

Consistent ecosystem functional response across precipitation extremes in a sagebrush steppe

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ABSTRACT

Precipitation is predicted to become more variable in the western United States, meaning years of above and below average precipitation will become more common. Periods of drought and deluge are major drivers of ecosystem functioning in water limited grasslands, but how ecosystems respond to precipitation may change as the duration of above and below average periods increases. Changes in ecosystem functional response could reflect underlying changes in species composition or species reaching physiological thresholds at extreme precipitation levels. Here we report the results of a five-year experiment where we used drought and irrigation treatments (approximately 50% decrease/increase) to investigate whether ecosystem functional response remains consistent at extreme precipitation levels. We also use species-level abundance data to link ecosystem functional response to community dynamics. Ecosystem functional response, defined as the relationship between soil moisture and aboveground net primary productivity, was surprisingly consistent across drought, control, and irrigation treatments. Aboveground net primary productivity (ANPP) responded positively to available soil moisture, but the response was weak and did not vary across treatments. There was also no evidence that treatment effects grew over time. The similarity of ecosystem functional response across treatments was not due to compensatory shifts at the plant community level, where species composition among treatments was similar and remarkably stable over the five years. At least in the short-term, ecosystem functional response and community composition in this sagebrush steppe system is resistant to increases and decreases in growing season precipitation, perhaps due to bet-hedging strategies of constituent species.

1 INTRODUCTION

Projected changes in precipitation regimes, associated with global climate change, will affect

Water availability is a strong driver of aboveground net primary productivity (ANPP) in grassland ecosystems. Mean annual precipitation (MAP) is strongly correlated with ANPP across sites (Huxman et al., 2004), and, in water-limited ecosystems, interannual variation in MAP is also correlated with interannual variation in ANPP (Hsu et al., 2012). These relationships suggest that chronic alterations in levels of precipitation, associated with global climate change, will impact ANPP.

At a given site, the functional response of ANPP to water availability (e.g., soil moisture) can be characterized by fitting a model to historical observations of ANPP and soil moisture. However, the fitted functional response may provide an incomplete picture because future conditions are likely to be outside the historical range of variability (Smith, 2011). For example, historical trends may show a linear soil moisture-ANPP relationship, but that relationship could potentially saturate if soil moisture is pushed far beyond typical levels. Saturating relationships are actually common (Hsu et al., 2012), perhaps because

46 other resources, like nitrogen, become more limiting in wet years than dry years. Knowing the curvature
47 of the soil moisture-ANPP relationship is critical for understanding how ecosystems will respond to
48 chronic alterations in water availability.

49 Another problem with relying on historical ecosystem functional responses is that they are not static.
50 Changes in species identities and abundances can alter an ecosystem's functional response to water
51 availability because different species have different physiological thresholds for producing biomass.
52 Smith et al. (2009) introduced the 'Hierarchical Response Framework' (HRF) for understanding the
53 interplay of community composition and ecosystem functioning in response to resource manipulations
54 over time. In the near term, ecosystem functioning such as annual net primary productivity (ANPP) will
55 reflect the physiological responses of individual species to the manipulated resource level. For example,
56 ANPP may decline under simulated drought because the initial community consisted of drought-intolerant
57 species (Hoover et al., 2014). Over longer time spans, ecosystem functioning may recover as new species
58 colonize or initial species reorder in relative abundance. For example, ANPP may initially decline,
59 but eventually rise back to pre-treatment levels once drought-tolerant species become more abundant
60 and compensate for drought-intolerant species (Hoover et al., 2014). It is also possible that ecosystem
61 functioning shifts to a new mean state, reflecting the suite of species in the new community (Knapp et al.,
62 2012).

63 Manipulating potentially limiting resources offers a route to understanding how ecosystems will
64 respond to resource levels that fall outside the historical range of variability (Avolio et al., 2015; Gherardi
65 and Sala, 2015; Knapp et al., 2017). Altering the amount of precipitation over many years should provide
66 insight into the time scales at which water-limited ecosystems respond to chronic resource alteration.
67 Following the HRF, we propose four alternative predictions for the effect of precipitation manipulation on
68 the ecosystem functional response to soil moisture, that is, the soil moisture-ANPP relationship (Figure
69 1). The four predictions are based on possible outcomes at the community (e.g., community composition)
70 and ecosystem (e.g., soil moisture-ANPP regression) levels.

