to be set on a copy of a slice from a DataFrame.
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  regression_df['is_'+str(year)] = (regression_df['year'] == year).astype(in
/var/folders/r0/zk7h1dpx693qc15c5y61js8w0000qn/T/ipykernel 89533/2185439882.
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t)
/var/folders/r0/zk7h1dpx693qc15c5y61js8w0000qn/T/ipykernel 89533/2185439882.
py:6: SettingWithCopyWarning:
A value is trying to be set on a copy of a slice from a DataFrame.
Try using .loc[row indexer,col indexer] = value instead
See the caveats in the documentation: https://pandas.pydata.org/pandas-docs/
stable/user_guide/indexing.html#returning-a-view-versus-a-copy
  regression df['is '+str(year)] = (regression df['year'] == year).astype(in
/var/folders/r0/zk7h1dpx693gc15c5y61js8w0000gn/T/ipykernel_89533/2185439882.
py:6: SettingWithCopyWarning:
A value is trying to be set on a copy of a slice from a DataFrame.
Try using .loc[row_indexer,col_indexer] = value instead
See the caveats in the documentation: https://pandas.pydata.org/pandas-docs/
stable/user guide/indexing.html#returning-a-view-versus-a-copy
  regression df['is '+str(year)] = (regression df['year'] == year).astype(in
t)
/var/folders/r0/zk7h1dpx693gc15c5y61js8w0000gn/T/ipykernel_89533/2185439882.
py:6: SettingWithCopyWarning:
A value is trying to be set on a copy of a slice from a DataFrame.
Try using .loc[row_indexer,col_indexer] = value instead
```

```
See the caveats in the documentation: https://pandas.pydata.org/pandas-docs/
stable/user_guide/indexing.html#returning-a-view-versus-a-copy
  regression_df['is_'+str(year)] = (regression_df['year'] == year).astype(in
t)
/var/folders/r0/zk7h1dpx693gc15c5y61js8w0000gn/T/ipykernel_89533/2185439882.
py:6: SettingWithCopyWarning:
A value is trying to be set on a copy of a slice from a DataFrame.
Try using .loc[row_indexer,col_indexer] = value instead

See the caveats in the documentation: https://pandas.pydata.org/pandas-docs/
stable/user_guide/indexing.html#returning-a-view-versus-a-copy
  regression_df['is_'+str(year)] = (regression_df['year'] == year).astype(in
t)
```

	:========	=======	=====	======	=======	=======	=======
== Dep. Vari	able:		rv	R-squa	red:		0.4
67 Model:			0LS	Adi. R	-squared:		0.4
67				_	·		
Method: 4.		Least Squa	res	F-stat	istic:		172
Date:	Su	n, 14 Apr 2	024	Prob (F-statisti	.c):	0.
00 Time:		18:01	.44	l oa-l i	kelihood:		4045
7.	wations.			-	Re cinoda.		
No. Obser	vacions:	70	756	AIC:			-8.084e+
Df Residu 04	ials:	70	719	BIC:			-8.050e+
Df Model:			36				
Covarianc	e Type:	nonrob	ust				
=======================================	=========	=======	=====	======	=======	========	=======
	coef	std err		t	P> t	[0.025	0.97
5]							
const 09	-1.026e+10	6.4e+09	-1	L.602	0.109	-2.28e+10	2.29e+
lag_rv	0.3522	0.003	110	589	0.000	0.346	0.3
	1.026e+10	6.4e+09	1	L.602	0.109	-2.29e+09	2.28e+
	1.026e+10	6.4e+09	1	L.602	0.109	-2.29e+09	2.28e+
	1.026e+10	6.4e+09	1	L.602	0.109	-2.29e+09	2.28e+
10 is_1983	1.026e+10	6.4e+09	1	L.602	0.109	-2.29e+09	2.28e+
10 is_1991 10	1.026e+10	6.4e+09	1	L.602	0.109	-2.29e+09	2.28e+
is_2000 10	1.026e+10	6.4e+09	1	L.602	0.109	-2.29e+09	2.28e+
is_1986 10	1.026e+10	6.4e+09	1	L.602	0.109	-2.29e+09	2.28e+
is_1987 10	1.026e+10	6.4e+09	1	L.602	0.109	-2.29e+09	2.28e+
is_1988 10	1.026e+10	6.4e+09	1	L.602	0.109	-2.29e+09	2.28e+
is_1989	1.026e+10	6.4e+09	1	L.602	0.109	-2.29e+09	2.28e+
10 is_1990	1.026e+10	6.4e+09	1	L.602	0.109	-2.29e+09	2.28e+
10 is_1992	1.026e+10	6.4e+09	1	L.602	0.109	-2.29e+09	2.28e+
10 is_1993 10	1.026e+10	6.4e+09	1	L.602	0.109	-2.29e+09	2.28e+
is_1994	1.026e+10	6.4e+09	1	L.602	0.109	-2.29e+09	2.28e+

=======================================	========	:========	======			======
Kurtosis: 13		16.379				7.69e+
Skew: 00		1.800	Prob(JB):		0.
