

# Micro- versus meso-scale population models for ecological forecasting under climate change

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## Abstract

The basic idea here is to do a head-to-head comparison of complex, individual-based integral projection models and less complex, “patch-based” models. The problem is this: as ecologists we want to forecast the impacts of global change, like climate change, at scales relevant to land management decisions and policy intervention. However, almost all studies of plant population dynamics rely on data at the sub 1 meter scale. This data is really hard to collect. Can we use average plot cover to forecast plant population dynamics just as well as when we use complex IPMs? Can “averaged-out” population models at the plot level adequately capture climate drivers? These are the questions we seek to answer here.

**Keywords:** population model, forecasting, integral projection model, etc.

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# 1 Introduction

Producing skillfull forecasts at spatial scales relevant to global environmental changes that play out at the landscape and regional scale has proven challenging for population ecologists. The major hurdle is that most population models, at least for plant species, are built using data from small, localized plots because parameterizing vital rate regressions requires tracking the fates of individuals. These models are difficult to scale up from the micro to meso-scales because the fitted parameters do not fully represent the spatial variation present at spatial scales beyond that at which the data are collected.

Recently, Freckelton et al., building on work by Hastings and NAME, have proposed density-structured population models that focus on the transition of populations among discrete states, rather than the traditional approach of modeling the transitions of individuals. Such an approach could be extremely valuable because the data needed to parameterize density-structured population models is much easier, and less costly, to collect (Queenborough et al. year). For example, using a density-structured approach, one could build a population model using a time series of annual plot-based censuses of species percent cover. However, a major assumption of the density-structured approach is that the aggregate dynamics of the population observed at coarse spatial resolution faithfully represent, and correspond to, the fates of individual plants. In other words, using a density-structured approach requires a leap of faith that important covariates (e.g., climate variables) at the level of the individual are captured adequately at the population level. If we seek to forecast the impacts of climate change on plant populations, then clearly this assumption requires testing.

## 2 Materials and Methods

### 2.1 Study site and data

Our demographic data comes from the Fort Keogh Livestock and Range Research Laboratory in eastern Montana’s northern mixed prairie near Miles City, Montana, USA (46 deg. 19’ N, 105 deg 48’ W). The dataset is freely available on Ecological Archives (CITE), and interested readers should refer to the metadata therein for a complete description. The site is about 800 m above sea level and mean annual precipitation (1878-2009) is 334 mm, with most annual precipitation falling from April through September (76). The site is grass dominated and, for the purposes of our study, we focus on the four most abundant graminoid species: *Bouteloua gracilis*, *Hesperostipa comata*, *Pascopyrum smithii*, and *Poa secunda*.

From 1932 to 1945 individual plants were identified and mapped annually in 44 1-m<sup>2</sup> quadrats using a pantograph. The quadrats were distributed in six pastures, each assigned a grazing treatment of light (1.24 ha/animal unit month), moderate (0.92 ha/aum), and heavy (0.76 ha/aum) stocking rates (two pastures per treatment). In this analysis we account for potential differences among the grazing treatments, but do not focus on grazing×climate interactions. The annual maps of the quadrats were digitized and the fates of individual plants tracked and extracted using a computer program. Daily climate data, which we aggregated into climate variables of interest, are available for the duration of the data collection period (1932 - 1945) from the Miles City airport, Wiley Field, 9 km from the study site.

## 2.2 Stastical models of vital rates

The first step in building our single-species population models was to fit statistical models of vital rates at both levels of inference: individual-level and quadrat-level. Here we describe the general statistical approach at the individual-level, but the same approach applies at the quadrat-level. We modeled survival and growth as functions of genet size (quadrat cover in the quadrat-level models) and climate covariates (described in more detail below). We maintained a consistent random effects structure for both models that included three terms: (1) a random year effect on the intercept, (2) a random year effect on the coefficient for plant cover (either individual or quadrate level), and (3) a random effect of group (see Data set description) on the intercept. Vital rates for each species are modeled separately.

We used logistic regression to model survival probability:

$$\text{logit}(s) = \beta_{0,t} + \beta_{s,t}x + \beta_Q + \beta_{c,1}\theta_{1,t} + \cdots + \beta_{c,i}\theta_{i,t} + \varepsilon_t \quad (1)$$

where  $x$  is the log of genet size (or log of quadrat areal cover),  $\beta_{0,t}$  is a year-specific intercept,  $\beta_{s,t}$  is the year-specific slope parameter for size,  $\beta_Q$  is the random effect of quadrat group location,  $\theta$  is a matrix of  $i$  climate effects over  $t$  years,  $\beta_{c,i}$  is the fixed parameter for the effect of the  $i$ th climate covariate, and  $\varepsilon_t$  is the error term.

