

$$\underline{Q2} \quad \phi(y, w_k^T x + b_k, \sigma_k^2) = \frac{1}{\sqrt{2\pi}} \times \frac{1}{\sigma_k^2} \times \exp\left(-\frac{(y - w_k^T x - b_k)^2}{2\sigma_k^2}\right)$$

where $y \in \mathbb{R}$ & $w_k^T x + b_k \in \mathbb{R}$.

Let $w_k^T x + b_k = \mu_k$.

$$\therefore \phi(y, \mu_k, \sigma_k^2) = \frac{1}{\sqrt{2\pi}} \times \frac{1}{\sigma_k^2} \times \exp\left(-\frac{(y - \mu_k)^2}{2\sigma_k^2}\right)$$

$$\therefore f(y|x; \theta) = \sum_{k=1}^K \epsilon_k \phi(y, \mu_k, \sigma_k^2)$$

$$\& \epsilon_k \geq 0 \& \sum \epsilon_k = 1$$

Let $S \in \{1, \dots, K\}$ be a discrete RV such that

$$\text{Pr}\{S = k\} = \epsilon_k$$

$$\therefore \text{Pr}(Y \in A) = \sum_{k=1}^K \left(\int_A \phi(y; \mu_k, \sigma_k^2) \times \epsilon_k dy \right)$$

Let $\underline{y} = (y_1, \dots, y_n)$ & $\underline{x} = (x_1, \dots, x_n)$ \therefore the likelihood is

$$L(\theta; \underline{y} | \underline{x}) = \prod_{i=1}^n f(y_i | x_i; \theta) = \prod_{i=1}^n \sum_{k=1}^K \epsilon_k \phi(y_i; \mu_k, \sigma_k^2)$$

\therefore log likelihood is.

$$l(\theta; \mathbf{y} | \mathbf{x}) = \sum_{i=1}^n \log \left(\sum_{k=1}^K e_k \phi(y_i; \mu_k, \sigma_k^2) \right)$$

Let $\underline{s} = (s_1 \dots s_n)$

$\underline{z} = (\underline{y}, \underline{x}, \underline{s})$. Also, let

$$\Delta_{ik} = \begin{cases} 1 & \text{if } s_i = k \\ 0 & \text{if } s_i \neq k \end{cases}$$

\therefore The complete data log likelihood is,

$$\log l(\theta; \underline{y} | \underline{x}, \underline{s}) = \sum_{i=1}^n \sum_{k=1}^K \Delta_{ik} \left[\log \theta_k + \log \phi(y_i; \mu_k, \sigma_k^2) \right]$$

Now, looking at E-step, we get.

$$Q(\theta, \theta^{(j)}) = E_{\underline{s} | \underline{y}, \underline{x}} \left[l(\theta; \underline{y} | \underline{x}, \underline{s}) | \underline{y} | \underline{x}; \theta^{(j)} \right]$$

$$\therefore Q(\theta, \theta^{(j)}) = \sum_{i=1}^n \sum_{k=1}^K \gamma_{i,k}^{(j)} \left[\log \theta_k + \log \phi(y_i; \mu_k, \sigma_k^2) \right]$$

$$\& \gamma_{i,k}^{(j)} = \frac{e_k^{(j)} \cdot \phi(y_i; \mu_k^{(j)}, \sigma_k^{(j)})}{\sum_{l=1}^K e_l^{(j)} \phi(y_i; \mu_l^{(j)}, \sigma_l^{(j)})}$$

\therefore In terms of $\theta, \theta^{(j)}$ & the data for $Q(\theta, \theta^{(j)})$:

$$\log l(\theta; \underline{y} | \underline{x}, \underline{s}) = \sum_{i=1}^n \sum_{k=1}^K \Delta_{i,k} \left[\log \epsilon_k + \log \phi(y_i, \omega_k^\top x_i + b_k, \sigma_k^2) \right]$$

$$\& Q(\theta, \theta^{(j)}) = \sum_{i=1}^n \sum_{k=1}^K \gamma_{i,k}^{(j)} \left[\log \epsilon_k + \log \phi(y_i, \omega_k^\top x_i + b_k, \sigma_k^2) \right]$$

$$\& \gamma_{(i,k)}^{(j)} = \frac{\epsilon_k^{(j)} \phi(y_i, \omega_k^\top x_i + b_k, \sigma_k^2)}{\sum_{d=1}^K \epsilon_d^{(j)} \phi(y_i, \omega_d^\top x_i + b_d, \sigma_d^2)}$$