Prob(Omnibu	S):	0.000	·	e-Bera (JB)	:	565891.8
Omnibus: 86	-1.	32775.551				1.9
10 ====================================	========	:========	======		:=======	=======
	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
10 is_1984	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
10 is_1995	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
10 is_2007	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
10 is_2006	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
is_2005	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
is_2004 10	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
is_2003 10	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
is_2014 10	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
is_2013 10	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
is_2012 10	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
is_2011 10	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
10						
10 is_2010	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
10 is_2009	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
10 is_2008	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
10 is_2002	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
10 is_2001	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
10 is_1999	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
10 is_1998	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
10 is_1997	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+
10 is_1996	1.026e+10	6.4e+09	1.602	0.109	-2.29e+09	2.28e+

Notes:

- [1] Standard Errors assume that the covariance matrix of the errors is correctly specified.
- [2] The smallest eigenvalue is 1.26e-23. This might indicate that there are strong multicollinearity problems or that the design matrix is singular.

```
In [11]: ind_fe_features = regression_df[['lag_rv'] + [('is_ind_' + str(ind)) for ind
    ind_fe_features = sm.add_constant(ind_fe_features)
    y = regression_df['rv'].fillna(0)
    ind_fe_model = OLS(y, ind_fe_features).fit()
    print(ind_fe_model.summary())
```

==		========	======	==========	======	
Dep. Variable: 89		rv	R-squa	red:		0.1
Model: 89		0LS	Adj. R	-squared:		0.1
Method:		Least Squares	F-stat	istic:		127
2. Date:	Sun	, 14 Apr 2024	Prob (F-statistic):		0.
00 Time:		18:01:44	Log-Li	kelihood:		2560
1. No. Observation	ons:	70756	AIC:			-5.117e+
04 Df Residuals:		70742	BIC:			-5.105e+
04		12				
<pre>Df Model: Covariance Typ</pre>	ne•	13 nonrobust				
=========			=======		======	=======
===						
751	coef	std err	t	P> t	[0.025	0.9
75] 						
const 094	-0.0413	0.069	-0.601	0.548	-0.176	0.
	0.3438	0.003	98.286	0.000	0.337	0.
is_ind_3.0 267	0.1318	0.069	1.915	0.055	-0.003	0.
	0.1821	0.069	2.646	0.008	0.047	0.
is_ind_6.0 349	0.2141	0.069	3.110	0.002	0.079	0.
is_ind_12.0 282	0.1474	0.069	2.142	0.032	0.013	0.
is_ind_11.0 251	0.1161	0.069	1.687	0.092	-0.019	0.
is_ind_1.0 248	0.1129	0.069	1.639	0.101	-0.022	0.
is_ind_9.0 278	0.1434	0.069	2.084	0.037	0.009	0.
is_ind_7.0 295	0.1595	0.069	2.315	0.021	0.024	0.
is_ind_2.0 267	0.1324	0.069	1.922	0.055	-0.003	0.
is_ind_5.0	0.1181	0.069	1.714	0.086	-0.017	0.
253 is_ind_8.0	0.0810	0.069	1.176	0.240	-0.054	0.
216 is_ind_4.0 292	0.1569	0.069	2.279	0.023	0.022	0.
======================================	=======	49472.965		======================================	======	2.0

```
48
Prob(Omnibus):
                                0.000
                                       Jarque-Bera (JB):
                                                                  1008799.9
98
                                3.142
                                       Prob(JB):
                                                                          0.
Skew:
00
Kurtosis:
                               20.398
                                       Cond. No.
                                                                          41
==
```

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

```
In [12]: joint_fe_features = regression_df[['lag_rv'] + [('is_' + str(year)) for year
    joint_fe_features.head()
    joint_fe_features = sm.add_constant(joint_fe_features)
    y = regression_df['rv'].fillna(0)
    joint_fe_model = OLS(y, joint_fe_features).fit()
    print(joint_fe_model.summary())
```

=======	========	:=======	=====	=====	========	========	=======
==							
Dep. Vari 67	able:		rv	R-squ	ared:		0.4
Model:			0LS	Adj.	R-squared:		0.4
67 Method:		Least Squ	ares	F-sta	tistic:		154
9.							