We modeled growth as gaussian process describing genet size (or quadrat cover) at time  $t + 1$  as a function of size at  $t$  and climate covariates:

$$x_{t+1} = \beta_{0,t} + \beta_{s,t}x_t + \beta_Q + \beta_{c,1}\theta_{1,t} + \cdots + \beta_{c,i}\theta_{i,t} + \varepsilon_t \quad (2)$$

where  $x$  is genet size and all other paramters are as described for the survival regression.

For the quadrat-level approach we modeled growth as a process describing proportional

cover within a quadrat at time  $t + 1$  as a function of proportional cover at time  $t$  and  
climate covariates. Thus, instead of a gaussian process, with a normal likelihood, we  
modeled growth at the quadrat level as above but with a beta likelihood and a logit link  
to the linear predictors.

While our approach for modeling survival and growth is similar at the individual and  
quadrat levels, the addition of new genets at the individual level or occupied sites at  
the quadrat level requires two separate approaches.

## 2.3 Selecting climate covariates

For both the individual-level IPM (IPM) and the quadrat-based IBM (QBM) we followed  
the same model selection approach for including climate covariates in vital rate regres-  
sions. Working within the random effects structure described previously, we fit a set of  
models with all possible combinations of four climate covariates: fall through spring pre-  
cipitation at  $t-1$  and  $t-2$  (ppt1 and ppt2, respectively) and mean spring temperature at  
 $t-1$  and  $t-2$  (TmeanSpr1 and TmeanSpr2, respectively), where  $t$  is the observation year.  
The combinations of climate covariates we fit included same-year interactions between  
precipitation and temperature (e.g., a ppt1  $\times$  TmeanSpr1 interaction). Finally, we also  
fit a model with no climate covariates included as fixed effects. In total, this resulted in  
a set of 25 candidate models.

We based model selection on the Deviance Information Criteria (DIC). However,  
for many species there were several models with  $\Delta\text{DIC} < 4$ , indicating that each model  
received generally the same amount of support. Given this ambiguity, we opted to simply  
select the candidate model with lowest DIC, regardless of the statistical nearness of other  
models, as the final model to use for simulating the population dynamics. In some cases  
this resulted in selecting a model with no climate covariates. We had to choose a single

statistical model for each vital rate for our population models, but when comparing  
106 climate covariate selection between the individual and quadrat levels we acknowledge  
climate covariates selected in all models with  $\Delta\text{DIC} < 4$ .

## 108 **2.4 Computing**

### **2.4.1 Model fitting**

110 All of our analyses (model selection, fitting, and simulating) were conducted in R. We  
used the analytical Bayesian software INLA to fit candidate vital rate models at the  
112 individual and quadrat levels for computational efficiency. This is the case except for  
the recruitment model at the individual level which required a more complex hierarchical  
114 approach. In that case we used the MCMC sampler in JAGS to estimate the posterior  
distributions of model parameters. After model selection, we used MCMC samplers,  
116 again in JAGS, to estimate the posterior distributions of model parameters for all vital  
rate models at the individual and quadrat levels. We obtained posterior distributions for  
118 all model parameters from three MCMC chains run for 10,000 iterations, after discarding  
the initial 5,000 iterations. We assessed convergence visually and using the Gelman  
120 diagnostic in the R package ‘coda.’

### **2.4.2 Population models**

122 With the posterior distribution of the vital rate statistical models in hand, it is straight-  
forward to simulate the population models.

### 3 Results

We assessed the statistical importance of the climate covariates included the final vital rate regressions by comparing the residual deviance of models with climate covariates and temporal random effects, climate covariates only, and temporal random effects only. When a model includes climate covariates, this comparison shows the relative contribution of the climate covariates in explaining the total interannual variability (Adler et al. 2012).

Table 1: The proportion of interannual variability in vital rates explained by the climate covariates. The contribution is defined as:  $(\text{Climate model} - \text{Constant Model}) / (\text{Full model} - \text{Constant model})$ . The values are residual deviance. Note that in some cases this will be slightly negative where DIC selected a model with climate covariates but a model without climate covariates had a very similar DIC value.

	N	Constant model	Climate model	Full model	Contribution of climate covariates
Growth					
BOGR	213	-881.25	-881.25	-938.35	0
HECO	171	-1365.35	-1372.06	-1381.61	0.41
PASM	217	-2107.02	-2116.98	-2164.8	0.17
POSE	197	-1503.69	-1502.32	-1551.44	-0.03
Survival					
BOGR	218	44.62	44.62	44.53	0
HECO	179	51.21	51.21	51.01	0
PASM	221	40.14	40.14	40.51	0
POSE	204	46.41	46.41	46.25	0
Colonization					
BOGR	29	38.32	39.24	44.53	0.15
HECO	68	75.78	75.78	51.01	0
PASM	26	32.35	32.35	40.51	0
POSE	43	54.05	54.05	46.25	0