x	1	2	3	4
$f_X(x)$	0.19	0.32	0.31	0.18
$f_{X Y}(x 1)$	10/19	5/19	2/19	2/19

y	1	2	3	4
$f_Y(y)$	0.19	0.32	0.31	0.18
$f_{Y X}(y 1)$	10/19	5/19	2/19	2/19

Table 1: Information for question 2.

Homework 2

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Question 1 Let X have a pdf of $f(x) = cx^2$ for $0 \le x \le 1$ and f(x) = 0 elsewhere.

(a) For this to be a valid pdf, it must integrate to 1 over the support. So

$$\int_0^1 cx^2 \, dx = \frac{cx^3}{3} \Big|_0^1 = \frac{c}{3} \stackrel{\text{set}}{=} 1,$$

which leads to c = 3. So $f(x) = 3x^2$ for $0 \le x \le 1$.

(b) The cdf is given by

$$F(x) = \int_{-\infty}^{x} f(t) dt = \int_{0}^{x} 3t^{2} dt = t^{3} \Big|_{0}^{x} = x^{3}$$

for $0 \le x \le 1$. We also have F(x) = 0 when x < 0 and F(x) = 1 when x > 1.

(c) We have

$$\Pr\left(\frac{1}{10} \le X \le \frac{1}{2}\right) = F\left(\frac{1}{2}\right) - F\left(\frac{1}{10}\right) = \frac{1}{2^3} - \frac{1}{10^3} = \frac{31}{250}.$$

Question 2 Two discrete random variables X and Y are jointly distributed.

- (a) The marginal pmf for X is obtained by summing over every value of Y for each value of X. For example, to find the marginal probability that X = 1, we have $f_X(x) = 0.10 + 0.05 + 0.02 + 0.02 = 0.19$. The other values are obtained in the same way. Finding the marginal pmf for Y is done the exact same way, and it turns out that $f_X(j) = f_Y(j)$ for $j = \{1, 2, 3, 4\}$; they are both found in Table 1.
- (b) X and Y are not independent. For two random variables to be independent, we need $f(x,y) = f_X(x) \cdot f_Y(y)$ for all possible (x,y) pairs. Here we have f(1,1) = 0.10 and $f_X(1) \cdot f_Y(1) = 0.19^2 \neq f(1,1)$, meaning X and Y are dependent.
- (c) To find the conditional pmf of X given that Y = 1, we take each value of f(x,1) and divide by $f_Y(1)$, i.e. $f_{X|Y}(x|1) = f(x,1)/f_Y(1)$. For example, we have $f_{X|Y}(1|1) = 0.10/0.19 = 10/19$. Finding the conditional pmf for Y given that X = 1 is done in a similar way, and again they are the same. Both pmfs can be found in be found in Table 1.

Question 3 We are considering points (x, y) uniformly selected within an ellipse given by the equation $(x/a)^2 + (y/b)^2 = 1$, where a, b > 0. Therefore, the probability of selecting a point (x, y) from this region is f(x, y) = c for $-a \le x \le a$, $-b \le y \le b$, and $(x/a)^2 + (y/b)^2 \le 1$, and f(x, y) = 0 elsewhere. To find c, we use the fact that a joint pdf must integrate to 1 over all values of x and y, so

$$1 = \int_{-a}^{a} \int_{-b\sqrt{1 - (x/a)^2}}^{b\sqrt{1 - (x/a)^2}} c \, dy \, dx = c \cdot 2 \int_{-a}^{a} b\sqrt{1 - (x/a)^2} \, dx = c \cdot \pi ab,$$

which implies that $c = 1/\pi ab$. The last equality is from the fact that the area of an ellipse is πab , which is what the integral is computing. To find the marginal density of X(Y), we integrate over all possible values of Y(X):

$$f_X(x) = \int_{-b\sqrt{1 - (x/a)^2}}^{b\sqrt{1 - (x/a)^2}} \frac{1}{\pi ab} \, \mathrm{d}y = \frac{y}{\pi ab} \Big|_{-b\sqrt{1 - (x/a)^2}}^{b\sqrt{1 - (x/a)^2}} = \frac{2\sqrt{1 - (x/a)^2}}{\pi a} \quad \text{for } -a \le x \le a,$$

$$f_Y(y) = \int_{-a\sqrt{1 - (y/b)^2}}^{a\sqrt{1 - (y/b)^2}} \frac{1}{\pi ab} \, \mathrm{d}x = \frac{x}{\pi ab} \Big|_{-a\sqrt{1 - (y/b)^2}}^{a\sqrt{1 - (y/b)^2}} = \frac{2\sqrt{1 - (y/b)^2}}{\pi b} \quad \text{for } -b \le y \le b.$$

These results make sense intuitively. For $f_X(x)$, we can see that the highest probability is obtained when x=0. At x=0, the height of the ellipse is the greatest, so there is more "width" for x=0 to be chosen. And as $x \to \pm a$, $f_X(x) \to 0$, meaning that the closer you get to the end of the ellipse, the less likely that point is to be chosen.