71 First, altered precipitation changes neither ecosystem functional response nor community composition
72 (Fig. 1, top left). In this case, changes in ANPP simply follow the soil moisture-ANPP relationship under
73 ambient conditions. This corresponds to the early phases of the HRF, where ecosystem response is due to
74 the physiological responses of individual species. Second, the ecosystem functional response changes but
75 community composition remains the same (Fig. 1, top right). A saturating soil moisture-ANPP response
76 fits this scenario, where individual species hit physiological thresholds or are limited by some other
77 resource. Third, the ecosystem functional response is consistent but underlying community composition
78 changes (Fig. 1, bottom left). In this case, changes in species' identities or abundances occur in
79 response to altered precipitation levels and species more suited to the new conditions compensate for
80 reduced function of initial species. Fourth, and last, both ecosystem functional response and community
81 composition change (Fig. 1, bottom right). New species, or newly abundant species, with different
82 physiological responses completely reshape the ecosystem functional response.

83 All four outcomes are possible in any given ecosystem, but the time scales at which the different
84 scenarios play out likely differ (Smith et al., 2009; Wilcox et al., 2016; Knapp et al., 2017). Thus, our
85 task is not to test the validity of the HRF, but rather to amass information on how quickly ecosystem
86 functional responses change in different ecosystems. Likewise, we need to understand whether changes at
87 the ecosystem level are driven community level changes or individual level responses.

88 To that end, here we report the results of a five-year precipitation manipulation experiment in a
89 sagebrush steppe grassland. We imposed drought and irrigation treatments (approximately $\pm 50\%$) and
90 measured ecosystem (ANPP) and community (species composition) responses. We focus on how the
91 drought and irrigation treatments affect the relationship between available soil moisture and ANPP, and if
92 community dynamics underly the ecosystem responses. In particular, we are interested in the consistency
93 of the soil moisture-ANPP relationship among treatments. Is the relationship steeper under the drought
94 treatment, at low soil moisture? Does the relationship saturate under the irrigation treatment, at high soil
95 moisture? To answer these questions we fit a random intercepts, random slopes model to test whether the
96 regressions differed among treatments. We also analyzed community composition over time, allowing us
97 to place our experimental results within the framework of the HRF and our competing predictions (Fig.
98 1).

99 2 METHODS

100 2.1 Study Area

101 We conducted our precipitation manipulation experiment at the United States Sheep Experimental Station
102 (USSES) near Dubois, Idaho (44.2° N, 112.1° W), 1500 m above sea level. The vegetation is typical of
103 high elevation sagebrush steppe. The plant community is dominated by the shrub *Artemisia tripartita* and
104 three perennial bunchgrasses, *Pseudoroegneria spicata*, *Poa secunda*, and *Hesperostipa comata*. During
105 the period of our experiment (2011 – 2015), average mean annual precipitation was 265 mm year⁻¹ and
106 mean monthly temperature ranged from -5.2°C in January to 21.8°C in July.

107 2.2 Precipitation Experiment

108 Between 1926 and 1932, range scientists at the USSES established 26 permanent 1 m² quadrats to track
109 vegetation change over time. In 2007, we (well, one of us [P. Adler]) relocated 14 of the original quadrats,
110 six of which were inside a large, permanent livestock enclosure. We use these six plots as control plots
111 that have recieved no treatment, just ambient precipitation. In spring 2011, we (well, two of us [A.
112 Kleinhesselink and P. Adler]) established 16 new 1 m² plots. We avoided areas on steep hill slopes, areas
113 with greater than 20% cover of bare rock, and areas with greater than 10% cover of the shrubs *Purshia*
114 *tridentata* and/or *Amelanchier utahensis*. We established the new plots in pairs and randomly assigned
115 each plot in a pair to receive a “drought” or “irrigation” treatment.

