Global Deaths prediction due to Covid-19

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Objective

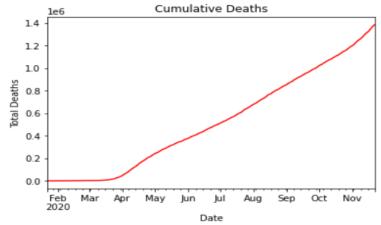
- Model the time series using standard time series models.
- Predict the future observations after fitting the model.
- Tune the parameters of the fitted model well to get good predictions

Motivation

- Machine learning and statistical methods, which time series forecasting is a subset of, have been successfully implemented in the past in the area of infectious diseases. For examplemodeling leptospirosis and its relationship to rainfall and temperature.
- Similar approaches have also been followed to model diseases that occur in cyclic or repeating patterns, such as the seasonal influenza, for which a number of studies that use time-series modeling to predict future outbreaks have been published.
- Regarding COVID-19 forecasting, there has been a surge in scientific work published during the last few months. The majority of these studies focus on predicting coronavirus-related metrics such as active cases and deaths across the world.

Data Set

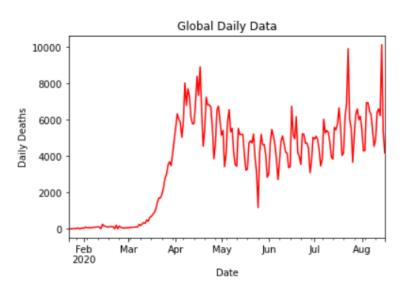
- We are using the data set available at John Hopkins University's Github Page.
- It contains the country wise data for cumulative deaths due to Covid-19.



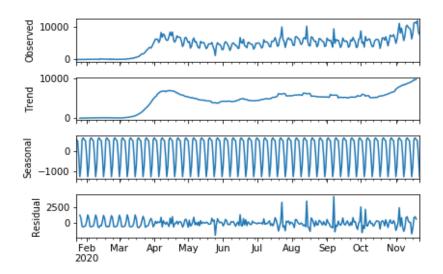
Methodology

- Organised the data and took the sum of data from all the countries.
- Took the first difference to get the daily data.
- Made the series stationary by taking an appropriate order of difference.
- Investigated the Auto-Correlations(ACF) and the Partial Auto-Correlations(PACF).
- Tried to fit different time series models and evaluated the model based on the Akaike Information Criterion(AIC).
- Compared the prediction performance of different model fits.

Daily Data

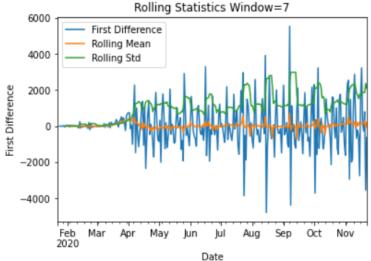


Seasonal Decomposition



Stationarity of first difference

 We tested the stationarity using the Augmented Dicky Fuller Test.



Dickey-Fuller Test for Trend on First Difference

 H_0 : Series has a trend. H_A : Series is stationary.

Results of Dickey-Fuller Test:

Test Statistic: -2.914002

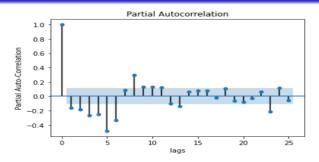
p-value: 0.043744 Lags Used: 12.000000

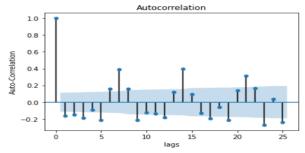
Number of Observations Used: 193.000000

Critical Value (1%): -3.464694 Critical Value (5%): -2.876635 Critical Value (10%): -2.574816

Here, Test Statistic is less than Critical Value at 5% and thus we reject the Null Hypothesis and thus at a confidence interval of 95%, Trend is not present in first difference.

Auto Correlation Plots







Parameter Estimation

- The auto-correlation plot suggests a seasonality of order 7, which is also evident from the plot.
- To begin with, we split the data into 2 parts(Test and Train).
 We tried fitting different models on the Train data and evaluated the performance on the test data.
- Then we started with the ARIMA Model.

ARIMA Model

•

- ARIMA offers a high level of interpretability, as, based on the assumptions of the model, the relationship between the independent variables and the dependent variables are well-understood and therefore easily explained.
- This enables researchers to gain a deep understanding not only of the relationship between the current state as a function of the past states (endogenous variables), but also of any influence inputs outside the state of the series might have (exogenous variables).
- ARIMA(Auto Regressive Integrated Moving Average) Model(X_t is the time series, ϵ_t s are i.i.d. residuals)

$$\phi(B)(1-B)^d X_t = \theta(B)\epsilon_t$$

where B denotes the lag operator and $\phi(B)$ and $\theta(B)$ are the AR and MA polynomials of order p and q respectively.

