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FINAL PROJECT REPORT

A Safety Signal Analysis with Three-Level Hierarchical Mixture

Model

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In this manuscript we follow the footsteps of others <sup>1</sup> and fit a three-level hierarchical mixture model into a dataset featuring adverse effects of trials across different body systems, and make inference on the comparison of treatment and control groups.

**KEYWORDS:** 

Safety Signal, Three-Level Hierarchical model, Adverse Effects(AEs), Body Systems, Multiplicities;

1 | INTRODUCTION

To our example we are going to apply a three-level hierarchical mixture model, proposed by Berry & Berry<sup>1</sup>. The reference literature was initially produced in order to address the multiplicity issues when considering potentially drug-caused adverse effects across and within different body systems.

The idea is to model different adverse effects during drug trials both across and within the body systems with a Bayesian approach through a three-level hierarchical model, that way we could make posteior inferences on the AEs simultaneously without losing power as the frequentists perspecitive would on this matter. For details we refer to our main reference <sup>1</sup>.

Here we are just going to introduce the model proposed by Berry&Berry.

2 | MODEL

Assuming in a drug trial we have control (total number  $N_c$ ) and treatment (total number  $N_t$ ) groups. Suppose there are B body systems and within body system b there are  $k_b$  types of AEs labeled  $A_{bj}$ , where b = 1, ..., B and  $j = 1, ..., k_b$ . Of the  $N_c$  controls,

 $X_{bj}$  experience  $A_{bj}$  and of the  $N_T$  patients in the treatment group,  $Y_{bj}$  experience  $A_{bj}$ . The probabilities of experiencing  $A_{bj}$  are  $c_{bj}$  and  $t_{bj}$  for control and treatment patients, respectively <sup>1</sup>. We use logistic transformations:

$$\gamma_{bj} = \log\left(\frac{c_{bj}}{1 - c_{bj}}\right)$$
 and  $\theta_{bj} = \log\left(\frac{t_{bj}}{1 - t_{bj}}\right) - \gamma_{bj}$ 

So given the data  $X_{bj}$  and  $Y_{bj}$ , there is:

$$X_{bj} \sim \text{Bin}(c_{bj}, N_c)$$
 and  $Y_{bj} \sim \text{Bin}(t_{bj}, N_t)$ 

Here  $\gamma_{bj}$  and  $\theta_{bj}$  are level 1 prior parameters and we assume:

$$\gamma_{bj} \sim N(\mu_{\gamma b}, \sigma_{\gamma}^2)$$
 for  $b = 1, ..., B$  and  $j = 1, ..., k_b$ .

For  $\theta_{bj}$  it represents the log odds ratio of AEs in treatment group versus in control group. It is modeled with a mixture prior distribution that

$$\theta_{bj} \sim \pi_b I_{[0]} + (1 - \pi_b) N(\mu_{\theta b}, \sigma_{\theta b}^2)$$
 for  $b = 1, ..., B; j = 1, ..., k_b$ .

The above expression can also be equivalently phrased as:

$$\theta_{bj} = (1 - \min_{1,bj}) \times \min_{2,bj}$$

with

$$\min_{1,bj} \sim \text{Bernoulli}(\pi_b) \text{ and } \min_{2,bj} \sim N(\mu_{\theta b}, \sigma^2_{\theta b})$$

We take this form in our OpenBUGS coding later on.

Now for the second level prior parameters we have  $\mu_{\gamma b}$ ,  $\sigma_{\gamma}^2$ ,  $\pi_b$ ,  $\mu_{\theta b}$ , and  $\sigma_{\theta b}^2$ . Assume:

$$\mu_{\gamma b} \sim N(\mu_{\gamma 0}, au_{\gamma 0}^2)$$
 for  $b=1,\ldots,B$  and  $\sigma_{\gamma}^2 \sim \mathrm{IG}(\alpha_{\sigma \gamma},\beta_{\sigma,\gamma})$ 

we also have

$$\pi_b \sim \text{Beta}(\alpha_{\pi}, \beta_{\pi}), b = 1, \dots, B$$

Here the  $\pi_b$  depends on the body systems, so the model considers different probabilities of having same AE effect between control and treatment group in different body systems.