Date: 00	S	Sun, 14 Apr	2024	Prob	(F-statisti	c):	0.
Time:		18:0	1:45	Log-L	ikelihood:		4043
8. No. Obser	vations:	70	0756	AIC:			-8.079e+
04 Df Residu	als:	7(0715	BIC:			-8.042e+
04							
Df Model:			40				
	e Type:						
=======================================	========	========	=====	=====	========	========	=======
	coef	std err		t	P> t	[0.025	0.97
5]						_	
			_				
const 11	-5.634e+10	8.78e+10	-0	.641	0.521	-2.28e+11	1.16e+
lag_rv 61	0.3542	0.004	98	. 888	0.000	0.347	0.3
	5.97e+09	1.51e+10	0	. 394	0.693	-2.37e+10	3.56e+
	1.353e+09	1.66e+10	0	.081	0.935	-3.12e+10	3.39e+
	5.258e+10	3.06e+10	1	.719	0.086	-7.36e+09	1.13e+
is_1983 11	3.705e+11	1.01e+11	3	658	0.000	1.72e+11	5.69e+
is_1991 11	7.668e+10	2.71e+10	2	.826	0.005	2.35e+10	1.3e+
is_2000 11	2.488e+10	4.77e+10	0	. 522	0.602	-6.86e+10	1.18e+
is_1986	-1.341e+10	7.33e+10	-0	. 183	0.855	-1.57e+11	1.3e+
11 is_1987	-4.516e+09	4.89e+10	-0	.092	0.926	-1e+11	9.13e+
10 is_1988	-6.716e+09	6.28e+10	-0	. 107	0.915	-1.3e+11	1.16e+
11 is_1989	-4.022e+10	3.48e+10	-1	. 155	0.248	-1.08e+11	2.8e+
10 is_1990	2.187e+10	6.62e+10	0	.330	0.741	-1.08e+11	1.52e+
11 is_1992	8.361e+08	2.33e+10	0	.036	0.971	-4.48e+10	4.65e+
10 is_1993	-2.242e+10	5.44e+10	-0	. 412	0.680	-1.29e+11	8.42e+
10 is_1994	1.522e+10	3.6e+10	0	. 423	0.673	-5.53e+10	8.58e+

1.0						
10 is_1996	8.207e+09	2.14e+10	0.384	0.701	-3.37e+10	5.01e+
10 is_1997 10	-1.039e+10	2.19e+10	-0.474	0.636	-5.33e+10	3.26e+
is_1998 11	4.011e+10	3.13e+10	1.282	0.200	-2.12e+10	1.01e+
is_1999 10	-4.821e+10	7.18e+10	-0.672	0.502	-1.89e+11	9.25e+
is_2001 11	6.471e+10	5.38e+10	1.202	0.229	-4.08e+10	1.7e+
is_2002 11	7.851e+10	7.75e+10	1.014	0.311	-7.33e+10	2.3e+
is_2008 11	1.419e+08	3.47e+11	0.000	1.000	-6.81e+11	6.81e+
is_2009 11	2.784e+10	4.45e+10	0.626	0.531	-5.93e+10	1.15e+
is_2010 11	2.758e+10	9.11e+10	0.303	0.762	-1.51e+11	2.06e+
is_2011 11	-2.569e+10	1.05e+11	-0.244	0.807	-2.32e+11	1.81e+
is_2012 11	2.903e+10	4.37e+10	0.664	0.506	-5.66e+10	1.15e+
is_2013 11	-3.896e+10	9.6e+10	-0.406	0.685	-2.27e+11	1.49e+
is_2014 10	-9.1e+10	9.52e+10	-0.956	0.339	-2.78e+11	9.56e+
is_2003 11	2.88e+10	4.36e+10	0.661	0.509	-5.67e+10	1.14e+
is_2004 10	-7.392e+10	6.34e+10	-1.166	0.244	-1.98e+11	5.04e+
is_2005 11	3.461e+11	2.39e+11	1.450	0.147	-1.22e+11	8.14e+
is_2006 11	3.038e+10	4.34e+10	0.701	0.484	-5.46e+10	1.15e+
is_2007 11	-2.544e+11	2.28e+11	-1.115	0.265	-7.02e+11	1.93e+
is_1995 11	-5.173e+10	1.06e+11	-0.488	0.625	-2.59e+11	1.56e+
is_1984 11	3.34e+10	4.45e+10	0.751	0.453	-5.38e+10	1.21e+
is_1985 11	2.803e+10	4.4e+10	0.637	0.524	-5.82e+10	1.14e+
is_1980 11	5.037e+10	8.04e+10	0.626	0.531	-1.07e+11	2.08e+
is_1981 11	5.498e+10	8.92e+10	0.617	0.538	-1.2e+11	2.3e+
is_1982 11	3.757e+09	7.9e+10	0.048	0.962	-1.51e+11	1.59e+
is_1983 09	-3.142e+11	1.65e+11	-1.902	0.057	-6.38e+11	9.58e+
is_1991 11	-2.035e+10	7.44e+10	-0.274	0.784	-1.66e+11	1.25e+