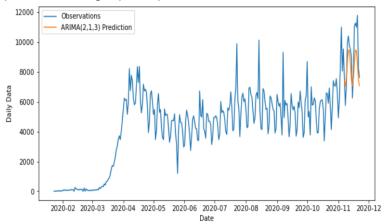
$$AIC = -2log\hat{L} + 2(p+q+d+1)$$

where \hat{L} denotes the likelihood and the second term contains the number of parameters being used.

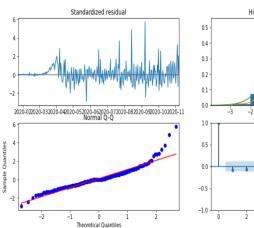


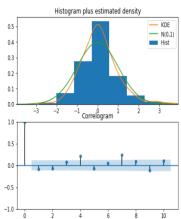
ARIMA Model

• Varying p and q, and using the AIC as the evaluation parameter, we get p = 2, q = 3, d = 1.



Residual Analysis





Results of ARIMA Model

SARIMAX Results

| Dep. Varia | able: | first | diff I | == ام. | Observations: | | 290 |
|-------------------------|-----------|------------------|----------|-----------|---------------|----------|-----------|
| Model: SAR | | SARIMAX(2, 1, 3) | | | Likelihood | | -2364.199 |
| | | hu, 26 Nov | | • | | 4740.398 | |
| Time: | | | 7:55 BIC | | 4762.397 | | |
| Sample: | | 01-23- | 2020 H | HQIC | | | 4749.213 |
| | | - 11-07- | 2020 | | | | |
| Covariance | Type: | | opg | | | | |
| ======= | coef | std err | | z | P> z | [0.025 | 0.975] |
| ar.L1 | 1.2457 | 0.005 | 226.5 | 592 | 0.000 | 1.235 | 1.256 |
| ar.L2 | -0.9977 | 0.005 | -191.8 | 315 | 0.000 | -1.008 | -0.987 |
| ma.L1 | -1.8813 | 0.038 | -48.9 | 952 | 0.000 | -1.957 | -1.806 |
| ma.L2 | 1.7177 | 0.057 | 30.1 | 170 | 0.000 | 1.606 | 1.829 |
| ma.L3 | -0.5955 | 0.041 | -14.6 | 557 | 0.000 | -0.675 | -0.516 |
| sigma2 | 7.506e+05 | 3.22e+04 | 23.3 | 314 | 0.000 | 6.88e+05 | 8.14e+05 |
| Ljung-Box (Q): | | | 187.0 | 90 | Jarque-Bera | (JB): | 486.70 |
| Prob(Q): | | | 0.0 | 90 | Prob(JB): | | 0.00 |
| Heteroskedasticity (H): | | | 3.1 | 18 | Skew: | | 1.37 |
| Prob(H) (two-sided): | | | 0.0 | 90 | Kurtosis: | | 8.74 |

SARIMA Model

• SARIMA Model for a time series X_t is given by,

$$\Phi(B^s)\phi(B)(1-B^s)^D(1-B)^dX_t = \Theta(B^s)\theta(B)\epsilon_t$$

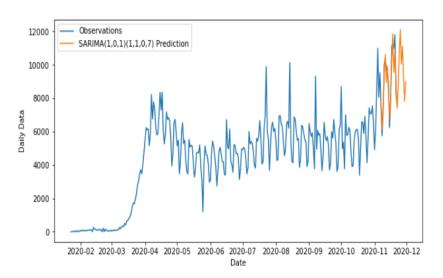
where B denotes the lag operator , $\phi(B)$ and $\theta(B)$ are the AR and MA polynomials of order p and q, and $\Phi(B)$ and $\Theta(B)$ are the seasonal AR and MA polynomials of order P and Q respectively, and d and D are the orders of difference and seasonal difference.

 $AIC = -2log\hat{L} + 2(p+q+P+Q+D+d+2)$

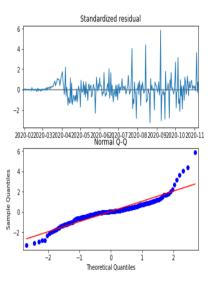
where \hat{L} denotes the likelihood and the second term contains the number of parameters being used.

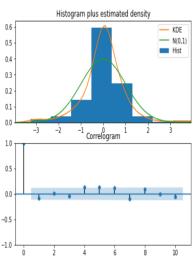
• Varying p and q, and using the AIC as the evaluation parameter, we get p = 1, d=0, q=1, P=1, D=1, Q=0, S=7.

