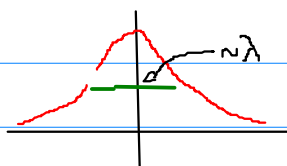

$$\gamma_z(t, t-\tau)$$

$$= \sum_{n \geq 0} \mathbb{E}[x_0 x_1] \frac{e^{-\lambda \tau} (\lambda \tau)^n}{n!} = e^{-\lambda \tau} \sum_{n \geq 0} \frac{\lambda^n \tau^n}{n!}$$

$$\chi_z(t, t-z) = \chi_z(t) = e^{\sigma^2 - \lambda t} = e^{\sigma^2} e^{-\lambda |t|}$$

$$* \quad \text{Spec } H^0(\omega) \text{ of } Z : \Gamma(\nu) = \int_{-\infty}^{\infty} \gamma_2(\tau) e^{-i2\pi\nu\tau} d\tau = \left\{ \int_{-\infty}^0 e^{-i2\pi\nu\tau + \lambda\tau} d\tau + \int_0^{\infty} e^{-i2\pi\nu\tau - \lambda\tau} d\tau \right\}$$



$$\frac{\sigma^2}{\lambda^2 + 4\pi^2 \nu^2} \sim \frac{\Gamma(\nu)}{z}$$

$$* \quad P(z_1 = z_1, z_2 = z_2) = P(z_1, z_2) = \sum_{k \geq 0} P(z_1, z_2 | N_z = k) P[N_z = k]$$

$\{X_n\}$ gaussian iid $N(0, \sigma^2)$

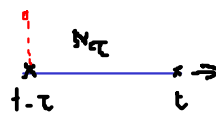
$$= \phi(r_1, r_2 | N_T > 0) P[N_T > 0] \quad (1)$$

$$+ P(\bar{z}_1, \bar{z}_2 | N_T = 0) P[N_T = 0] \quad (2)$$

$$\textcircled{1} = \underbrace{P\{Z_t = z_2 \mid Z_{t-\tau} = z_1, N_\tau > 0\}}_{P_X(x_0 = z_2)} \underbrace{P\{Z_{t-\tau} = z_1 \mid N_\tau > 0\}}_{P\{Z_{t-\tau} = z_1\}} \underbrace{P\{N_\tau > 0\}}_{(1 - e^{-\lambda \tau})}$$

$$= \frac{(1 - e^{-\lambda \tau})}{2\pi\sigma^2} e^{-\frac{\mu_1^2 + \mu_2^2}{2\sigma^2}}$$

$$① = \frac{-\lambda \tau}{2\pi} e^{-\lambda \tau} - \frac{\sigma_1^2 + \sigma_2^2}{2\sigma^2}$$



$$② = \underbrace{P(z_1 = z_2 | z_{t-\tau} = z_1, N_t = 0)}_{\delta(z_2 - z_1)} \underbrace{P(z_{t-\tau} = z_1 | N_t = 0)}_{p_{x_0}(z_1)} \underbrace{P(N_t = 0)}_{e^{-\lambda \tau}} = \frac{e^{-\lambda \tau}}{\sqrt{2\pi\sigma^2}} e^{-\frac{z_1^2}{2\sigma^2}} \delta(z_2 - z_1)$$

③

$$\gamma_z(\tau) = e^{-\lambda \tau} \sum_{k=0}^{\infty} \gamma_x^{(k)} \frac{(\lambda \tau)^k}{k!}$$

$$x_n = \alpha x_{n-1} + u_n$$

$$\{u_n\} \text{ iid. } E[u_n] = 0, E[u_n^2] = \beta$$

$$|\alpha| < 1$$

$$\gamma_x^{(n)} = E[x_n \cdot x_{n-k}]$$

$$\gamma_x^{(0)} = E[x_n^2] = \underbrace{\alpha^2 E[x_{n-1}^2]}_{\gamma_x^{(0)}} + \underbrace{E[u_n^2]}_{\beta} + \underbrace{2\alpha E[x_{n-1} \cdot u_n]}_0 \Rightarrow \gamma_x^{(0)} = \frac{\beta}{1-\alpha^2}$$

$$E[x_n \cdot x_{n-k}] = \alpha E[x_{n-1} \cdot x_{n-k}] + \underbrace{E[u_n \cdot x_{n-k}]}_0 \Rightarrow \gamma_x^{(k)} = \alpha \gamma_x^{(k-1)}$$

$$\gamma_x^{(k)} = \frac{\beta}{1-\alpha^2} \cdot \alpha^k \quad |\alpha| < 1$$

$$\gamma_z(\tau) = e^{-\lambda \tau} \underbrace{\frac{\beta}{1-\alpha^2}}_{\sigma_x^2} \sum_{k=0}^{\infty} \frac{(\lambda \tau)^k}{k!} \alpha^k = \underbrace{\frac{\sigma_x^2}{\sigma_x^2} e^{-\lambda(1-\alpha)\tau}}_{1}$$