is_2000 11	3.146e+10	4.21e+10	0.748	0.455	-5.1e+10	1.14e+
is_1986	6.975e+10	1.05e+11	0.663	0.508	-1.37e+11	2.76e+

11						
11 is_1987 11	6.085e+10	1.09e+11	0.561	0.575	-1.52e+11	2.74e+
is_1988 11	6.305e+10	1.35e+11	0.467	0.641	-2.02e+11	3.28e+
is_1989 11	9.656e+10	7.75e+10	1.246	0.213	-5.53e+10	2.48e+
is_1990 11	3.446e+10	8.11e+10	0.425	0.671	-1.25e+11	1.93e+
is_1992 11	5.55e+10	9.98e+10	0.556	0.578	-1.4e+11	2.51e+
is_1993 11	7.876e+10	8.55e+10	0.921	0.357	-8.89e+10	2.46e+
is_1994 11	4.112e+10	8.05e+10	0.511	0.609	-1.17e+11	1.99e+
is_1996 11	4.813e+10	7.96e+10	0.605	0.545	-1.08e+11	2.04e+
is_1997 11	6.672e+10	9.34e+10	0.714	0.475	-1.16e+11	2.5e+
is_1998 11	1.623e+10	8.67e+10	0.187	0.852	-1.54e+11	1.86e+
is_1999 11	1.045e+11	9.53e+10	1.097	0.272	-8.22e+10	2.91e+
is_2001 11	-8.372e+09	8.84e+10	-0.095	0.925	-1.82e+11	1.65e+
is_2002 11	-2.217e+10	1.1e+11	-0.202	0.840	-2.37e+11	1.93e+
is_2008 11	5.619e+10	3.48e+11	0.161	0.872	-6.26e+11	7.39e+
is_2009 11	2.85e+10	4.34e+10	0.656	0.512	-5.66e+10	1.14e+
is_2010 11	2.876e+10	1.14e+11	0.253	0.800	-1.94e+11	2.51e+
is_2011 11	8.202e+10	1.31e+11	0.628	0.530	-1.74e+11	3.38e+
is_2012 11	2.73e+10	4.42e+10	0.618	0.536	-5.93e+10	1.14e+
is_2013 11	9.529e+10	1.28e+11	0.745	0.456	−1 . 55e+11	3.46e+
is_2014 11	1.473e+11	1.32e+11	1.120	0.263	-1.1e+11	4.05e+
is_2003 11	2.753e+10	4.43e+10	0.622	0.534	-5 . 93e+10	1.14e+
is_2004 11	1.303e+11	1.05e+11	1.242	0.214	-7 . 53e+10	3.36e+
is_2005 11	-2.898e+11	2.42e+11	-1.196	0.232	-7 . 65e+11	1.85e+
is_2006 11	2.596e+10	4.51e+10	0.576	0.565	-6.24e+10	1.14e+
is_2007 11	3.107e+11	2.31e+11	1.344	0.179	-1.42e+11	7.64e+
is_1995 11	1.081e+11			0.402	-1.45e+11	3.61e+
is_1984 11			0.519	0.604	-6.37e+10	1.1e+
is_1985	2.83e+10	4.39e+10	0.645	0.519	-5.77e+10	1.14e+

11 --Omnibus: 32923.076 Durbin-Watson: 1.9 91 Prob(Omnibus): 0.000 Jarque-Bera (JB): 570589.7 Skew: 1.810 Prob(JB): 0. 00 Kurtosis: 16.433 Cond. No. 1.71e+ 16

Notes:

- [1] Standard Errors assume that the covariance matrix of the errors is correctly specified.
- [2] The smallest eigenvalue is 2.63e-28. This might indicate that there are strong multicollinearity problems or that the design matrix is singular.

Across regressions, we see pretty consistent t+1 autocorrelation of returns represented by the lag_rv coef which hovers around 35% (just under). Additionally, the is_year variables help in extracting out the realized mean bump to vol in that year with the is_ind_x variable doing the same for industry effects. Which we want to keep comes down to how parsimonious we would like our model to be and is up to the user

2. Also try forecasting at the 5-year horizon (rv in 5 years). How do the results change?

Can we predict return variance 5-years ahead? Is the 5-year lagged rv significant, or are other variables more important?

