

# MAX. ENTROPY ESTIMATION OF SPECTRUM

Stationary 2nd order process

$$\{x_1, x_2, \dots\}$$

$$E x_i = 0 ; E |x_i|^2 < \infty$$

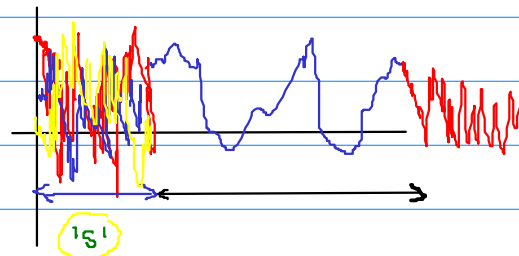
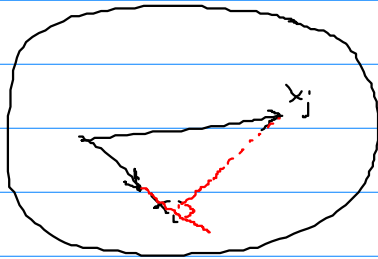
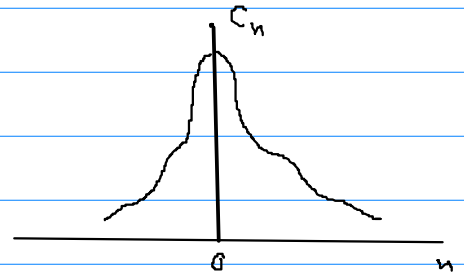
$$E \{x_i x_j\}^2 \leq E \{x_i^2\} E \{x_j^2\}$$

$< \infty$

(Schwarz  
Inequality)

$$E \{x_i x_{i+n}\} = c_n$$

$$|c_n|^2 < \infty$$



$$\{x_1, \dots, x_N\}$$

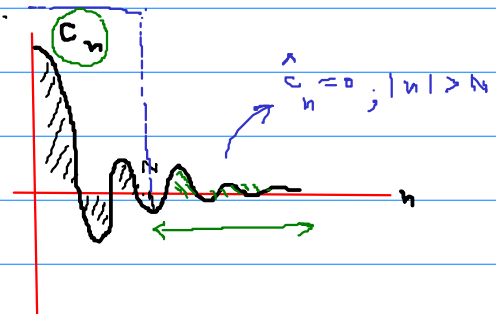
Reminder:  $c_n = E \{x_k x_{k+n}\} \rightarrow \begin{cases} \Gamma(\lambda) = \sum_{n=-\infty}^{\infty} c_n e^{-i2\pi\lambda n} = \text{spectral power density} \\ \text{or spectrum} \end{cases}$

$$\hat{\Gamma}(\lambda) = f(x_1, \dots, x_N)$$

$$\hat{\Gamma}(\lambda) = \sum_{n=-\infty}^{\infty} \hat{c}_n e^{-i2\pi\lambda n}$$

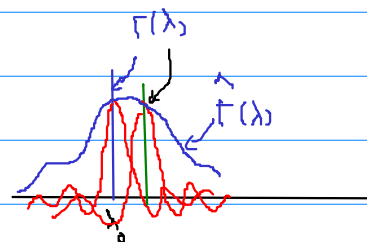
$$\hat{c}_n = \frac{1}{N} \sum_{k=1}^{N-|n|} x_k x_{k+n}$$

$$N > N$$



$$\hat{c}_n = c_n \cdot \phi_n \xrightarrow{FT} \Gamma(\lambda) \otimes \hat{\phi}(\lambda)$$

Rect.  
width = N



$$\left. \begin{aligned} f(\lambda) &\geq 0 \\ \Phi[f] &= 0 \\ \dots \\ \Phi_N[f] &= 0 \end{aligned} \right\} \rightarrow f(\lambda)$$

$$\begin{aligned} 0 &\leq E \{(X - \alpha Y)^2\} = E \{X^2 - 2\alpha E \{XY\} + \alpha^2 E \{Y^2\}\} \\ &= C - 2\alpha B + \alpha^2 A \\ &\rightarrow B^2 - AC \geq 0 \\ &\rightarrow E \{X^2\} E \{Y^2\} \geq E \{XY\}^2 \end{aligned}$$

# Max. Entropy Spectrum Estimation

Inf. theory  $\rightarrow$  Coding

$\{x_i\}_{i=1}^{\infty}$  stoch. proc.

