

Homework 3

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Question 1

Suppose that $\mathbf{X} \stackrel{\text{iid}}{\sim} \chi^2(\theta)$, where $\theta \in \mathbb{N}$ is unknown. We would like to test $H_0 : \theta \leq 8$ against $H_A : \theta > 8$, using a UMP test δ^* with a specified significance $\alpha_* \in (0, 1)$. The joint density of \mathbf{X} is given by

$$f(\mathbf{x} | \theta) = \prod_{i=1}^n \frac{x_i^{\theta/2-1} e^{-x_i/2}}{2^{\theta/2} \Gamma(\theta/2)} = 2^{-n\theta/2} \cdot \Gamma^{-n}(\theta/2) \cdot \left(\prod_{i=1}^n x_i \right)^{n(\theta/2-1)} \cdot \exp \left(-\frac{1}{2} \sum_{i=1}^n x_i \right)$$

To determine δ^* , we will look at the likelihood ratio. If we have two values θ_1, θ_2 such that $\theta_1 < \theta_2$, then then likelihood ratio is

$$\begin{aligned} \frac{f(\mathbf{x} | \theta_2)}{f(\mathbf{x} | \theta_1)} &= \frac{2^{-n\theta_1/2} \cdot \Gamma^{-n}(\theta_2/2) \cdot (\prod_i x_i)^{n(\theta_2/2-1)} \cdot \exp(-\frac{1}{2} \sum_i x_i)}{2^{-n\theta_1/2} \cdot \Gamma^{-n}(\theta_1/2) \cdot (\prod_i x_i)^{n(\theta_1/2-1)} \cdot \exp(-\frac{1}{2} \sum_i x_i)} \\ &= 2^{n(\theta_1-\theta_2)/2} \left(\frac{\Gamma(\theta_1/2)}{\Gamma(\theta_2/2)} \right)^n \left(\prod_{i=1}^n x_i \right)^{n(\theta_2-\theta_1)/2}, \end{aligned}$$

which is a monotone increasing function of the test statistic $\prod_{i=1}^n x_i$. Therefore, the UMP test is δ^* : reject H_0 if $\prod_{i=1}^n X_i \geq c_*$, where c_* is chosen such that the test has a significance level α_* . Taking the log of both sides gives us $\sum_{i=1}^n \log X_i \geq \log c_* := k$.

Question 2

Question 8

By definition, the least-squares estimate for the intercept is $\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$. Rearranging gives us $\bar{y} = \hat{\beta}_0 + \hat{\beta}_1 \bar{x}$, and so the least-squares line $y = \hat{\beta}_0 + \hat{\beta}_1 x$ will always pass through the point (\bar{x}, \bar{y}) .

Question 9

(a) The least-squares coefficients for the model are given by $\hat{\beta}_0 = 40.9$ and $\hat{\beta}_1 = 0.548$.

Question 10

Let (\mathbf{x}, \mathbf{y}) be the vectors of observations for the predictor x and the response Y . From the data, we have $\bar{x} = 2.33$ and $\bar{y} = 0.81$. Here we are assuming that $Y = \beta_0 + \beta_1 x + \epsilon$, where $\epsilon \sim N(0, \sigma^2)$.

(a) The MLEs for β_0 , β_1 , and σ^2 are given by

$$\hat{\beta}_1 = \frac{(\mathbf{y} - \bar{y}\mathbf{1})^T(\mathbf{x} - \bar{x}\mathbf{1})}{\|\mathbf{x} - \bar{x}\mathbf{1}\|^2} = 0.685, \quad \hat{\beta}_0 = \bar{y} - \hat{\beta}_1\bar{x} = -0.786, \quad \text{and} \quad \hat{\sigma}^2 = \frac{\|\mathbf{y} - \hat{\beta}_0\mathbf{1} - \hat{\beta}_1\mathbf{x}\|^2}{n} = 0.938.$$

(b) The variance of $\hat{\beta}_0$ and $\hat{\beta}_1$ is given by

$$\text{Var}[\hat{\beta}_1] = \frac{\sigma^2}{\|\mathbf{x} - \bar{x}\mathbf{1}\|^2} = 0.0277\sigma^2 \quad \text{and} \quad \text{Var}[\hat{\beta}_0] = \sigma^2 \left(\frac{1}{n} + \frac{\bar{x}^2}{\|\mathbf{x} - \bar{x}\mathbf{1}\|^2} \right) = 0.25\sigma^2.$$

(c) The covariance between $\hat{\beta}_0$ and $\hat{\beta}_1$, and therefore the correlation, is

$$\text{Cov}[\hat{\beta}_0, \hat{\beta}_1] = -\frac{\bar{x}\sigma^2}{\|\mathbf{x} - \bar{x}\mathbf{1}\|^2} = -0.0646\sigma^2 \quad \text{and} \quad \text{Cor}[\hat{\beta}_0, \hat{\beta}_1] = \frac{\text{Cov}[\hat{\beta}_0, \hat{\beta}_1]}{\sqrt{\text{Var}[\hat{\beta}_0] \cdot \text{Var}[\hat{\beta}_1]}} = -0.775.$$