M step:

$$\begin{aligned} \theta^{(j+1)} &= \arg \max_{\theta} Q(\theta, \theta^{(j)}) \\ &= \sum_{i=1}^n \sum_{k=1}^K \gamma_{i,k}^{(j)} \left[\log \phi(y_i, \omega_k^\top x_i + b_k, \sigma_k^2) + \log \epsilon_k \right] \end{aligned}$$

$$\log \phi(y_i, \omega_k^\top x_i + b_k, \sigma_k^2) =$$

$$\log \left(\frac{1}{\sqrt{2\pi}} \times \frac{1}{\sqrt{\sigma_k^2}} \times \exp \left(- \frac{(y_i - \omega_k^\top x_i - b_k)^2}{2\sigma_k^2} \right) \right)$$

$$= \underbrace{\log\left(\frac{1}{\sqrt{2\pi}}\right)}_C - \frac{1}{2} \log(\sigma_k^2) - \frac{(y_i - \hat{w}_k^T x_i - b_k)^2}{2\sigma_k^2}$$

$$= \frac{-1}{2} \log(\sigma_k^2) - \frac{(y_i - \hat{w}_k^T x_i - b_k)^2}{2\sigma_k^2}$$

$$\therefore Q(\theta, \theta^{(j)}) = \sum_{i=1}^n \sum_{k=1}^K \gamma_{i,k}^{(j)} \left[\log \epsilon_k - \frac{1}{2} \log(\sigma_k^2) - \frac{(y_i - \hat{w}_k^T x_i - b_k)^2}{2\sigma_k^2} \right]$$

$$\therefore \frac{\partial Q(\theta, \theta^{(j)})}{\partial b_k} = \sum_{i=1}^n \gamma_{i,k}^{(j)} \times \frac{1}{\sigma_k^2} \times (y_i - \hat{w}_k^T x_i - b_k) = 0$$

$$\text{or } \sum y_i - \hat{w}_k^T \sum x_i = b_k$$

$$\text{or } b_k = \bar{y} - \hat{w}_k^T \bar{x}$$

$$\frac{\partial Q(\theta, \theta^{(j)})}{\partial \hat{w}_k} =$$

$$\frac{\partial Q(\theta, \theta^{(j)})}{\partial b_k} = \sum_{i=1}^n \gamma_{i,k}^{(j)} \left[\frac{1}{\sigma_k^2} \times (y_i - \hat{w}_k^T x_i - b_k) \right] = 0$$

$$\Rightarrow b_k = \frac{\sum \gamma_{i,k}^{(j)} y_i}{\sum \gamma_{i,k}^{(j)}} - \hat{w}_k^T \frac{\sum \gamma_{i,k}^{(j)} x_i}{\sum \gamma_{i,k}^{(j)}} = \bar{y} - \hat{w}_k^T \bar{x}$$

Similarly,

$$\frac{\partial \mathcal{O}(\theta, \theta^{(j)})}{\partial \omega_k} = \frac{\partial}{\partial \omega_k} \left(\sum_{i=1}^n \frac{\gamma_{i,k}^{(j)}}{2\sigma_k^2} \left[- \frac{(y_i - \hat{\omega}_k x_i - \frac{\bar{y}_y - \hat{\omega}_k \bar{x}_y}{\gamma_{i,k}^{(j)}})^2}{\gamma_{i,k}^{(j)}} \right] \right)$$

$$= \frac{\partial}{\partial \omega_k} \left(\sum_{i=1}^n - \frac{\gamma_{i,k}^{(j)}}{2\sigma_k^2} \left[\frac{y_i - \bar{y}_y - \hat{\omega}_k (\frac{\gamma_{i,k}^{(j)}}{\gamma_{i,k}^{(j)}} (x_i - \bar{x}_y))}{\gamma_{i,k}^{(j)}} \right]^2 \right)$$

$$= + \frac{\gamma_{i,k}^{(j)}}{\sigma_k^2} \left(y_i - \hat{\omega}_k x_i - \frac{\bar{y}_y - \hat{\omega}_k \bar{x}_y}{\gamma_{i,k}^{(j)}} \right) \times \left(\frac{1}{\gamma_{i,k}^{(j)}} \right)$$

$$= \frac{\partial}{\partial \omega_k} \left(\sum_{i=1}^n - \frac{\gamma_{i,k}^{(j)}}{2\sigma_k^2 \gamma_{i,k}^{(j)}} \left[\gamma_{i,k}^{(j)} \right] \right)$$

