# Microeconometrics Preliminaries

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#### Microeconometric Analysis

- Microeconometric Analysis: "The analysis of individual-level data on the economic behavior of individuals or firms," Cameron and Trivedi, 2005
- Microeconomic Data: cross-sections or panel data
- Cross-sectional Data: "consists of a sample of individuals taken at a given point in time," Wooldrige, 2013
- ▶ Panel Data: "A panel data (or longitudinal data) set consists of a time series for each cross-sectional member in the data set," Wooldrige, 2013
- ▶ In this course: Cross-sectional data

#### Cross-sectional Data

- Cross-sectional data. Examples:
- ► California Test Score: (Stock and Watson, 2012)

 $Data: tscore_i, str_i, expen_i, eng_i$ 

▶ **Wage Equations**: (Wooldrige, 2013)

 $Data: w_i$ ,  $educ_i$ ,  $exper_i$ ,  $female_i$ ,  $married_i$ 

► **Labor Force Participation**: (Wooldrige, 2013)

 $Data:inlf_i, nwifeinc_i, educ_i, exper_i, age_i, kidslt6_i, kidsge6_i$ 

► Crime: (Wooldrige, 2013)

 $Data: crime_i, wage_i, othinc_i, freqarr_i, freqconv_i, avgsen_i, age_i$ 



#### Cross-sectional Data: California Test Score

- ► California Test Score: Data: Stock and Watson (p. 51)
  - ► *tscore*<sub>i</sub>: average of the math and science test scores for all fifth grades in 1999 in district i
  - $ightharpoonup str_i$ : average student-teacher ratio in district i
  - $\exp en_i$ : average expenditure per pupil
  - *eng*<sub>i</sub>: percentage of students still learning English

#### Cross-sectional Data: Wage Equations

- ▶ **Wage Equations**: Data: Wooldrige (p. 218)
  - $w_i$ : hourly wage
  - $educ_i$ : years of formal education
  - ightharpoonup exp  $er_i$ : years of workforce experience
  - ► *female*<sub>i</sub>: 1 if person *i* is female, otherwise
  - $married_i$ : 1 if person i is married, otherwise

#### Cross-sectional Data: Labor Force Participation

- ► Labor Force Participation: Data: Wooldrige (p. 239)
  - *inlf<sub>i</sub>* : 1 if woman *i* reports working for a wage outside the home, 0 otherwise
  - ► *nwifeinc<sub>i</sub>* : husband's earnings
  - educ<sub>i</sub>: years of education
  - $\exp er_i$ : past years of labor market experience
  - ► *kidslt6<sub>i</sub>*: number of children less than six years old
  - ► *kidsge*6<sub>i</sub>: number of kids between 6 and 18 years of age

#### Cross-sectional Data: Crime Data

- ► **Crime**: Data: Wooldrige (p. 4, 78,172, 295, 583)
  - ightharpoonup *crime*<sub>i</sub>: some measure of the frequency of criminal activity
  - ► Ex: *narr*86<sub>i</sub>: number of times a man was arrested
  - $pcnv_i$ : proportion of prior arrests leading to conviction
  - ► *tottime*<sub>i</sub>: total time the man has spent in prison prior to 1986 since reaching the age of 18
  - ▶ *ptime*86<sub>i</sub>: months spent in prison in 1986
  - qemp86<sub>i</sub>: number of quarters in 1986 during which the man was legally employed

#### **Databases**

#### **Some Sources of Microdata:** Cameron and Trivedi (2005) p.58

- Panel Study in Income Dynamics (PSID)
- Current Population Survey (CPS)
- National Longitudinal Survey (NLS)
- National Longitudinal Surveys of Youth (NLSY)
- Survey of Income and Program Participation (SIPP)
- ► Health and Retirement (HRS)
- World Bank's Living Standards Measurement Study (LSMS)
- Data clearinghouses
- ► Journal data archives

#### Students Resources

#### **Some Students Resources:**

- Stock and Watson:
  - Stock: http://scholar.harvard.edu/stock/home
  - Watson: http://www.princeton.edu/~mwatson/
- Wooldrige: http://econ.msu.edu/faculty/wooldridge/
- ► Greene: http://people.stern.nyu.edu/wgreene/
- Cameron and Trivedi:
  - Cameron: http://cameron.econ.ucdavis.edu/
  - ► Trivedi: http://pages.iu.edu/~trivedi/

#### Microeconometric Analysis: Regressions

What is it usually done with a dataset,  $y_i$ ,  $x_{1i}$ ,  $x_{2i}$ , ...,  $x_{ki}$ , in microeconometrics?