We also have

$$\mu_{\theta b} \sim N(\mu_{\theta 0}, \tau_{\theta 0}^2)$$
 for  $b = 1, ..., B$  and  $\sigma_{\theta b}^2 \sim \text{IG}(\alpha_{\theta}, \beta_{\theta})$ .

Finally for the third level prior parameters, we have  $\mu_{\gamma 0},\, au_{\gamma 0}^2,\, lpha_\pi,\, eta_\pi,\, \mu_{\theta 0}$  and  $au_{\theta 0}^2$ . We assume:

$$\mu_{\gamma 0} \sim N(\mu_{\gamma 00}, \tau_{\gamma 00}^2)$$
 and  $\tau_{\gamma 0}^2 \sim \mathrm{IG}(\alpha_{\tau \gamma}, \beta_{\tau \gamma})$ .

For  $\alpha_{\pi}$  and  $\beta_{\pi}$  as the hyper parameter of  $\pi_b$  in the beta distribution, we restrict them to be \$ > 1\$ and assume a left truncated exponential distribution:

$$\alpha_{\pi} \sim \frac{\lambda_{\alpha} \exp(-\alpha \lambda_{\alpha})}{\exp(-\lambda_{\alpha})} I_{[\alpha > 1]} \text{ and } \beta_{\pi} \sim \frac{\lambda_{\beta} \exp(-\beta \lambda_{\beta})}{\exp(-\lambda_{\beta})} I_{[\beta > 1]}$$

The reason of doing so is explained in detail, see (Berry & Berry 2004)<sup>1</sup>.

For constants in the model it is assumed as  $\mu_{\theta 00} = 0$ ,  $\tau_{\theta 00}^2 = 10$ ,  $\alpha_{\theta} = 3$ ,  $\beta_{\theta} = 1$ ,  $\alpha_{\theta 0} = 3$ ,  $\beta_{theta0} = 1$ ,  $\alpha_{\tau \gamma} = 3$ ,  $\beta_{\tau \gamma} = 1$ ,  $\alpha_{\sigma \gamma} = 3$ ,  $\beta_{\sigma \gamma} = 1$  and  $\lambda_{\gamma} = \lambda_{\beta} = 1$ .

Figure 1 provides an OpenBUGS doodle that helps with seeing the hierarchical structure. The variables inside eclipses are either stochastic or logical node, and those in rectangles are constants.

The notation we use in OpenBUGs is consistent with our reference literature, except that to avoid transform between precision and variance, everywhere where there is notation for variance in the paper, we use it as precision in our code. For example, when the paper mention  $\tau_{00}^2 = 10$ , we put it as  $\tau_{0.0}^2 = 0.1$  in the OpenBUGs code

#### 3 | ANALYSIS

We proceed the analysis with a few steps. We first fit the data presented by our reference <sup>1</sup> in OpenBUGs. In the meantime, to check consistency, we also utilize an R package to do the same job, and we compare between both results, and also with those

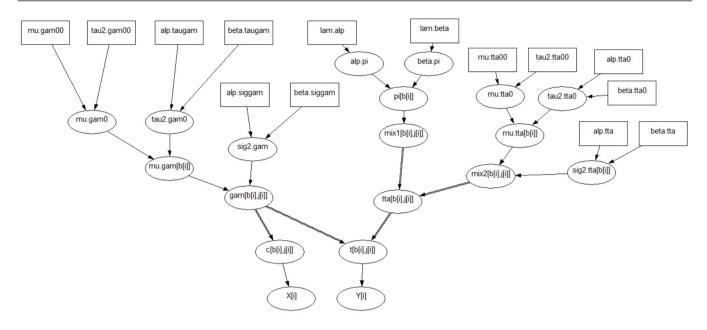


FIGURE 1 OpenBUGS doodle for the hierarchical model

presented by the paper. Then we fit the three-level hierarchical model to a different set of data in winBUGS, and we compare the results with an independent model.