$f_n(X_{1:n})$ ,  $n=1,2,\dots$  partial distributions

$$h_n(x_1, \dots, x_n) = h_n(X_{1:n}) = - \int f(x_{1:n}) \log f(x_{1:n}) dx^n = h[f] \quad (1)$$

JAYNES  
SHORE SOLUTION

$$h = \lim_{n \rightarrow \infty} \frac{h_n(X_{1:n})}{n} \quad \text{exists for stat. sources. (2)} \quad f(x_{1:n})$$

Shannon-McMillan-Breiman

\* For Gaussian process

$$h = \frac{1}{2} \log[2\pi e] + \frac{1}{4\pi} \int_{-\pi}^{\pi} \log[2\pi \Gamma(\lambda)] d\lambda \quad (2) \quad C_{ij} = E[x_i x_j]$$

BURG'S ALGORITHM

$\downarrow$   
Covers, Thomas,  
Haykin

$$\Gamma(\lambda) = \frac{1}{2\pi} \sum_{l=-\infty}^{\infty} \sigma(l) e^{-i\lambda l} = \text{spectral density}$$

$$\sigma(l) = E[x_n x_{n+l}] \equiv C_l \quad (4)$$

Problem: Estimate  $\Gamma(\lambda)$  or  $\sigma(l)$  for all  $l=0,1,2,\dots$

Subject to constraints:

$$\begin{cases} \sigma(0) = \alpha_0, \dots, \sigma(p) = \alpha_p, & \hat{\sigma}(p+1), \hat{\sigma}(p+2), \dots = ?? \\ \{\sigma(l)\}_{l=-\infty}^{\infty} \text{ Auto-Covariance} \end{cases}$$

## THEOREM

The stoch. proc.  $\{x_i\}$  that maximizes the diff. Entropy subject to  
corr. constraints  $E[x_i x_{i+k}] = \alpha_k$   $k=0:p$   $i=1,2,\dots$  (5)  
is the  $p$ th order Gauss. Markov process satisfying these constraints

Rem.  $\{x_i\}$  is not assumed Gaussian, nor stationary

Proof:

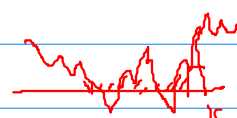
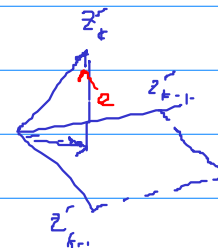
$X_{1:n}$  any G.M. of order  $p$  satisfy (5)

$Z_{1:n}$  Zero mean, Normal with cov. given by (5)

$Z'_{1:n}$   $p$ th order Gauss-Markov process. Satisfying (5)

$$Z'_k = g_1 Z'_{k-1} + \dots + g_p Z'_{k-p} + u_k$$

$u_k$  i.i.d



Then for  $n \geq p$

$$h(X_{1:n}) \stackrel{(a)}{\leq} h(Z_{1:n})$$

$$\leq h(Z_{1:p}) + \sum_{k=p+1}^n h(Z_k | Z_{1:k-1}) \quad (b) \quad \text{Chain Rule for Entropy}$$

$$\leq h(Z_{1:p}) + \sum_{k=p+1}^n h(Z_k | Z_{k-p:k-1}) \quad (c) \quad h(A|B,C) \leq h(A|B)$$

$$= h(Z'_{1:p}) + \sum_{k=p+1}^n h(Z'_k | Z'_{k-p:k-1}) \quad \leftarrow \text{Markov property of SRF}$$

$$\stackrel{(d)}{=} h(Z'_{1:n}) + \sum_{k=p+1}^n h(Z'_k | Z'_{1:k-1}) = h(Z'_{1:n})$$