#### REGRESSIONS

"In modern microeconometrics the term regression refers to a bewildering range of procedures for studying the relationship between an outcome variable y and a set of regressors x." Cameron and Trivedi (2005) p.66

#### **Motivating Regressions**

#### **Conditional Prediction of** *y* **given** *x***.** (CT, 2005 p.66)

▶ Loss function

$$L\left( e\right) =L\left( y-h\left( x\right) \right)$$

where h(x) denotes the predictor defined as a function of x, e = y - h(x) is the prediction error and L(e) is the loss associated with the error e.

Expected Loss:

$$E\left[L\left(y-h\left(x\right)\right)|x\right]$$

Optimal Predictor

$$\min_{h(x)} E\left[L\left(y - h\left(x\right)\right) | x\right]$$



#### Motivating Regressions: Mean-square error loss

- ► The choice of the loss function depends on the nature of the problem being studied
- ► The quadratic loss function is often used in econometrics:

$$E[L(y - h(x)) | x] = E[e^{2} | x]$$

▶ **Important**: For the mean-square error loss function the optimal predictor is the conditional expectation E[y|x], i.e. if

$$\min_{h(x)} E\left[ \left( y - h\left( x \right) \right)^2 | x \right]$$

then 
$$h(x) = E[y|x]$$

# Motivating Regressions: Mean-square error loss

- ▶ Two approaches: Nonparametric or Parametric E[y|x]
- ▶ In this course, we will specify a parametric model for  $E[y|x] = g(x, \beta)$  where  $\beta$  needs to be estimated

#### Motivating Regressions: Mean-square error loss

▶ Sample Analog

$$\frac{1}{n}\sum_{i=1}^{n}L\left(e_{i}\right)$$

► For the mean-square error loss function:

$$\frac{1}{n}\sum_{i=1}^{n}L(e_i) = \frac{1}{n}\sum_{i=1}^{n}e_i^2 = \frac{1}{n}\sum_{i=1}^{n}(y_i - g(x_i, \beta))^2,$$

and the  $\beta$  that minimizes it is known as least squares. If g is linear, then it known as ordinary least squares

#### Motivating Regressions: Absolute error loss

- ▶ Absolute error loss: L(e) = |e|
- ▶ Optimal predictor: med[y|x]
- ▶ If  $med[y|x] = x\beta$ , then

$$\sum_{i=1}^{n} L(e_i) = \sum_{i=1}^{n} |y_i - x_i \beta|$$

and the  $\beta$  that minimizes it is known as the least absolute deviations estimator

► Robustness (outliers)

# Motivating Regressions: Asymmetric absolute error loss

- ► Asymmetric absolute error loss: penalty of  $(1 \alpha) |e|$  on overprediction and  $\alpha |e|$  on underprediction
- ▶  $\alpha \in (0,1)$  and  $\alpha = 0.5$  implies symmetry
- ▶ Optimal predictor: Conditional quantile:  $q_{\alpha}[y|x]$
- ► Basis for Quantile Regressions:

$$\sum_{i=1}^{n} L(e_i) = \sum_{i:y_i \ge x_i \beta}^{n} \alpha |y_i - x_i \beta_{\alpha}| + \sum_{i:y_i < x_i \beta}^{n} (1 - \alpha) |y_i - x_i \beta_{\alpha}|$$

and the  $\beta_{\alpha}$  that minimizes it is known as the  $\alpha^{th}$  quantile regression estimator. For  $\alpha=0.5$ , we get the median regression estimator or least absolute deviations estimator described above.