# 3.1 | Fit data from Berry(2004)1 by WinBUGS and R separately

The reference paper 1 present the data as shown in Figure 2:

and the results are partially presented in the paper (we are mainly interested in the posterior  $\theta > 0$  probability since it reflects drug effects on AEs) as shown in Figure 3

The original authors are particularly interested in the 4 AEs that are significant in the Fisher-exact test since we want to know how the Bayesian hierarchical model concludes differently. Table 1 presents the results:

AEs	Index	$\theta > 0$	Fisher exact p
Irritability	$\theta_{8,3}$	0.780	0.003
Diarrhea	$\theta_{3,4}$	0.231	0.029
Rash	$\theta_{10,4}$	0.19	0.021
Rash, measles /rub-like	$\theta_{10,6}$	0.126	0.039

**TABLE 1** Fisher exact test vs Posterior probability

The key argument made by Berry & Berry is that the order of Fisher exact p values on these four AEs are not the same as the order of the posterior probability  $\theta > 0$ , and that is due to the reason that Fisher exact test only focus on individual AE while the hierarchicao model takes into account the body system factor.

				eatment = 148)	_	ontrol = 132)	22.005-1.00
b	j	Type of AE $A_{bj}$	$Y_{bj}$	Rate	$X_{bj}$	Rate	Fisher's exact p
1	1	Astenia/fatigue	57	0.385	40	0.303	0.167
1	2	Fever	34	0.230	26	0.197	0.561
1	3	Infection, fungal	2	0.014	0	0.000	0.500
1	4	Infection, viral	3	0.020	1	0.008	0.625
1	5	Malaise	27	0.182	20	0.152	0.525
3	1	Anorexia	7	0.047	2	0.015	0.179
3	2	Cendisiasis, oral	2	0.014	0	0.000	0.500
3	3	Constipation	2	0.014	0	0.000	0.500
3	4	Diarrhea	24	0.162	10	0.076	0.029*
3	5	Gastroenteritis	3	0.020	1	0.008	0.625
3	6	Nausea	2	0.014	7	0.053	0.089
3	7	Vomiting	19	0.128	19	0.144	0.730
5	i	Lymphadenopathy	3	0.020	2	0.015	1.000
6	1	Dehydration	0	0.000	2	0.015	0.221
8	1	Crying	2	0.014	0	0.000	0.500
8	2	Insomnia	2	0.014	2	0.005	1.000
8	3	Irritability	75	0.507	43	0.326	0.003*
9	1	Bronchitis	4	0.027	1	0.008	0.375
9	2	Congestion, nasal	4	0.027	2	0.015	0.375
9	3	Congestion, respiratory	1	0.007	2	0.015	0.603
9	4	Cough	13	0.088	8	0.061	0.497
9	5	Infection, upper respiratory	28	0.189	20	0.152	0.431
9	6	Laryngotracheobronchitis	2	0.014	1	0.008	1.000
9	7	Pharyngitis	13	0.014	8	0.061	0.497
9	8	Rhinorrhea	15	0.101	14	0.106	1.000
9	9	Sinusitis	3	0.020	1	0.100	0.625
9	10	Tonsillitis	2	0.020	1	0.008	1.000
9	11	Wheezing	3	0.020	1	0.008	0.625
10	1	Bite/sting	4	0.027	0	0.000	0.125
10	2	Eczema	2	0.014	0	0.000	0.500
10	3	Pruritis	2	0.014	1	0.008	1.000
10	4	Rash	13	0.014	3	0.023	0.021*
10	5	Rash, diaper	6	0.000	2	0.015	0.288
10	6	Rash, measles/rubella-like	8	0.054	1	0.013	0.239*
10	7	Rash, varicella-like	4	0.034	2	0.005	0.687
10	8	Urticaria	0	0.027	2	0.015	0.221
10	9	Viral exanthema	1	0.007	2	0.015	0.603
11	1	Conjunctivitis	0	0.007	2	0.015	0.003
11	2	Otitis media	18	0.000	14	0.015	0.221
							1.000
11	3	Otorrhea	2	0.014	1	0.008	1.00

FIGURE 2 Data from Berry & Berry's paper

While I agree with the authors, our OpenBugs and R both return similar results to each other but different than those presented in the paper.