=========	=======	========	=====	======	========	======	=======
== Dep. Variable:			rv	R-squ	ared:		0.0
13 Model:			0LS	Adj. I	R-squared:		0.0
13				-	·		
Method: 1.7		Least Squa	res	F–sta	tistic:		90
Date: 97	Sı	ın, 14 Apr 2	024	Prob	(F-statistic):		7.30e-1
Time:		18:01	:45	Log-L	ikelihood:		1862
0.		70	756	ATC.			-3.724e+
No. Observatio 04	ns:	70	756	AIC:			-3.724e+
Df Residuals: 04		70	754	BIC:			-3.722e+
Df Model:			1				
Covariance Type	e:	nonrob					
=======================================	=======	:=======	=====	=====	=========	======	=======
==							
_	coef	std err		t	P> t	[0.025	0.97
5]							
	0.1407	0.001	153	3.553	0.000	0.139	0.1
42	011107	0.001	100		0.000	0.155	0.1
5y_lag_rv 19	0.1122	0.004	30	0.028	0.000	0.105	0.1
=======================================	=======			======		======	=======
==		40004	4 - 4	Б. 1.			4.2
Omnibus:		48804.	154	Durbi	n-Watson:		1.2
<pre>Prob(Omnibus):</pre>		0.	000	Jarqu	e-Bera (JB):		740413.3
23							
Skew:		3.	205	Prob(JB):		0.
00 Kurtosis:		17.	493	Cond.	No.		5.
48		1/1		Condi			J.
	=======	=======	=====	:====:	=========	======	=======
==							

Notes:

 $\[1\]$ Standard Errors assume that the covariance matrix of the errors is correctly specified.

In [14]: print(year_fe_model_5y.summary())

========	========		=====	=====	========		=======
==							
Dep. Varia 91	ble:		rv	R-squ	ared:		0.3
Model: 91			0LS	Adj.	R-squared:		0.3
Method:		Least Squa	ires	F-sta	tistic:		126
3.	_						
Date: 00	S	Sun, 14 Apr 2	2024	Prob	(F—statisti	c):	0.
Time:		18:01	: 45	Log-L	ikelihood:		3573
9. No. Observ	ations:	70	756	AIC:			-7.140e+
04 Df Residua	ls:	70	719	BIC:			-7.107e+
04		, ,	,, 13	510.			,120,0
Df Model:			36				
		nonrob					
=======================================	=======	-=======	=====	=====	:=======	=======	=======
	coef	std err		t	P> t	[0.025	0.97
5]							
	1.452e+10	6.15e+09	2	362	0.018	2.47e+09	2.66e+
	0.1349	0.003	43	275	0.000	0.129	0.1
	-1.452e+10	6.15e+09	-2	362	0.018	-2.66e+10	-2.47e+
	-1.452e+10	6.15e+09	-2	362	0.018	-2.66e+10	-2.47e+
	-1.452e+10	6.15e+09	-2	362	0.018	-2.66e+10	-2.47e+
09 is_1983	-1.452e+10	6.15e+09	-2	362	0.018	-2.66e+10	-2.47e+
09 is_1991	-1.452e+10	6.15e+09	-2	362	0.018	-2.66e+10	-2.47e+
09 is_2000	-1.452e+10	6.15e+09	-2	362	0.018	-2.66e+10	-2.47e+
09 is_1986	-1.452e+10	6.15e+09		. 362	0.018	-2.66e+10	-2.47e+
09 is_1987	-1.452e+10	6.15e+09		.362	0.018	-2.66e+10	-2.47e+
09	-1.4326+10	0.136+09	-2	. 302	0.010	-2:000+10	-2.4/6+
is_1988 09	-1.452e+10	6.15e+09	-2	362	0.018	-2.66e+10	-2.47e+
is_1989 09	-1.452e+10	6.15e+09	-2	362	0.018	-2.66e+10	-2.47e+
is_1990	-1.452e+10	6.15e+09	-2	362	0.018	-2.66e+10	-2.47e+
09 is_1992	-1.452e+10	6.15e+09	-2	362	0.018	-2.66e+10	-2.47e+
09 is_1993	-1.452e+10	6.15e+09	-2	362	0.018	-2.66e+10	-2.47e+
09 is_1994	-1.452e+10	6.15e+09	-2	362	0.018	-2.66e+10	-2.47e+

09						
is_1996 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
is_1997 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
is_1998 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
is_1999 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
is_2001 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
is_2002 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
is_2008 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
is_2009 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
is_2010 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
is_2011 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
is_2012 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
is_2013 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
is_2014 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
is_2003 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
is_2004 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
is_2005 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
is_2006 09	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
	-1.452e+10	6.15e+09	-2.362	0.018	-2.66e+10	-2.47e+
=======================================	=========	=========	======	========	=======	=======
Omnibus: 18		30146.1	L56 Dur	bin-Watson:		1.2
Prob(Omnib	us):	0.0	000 Jar	que-Bera (JB	;):	336228.3
Skew:		1.7	747 Pro	b(JB):		0.