$$(a) \quad D[f||q] \triangleq \int f \log \frac{f}{q} \geq 0$$

$f$ : density of  $\{X\}$

$$\phi(x) = N(x|0, R), \quad R = E[XX^T]$$

Conclusion: the  $p$ th order Gauss-Markov Process with covariance  $\alpha_{0,p}$  has higher entropy  $h(Z'_{1:n})$  than any other process satisfying the auto-corr. constraints

$$\Rightarrow \lim_{n \rightarrow \infty} \frac{1}{n} h(X_{1:n}) \leq \lim_{n \rightarrow \infty} h(Z'_{1:n}) = h$$

for all stoch. process - satisfying constraints.

EQUIVALENT CHARACTERISATION:

$$\begin{cases} X_n = \sum_{i=1}^p \alpha_i X_{n-i} + u_n \end{cases} \quad \{u_n\} \text{ iid}, \quad u_n \sim N(0, \sigma_u^2)$$

$$\sigma_l^2 = E[X_n \cdot X_{n-l}]$$

$$\sigma_l^2 = \overline{X_n X_{n-l}} = \sum_{j=1}^p \alpha_j \underbrace{\overline{X_{n-i} X_{n-l}}}_{\sigma_{l-i}} + \underbrace{\overline{u_n X_{n-l}}}_{\sigma_{l-p}} \quad l > 0$$

$u_{n-p}, u_{n-p-1}, \dots$   
 $\overline{u_n \cdot X_{n-p}} = 0$

$$\begin{cases} \sigma_l^2 = \sum_{j=1}^p \alpha_j \sigma_{l-j}^2 & l > 0 \\ \sigma_0^2 = \sum_{i,j} \alpha_i \alpha_j \sigma_{i-j}^2 + \sigma_u^2 \end{cases}$$

FOUR-WALKER GR.

$$H(\lambda) = \frac{1}{1 - \sum_{j=1}^p \alpha_j e^{i2\pi j \lambda}}$$

$$\Gamma(\lambda) = \frac{\sigma_u^2}{\left| 1 - \sum_{j=1}^p \alpha_j e^{-i2\pi j \lambda} \right|^2}$$

$$\begin{cases} \xrightarrow{u_n} H(\lambda) \rightarrow x_n \\ \Gamma(\lambda) = \sigma_u^2 \cdot |H(\lambda)|^2 \end{cases}$$

→ Fitting a Gauss-Markov model to data (Linear Prediction)

$\{x_n\}$  stat. process.

$$E_n = x_n - L_{nm} x_{n-1:n-m}$$

$L_{nm}$  = opt. projection of  $x_n$  /  $x_{n-1:n-m}$

$$L_{nm} = E[x_n \cdot y_{nm}^T] \cdot R_{y_{nm}}^{-1}$$

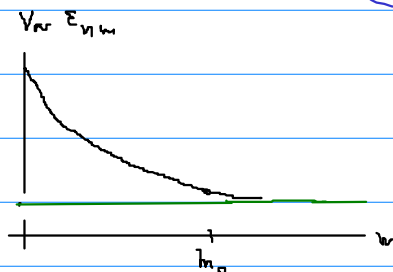
$\text{Var}(E_n)$  decreases with  $n$

$$E_{nm} = x_n - L_{nm} y_n$$

$$\text{Var } E_{n,m+1} \leq \text{Var } E_{n,m}$$

Wold's theorem

$$= \sigma_{min}^2$$



$m_0 \approx$  order at which

$x_n \approx$  Gauss-Markov



{ Estimate parameters  $\hat{a}_k$   
By M.S. method

Good Ref for Stoch. Process:

- 1/ GIKHMAN, SKOROKHOV Random Processes (Dover)
- 2/ TOPOLOVIC Intro St. proc. & Applic (Springer)

Next Lecture

- 21/04  $\mathcal{H}_{-1}^H$
- 22/04  $\mathcal{H}_{+5}^H - \mathcal{H}_{-5}^H$

(Libgen...) § 6-7

$$L = E[x \cdot y^T] E[y \cdot y^T]^{-1}$$