Conditioning on Gaussian Vectors

Linear square estimation of parameters

→ Antenna Processing

→ Spectrum Estimation

* GAUSSIAN CONDITIONING

$$\underline{z} = \begin{bmatrix} \underline{x} \\ \underline{y} \end{bmatrix}$$

$$\underline{z} \in \mathbb{R}^{n+m}$$

$$E[\underline{z}] = 0$$

$$R_z = E[\underline{z} \underline{z}^T]$$

$$R_z = \begin{bmatrix} E[\underline{x} \underline{x}^T] & E[\underline{x} \underline{y}^T] \\ E[\underline{y} \underline{x}^T] & E[\underline{y} \underline{y}^T] \end{bmatrix} = \begin{bmatrix} R_x & R_{xy} \\ R_{yx} & R_y \end{bmatrix}$$

$$R_{yx} = R_{xy}^T$$

$$\underline{z} \sim N(0, R_z) \rightarrow p(\underline{z}) = \frac{1}{|2\pi R_z|^{1/2}} \exp -\frac{1}{2} \underline{z}^T R_z^{-1} \underline{z}$$

$\hat{X}(y) = E[X|Y] =$ best. m.s. estm of X observing Y

$$= \int x p(x|y) dx = \frac{\int x p_2(x, y) dx}{p_y(y)}$$

→ calculate R_z^{-1}

$$\text{let } M = R_z^{-1} = \begin{bmatrix} A & B^T \\ B & C \end{bmatrix}; \begin{bmatrix} R_x & R_{xy} \\ R_{yx} & R_y \end{bmatrix} \begin{bmatrix} A & B^T \\ B & C \end{bmatrix} = \begin{bmatrix} I_n & 0 \\ 0 & I_m \end{bmatrix}$$

$$R_{yx} A + R_y B = 0 \Rightarrow B = -R_y^{-1} R_{yx} A \Rightarrow B^T = -A^T R_{xy} R_y^{-1}$$

$$R_x A + R_{xy} B = I$$

$$x^T R_z^{-1} \underline{z} - y^T R_y^{-1} y = [x^T \ y^T] \begin{bmatrix} A & B^T \\ B & C \end{bmatrix} \begin{bmatrix} x \\ y \end{bmatrix} - y^T R_y^{-1} y = x^T A x + x^T B^T y + y^T C y - y^T R_y^{-1} y$$

$$= (x + A^{-1} B^T y)^T A (x + A^{-1} B^T y) - y^T [B A^{-1} B^T + R_y^{-1}] y$$

Extremum at

$$x^* = -A^{-1} B^T y = \underbrace{R_{xy} R_y^{-1}}_{L_{xy}} y = \hat{X}(y) = E[X|Y]$$

$$\text{Cov}(X|Y) = \text{Cov}(X - \hat{X}(y)) = A^{-1}$$

$$A^{-1} = R_x + R_{xy} B A^{-1}$$

$$= R_x - R_{xy} R_y^{-1} R_{yx} = R_x - L_{xy} R_y L_{xy}^T$$

* Now minimize directly

$$\min_L \mathbb{E} \left\{ \underbrace{\| \underline{x} - L \underline{y} \|^2}_{\varepsilon} \right\}$$

for any (n-u necessarily joint gaussian)
random vectors $\underline{x}, \underline{y}$

Reminder: Many square matrix M $\text{tr}(M) = \sum_{k=1}^n M_{kk} = \sum \text{eigen values} = \sum_{k=1}^n \lambda_k$

$\underline{x}, \underline{y}$ vectors

$$\frac{\underline{x}^T \cdot \underline{y}}{\sum x_k y_k} = \text{tr}(\underline{x} \cdot \underline{x}^T)$$

$$\text{Var}(\underline{\varepsilon}) = \mathbb{E}\{\|\underline{\varepsilon}\|^2\} = \mathbb{E}(\underline{\varepsilon}^T \cdot \underline{\varepsilon}) = \mathbb{E}[\text{tr}(\underline{\varepsilon} \cdot \underline{\varepsilon}^T)] = \text{tr} \mathbb{E}[\underline{\varepsilon} \underline{\varepsilon}^T] \quad \mathbb{E}[\underline{w}] = \bar{\underline{w}}$$