- **Question 4** The joint cdf of X and Y is given by $F(x,y) = (1 e^{-\alpha x})(1 e^{-\beta y})$ for $x,y \ge 0$ and F(x,y) = 0 elsewhere, where $\alpha, \beta > 0$ are fixed constants.
 - (a) X and Y are independent. This is because we can express the joing cdf as $F(x,y) = G(x) \cdot H(y)$, where $G(x) = 1 e^{-\alpha x}$ and $H(y) = 1 e^{-\beta y}$.
 - (b) The joint pdf is given by

$$f(x,y) = \frac{\partial^2 F}{\partial x \partial y} = \frac{\partial}{\partial x} (1 - e^{-\alpha x}) \cdot \frac{\partial}{\partial y} (1 - e^{-\beta y}) = \alpha e^{-\alpha x} \cdot \beta e^{-\beta y}$$

for $x, y \ge 0$ and f(x, y) = 0 elsewhere. The marginal densities are given by

$$f_X(x) = \lim_{\phi \to \infty} \int_0^{\phi} \alpha e^{-\alpha x} \cdot \beta e^{-\beta y} dy = \alpha e^{-\alpha x} \cdot \lim_{\phi \to \infty} \left[-e^{-\beta y} \right]_0^{\phi} = \alpha e^{-\alpha x},$$

$$f_Y(y) = \lim_{\phi \to \infty} \int_0^{\phi} \alpha e^{-\alpha x} \cdot \beta e^{-\beta y} dx = \beta e^{-\beta y} \cdot \lim_{\phi \to \infty} \left[-e^{-\alpha x} \right]_0^{\phi} = \beta e^{-\beta y}.$$

Here we see that $f_X(x) \cdot f_Y(y) = \alpha e^{-\alpha x} \cdot \beta e^{-\beta y} = f(x, y)$, another indication that X and Y are independent.

Question 5 Let X and Y have a joint pdf $f(x,y) = c(x^2 - y^2)e^{-x}$ for $x \le 0$ and $-x \le y \le x$. To help with some of the integrations, this question makes use of the following result:

$$\Omega(\theta, n) := \int_{\theta}^{\infty} x^n e^{-x} dx = n! \cdot e^{-\theta} \sum_{k=0}^{n} \frac{\theta^k}{k!}$$

for all $\theta \in \mathbb{R}$ and $n \in \mathbb{Z}_{\geq 0}$. This specific integral came up so many times in this question that I thought it was worth figuring out a general formula for it. A derivation of this result can be found in Appendix A.

(a) Using the fact that the joint pdf must integrate to 1, we have

$$1 = \int_0^\infty \int_{-x}^x c(x^2 - y^2) e^{-x} dy dx = c \int_0^\infty e^{-x} \left[x^2 y - \frac{y^3}{3} \right]_{-x}^x dx$$
$$= \frac{4c}{3} \cdot \int_0^\infty x^3 e^{-x} dx = \frac{4c \cdot \Omega(0, 3)}{3} = \frac{4c \cdot 3!}{3} = 8c,$$

and so c = 1/8.

- (b) X and Y are not independent since the joint pdf does not have rectangular support.
- (c) When considering all possible $x \ge 0$ values, we see that we must have $-x \le y \le x$. So the marginal density of X is given by

$$f_X(x) = \int_{-x}^{x} \frac{(x^2 - y^2)e^{-x}}{8} dx = \frac{e^{-x}}{8} \left[x^2 y - \frac{y^3}{3} \right]_{-x}^{x} = \frac{e^{-x}}{8} \cdot \frac{4x^3}{3} = \frac{x^3 e^{-x}}{6}$$

for $x \ge 0$. When considering all possible $-\infty < y < \infty$, we must split the support into two regions: where $y \ge 0$ and where y < 0. When $y \ge 0$, we have $0 \le y \le x < \infty$, so we have $y \le x < \infty$. In this region, the marginal density is given by

$$f_Y(y) = \int_y^\infty \int_{-x}^x \frac{(x^2 - y^2)e^{-x}}{8} dx = \frac{1}{8} \left(\int_y^\infty x^2 e^{-x} dx - y^2 \int_y^\infty e^{-x} dx \right)$$
$$= \frac{1}{8} \left(\Omega(y, 2) - y^2 \cdot \Omega(y, 0) \right) = \frac{1}{8} \left(e^{-y} (2 + 2y + y^2) - y^2 e^{-y} \right) = \frac{(1 + y)e^{-y}}{8}.$$

In the second region, where y < 0, we have $0 > y \ge -x > -\infty$, and so we can have $-y < x < \infty$. The marginal density is this region is given by

$$f_Y(y) = \int_{-y}^{\infty} \frac{(x^2 - y^2)e^{-x}}{8} dx = \frac{1}{8} \left(\int_{-y}^{\infty} x^2 e^{-x} dx - y^2 \int_{y}^{\infty} e^{-x} dx \right)$$
$$= \frac{1}{8} \left(\Omega(-y, 2) - y^2 \cdot \Omega(-y, 0) \right) = \frac{1}{8} \left(e^y (2 - 2y + y^2) - y^2 e^y \right) = \frac{(1 - y)e^y}{8}.$$

As a whole, the joint density is given by

$$f_Y(y) = \begin{cases} (1+y)e^{-y}/8 & \text{when } y \ge 0, \\ (1-y)e^{y}/8 & \text{when } y < 0. \end{cases}$$

(d) The conditional distribution $f_{X|Y}(x|y)$ is found by dividing f(x,y) by $f_Y(y)$:

$$f_{X|Y}(x|y) = \begin{cases} \frac{(x^2 - y^2)e^{y-x}}{2(1+y)}, & y \ge 0, \ y \le x < \infty \\ \frac{(x^2 - y^2)e^{-(y+x)}}{2(1-y)}, & y < 0, \ -y < x < \infty \end{cases}$$

Similarly,

$$f_{Y|X}(y|x) = \frac{3(x^2 - y^2)}{4x^2}, -x \le y \le x.$$