# Motivating Regressions: Conditional Expectation

▶ **Main focus of this course**: Conditional Expectation: E[y|x]

$$y = E[y|x] + u$$

► E[y|x] linear: Example: Linear wage equation

$$E[wage|x] = \beta_0 + \beta_1 educ + \beta_2 exper + \beta_3 female$$

▶ E[y|x] is nonlinear. Example: Poisson Regression for number of arrests

$$E[narr86|x] = \exp(\beta_0 + \beta_1 pcnv + \beta_2 avgsen + \beta_3 tottime)$$



**Definition**: *Conditional Expectation (Bivariate case)*: Let Y and X be random variables with joint density function f(x,y). Let the conditional density function of Y given  $x \in B$  be  $f(y|x \in B)$ . Let g(Y) be a real-valued function of Y. Then the conditional expectation of g(Y) given  $x \in B$ , is defined as

(i) Discrete case

$$E\left[g\left(Y\right)|x\in B\right] = \sum_{y\in R(Y)} g\left(Y\right) f\left(y|x\in B\right)$$

(ii) Continuous case

$$E[g(Y)|x \in B] = \int_{-\infty}^{\infty} g(Y)f(y|x \in B) dy$$

(Mittelhammer (2013) p. 125)



**Definition**: Conditional Density Function (Bivariate case): Let Y and X be random variables with joint density function f(x,y) and let  $f_X(x)$  be the marginal density function of X. The conditional density of Y given  $x \in B$  is

$$f(y|x \in B) = \frac{f(x \in B, y)}{f_X(x \in B)}$$

**Definition**: *Marginal Density Function (Bivariate case)*: Let Y and X be random variables with joint density function f(x,y). The marginal density function of X is

$$f_{X}(x) = \begin{cases} \sum_{y \in R(Y)} f(x,y) & \text{discrete case} \\ \int_{-\infty}^{\infty} f(x,y) \, dy & \text{continuous case} \end{cases}$$

#### Example: (Mittelhammer, p. 82)

► A company has two processing plants, plant 1 and plant 2. The proportion of processing capacity at which each of the plants operates on any given day is the outcome of a bivariate random variable

Joint density function:

$$f(x_1, x_2) = (x_1 + x_2) I_{[0,1]}(x_1) I_{[0,1]}(x_2)$$



► Marginal for  $X_1$ : Integrate out  $x_2$  from  $f(x_1, x_2)$  as

$$f_{1}(x_{1}) = \int_{-\infty}^{\infty} f(x_{1}, x_{2}) dx_{2}$$

$$= \int_{-\infty}^{\infty} (x_{1} + x_{2}) I_{[0,1]}(x_{1}) I_{[0,1]}(x_{2}) dx_{2}$$

$$= \int_{0}^{1} (x_{1} + x_{2}) I_{[0,1]}(x_{1}) dx_{2}$$

$$= \left(x_{1}x_{2} + \frac{x_{2}^{2}}{2}\right) I_{[0,1]}(x_{1}) \Big|_{0}^{1}$$

$$= \left(x_{1} + \frac{1}{2}\right) I_{[0,1]}(x_{1})$$

► Conditional density function of plan 1's capacity given that plant 2 operates at less than half of capacity

$$f(x_1|x_2 \le 0.5) = \frac{\int_{-\infty}^{0.5} f(x_1, x_2) dx_2}{\int_{-\infty}^{0.5} f_2(x_2) dx_2}$$

$$= \frac{\int_{0}^{0.5} (x_1 + x_2) I_{[0,1]}(x_1) dx_2}{\int_{0}^{0.5} (x_2 + \frac{1}{2}) dx_2}$$

$$= \left(\frac{4}{3}x_1 + \frac{1}{3}\right) I_{[0,1]}(x_1)$$

▶ What about: Conditional density function for plant 1's capacity given that plant 2's capacity proportion is  $x_2 = 0.75$ ?

$$f(x_1|x_2 = 0.75) = \frac{\int_{0.75}^{0.75} f(x_1, x_2) dx_2}{\int_{0.75}^{0.75} f_2(x_2) dx_2} = \frac{0}{0}$$

 In that case, by an approximation argument (see Mittelhammer p.88),

$$f(x_1|x_2 = 0.75) = \frac{f(x_1, 0.75)}{f_2(0.75)}$$

$$= \frac{(x_1 + 0.75) I_{[0,1]}(x_1)}{1.25}$$

$$= \left(\frac{4}{5}x_1 + \frac{3}{5}\right) I_{[0,1]}(x_1)$$

► Conditional Expectation of  $X_1$  given  $x_2 = 0.75$ ?