Let $(\gamma_{i,k}^{(j)} y_i - \bar{y}_y) / \sqrt{\gamma_{i,k}^{(j)}} = y_{\alpha i}$

& $(\gamma_{i,k}^{(j)} x_i - \bar{x}_y) / \sqrt{\gamma_{i,k}^{(j)}} = x_{\alpha i}$

$$= \frac{\partial}{\partial \omega_k} \left(\sum_{i=1}^n - \frac{1}{2\sigma_k^2 \gamma_{i,k}^{(j)}} \left[(y_{\alpha i} - \hat{\omega}_k x_{\alpha i})^2 \right] \right)$$

$$= \frac{\partial}{\partial \omega_k} \left(\sum_{i=1}^n - \frac{1}{2\sigma_k^2} (y_{\alpha i} - \hat{\omega}_k x_{\alpha i})^2 \right)$$

$$= \frac{\partial}{\partial \omega_k} \left[\sum_{i=1}^n (y_{\alpha i} - \hat{\omega}_k x_{\alpha i})^2 \right]$$

$$\Rightarrow \omega_k = (X_{\alpha}^T X_{\alpha})^{-1} X_{\alpha}^T Y_{\alpha}$$

where $X_{\alpha} = \begin{bmatrix} x_{\alpha 1} \\ \vdots \\ x_{\alpha n} \end{bmatrix}$ $Y_{\alpha} = \begin{bmatrix} y_{\alpha 1} \\ \vdots \\ y_{\alpha n} \end{bmatrix}$

$$\frac{\partial Q(\theta, \theta^{(j)})}{\partial \sigma_k^2} = \sum_{i=1}^n \left(-\frac{1}{2} \times \frac{1}{\sigma_k^2} + \frac{(y_i - w_k^T x_i - b_k)^2}{2(\sigma_k^2)^3} \right) \gamma_{i,k}^{(j)} = 0$$

$$\text{or } -\frac{n}{\sigma_k^2} + \frac{\sum \gamma_{i,k}^{(j)} (y_i - w_k^T x_i - b_k)^2}{(\sigma_k^2)^3} = 0$$

$$\text{If } \sigma_k^2 \neq 0,$$

$$\Rightarrow \sigma_k^2 \times n = \sum \gamma_{i,k}^{(j)} (y_i - w_k^T x_i - b_k)^2$$

$$\text{or } \sigma_k^2 = \frac{1}{n} \sum \gamma_{i,k}^{(j)} (y_i - w_k^T x_i - b_k)^2$$

Now, for ϵ_k

$$\text{Let } J(\theta, \theta^{(j)}) = \sum_{i=1}^n \sum_{k=1}^K \gamma_{i,k}^{(j)} \log(\epsilon_k) + \beta \left(\sum_{k=1}^K \epsilon_k - 1 \right)$$

$$\frac{\partial J}{\partial \epsilon_k} = \sum_{i=1}^n \frac{\gamma_{i,k}^{(j)}}{\epsilon_k} + \beta = 0$$

$$\therefore \epsilon_k = \frac{\sum \gamma_{i,k}^{(j)}}{-\beta}$$

Now, using

$$\sum_{k=1}^K \epsilon_k = 1, \text{ we get}$$

$$-\beta = \sum_{i=1}^n \sum_{k=1}^K \gamma_{i,k}^{(j)}$$

$$-\beta = n$$

$$\therefore \epsilon_k = \frac{\sum_{i=1}^n \gamma_{i,k}^{(j)}}{n}$$

Solving for b_k & w_k

Let $x_i = \begin{bmatrix} 1 & x_i \end{bmatrix}$
 $\beta_k = \begin{bmatrix} b_k & w_k \end{bmatrix}$

$$\therefore \frac{\partial Q(0,0^{(j)})}{\partial \beta_k} = \frac{\partial}{\partial \beta_k} \left(\sum_{i=1}^n \frac{\gamma_{i,k}^{(j)}}{\sigma_k^2} (y_i - \beta_k^T x_i)^2 \right)$$