$$= \text{tr} \left\{ \underbrace{\mathbb{E}[\underline{x} \underline{x}^T]}_{R_x} - L \underbrace{\overline{\underline{y} \underline{x}^T}}_{R_{yx}} - \underbrace{\overline{\underline{x} \underline{y}^T}}_{R_{xy}} L^T + L \underbrace{\overline{\underline{y} \underline{y}^T}}_{R_y} L^T \right\}$$

$$= \text{tr} \left\{ (L - R_{xy} R_y^{-1}) \underbrace{R_y}_{\text{positive definite}} (L - R_{xy} R_y^{-1})^T \right\} + \text{tr} \{ R_x - R_{xy} R_y^{-1} R_{yx} \}$$

positive definite: $\forall \underline{w} : \underline{w}^T R_y \underline{w} > 0$

$\text{Var}(\varepsilon)$ is minimized by

$$L = R_{xy} \cdot R_y^{-1}$$

$$\min_L \text{Var}(\varepsilon) = \text{tr} \left\{ R_x - \underbrace{R_{xy} R_y^{-1} R_{yx}}_{L R_y L^T} \right\}$$

geometrical Interpretation:

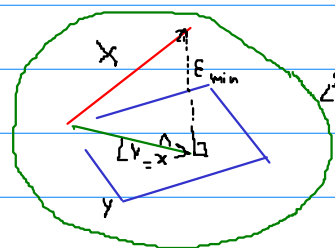
$$\mathbb{E}[\underline{\varepsilon}^T \cdot \underline{y}] = 0$$

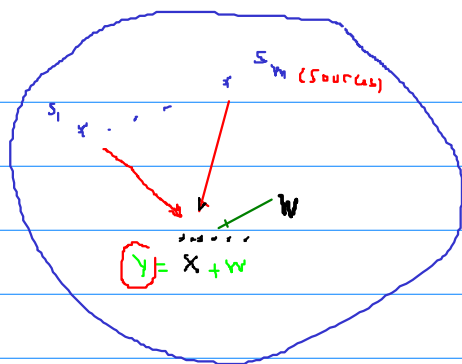
$$\text{tr} \mathbb{E}[(\underline{x} - L \underline{y}) \underline{y}^T] = 0$$

$$\underbrace{\mathbb{E}[\underline{x} \cdot \underline{y}^T]}_{R_{xy}} - L \underbrace{\mathbb{E}[\underline{y} \underline{y}^T]}_{R_y} = 0$$

(Normal Equations)

$$\text{Cov}(\underline{\varepsilon}) = \text{Cov}(\underline{x}) - \text{Cov}(L \underline{y})$$





VAN TREES : Estimation, Detection & Modulation (2 vol.)

LIBGEN

$$\begin{cases} \vec{x} = \sum_{i=1}^m \theta_i \vec{s}_i & \text{Antenna} = \vec{s} \cdot \vec{\theta} \\ \vec{y} = \vec{x} + \vec{w} & \text{Random Noise} \end{cases}$$

Annotations: θ_i is circled in red and labeled "Deterministic". \vec{s}_i is circled in green and labeled "Random".

Optimal Receiver

$$\begin{aligned} L = R_{xx} R_y^{-1} &= \overline{x(x+w)^T} \left\{ \overline{(x+w)(x+w)^T} \right\}^{-1} \\ &= R_x [R_x + R_w]^{-1} = \underbrace{\vec{s} \underbrace{\vec{\theta} \vec{\theta}^T}_{C_\theta} \vec{s}}_{C_\theta} [\underbrace{\vec{s} C_\theta \vec{s}^T + R_w}_{C_\theta}]^{-1} \end{aligned}$$

$$\vec{s} = [\vec{s}_1, \dots, \vec{s}_m], \quad \vec{\theta} = \begin{bmatrix} \theta_1 \\ \vdots \\ \theta_m \end{bmatrix}$$

* Hypothesis: Noise is 'white', isotropic $\rightarrow R_w = \beta I$

$$\rightarrow L = R_x R_y^{-1} = (R_y - R_w) R_y^{-1} = I - R_w R_y^{-1} = I - \beta R_y^{-1}$$