SARIMA Model



Residual Analysis





Results of SARIMA Model

SARIMAX Results

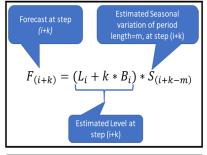
| Dep. Varia Model: Date: Time: Sample: | | ARIMAX(1, 0, | 1)x(1, 1, [Thu, 26 Nov |], 7) Log 2020 AIC 12:36 BIC -2020 HQI | | | 290 -2323.928 4655.857 4670.439 4661.704 |
|---|-------------|--------------|----------------------------|---|---------------------|----------|--|
| Covariance | Type: | | | opg | | | |
| ======= | coe | f std err | Z | P> z | [0.025 | 0.975] | |
| ar.L1 | 0.961 | 0.022 | 44.385 | 0.000 | 0.919 | 1.004 | |
| ma.L1 | -0.697 | 0.044 | -15.830 | 0.000 | -0.784 | -0.611 | |
| ar.S.L7 | -0.450 | 0.029 | -15.470 | 0.000 | -0.507 | -0.393 | |
| sigma2 | 7.927e+0 | 3.19e+04 | 24.838 | 0.000 | 7.3e+05 | 8.55e+05 | |
| Ljung-Box | (Q): | | 110.79 | Jarque-Ber | ======== a (JB): | 578 | 3.91 |
| Prob(Q): | , | | 0.00 | Prob(JB): | | 6 | 0.00 |
| Heteroske | dasticity (| H): | 3.88 | Skew: | | 6 | .91 |
| Prob(H) (t | two-sided): | • | 0.00 | Kurtosis: | | 9 | .77 |

Holt-Winters Exponential Smoothing

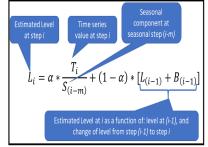
 Holt-Winters Exponential Smoothing is used for forecasting time series data that exhibits both a trend and a seasonal variation. The Holt-Winters technique is made up of the following four forecasting techniques stacked one over the other:

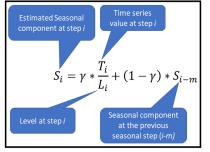


Holt-Winters Exponential Smoothing

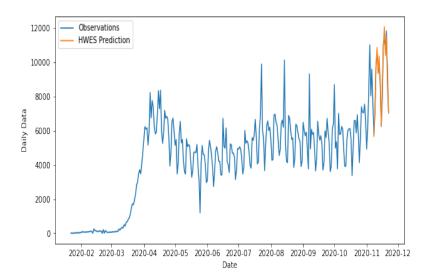


Estimated Trend at step
$$i$$
 at step i $B_i = \beta * [L_i - L_{(i-1)}] + (1 - \beta) * B_{(i-1)}$ Estimated rate of change of level from step $(i-1)$ to step i





Holt-Winters Exponential Smoothing



Results of HWES

ExponentialSmoothing Model Results

| Dep. Variable: | 0 | No. Observations: | 291 |
|-------------------|----------------------|-------------------|------------------|
| Model: | ExponentialSmoothing | SSE | 314026579.701 |
| Optimized: | True | AIC | 4064.474 |
| Trend: | Multiplicative | BIC | 4104.881 |
| Seasonal: | Multiplicative | AICC | 4065.789 |
| Seasonal Periods: | 7 | Date: | Sun, 29 Nov 2020 |
| Box-Cox: | False | Time: | 15:48:04 |
| Box-Cox Coeff.: | None | | |

| | coeff | code | optimized | |
|--------------------|-----------|-------|-----------|--|
| | | | | |
| smoothing_level | 0.2878571 | alpha | True | |
| smoothing_trend | 0.0575714 | beta | True | |
| smoothing_seasonal | 0.3560714 | gamma | True | |
| initial_level | 4831.4762 | 1.0 | True | |
| initial_trend | 1.1558217 | b.0 | True | |
| initial_seasons.0 | 0.0035186 | 5.0 | True | |
| initial_seasons.1 | 0.0002070 | 5.1 | True | |
| initial_seasons.2 | 0.0016558 | 5.2 | True | |
| initial_seasons.3 | 0.0033116 | s.3 | True | |
| initial_seasons.4 | 0.0028977 | 5.4 | True | |
| initial_seasons.5 | 0.0053814 | 5.5 | True | |
| initial_seasons.6 | 0.0101418 | 5.6 | True | |

Comparison of Models Based on Root Mean Squared Error

| | 15 Days RMSE | 30 Days RMSE |
|--------------|--------------|--------------|
| ARIMA Model | 1414.776 | 3056.512 |
| SARIMA Model | 674.602 | 2883.674 |
| HWES Model | 465.38 | 2446.559 |

- We found out that SARIMA Model gives us a better fit for the data as compared to the ARIMA Model as our data has strong seasonality(evident from ACF plots).
- We found that HWES Model given even better fit than SARIMA Model as our data has noise which gets reduced by the exponential smoothing
- HWES also takes into account the seasonality in the data and gives more importance to recent observations.



References

- (Data source) Github page of John Hopkins University
- (Modeling seasonal leptospirosis transmission) [1]
- (Holt-Winters Exponential Smoothing) [2]

Thank You!