Kurtosis:		13.0	991 Con	d. No.		6.90e+
	========	-=======		========	========	.======
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Notes:

- [1] Standard Errors assume that the covariance matrix of the errors is correctly specified.
- [2] The smallest eigenvalue is 1.57e-23. This might indicate that there are strong multicollinearity problems or that the design matrix is singular.

In [15]: print(ind_fe_model_5y.summary())

=======================================	=======	=========	=======	========	=======	======
Dep. Variable	:	rv	R-squa	red:		0.0
82 Model:		0LS	Adj. R	-squared:		0.0
82 Method:		Least Squares	F-stat:	istic:		48
5.2 Date:	Sun	, 14 Apr 2024	Proh (1	F_statistic):		0.
00	Juli	•				
Time: 4.		18:01:45	Log-Lil	kelihood:		2119
No. Observation 04	ons:	70756	AIC:		-	-4.236e+
Df Residuals:		70742	BIC:		-	-4.223e+
04 Df Model:		13				
Covariance Ty	pe:					
=======================================	=======	========	=======	========	=======	======
	coef	std err	t	P> t	[0.025	0.9
75]						
const	-0.0049	0.073	-0.067	0.947	-0.148	0.
. –	0.0573	0.004	15.569	0.000	0.050	0.
065 is_ind_3.0	0.1357	0.073	1.853	0.064	-0.008	0.
279 is_ind_10.0	0.2055	0.073	2.805	0.005	0.062	0.
349 is_ind_6.0	0.2483	0.073	3.390	0.001	0.105	0.
392						
is_ind_12.0 302	0.1583	0.073	2.161	0.031	0.015	0.
is_ind_11.0 257	0.1139	0.073	1.556	0.120	-0.030	0.
is_ind_1.0	0.1090	0.073	1.487	0.137	-0.035	0.
253 is_ind_9.0	0.1534	0.073	2.094	0.036	0.010	0.
297 is_ind_7.0	0.1715	0.073	2.340	0.019	0.028	0.
315 is_ind_2.0	0.1368	0.073	1.865	0.062	-0.007	0.
281 is_ind_5.0	0.1157	0.073	1.579	0.114	-0.028	0.
259						
is_ind_8.0 206	0.0626	0.073	0.855	0.393	-0.081	0.
is_ind_4.0 316	0.1719	0.073	2.345	0.019	0.028	0.
=======================================	========	=========	=======	=========	=======	======
Omnibus:		48570.035	Durbin-	-Watson:		1.3

Notes:

 $\[1\]$ Standard Errors assume that the covariance matrix of the errors is correctly specified.

In [16]: print(joint_fe_model_5y.summary())

=======	========	:=======	=====	=====	========	=======	=======
==							
Dep. Varia 91	ble:		rv	R-squ	ared:		0.3
Model: 91			0LS	Adj.	R-squared:		0.3
Method:		Least Squ	ares	F-sta	tistic:		116
4.					/	,	_
Date: 00	S	Sun, 14 Apr	2024	Prob	(F-statisti	c):	0.