Particularly, we did our own OpenBugs code based on the model, while the R result actually comes from a package named "c212", which is developed by Dr Raymond Carragher from Strathclyde Institute of Pharmacy and Biomedical Sciences.

For OpenBugs we burned 1,000 iterations and take the next 10,000 updates, then we extract the coda information of  $\theta$  and compute the mean number of  $\theta$ s of interest that is positive to estimate the posterior  $P(\theta > 0)$ . For R the package will also generate point estimate for  $\theta$  and we compute the posterior  $P(\theta > 0)$  in the same way.

Figure 4 and Figure 5 shows the trace and history plot from OpenBUGS iteration. Graphs show that our model converges all right.

Table 2 presents the comparison among the three for the four AEs:

			Post probability					Post probability	
b	j	Type of AE	$\theta = 0$	$\theta > 0$	b	j	Type of AE	$\theta = 0$	$\theta > 0$
1	1	Astenia/fatigue	0.762	0.211	9	4	Cough	0.906	0.062
1	2	Fever	0.827	0.122	9	5	Infection, respiratory	0.897	0.083
1	3	Infection, fungal	0.796	0.101	9	6	Bronchitis	0.898	0.047
1	4	Infection, viral	0.813	0.100	9	7	Pharyngitis	0.906	0.061
1	5	Malaise	0.826	0.116	9	8	Rhinorrhea	0.904	0.051
3	1	Anorexia	0.821	0.117	9	9	Sinusitis	0.903	0.051
3	2	Cendisiasis, oral	0.835	0.083	9	10	Tonsillitis	0.905	0.042
3	3	Constipation	0.812	0.101	9	11	Wheezing	0.907	0.050
3	4	Diarrhea	0.743	0.231	10	1	Bite/sting	0.859	0.087
3	5	Gastroenteritis	0.823	0.093	10	2	Eczema	0.860	0.070
3	6	Nausea	0.805	0.050	10	3	Pruritis	0.868	0.062
3	7	Vomiting	0.849	0.076	10	4	Rash	0.784	0.190
5	1	Lymphadenopathy	0.717	0.136	10	5	Rash, diaper	0.852	0.099
6	1	Dehydration	0.666	0.087	10	6	Rash, measles/rub-like	0.836	0.126
8	1	Crying	0.655	0.185	10	7	Rash, varicella-like	0.862	0.076
8	2	Insomnia	0.661	0.153	10	8	Urticaria	0.852	0.048
8	3	Irritability	0.214	0.780	10	9	Viral exanthema	0.855	0.055
9	1	Bronchitis	0.900	0.059	11	1	Conjunctivitis	0.721	0.079
9	2	Congestion, nasal	0.901	0.058	11	2	Otitis media	0.757	0.102
9	3	Congestion, respiratory	0.896	0.040	11	3	Otorrhea	0.749	0.121

FIGURE 3 Berry & Berry's paper results

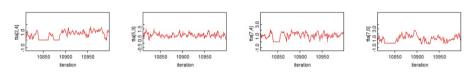


FIGURE 4 Trace Plot from OpenBugs

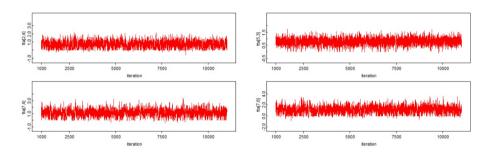


FIGURE 5 Trace Plot from OpenBugs

From table 2 we can conclude the following:

- 1. OpenBUGs and R return similar estiamtes on posterior probability of treatment effects on AEs, but different than the original reference paper. Our analysis suggests that for all four adverse effects, there is a significant effect from the treatment (high probability of  $\theta > 0$ ).
- 2. Posterior probability of treatment effect on irritability is strong, consistent with Fisher exact test. In the mean time, due to body system effect the extent of treatment effect for these adverse effects are not completely in the same order as the p values from Fisher exact tests, although the difference is nearly ignorable.