116 Drought and irrigation treatments were designed to decrease and increase the amount of ambient
117 precipitation by 50%, respectively. To achieve this, we used a system of rain-out shelters and automatic
118 irrigation (Gherardi and Sala, 2013). The rain-out shelters consisted of transparent acrylic shingles 1-1.5
119 m above the ground that covered an area of 2.5 × 2 m. The shingles intercepted approximately 50% of
120 incoming rainfall, which was channeled into 75 liter containers. Captured rainfall was then pumped out
121 of the containers and sprayed on to the adjacent irrigation plot via two suspended sprinklers. Pumping
122 was triggered by float switches once water levels reached about 20 liters. We disconnected the irrigation
123 pumps each fall and reconnected them, often with difficulty, each spring. The rain-out shelters remained
124 in place throughout the year.

125 To make sure the treatments were having the desired effects, we monitored soil moisture in four of
126 the drought-irrigation pairs using Decagon Devices (Pulman, Washington) 5TM and EC-5 soil moisture
127 sensors. We installed four sensors in each plot, two at 5 cm soil depth and two at 25 cm soil depth. We
128 also installed four sensors in areas nearby the four selected plot pairs to measure ambient soil moisture at
129 the same depths. Soil moisture measurements were automatically logged every four hours. We coupled
130 this temporally intensive soil moisture sampling with spatially extensive readings taken at six points
131 within all 16 plots and associated ambient measurement areas. These snapshot data were collected on
132 06/06/2012, 04/29/2015, 05/07/2015, 06/09/2015, and 05/10/2016 using a handheld EC-5 sensor.

133 **ANDY:** add paragraph on the statistical modeling of soil moisture to produce Figure 2B (same as your
134 Figure 2 in the precip-demography-prediction dissertation chapter).

135 2.3 Data Collection

136 We estimated aboveground net primary productivity (ANPP) using a radiometer to relate ground re-
137 flectance to plant biomass (see Byrne et al., 2011, for a review). We recorded ground reflectance at
138 four wavelengths, two associated with red reflectance (626 nm and 652 nm) and two associated with
139 near-infrared reflectance (875 nm and 859 nm). At each plot in each year, we took four readings of ground
140 reflectances at the above wavelengths. We also took readings in 12 (2015), 15 (2012, 2013, 2014), or 16
141 (2016) calibration plots adjacent to the experimental site, in which we harvested all aboveground biomass,
142 dried it to a constant weight at 60°C, and weighed it to estimate ANPP.

143 For each plot and year, we averaged the four readings for each wavelength and then calculated a
144 greenness index based on the same bands used to calculate NDVI using the MODIS and AVHRR bands for
145 NDVI. To convert the greenness index to ANPP we regressed the greenness index against the dry biomass
146 weight from the ten calibration plots. We fit regressions to MODIS-based index and AVHRR-based index
147 for each year and retained the regression with the better fit. Using the best regression equation for each
148 year, we predicted ANPP (Appendix 1).

Did I say this
right?

Species composition data came from two sources: yearly census maps for each plot made using a pantograph (Hill, 1920) and yearly counts of annual species in each plot. From those maps, we have data on the density of all annuals and perennials forbs, basal cover of perennial grasses, and canopy cover of shrubs. We made a large plot-treatment-year by species matrix, where columns were filled with either basal cover or density, depending on the measurement made for the particular species. So we could analyze the different types of data together, we standardized the values in each column. This puts all abundance values on the same scale, meaning that common and rare species are weighted equally. Nonetheless, if we assume that rare species will respond more than common ones, then this approach is anti-conservative for detecting community change. This means that our approach is biased toward detecting compositional changes.

2.4 Data Analysis

Our main goal was to test whether the relationship between ANPP and soil moisture differed among the drought, control, and irrigation treatments. Based on our own observations and previous work at our study site (Blaisdell, 1958; Dalglish et al., 2011; Adler et al., 2012), we chose to use cumulative volumetric water content from March through June as our metric of soil moisture (hereafter referred to simply as ‘soil moisture’). To achieve this goal, we fit a multi-level random intercept and random slope regression with $\log(\text{ANPP})$ as the response variable and soil moisture as the sole predictor. We log-transformed ANPP to account for heteroscedasticity. We fit the model under a Bayesian framework, allowing us to test for treatment differences by comparing the posterior distributions of the treatment-level coefficients (e.g., Tredennick et al., 2013). Both $\log(\text{ANPP})$ and soil moisture were standardized to have mean 0 and unit variance before fitting the model [i.e., $(x_i - \bar{x})/\sigma_x$].