Time:		18:0	1:45	Log-L	ikelihood:		3571
9. No. Observ	ations:	70	0 756	AIC:			-7 . 136e+
04 Df Residua	ls:	7(0716	BIC:			-7 . 099e+
04							
Df Model:			39				
		nonrol					
=======================================	========	========	=====	=====	========	=======	=======
	coef	std err		t	P> t	[0.025	0.97
5]							
	1 72/0:10	5 17 ₀₊ 10	0	225	0 727	0 410+10	1 1001
11	1.7340+10	5.17e+10	U	. 333	0.737	-0.416+10	1.196+
	0.1347	0.003	40	.717	0.000	0.128	0.1
is_1980	5.213e+10	9.85e+10	0	.530	0.596	-1.41e+11	2.45e+
	-1.265e+09	1.07e+10	-0	.119	0.905	-2.21e+10	1.96e+
	-2.149e+10	1.42e+10	-1	.516	0.129	-4.93e+10	6.29e+
09 is_1983	-8.792e+10	1.89e+10	-4	.651	0.000	-1.25e+11	-5.09e+
10	-2 497e+10	3.8e+10	_0	.658	0.511	-9.94e+10	4.94e+
10	214576110	3100110	Ū	1050	01511	31340110	413461
is_2000 11	1.321e+10	7.72e+10	0	.171	0.864	-1.38e+11	1.64e+
is_1986 11	-2.513e+10	6.75e+10	-0	.372	0.710	-1.57e+11	1.07e+
is_1987	9.591e+09	3.78e+10	0	.254	0.799	-6.44e+10	8.36e+
10 is_1988	-4.483e+10	7.38e+10	-0	.607	0.544	-1 . 9e+11	9.99e+
10 is_1989	-3.389e+10	2.02e+10	-1	.681	0.093	-7.34e+10	5.62e+
09 is_1990	2.743e+10	5.77e+10	0	. 475	0.635	-8.58e+10	1.41e+
11		- - •	·	_		- •	_ ·
is_1992 10	-1.512e+10	2.84e+10	-0	. 532	0.595	-7.09e+10	4.06e+
is_1993 10	2.307e+10	3.01e+10	0	.766	0.444	-3.6e+10	8.21e+
is_1994	1.227e+10	7.17e+10	0	.171	0.864	-1.28e+11	1.53e+

11						
11 is_1996	1.056e+10	4.59e+10	0.230	0.818	-7 . 93e+10	1e+
11 is_1997	1.251e+10	2.96e+10	0.423	0.672	-4.54e+10	7.05e+
10 is_1998 10	3.005e+10	2.54e+10	1.182	0.237	-1.98e+10	7 . 99e+
is_1999 11	5.959e+10	5.03e+10	1.185	0.236	-3.9e+10	1.58e+
is_2001 11	1.368e+11	1.42e+11	0.964	0.335	-1.41e+11	4.15e+
is_2002 10	-7.705e+10	5.52e+10	-1.395	0.163	-1.85e+11	3.12e+
is_2008 11	1.915e+10	7.46e+10	0.257	0.797	-1.27e+11	1.65e+
is_2009 10	-6.881e+10	7.84e+10	-0.878	0.380	-2.22e+11	8.49e+
is_2010 11	2.621e+10	1.6e+11	0.164	0.870	-2.87e+11	3.4e+
is_2011 10	-9.4e+09	2.58e+10	-0.364	0.716	-6e+10	4.12e+
is_2012 11	7.275e+09	1.39e+11	0.052	0.958	-2.66e+11	2.8e+
is_2013 11	1.096e+10	1.18e+11	0.093	0.926	-2.21e+11	2.43e+
is_2014 11	-1.452e+11	2.32e+11	-0.626	0.531	-6e+11	3.09e+
is_2003 11	-1.845e+11	2.1e+11	-0.879	0.380	-5.96e+11	2.27e+
is_2004 10	-6.672e+10	7.4e+10	-0.902	0.367	-2.12e+11	7.83e+
is_2005 11	3.413e+11	1.82e+11	1.876	0.061	-1.53e+10	6.98e+
is_2006 11	-9.533e+09	1.44e+11	-0.066	0.947	-2.92e+11	2.73e+
is_2007 11	-1.546e+11	1.59e+11	-0.975	0.330	-4.65e+11	1.56e+
is_1995 11	-1.374e+11	1.94e+11	-0.710	0.478	-5.17e+11	2.42e+
is_1984 10	-1.611e+10	2.77e+10	-0.581	0.561	-7.05e+10	3.83e+
is_1985 10	-8.413e+09	2.59e+10	-0.325	0.745	-5.91e+10	4.23e+
is_1980 11	-6.948e+10	9.63e+10	-0.721	0.471	-2.58e+11	1.19e+
is_1981 10	-1.608e+10	4.98e+10	-0.323	0.747	-1.14e+11	8.14e+
is_1982 10	4.144e+09	4.65e+10	0.089	0.929	-8.7e+10	9.53e+
is_1983 11	7.058e+10	4.96e+10	1.423	0.155	-2.67e+10	1.68e+
is_1991 11	7.627e+09	5.23e+10	0.146	0.884	-9.49e+10	1.1e+
is_2000 11	-3.055e+10	8.03e+10	-0.380	0.704	-1.88e+11	1.27e+
is_1986	7.786e+09	7.48e+10	0.104	0.917	-1.39e+11	1.54e+

11 is_1987	-2.693e+10	3.58e+10	-0.753	0.451	-9.7e+10	4.32e+
10 is_1988 11	2.749e+10	9.89e+10	0.278	0.781	-1.66e+11	2.21e+
is_1989 11	1.655e+10	5.47e+10	0.302	0.762	-9.07e+10	1.24e+
is_1990 10	-4.477e+10	6.78e+10	-0.660	0.509	-1.78e+11	8.82e+
