Homework 3

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Question 1

Suppose that $X \stackrel{\text{iid}}{\sim} \chi^2(\theta)$, where $\theta \in \mathbb{N}$ is unknown. We would like to test $H_0: \theta \leq 8$ against $H_A: \theta > 8$, using a UMP test δ^* with a specified significance $\alpha_* \in (0,1)$. The joint density of X is given by

$$f(\mathbf{x} \mid \theta) = \prod_{i=1}^{n} \frac{x_i^{\theta/2 - 1} e^{-x/2}}{2^{\theta/2} \Gamma(\theta/2)} = 2^{-n\theta/2} \cdot \Gamma^{-n}(\theta/2) \cdot \left(\prod_{i=1}^{n} x_i\right)^{n(\theta/2 - 1)} \cdot \exp\left(-\frac{1}{2} \sum_{i=1}^{n} x_i\right)$$

To determine δ^* , we will look at the likelihood ratio. If we have two values θ_1, θ_2 such that $\theta_1 < \theta_2$, then then likelihood ratio is

$$\begin{split} \frac{f(\mathbf{x} \mid \theta_2)}{f(\mathbf{x} \mid \theta_1)} &= \frac{2^{-n\theta_1/2} \cdot \Gamma^{-n}(\theta_2/2) \cdot \left(\prod_i x_i\right)^{n(\theta_2/2-1)} \cdot \exp\left(-\frac{1}{2} \sum_i x_i\right)}{2^{-n\theta_1/2} \cdot \Gamma^{-n}(\theta_1/2) \cdot \left(\prod_i x_i\right)^{n(\theta_1/2-1)} \cdot \exp\left(-\frac{1}{2} \sum_i x_i\right)} \\ &= 2^{n(\theta_1 - \theta_2)/2} \left(\frac{\Gamma(\theta_1/2)}{\Gamma(\theta_2/2)}\right)^n \left(\prod_{i=1}^n x_i\right)^{n(\theta_2 - \theta_1)/2}, \end{split}$$

which is a monotone increasing function of the test statistic $\prod_{i=1}^n x_i$. Therefore, the UMP test is δ^* : reject H_0 if $\prod_{i=1}^n X_i \ge c_*$, where c_* is chosen such that the test has a significance level α_* . Taking the log of both sides gives us $\sum_{i=1}^8 \log X_i \ge \log c_* := k$.

Question 2

Question 8

By definition, the least-squares estimate for the intercept is $\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$. Rearranging gives us $\bar{y} = \hat{\beta}_0 + \hat{\beta}_1 \bar{x}$, and so the least-squares line $y = \hat{\beta}_0 + \hat{\beta}_1 x$ will always pass through the point (\bar{x}, \bar{y}) .

Question 9

(a) The least-squares coefficients for the model are given by $\hat{\beta}_0 = 40.9$ and $\hat{\beta}_1 = 0.548$.

Question 10

Let (\mathbf{x}, \mathbf{y}) be the vectors of observations for the predictor x and the response Y. From the data, we have n = 10, $\bar{x} = 2.33$, $\bar{y} = 0.81$, $\|\mathbf{x} - \bar{x}\mathbf{1}\|^2 = 36.081$, and $(\mathbf{y} - \bar{y}\mathbf{1})^T(\mathbf{x} - \bar{x}\mathbf{1}) = 24.717$. Here we are assuming that $Y = \beta_0 + \beta_1 x + \epsilon$, where $\epsilon \sim N(0, \sigma^2)$.

(a) The MLEs for β_0 , β_1 , and σ^2 are given by

$$\hat{\beta}_1 = \frac{(\mathbf{y} - \bar{y}\mathbf{1})^T(\mathbf{x} - \bar{x}\mathbf{1})}{\|\mathbf{x} - \bar{x}\mathbf{1}\|^2} = 0.685, \quad \hat{\beta}_0 = \bar{y} - \hat{\beta}_1\bar{x} = -0.786, \quad \hat{\sigma}^2 = \frac{\|\mathbf{y} - \hat{\beta}_0\mathbf{1} - \hat{\beta}_1\mathbf{x}\|^2}{n} = 0.938.$$

(b) The variance of $\hat{\beta}_0$ and $\hat{\beta}_1$ is given by

$$\operatorname{Var}[\hat{\beta}_1] = \frac{\sigma^2}{\|\mathbf{x} - \bar{x}\mathbf{1}\|^2} = 0.0277\sigma^2, \quad \operatorname{Var}[\hat{\beta}_0] = \sigma^2 \left(\frac{1}{n} + \frac{\bar{x}^2}{\|\mathbf{x} - \bar{x}\mathbf{1}\|^2}\right) = 0.25\sigma^2.$$

(c) The covariance between $\hat{\beta}_0$ and $\hat{\beta}_1$, and therefore the correltaion, is

$$\operatorname{Cov}\left[\hat{\beta}_{0},\hat{\beta}_{1}\right] = -\frac{\bar{x}\sigma^{2}}{\|\mathbf{x} - \bar{x}\mathbf{1}\|^{2}} = -0.0646\sigma^{2}, \quad \operatorname{Cor}\left[\hat{\beta}_{0},\hat{\beta}_{1}\right] = \frac{\operatorname{Cov}\left[\hat{\beta}_{0},\hat{\beta}_{1}\right]}{\sqrt{\operatorname{Var}\left[\hat{\beta}_{0}\right] \cdot \operatorname{Var}\left[\hat{\beta}_{1}\right]}} = -0.775.$$

Question 11

Suppose β_0, β_1 are the coefficients from the linear model in question 10, and we want to estimate $\theta = 3\beta_0 - 2\beta_1 + 5$. Because $\hat{\beta}_0$ and $\hat{\beta}_1$ are unbiased estimators for the coefficients, we can estimate θ with $\hat{\theta} = 3\hat{\beta}_0 - 2\hat{\beta}_1 + 5$. The MSE of $\hat{\theta}$, which is just its variance, is given by

$$\mathbb{E}\left[(\hat{\theta}-\theta)^2\right] = \operatorname{Var}\left[3\hat{\beta}_0 - 2\hat{\beta}_1 + 5\right] = 9\operatorname{Var}\left[\hat{\beta}_0\right] + 4\operatorname{Var}\left[\hat{\beta}_1\right] - 6\operatorname{Cov}\left[\hat{\beta}_0, \hat{\beta}_1\right] = 10.549\sigma^2.$$

Question 12

We know that the MLE and least-squares estimates of β_0 and β_1 are the same, so $\hat{\beta}_0 = -0.786$ and $\hat{\beta}_1 = 0.685$. Using this linear model, when x = 2, we predict $\hat{Y} = -0.786_0.685 \cdot 2 = 0.584$. The variance of \hat{Y} is given by

$$\mathbb{E}\Big[(\hat{Y} - Y)^2\Big] = \sigma^2 \left(1 + \frac{1}{n} + \frac{(2 - \bar{x})^2}{\|\mathbf{x} - \bar{x}\mathbf{1}\|^2}\right) = 1.103\sigma^2.$$

Question 13