$$= \frac{\partial}{\partial \beta_k} \sum_{i=1}^n \gamma_{i,k}^{(j)} (y_i - \beta_k^T x_i)^2$$

$$\therefore \beta_k = (X^T W_k^* X)^{-1} X^T W_k^* y$$

where $W_k^* = \begin{bmatrix} \gamma_{1,k}^{(j)} & & \\ & \gamma_{2,k}^{(j)} & \\ & & \ddots \\ & & & \gamma_{n,k}^{(j)} \end{bmatrix}$

d) The # iteration = 10

$b =$	$[-0.0192 \quad 0.3160]$
$w =$	$[0.5328 \quad -1.3234]$
$V =$	$[3.7 \times 10^{-11} \quad 0.3468]$
$e =$	$[3.2 \times 10^{-5} \quad 1]$

ω^* /

Q1 a) # iterations = 225

The portions which are present at a lot of pixels in the image & there there's not sudden change in intensity.

~~Comp~~ Compression ratio = 0.0707

R M Abs. error = 0.0232

$$b) \text{ app} = \frac{k \times m^2 \times 24 + \log_2 k \times N^2 / m^2}{N^2}$$

$$\text{Compression ratio} = \frac{k \times m^2 \times 24 + \log_2 k \times N^2 / m^2}{N^2 \times 24}$$

=


```

clc;
clear;

% load image
y = imread('mandrill.tiff');
% extract blocks
M = 2; % block side-length
k = 100;
n = numel(y)/(3*M*M); % number of blocks
d = size(y,1); % image length/width

c=0; % counter
x = zeros(n,3*M*M);
for i=1:M:d % loop through blocks
    for j=1:M:d
        c = c+1;
        x(c,:) = reshape(y(i:i+M-1,j:j+M-1,:), [1,M*M*3]);
    end
end

% K means algorithm
rng(0);
perm = randperm(n);
centroids = x(perm(1:k), :);
obj_vec = zeros(500, 1);
for iter = 1:500
    dist_frm_centroids = dist2(x, centroids);
    [min_dist, cluster_map] = min(dist_frm_centroids, [], 2);
    obj = mean(min_dist);
    obj_vec(iter) = obj;
    centroids_old = centroids;
    centroids = zeros(100, 12);
    for i = 1:k
        centroids(i, :) = mean(x(find(cluster_map == i),:));
    end
    if norm(centroids - centroids_old) < 0.01
        break
    end
end

obj_vec = obj_vec(1:iter);
plot(1:iter, obj_vec)

x_new = centroids(cluster_map,:);

y_new = zeros(512, 512, 3);
c = 0;
for i=1:M:d % loop through blocks
    for j=1:M:d
        c = c+1;
        y_new(i:i+M-1,j:j+M-1,:) = reshape(x_new(c, :), M, M, 3);
    end
end

y_new = uint8(y_new);

subplot(1,2, 1)