We argue that since the original paper was published in 2004 and over the time the computational software has evolved tremendously, the result from the original paper may be subject to further verification.

		po			
AEs	Type	Reference	OpenBugs	R	Fisher exact p
Irritability	$\theta_{8,3}$	0.78	0.978	0.984	0.003
Diarrhea	$\theta_{3,4}$	0.231	0.847	0.853	0.029
Rash	$\theta_{10,4}$	0.19	0.945	0.993	0.021
Rash, measles/rub-like	$\theta_{10,6}$	0.126	0.890	0.946	0.039

**TABLE 2** OpenBugs and R results compare with reference

### 3.2 | Fit the new example with both OpenBUGs and R and compare with an independent model

Figure 6 presents a new set of data, similar to the one from our main reference 1:

			Treatme	ent (N <sub>7</sub> =320)	Control	$(N_c=320)$
b	j	Type of SAE A <sub>bj</sub>	$Y_{bj}$	Rate	$X_{bj}$	Rate
1	1	Arrhythmia	0	0.000	1	0.003
1	2	Increased BP	3	0.009	3	0.009
1	3	Other CB AEs	0	0.000	2	0.006
1	4	Pre-eclampsia	8	0.025	6	0.019
2	1	Emesis	0	0.000	1	0.003
2	2	Other GI AEs	0	0.000	1	0.003
3	1	Depression	0	0.000	1	0.003
3	2	Headache	0	0.000	1	0.003
3	3	Other HNMB AEs	1	0.003	0	0.000
4	1	Gestational Diabetes Mellitus	2	0.006	0	0.000
4	2	Other MAN AEs	0	0.000	1	0.003
5	1	Chorioamnionitis	0	0.000	2	0.006
5	2	Decreased Fetal Movement	1	0.003	2	0.006
5	3	Endomyometritis	1	0.003	0	0.000
5	4	Miscarriage	0	0.000	1	0.003
5	5	Other PD AEs	3	0.009	1	0.003
5	6	Postpartum Hemorrhage	1	0.003	0	0.000
5	7	Premature Delivery	5	0.016	5	0.016
5	8	Premature ROM	2	0.006	10	0.031
5	9	Preterm Contractions	2	0.006	4	0.013
6	1	Other RESP AEs	0	0.000	1	0.003
6	2	Shortness of Breath	0	0.000	1	0.003
7	1	Other UG AEs	1	0.003	0	0.000
7	2	Pyelnephritis	1	0.003	4	0.013
7	3	Urinary Tract Infection	1	0.003	1	0.003
7	4	Vaginal Bleeding	1	0.003	1	0.003
8	1	Abdominal Pain	1	0.003	2	0.006
8	2	Other BODY AEs	0	0.000	3	0.009
8	3	PD012	2	0.006	0	0.000
8	4	Pelvic Pain	1	0.003	0	0.000
8	5	Polyhydramnios	0	0.000	1	0.003

FIGURE 6 Safety Signal Data

We fit the data with winBUGS and R. For the winBUGS part this time we also fit an independent model, which means that every  $\gamma_{bj}$  and  $\theta_{bj}$  has its own normal prior distribution. Figure 7 shows an OpenBUGS doodle version of the model:

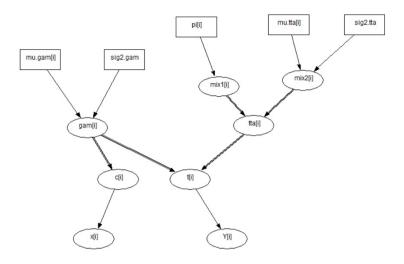


FIGURE 7 OpenBugs doodle for Independent Model

Table 3 compares the posterior probability  $P(\theta > 0)$  between winBUGS results for hierarchical and independent model, and also the R package output for the hierarchical model.