Our model is defined as follows:

$$\mu_{j(k(t))} = \beta_{0,j(k)} + \beta_{1,j(k)}x_{k(t)} + \gamma_t, \quad (1)$$

$$y_{j(k(t))} \sim \text{Normal}(\mu_{j(k(t))}, \sigma_k^2), \quad (2)$$

where $\mu_{j(k(t))}$ is the deterministic prediction from the regression model for plot j associated with treatment k in year t , $\beta_{0,j(k)}$ is the intercept for plot j associated with treatment k , $\beta_{1,j(k)}$ is the slope term for the effect of soil moisture for plot j associated with treatment k , γ_t is the intercept offset for year t , and σ_k^2 is the process variance for treatment k . Data include the standardized $\log(\text{ANPP})$ observations ($y_{j(k(t))}$) and soil moisture ($x_{k(t)}$).

The intercept and slope terms are modeled hierarchically to account for the non-independence of observations across years within plots and to allow us to test the hypothesis that our treatments alter the ANPP-soil moisture relationship. Thus, our multi-level model has three grouping levels for coefficients, representing the nested structure of the data: (i) overall coefficients, (ii) treatment coefficients, and (iii) plot coefficients. Each subsequent level is drawn from the distribution of coefficients at the previous level, meaning plot-level coefficients are drawn from treatment-level coefficients, which are drawn from overall coefficients. We also include a covariance structure among the intercept and slope at each level. Formally, our hierarchical structure is as follows, where we drop the intercept (0) and slope (1) subscripts and instead refer to a vector of coefficients, β :

$$\beta_{j(k)} \sim \text{MVN}(\beta_k, \Sigma(k)), \quad (3)$$

$$\beta_k \sim \text{MVN}(\beta, \Sigma), \quad (4)$$

$$\beta \sim \text{Normal}(0, 1), \quad (5)$$

where $\beta_{j(k)}$ is the vector of regression coefficients (intercept and slope) for plot j associated with treatment k , β_k is the vector of coefficients for each treatment, and β is the vector of overall coefficients. The plot- and treatment-level coefficients are drawn from multivariate normal distributions with covariance matrix Σ . For the plot-level coefficients, each treatment has its own variance-covariance matrix (i.e., $\Sigma(k)$). The overall coefficients are drawn from a normal prior with mean 0 and standard deviation 1. The random year effects (γ) are drawn from a normal prior with mean 0 and standard deviation σ_{year} , which was drawn from a weibull distribution. A full description of model is in Appendix 2.

We fit the model using a Bayesian approach, obtaining posterior estimates of all unknowns via the No-U-Turn Hamiltonian Monte Carlo sampler in Stan (Stan Development Team, 2016b). We used the R

194 package ‘rstan’ (Stan Development Team, 2016a) to link R (R Core Team, 2016) to Stan. We obtained
195 samples from the posterior distribution for all model parameters from four parallel MCMC chains run for
196 10,000 iterations, saving every 10th sample. Traceplots of all parameters were visually inspected to ensure
197 well-mixed chains and convergence. We also made sure all scale reduction factors (\hat{R}) were less than 1.1.

198 To see if community composition differed among treatments through time, we used non-dimensional
199 multivariate scaling (NMDS) based on Bray-Curtis distances. For each year of the experiment, we first
200 calculated Bray-Curtis distances among all plots, and then extracted those distances for use in the NMDS.
201 Because we standardized species’ abundances, some values were negative, which is not allowed for
202 calculating Bray-Curtis distances. We simply added ‘2’ to each abundance value to ensure all values
203 were greater than zero. We plotted the first two axes of NMDS scores to see if community composition
204 overlapped, or not, among treatments in each year. We used the ‘metaMDS()’ function in the R package
205 ‘vegan’ (Oksanen, 2016) to calculate Bray-Curtis distances and then to run the NMDS analysis. We used
206 the ‘vegan::adonis()’ function (Oksanen, 2016) to perform permutational multivariate analysis of variance
207 to test whether treatment plots formed distinct groupings. To test whether treatment plots were equally
208 dispersed, or not, we used the ‘vegan::betadisper’ function (Oksanen, 2016).