```



```

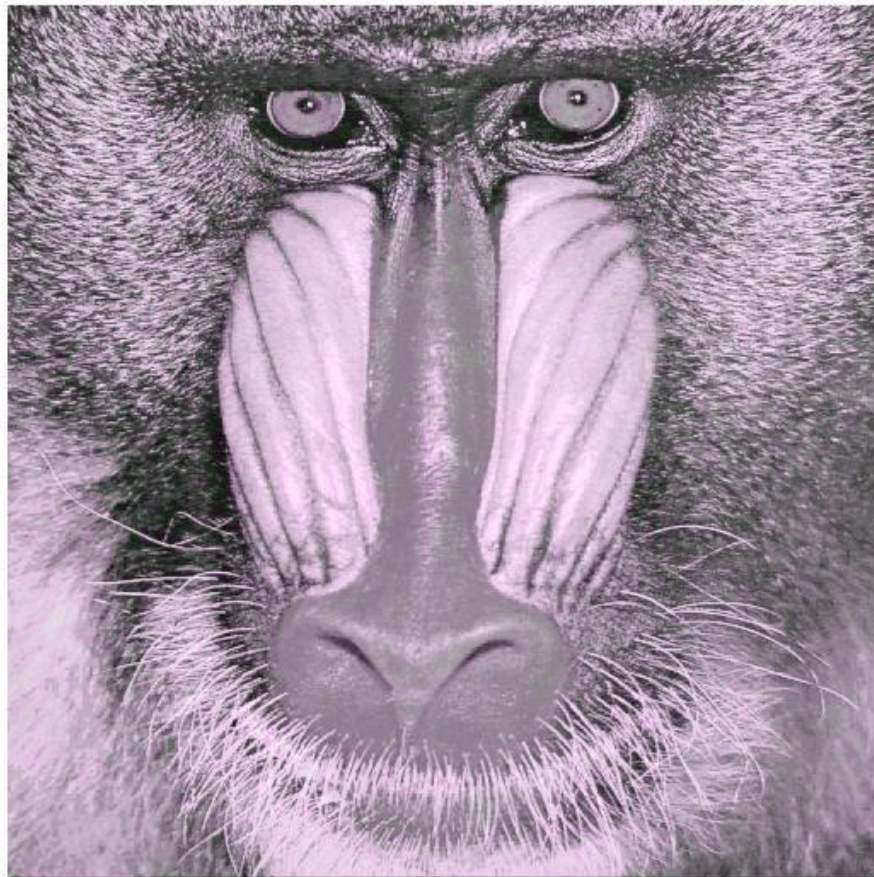
imagesc(y)
subplot(1,2,2)
imagesc(y_new)
drawnow

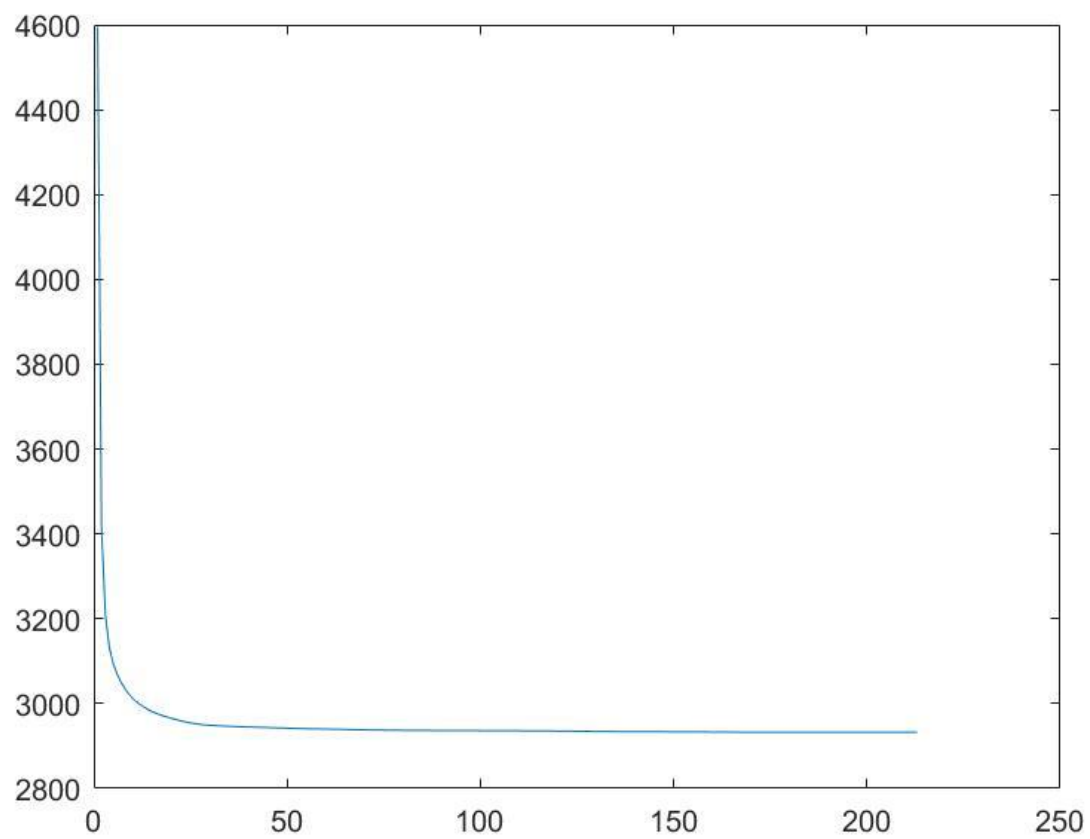
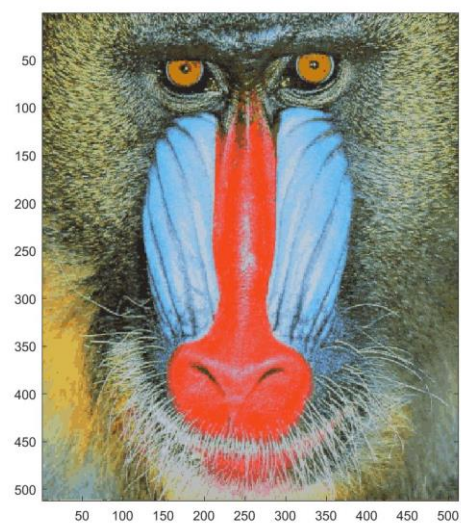
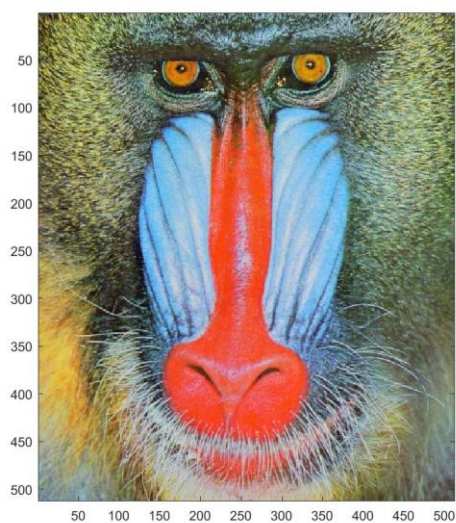
subplot(1, 1, 1)
imshowpair(y,y_new);

rel_mean_abs_error = sum(abs(y_new(:) - y(:))) / (3 * 512 * 512 * 256);
% 0.0232

N = 512;
compression_ratio = (k * M*M*24 + log2(k) * (N*N)/(M*M)) / (N*N*24)