is_1992 11	-2.221e+09	5.45e+10	-0.041	0.967	-1.09e+11	1.05e+
is_1993 10	-4.041e+10	5.4e+10	-0.748	0.455	-1.46e+11	6.55e+
is_1994 11	-2.961e+10	9.06e+10	-0.327	0.744	-2.07e+11	1.48e+
is_1996 11	-2.79e+10	6.88e+10	-0.406	0.685	-1.63e+11	1.07e+
is_1997 10	-2.985e+10	5.89e+10	-0.506	0.613	-1.45e+11	8.57e+
is_1998 10	-4.74e+10	5.36e+10	-0.884	0.377	-1.52e+11	5.76e+
is_1999 10	-7.693e+10	6.41e+10	-1.201	0.230	-2.02e+11	4.86e+
is_2001 11	-1.542e+11	1.46e+11	-1.056	0.291	-4.4e+11	1.32e+
is_2002 11	5.971e+10	6.89e+10	0.867	0.386	-7.54e+10	1.95e+
is_2008 11	-3.65e+10	8.57e+10	-0.426	0.670	-2.04e+11	1.31e+
is_2009 11	5.146e+10	8.99e+10	0.572	0.567	-1.25e+11	2.28e+
is_2010 11	-4.355e+10	1.61e+11	-0.270	0.787	-3.6e+11	2.73e+
is_2011 10	-7 . 942e+09	2.6e+10	-0.305	0.760	-5.89e+10	4.3e+
is_2012 11	-2.462e+10	1.42e+11	-0.174	0.862	-3.02e+11	2.53e+
is_2013 11	-2.83e+10	1.23e+11	-0.230	0.818	-2.69e+11	2.13e+
is_2014 11	1.279e+11	2.33e+11	0.550	0.582	-3.28e+11	5.84e+
is_2003 11	1.672e+11	2.11e+11	0.793	0.428	-2.46e+11	5.8e+
is_2004 11	4.937e+10	8.56e+10	0.577	0.564	-1.18e+11	2.17e+
is_2005 08	-3.586e+11	1.83e+11	-1.965	0.049	-7.16e+11	-8.81e+
is_2006 11	-7.81e+09	1.46e+11	-0.053	0.957	-2.94e+11	2.79e+
is_2007 11	1.372e+11	1.6e+11	0.858	0.391	-1.76e+11	4.51e+
is_1995 11	1.2e+11	1.94e+11	0.618	0.536	-2.61e+11	5.01e+
is_1984 10	-1.235e+09	2.87e+10	-0.043	0.966	-5.75e+10	5.5e+
is_1985	-8.929e+09	2.59e+10	-0.345	0.730	-5.97e+10	4.18e+

==			
Omnibus:	30186.031	Durbin-Watson:	1.2
18			
<pre>Prob(Omnibus):</pre>	0.000	Jarque-Bera (JB):	336709.2
88			
Skew:	1.750	<pre>Prob(JB):</pre>	0.
00			
Kurtosis:	13.097	Cond. No.	2.50e+
17			
	:=========		=======================================

Notes:

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- [1] Standard Errors assume that the covariance matrix of the errors is correctly specified.
- [2] The smallest eigenvalue is 1.23e-30. This might indicate that there are strong multicollinearity problems or that the design matrix is singular.

As we can see the clustering effects of volatility still translate into a statistically significant 13-14% 5y autocorrelation as demonstrated by these various regressions.

3. What are the benefits of the panel approach, versus simply running one regression for

each firm? What are the potential costs?

The more variables you add to the regression, the more risk of multi-colinearity (or just generally losing track of causality) you introduce. You also could start having colinearity between the effects of individual explanatory variables, reducing the explanatory power of the individual betas.

The positive effects on the other hand are that these panel regressions can allow the individual beta's of each variable to store exclusively the information related to that variable by assigning the relative noise of external effects to their respective variables. In our case for example, asigning industry and mean-annual effects to panel variables allowed us to extract a more expressive value for the autocorrelation of variances.

It thus comes down to a judgement call when running the regression to find the optimal linearly independent set of risk factors to which the regression will "assign" linearly independent risk-factors in-line with those we are trying to explain.