From Table 3 we conclude that when using hierarchical model we tends to be more conservative about claiming that there is a treatment effect on the adverse effects. Both OpenBUGS and R gave similar results to the posterior probability  $P(\theta > 0)$ .

However the independent model gives in general larger values of this probability, if we focus on comparing row by row. When using independent model, we are considering each AE without taking into consideration of body system, so this is essentially in the same spirit of making inference with fisher exact test on each individual adverse effect.

Take some observations for example, for "Other HNMB AEs", the hierarchical models give a posteior probability fairly small (0.046 from R and 0.053 from OpenBUGS), while the independent model gives a value up to 0.20.

Should we claim a treatment effect on this particular AE? In this case the posteior probability is pretty small even for independent model. So the answer is probably no.

Suppose the posterior probability is quite large for independent model, say pass over 0.5, but the value given by hierarchical model is pretty small, then we need to think over before making a decision. What seems to be a treatment effect might actually be caused by the body system itself.

### 4 | CONCLUSION

We have explored the three-level hierarchical model proposed by Berry & Berry, and tried to fit different datasets into the model.

Both R and OpenBUGS give similar outcomes when it comes to estimating the posterior probabilities.

Although we did not show for every dataset in every model, the model generally converges quite well. (The R package does not provide convergence check).

	post $P(\theta > 0)$				
AE	R(Hierarchical)	OpenBUGS(Hierarchical)	OpenBUGS(Independent)		
Arrhythmia	0.0207	0.0267	0.0496		
Increased BP	0.0487	0.0591	0.0939		
Other CB AEs	0.0136	0.0237	0.0227		
Pre-eclampsia	0.0763	0.1102	0.1058		
Emesis	0.0258	0.0264	0.0457		
Other GI AEs	0.022	0.0249	0.0463		
Depression	0.0229	0.0275	0.0464		
Headache	0.0232	0.0301	0.0508		
Other HNMB AEs	0.046	0.0529	0.2012		
Gestational Diabetes Mellitus	0.0763	0.0961	0.2920		
Other MAN AEs	0.0331	0.0341	0.0476		
Chorioamnionitis	0.0061	0.0156	0.0240		
Decreased Fetal Movement	0.0199	0.0281	0.0661		
Endomyometritis	0.037	0.0413	0.1984		
Miscarriage	0.0062	0.0194	0.0493		
Other PD AEs	0.0634	0.0718	0.2238		
Postpartum Hemorrhage	0.0309	0.0406	0.1871		
Premature Delivery	0.0486	0.0671	0.0708		
Premature ROM	0.0059	0.0109	0.0045		
Preterm Contractions	0.0186	0.0331	0.0434		
Other RESP AEs	0.0113	0.0232	0.0472		
Shortness of Breath	0.0105	0.0251	0.0474		
Other UG AEs	0.0466	0.0508	0.1887		
Pyelnephritis	0.0254	0.0313	0.0272		
Urinary Tract Infection	0.0466	0.0452	0.1047		
Vaginal Bleeding	0.0402	0.0473	0.1046		
Abdominal Pain	0.0335	0.0434	0.0616		
Other BODY AEs	0.0125	0.0213	0.0118		
PD012	0.0502	0.074	0.2885		
Pelvic Pain	0.0414	0.0542	0.1897		
Polyhydramnios	0.0198	0.0268	0.0483		

TABLE 3 hierarchical model vs independent model, and winBUGS vs R

In the mean time, we particularly compared the results between fitting the hierarchical model and the independent model. We conclude that the hierarchical model is more conservative when it comes to attributing the adverse effects to the effect of treatment, due to the consideration of body system effect.

Finally, it seems that we could not reproduce the exact same estimate on the data provided in the reference paper. However both our OpenBUGS and R program return similar results to each other. We argue that the difference than the original publication might due to the change of computational software over a time period of more than a decade.

## References

1. Berry Scott M., Berry Donald A.. Accounting for Multiplicities in Assessing Drug Safety: A Three-Level Hierarchical Mixture Model. *Biometrics*. 2004;:418-426.