

MA 3140: Statistical Inference

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Method 3: Conditioning

Example 1: Revisiting Poisson Distribution

Let $X_1, \dots, X_n \sim P(\theta)$. Suppose we are interested in finding the UMVUE for $g(\theta) = e^{-\theta}$.

Solution: We know that $T = \sum X_i \sim P(n\theta)$ is a CSS.

Let

$$\delta(X_1) = \begin{cases} 1, & \text{if } X_1 = 0 \\ 0, & \text{if } X_1 \neq 0 \end{cases}$$

Since $E[\delta(X_1)] = P(X_1 = 0) = e^{-\theta}$, we find that $\delta(X_1)$ is unbiased for $g(\theta)$. Consider

Method 3: Conditioning cont'd

$$\begin{aligned}\eta(t) &= E[\delta(X_1)|T = t] = 1.P(X_1 = 0|T = t) + 0.P(X_1 \neq 0|T = t) \\&= P(X_1 = 0 | \sum_{i=1}^n X_i = t) \\&= \frac{P(X_1 = 0, \sum_{i=2}^n X_i = t)}{P(\sum_{i=1}^n X_i = t)} \\&= \frac{P(X_1 = 0)P(\sum_{i=2}^n X_i = t)}{P(\sum_{i=1}^n X_i = t)} \\&= \frac{e^{-\theta} \frac{e^{-(n-1)\theta} [(n-1)\theta]^t}{t!}}{\frac{e^{-n\theta} (n\theta)^t}{t!}} = \left(1 - \frac{1}{n}\right)^t\end{aligned}$$

So, $\left(1 - \frac{1}{n}\right)^T$ is UMVUE for $e^{-\theta}$.

Method 3: Conditioning cont'd

Example 2: Hypergeometric Distribution

Let X have Hypergeometric distribution with pmf

$$P(X = x) = \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{n}}, \quad x = 0, 1, \dots, n, x \leq M, n-x \leq N-M,$$

where N is known and M is unknown. Suppose you are interested in estimating M .

Solution: Here, we know that X is sufficient.

To check completeness of X , let $E[g(X)] = 0$.

$$\Rightarrow \sum_{x=0}^n g(x) \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{n}} = 0, \quad \forall M$$

Method 3: Conditioning cont'd

Now,

when $M = 0 \implies g(0) = 0$.

when $M = 1 \implies g(0)\binom{N-1}{n} + g(1)\binom{N-1}{n-1} = 0 \implies g(1) = 0$.

By induction, it can be shown that X is complete.

$$EX = \frac{n}{N}M \implies E\left(\frac{N}{n}X\right) = M$$

Thus, we conclude that $\frac{N}{n}X$ is UMVUE of M .

Method 3: Conditioning cont'd

Example 3: Exponential UMVUE

Let $X_1, \dots, X_n \sim \frac{1}{\lambda} e^{-\frac{x}{\lambda}}, \lambda > 0$.

Suppose we are interested to find the UMVUE of the reliability function $e^{-m/\lambda}$.

Solution: Define

$$\delta(X_1) = \begin{cases} 1, & \text{if } X_1 > m \\ 0, & \text{otherwise} \end{cases}$$

Thus, $E[\delta(X_1)] = e^{-m/\lambda}$.

We also know that $T = \sum X_i$ or \bar{X} is CSS.

Method 3: Conditioning cont'd

Now,

$$\begin{aligned}\eta(t) &= E[\delta(X_1)|T] = 1.P[X_1 > m|\bar{X}] + 0 \\&= P\left[\frac{X_1}{\bar{X}} > \frac{m}{\bar{X}}|\bar{X}\right] \\&= P\left[\frac{X_1}{\bar{X}} > \frac{m}{\bar{X}}\right] \quad \left(\because \frac{X_1}{\bar{X}} \text{ is ancillary}\right) \\&= P\left[\frac{X_1}{\frac{1}{n}\sum X_i} > \frac{m}{\frac{1}{n}\sum x_i}\right] \\&= P\left[\frac{X_1}{X_1 + \sum_{i=2}^n X_i} > \frac{m}{x_1 + \sum_{i=2}^n x_i}\right] \\&= P\left[Z > \frac{m}{n\bar{X}}\right],\end{aligned}$$

where $Z = \frac{X_1}{X_1 + \sum_{i=2}^n X_i} \sim Be(1, n-1)$.

Method 3: Conditioning cont'd

$$\begin{aligned}P\left[Z > \frac{m}{n\bar{X}}\right] &= \frac{1}{Be(1, n-1)} \int_{\frac{m}{n\bar{X}}}^1 (1-z)^{n-2} z^{1-1} dz \\&= \frac{\Gamma_n}{\Gamma_1 \Gamma(n-1)} \int_{\frac{m}{n\bar{X}}}^1 (1-z)^{n-2} dz \\&= \left(1 - \frac{m}{n\bar{X}}\right)^{n-1}\end{aligned}$$

So, $\left(1 - \frac{m}{n\bar{X}}\right)^{n-1}$ is UMVUE for $e^{-m/\lambda}$.

Method 3: Conditioning cont'd

- **Beta Distribution:** An absolutely continuous random variable X is said to follow Beta distribution with shape parameters $a > 0$ and $b > 0$ (written as $X \sim Be(a, b)$) if its pdf is given by

$$f_X(x) = \begin{cases} \frac{x^{a-1} (1-x)^{b-1}}{Be(a,b)}, & \text{if } 0 < x < 1 \\ 0, & \text{otherwise} \end{cases}$$

- Consider two independent random variables X and Y .
 - If $X \sim Gamma(a, b_1)$ and $Y \sim Gamma(a, b_2)$ then $\frac{X}{X+Y} \sim Be(b_1, b_2)$.
 - If $X \sim Gamma(a_1, b)$ and $Y \sim Gamma(a_2, b)$ then $\frac{X}{X+Y} \sim Be(a_1, a_2)$.

Thanks for your patience!