* $R_y = R_x + \beta I \rightarrow$ \forall Eigenvectors of R_x are eigenvectors of R_y
 \mathcal{E} and \mathcal{E}^\perp are eigenspaces of R_y

$$\bullet R_y = \beta I + \sum_{n=1}^m C_{\theta_n} \vec{s}_n \vec{s}_n^T$$

$$\bullet R_y \vec{s}_i = \beta \vec{s}_i + \sum_{n=1}^m C_{\theta_n} \vec{s}_n \cdot \vec{s}_i \in \mathcal{E} \quad : \quad R_y \mathcal{E} \subseteq \mathcal{E}$$

$$\bullet \text{ If } \vec{v} \in \mathcal{E}^\perp, \quad R_y \vec{v} = \beta \vec{v} + \sum_{n=1}^m C_{\theta_n} \vec{s}_n \cdot \vec{v} = \beta \vec{v} \in \mathcal{E}^\perp$$

$$\bullet \text{ Let } R_x = \overline{xx^T} = \sum_{i=1}^k \lambda_i \vec{u}_i \vec{u}_i^T \quad : \begin{cases} \{\vec{u}_i\} \text{ are orthogonal for } R_x = R_x^T \\ \lambda_i \geq 0 \quad (R_x \text{ is positiv. definit.}) \end{cases}$$

Annotations: $\vec{u}_i \perp \vec{u}_j \quad i \neq j$ and $\{\vec{u}_i\}$ span \mathcal{E} are written in red.

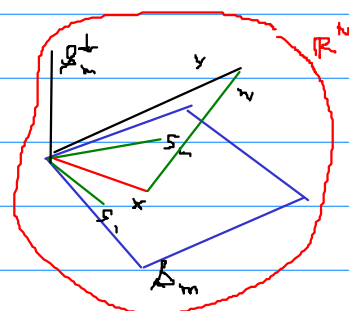
$$R = \underline{U} \underline{\Lambda} \underline{U}^T = \underline{U} \underline{\Lambda} \underline{U}^+ = [\vec{u}_1, \dots, \vec{u}_k] \begin{bmatrix} \lambda_1 & & 0 \\ & \ddots & \\ 0 & & \lambda_k \end{bmatrix} \begin{bmatrix} \vec{u}_1^T \\ \vdots \\ \vec{u}_k^T \end{bmatrix}$$

* Let $\vec{u}_{k+1}, \dots, \vec{u}_N$ span \mathcal{E}^\perp and $\vec{u}_i \perp \vec{u}_j = 0 \quad i \neq j, \quad i, j = k+1 : N$

$$I = \sum_{i=1}^N \vec{u}_i \vec{u}_i^T \Rightarrow R_w = \beta I = \underbrace{\beta \sum_{i=1}^k \vec{u}_i \vec{u}_i^T}_{\text{Proj on } \mathcal{E}} + \underbrace{\beta \sum_{i=k+1}^N \vec{u}_i \vec{u}_i^T}_{\text{Proj on } \mathcal{E}^\perp}$$

Annotation: The second sum is circled in red and labeled "projector in direction \vec{u}_i ".

$$R_y = \sum_{i=1}^k \mu_i \vec{u}_i \vec{u}_i^T + \beta \sum_{i=k+1}^N \vec{u}_i \vec{u}_i^T$$



$$* \quad R_Y = \sum_{i=1}^m \underbrace{(\lambda_i + \beta)}_{\mu_i} \underline{u}_i \underline{u}_i^T + \beta \sum_{i=m+1}^N \underline{u}_i \underline{u}_i^T, \quad \lambda_i \geq 0 \Rightarrow \beta = \underline{\text{least eigenvalue of } R_Y}$$

μ_i = 'Dominant' eigenvalues of R_Y

$$R_Y^{-1} = \sum \mu_i^{-1} \underline{u}_i \underline{u}_i^T + \frac{1}{\beta} \sum \underline{u}_i \underline{u}_i^T$$

$$\underline{L} = \underline{I} - \beta R_Y^{-1} = \sum_{i=1}^m \left(1 - \frac{\beta}{\mu_i}\right) \underline{u}_i \underline{u}_i^T$$

$$1 - \frac{\beta}{\mu_i} = 1 - \frac{\beta}{\lambda_i + \beta} = \frac{\lambda_i}{\lambda_i + \beta} = \frac{\lambda_i/\beta}{1 + \lambda_i/\beta}$$

$$\text{if } \beta/\mu_i \downarrow 0 : \underline{L} \rightarrow \underline{P}_{\text{reg}} = \sum_{i=1}^m \underline{u}_i \underline{u}_i^T$$

* EXERCISE

$$\text{Var}(\underline{E}_{\text{min}}) = \text{tr} \{ \underline{R}_x - \underline{R}_{xy} \underline{R}_y^{-1} \underline{R}_{yx} \}$$

$$= \beta \sum_{i=1}^m \frac{\lambda_i}{\beta + \lambda_i} \leq m\beta$$

$$\text{Var}(\underline{w}) = N\beta$$

$$\frac{\text{Var}(\underline{E}_{\text{min}})}{\text{Var}(\underline{w})} \leq \frac{m}{N}$$

* VAN TREES

* SHELDON ROSS

Applied probability