$$E[X_1|x_2 = 0.75] = \int_{-\infty}^{\infty} x_1 f(x_1|x_2 = 0.75) dx_1$$
$$= \int_{0}^{1} x_1 \left(\frac{4}{5}x_1 + \frac{3}{5}\right) dx_1$$
$$= \frac{17}{30}$$

► Conditional Expectation of  $X_1$  as a function of  $x_2$ :

$$E[X_{1}|x_{2}] = \int_{-\infty}^{\infty} x_{1}f(x_{1}|x_{2}) dx_{1}$$

$$= \int_{-\infty}^{\infty} x_{1}\frac{f(x_{1},x_{2})}{f_{2}(x_{2})} dx_{1}$$

$$= \int_{0}^{1} x_{1}\frac{(x_{1}+x_{2})I_{[0,1]}(x_{2})}{(x_{2}+\frac{1}{2})I_{[0,1]}(x_{2})} dx_{1}$$

$$= \left[\frac{\frac{1}{2}x_{2}+\frac{1}{3}}{x_{2}+\frac{1}{2}}\right] \text{ for } x_{2} \in [0,1]$$

▶ When evaluated at  $x_2 = 0.75$  same result as above

### Motivating Regressions: Conditional Expectation

▶ **Main focus of this course**: Conditional Expectation: E[y|x]

$$y = E\left[y|x\right] + u$$

► First part of the course E[y|x] linear: Example: Linear wage equation

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▶ Second part of the course E[y|x] is nonlinear. Example: Poisson Regression for number of arrests

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# **Conditional Expectation,** E[y|x]**, Properties**: (Wooldrige 2010, p. 30)

1. The conditional expectation is a linear operator: Let x and y be two random scalars and a(x) and b(x) two scalar functions of x. Then,

$$E[a(x) y + b(x) | x] = E[a(x) y | x] + E[b(x) | x] = a(x) E[y | x] + b(x)$$
  
provided that  $E(|y|) < \infty$ ,  $E(|a(x) y|) < \infty$ , and  $E(|b(x)|) < \infty$ .

**1.** (General) The conditional expectation is a linear operator: Let  $a_1(x)$ , ...,  $a_G(x)$  and b(x) be scalar functions of x, and let  $y_1$ , ...,  $y_G$  be random scalars. Then,

$$E\left[\sum_{j=1}^{G} a_{j}(x) y_{j} + b(x) | x\right] = \sum_{j=1}^{G} a_{j}(x) E[y_{j}|x] + b(x)$$

provided that  $E(|y_j|) < \infty$ ,  $E(|a_j(x)y_j|) < \infty$ , and  $E(|b(x)|) < \infty$ .

**2**. Law of Iterated Expectations: (Simplest case):

$$E(y) = E[E(y|x)]$$

**3**. Law of Iterated Expectations: (General case):

$$E[y|x] = E[E(y|w)|x]$$

where x and w are vectors with x = f(w) for some nonstochastic function f(.).

**4**. If  $f(x) \in \mathbb{R}^{J}$  is a function of x such that E[y|x] = g(f(x)) for some scalar function g(.), then

$$E\left[y|f\left(x\right)\right] = E\left[y|x\right]$$

5. If the vector (u, v) is independent of the vector x, then

$$E\left[u|x,v\right] = E\left[u|v\right]$$

**6**. If  $u \equiv y - E[y|x]$ , then

$$E\left[g\left(x\right)u\right]=0$$

for any function g(x), provided that  $E(|g_j(x)u|) < \infty$ , j = 1, ..., J, and  $E(|u|) < \infty$ . In particular, E(u) = 0 and  $Cov(x_j, u) = 0$ , j = 1, ..., K.

7. Conditional Jensen's Inequality: If  $c : \mathbb{R} \to \mathbb{R}$  is a convex function defined on  $\mathbb{R}$  and  $E(|y|) < \infty$ , then

$$c\left(E\left[y|x\right]\right) \le E\left[c\left(y\right)|x\right]$$

**8**. If  $E(y^2) < \infty$  and  $\mu(x) \equiv E[y|x]$ , then  $\mu$  is a solution to

$$\min_{m \in M} E\left[ \left( y - m\left( x \right) \right)^2 \right]$$

where M is the set of functions  $m : \mathbb{R}^K \to \mathbb{R}$  such that  $E\left[m\left(x\right)^2\right] < \infty$ . (That is  $E\left[y|x\right]$  is the best mean square predictor of y given x)

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