

# Homework 3

Aiden Kenny

STAT GR5204: Statistical Inference

Columbia University

December 14, 2020

## Question 1

Suppose that  $\mathbf{X} \stackrel{\text{iid}}{\sim} \chi^2(\theta)$ , where  $\theta \in \mathbb{N}$  is unknown. We would like to test  $H_0 : \theta \leq 8$  against  $H_A : \theta > 8$ , using a UMP test  $\delta^*$  with a specified significance  $\alpha_* \in (0, 1)$ . The joint density of  $\mathbf{X}$  is given by

$$f(\mathbf{x} | \theta) = \prod_{i=1}^n \frac{x_i^{\theta/2-1} e^{-x_i/2}}{2^{\theta/2} \Gamma(\theta/2)} = 2^{-n\theta/2} \cdot \Gamma^{-n}(\theta/2) \cdot \left( \prod_{i=1}^n x_i \right)^{n(\theta/2-1)} \cdot \exp \left( -\frac{1}{2} \sum_{i=1}^n x_i \right)$$

To determine  $\delta^*$ , we will look at the likelihood ratio. If we have two values  $\theta_1, \theta_2$  such that  $\theta_1 < \theta_2$ , then then likelihood ratio is

$$\begin{aligned} \frac{f(\mathbf{x} | \theta_2)}{f(\mathbf{x} | \theta_1)} &= \frac{2^{-n\theta_1/2} \cdot \Gamma^{-n}(\theta_2/2) \cdot (\prod_i x_i)^{n(\theta_2/2-1)} \cdot \exp(-\frac{1}{2} \sum_i x_i)}{2^{-n\theta_1/2} \cdot \Gamma^{-n}(\theta_1/2) \cdot (\prod_i x_i)^{n(\theta_1/2-1)} \cdot \exp(-\frac{1}{2} \sum_i x_i)} \\ &= 2^{n(\theta_1-\theta_2)/2} \left( \frac{\Gamma(\theta_1/2)}{\Gamma(\theta_2/2)} \right)^n \left( \prod_{i=1}^n x_i \right)^{n(\theta_2-\theta_1)/2}, \end{aligned}$$

which is a monotone increasing function of the test statistic  $\prod_{i=1}^n x_i$ . Therefore, the UMP test is  $\delta^*$  : reject  $H_0$  if  $\prod_{i=1}^n X_i \geq c_*$ , where  $c_*$  is chosen such that the test has a significance level  $\alpha_*$ . Taking the log of both sides gives us  $\sum_{i=1}^n \log X_i \geq \log c_* := k$ .

## Question 2

## Question 8

By definition, the least-squares estimate for the intercept is  $\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$ . Rearranging gives us  $\bar{y} = \hat{\beta}_0 + \hat{\beta}_1 \bar{x}$ , and so the least-squares line  $y = \hat{\beta}_0 + \hat{\beta}_1 x$  will always pass through the point  $(\bar{x}, \bar{y})$ .

## Question 9

(a) The least-squares coefficients for the model are given by  $\hat{\beta}_0 = 40.9$  and  $\hat{\beta}_1 = 0.548$ .

## Question 10

Let  $(\mathbf{x}, \mathbf{y})$  be the vectors of observations for the predictor  $x$  and the response  $Y$ . From the data, we have  $\bar{x} = 2.33$  and  $\bar{y} = 0.81$ . Here we are assuming that  $Y = \beta_0 + \beta_1 x + \epsilon$ , where  $\epsilon \sim N(0, \sigma^2)$ .

(a) The MLEs for  $\beta_0$ ,  $\beta_1$ , and  $\sigma^2$  are given by

$$\hat{\beta}_1 = \frac{(\mathbf{y} - \bar{y}\mathbf{1})^T(\mathbf{x} - \bar{x}\mathbf{1})}{\|\mathbf{x} - \bar{x}\mathbf{1}\|^2} = 0.685, \quad \hat{\beta}_0 = \bar{y} - \hat{\beta}_1\bar{x} = -0.786, \quad \hat{\sigma}^2 = \frac{\|\mathbf{y} - \hat{\beta}_0\mathbf{1} - \hat{\beta}_1\mathbf{x}\|^2}{n} = 0.938.$$

(b) The variance of  $\hat{\beta}_0$  and  $\hat{\beta}_1$  is given by

$$\text{Var}[\hat{\beta}_1] = \frac{\sigma^2}{\|\mathbf{x} - \bar{x}\mathbf{1}\|^2} = 0.0277\sigma^2, \quad \text{Var}[\hat{\beta}_0] = \sigma^2 \left( \frac{1}{n} + \frac{\bar{x}^2}{\|\mathbf{x} - \bar{x}\mathbf{1}\|^2} \right) = 0.25\sigma^2.$$

(c) The covariance between  $\hat{\beta}_0$  and  $\hat{\beta}_1$ , and therefore the correlation, is

$$\text{Cov}[\hat{\beta}_0, \hat{\beta}_1] = -\frac{\bar{x}\sigma^2}{\|\mathbf{x} - \bar{x}\mathbf{1}\|^2} = -0.0646\sigma^2, \quad \text{Cor}[\hat{\beta}_0, \hat{\beta}_1] = \frac{\text{Cov}[\hat{\beta}_0, \hat{\beta}_1]}{\sqrt{\text{Var}[\hat{\beta}_0] \cdot \text{Var}[\hat{\beta}_1]}} = -0.775.$$

### Question 11

Suppose  $\beta_0, \beta_1$  are the coefficients from the linear model in question 10, and we want to estimate  $\theta = 3\beta_0 - 2\beta_1 + 5$ . Because  $\hat{\beta}_0$  and  $\hat{\beta}_1$  are unbiased estimators for the coefficients, we can estimate  $\theta$  with  $\hat{\theta} = 3\hat{\beta}_0 - 2\hat{\beta}_1 + 5$ . The MSE of  $\hat{\theta}$ , which is just its variance, is given by

$$\mathbb{E}[(\hat{\theta} - \theta)^2] = \text{Var}[3\hat{\beta}_0 - 2\hat{\beta}_1 + 5] = 9\text{Var}[\hat{\beta}_0] + 4\text{Var}[\hat{\beta}_1] - 6\text{Cov}[\hat{\beta}_0, \hat{\beta}_1] = 10.549\sigma^2.$$

### Question 12

We know that the MLE and least-squares estimates of  $\beta_0$  and  $\beta_1$  are the same, so  $\hat{\beta}_0 = -0.786$  and  $\hat{\beta}_1 = 0.685$ . Using this linear model, when  $x = 2$ , we predict  $\hat{Y} = -0.786 + 0.685 \cdot 2 = 0.584$ . The variance of  $\hat{Y}$  is given by

$$\mathbb{E}[(\hat{Y} - Y)^2] = \sigma^2 \left( 1 + \frac{1}{n} + \frac{(2 - \bar{x})^2}{\|\mathbf{x} - \bar{x}\mathbf{1}\|^2} \right) = 1.103\sigma^2.$$