```






```

clear all
close all
rng(0);
n = 200; % sample size
K = 2; % number of lines
e = [.7 .3]; % mixing weights
w = [-2 1]; % slopes of lines
b = [.5 -.5]; % offsets of lines
v = [.2 .1]; % variances
for i=1:n
x(i) = rand;
if rand < e(1);
y(i) = w(1)*x(i) + b(1) + randn*sqrt(v(1));
else
y(i) = w(2)*x(i) + b(2) + randn*sqrt(v(2));
end
end
plot(x,y,'bo')
hold on
t=0:0.01:1;
plot(t,w(1)*t+b(1),'k')
plot(t,w(2)*t+b(2),'k')

% EM algorithm
e = [.5 .5];
w = [1 -1]; % slopes of lines
b = [0 0]; % offsets of lines
v = repmat(var(y), 1, 2); % variances
log_likelihood = zeros(500, 1);
for iter = 1:10
    u = w' * x + b' * ones(n, 1)';
    w_mat = repmat(w', 1, 200);
    x_mat = vertcat(x, x);
    y_mat = vertcat(y, y);
    v_mat = v' * ones(1, n);

    gaussian_pdf = normpdf(y_mat, u, sqrt(v_mat));
    e_mat = repmat(e', 1, 200);

```



```

    phi_num = e_mat .* gaussian_pdf;
    phi_dem = sum(phi_num, 1);
    phi = phi_num ./ repmat(phi_dem, 2, 1);
    log_likelihood(iter,:) =
sum(log(sum(gaussian_pdf .* e_mat,1)));
    if iter ~= 1
        if abs(log_likelihood(iter-1,:) -
log_likelihood(iter,:)) <= (10 ^ (-4))
            break
        end
    end
    x_new = [ones(1,n); x];
    phi_diag_1 = diag(phi(1,:));
    a1 = (inv(x_new * phi_diag_1 * x_new') * x_new
* phi_diag_1 * y')';

    phi_diag_2 = diag(phi(2,:));
    a2 = (inv(x_new * phi_diag_2 * x_new') * x_new
* phi_diag_2 * y')';

    b = [a1(1) a2(1)];
    w = [a1(2) a2(2)];
    u = w' * x + b' * ones(n, 1)';
    v = sum(phi .* ((y_mat - u) .^ 2), 2)' / n;
    e = sum(phi, 2)' / n;
end

hold on
plot(t,w(1)*t+b(1),':')
plot(t,w(2)*t+b(2),':')

hold off
subplot(1, 1, 1);
plot(1:iter, log_likelihood(1:iter,:))

```