209 All R code and data necessary to reproduce our analysis has been archived on Figshare (*link here after*
210 *acceptance*) and released on GitHub ([https://github.com/atredennick/usses_water/](https://github.com/atredennick/usses_water/releases/v0.1)
211 [releases/v0.1](https://github.com/atredennick/usses_water/releases/v0.1)). We also include annotated Stan code in our model description in Appendix 2.

212 3 RESULTS

213 Three of our five treatment years fell in years of below average rainfall (Fig. 2A). Thus, those three years
214 represent a lower magnitude of absolute change in precipitation experienced by the treatments. Averaged
215 across treatments, ANPP varied from a minimum of 74.5 g m⁻² in 2014 to a maximum of 237.1 g m⁻² in
216 2016 (Fig. 2C). ANPP was slightly higher in irrigation plots and slightly lower in drought plots (Fig. 2C),
217 corresponding to estimated soil volumetric water content (VWC) differences among treatments (Fig. 2B).
218 Such differences in soil VWC indicate our treatment infrastructure was successful.

219 Cumulative March-June soil moisture had a weak positive effect on ANPP (mean of $\beta_1 = 0.41$), but
220 the effect is associated with high uncertainty, with credible intervals that overlap zero (80% BCI = -0.08,
221 0.86; 95% BCI = -0.45, 1.13) (Fig. 3). Ecosystem functional response was similar among treatments, with
222 treatment level regressions between soil moisture and log(ANPP) having similar intercepts and slopes
223 (Fig. 3A). There was also no evidence that the treatment effects became more important over time because
224 there was no directional trend in the random year effects (Fig. 4).

225 Community composition was similar among treatments. In no year did community composition among
226 treatments not overlap, and they were equally dispersed in all years (Table 1; Fig. 5). Likewise, community
227 composition was remarkably stable over time, with no evidence of divergence among treatments (Table 1;
228 Fig. 5).

229 4 DISCUSSION

230 Ecosystem response to chronic resource alteration is expected to follow a temporal trend. Initially,
231 ecosystem response will be modest and reflect the physiological responses of constituent species. Over
232 longer time periods, species reordering will cause greater responses as species better able to take advantage
233 of, or cope with, new resource levels become more abundant. This temporal trend is formalized by the
234 ‘Hierarchical Response Framework’ (HRF, Smith et al., 2009), and has been empirically supported (Knapp
235 et al., 2012; Wilcox et al., 2016). A lingering question is not whether the HRF is a reasonable model
236 of ecosystem dynamics, but rather how the time scales of ecosystem response and community change
237 differ among ecosystems. Indeed, previous work has shown that community compositional shifts can
238 be both rapid, on the order of years (Hoover et al., 2014), and slow, on order of decades (Knapp et al.,
239 2012; Wilcox et al., 2016). To add to this growing body of knowledge, we performed a precipitation
240 manipulation experiment in a sagebrush steppe ecosystem.

241 Ecosystem functional response (i.e., the relationship between soil moisture and ANPP) and community
242 composition in this sagebrush steppe ecosystem was consistent across drought, control, and irrigation
243 treatments (Fig. 2D, Fig. 3). In other words, functional responses estimated from any precipitation level

sampled apply to any other level of precipitation, and the consistent functional response was not due to species reordering and compensation. Thus, our results conform to the first of our four predictions: neither ecosystem functional response nor community composition changed with chronic alteration in water availability (Fig. 1, top left).

This is surprising because grasslands generally, and sagebrush steppe specifically, are considered water-limited systems. Thus, our naive expectations were that either (i) ecosystem functional response would differ among treatments as species reach physiological thresholds and resource limitations (Fig. 1, top right), or (ii) ecosystem functional response stays the same due to species reordering and compensation (Fig. 1, bottom left).

In the absence of community composition data, an obvious hypothesis to explain the consistent ecosystem functional response is that species reordered rapidly. In so doing, newly abundant species could compensate for the loss of ecosystem functioning of previously abundant species. We did not find this to be the case, as species composition was remarkably stable over the course of the experiment (Fig. 5).

In combination, the lack of ecosystem and community response to drought or irrigation shows that this sagebrush steppe ecosystem is resistant to chronic alterations in water availability. There are three possible explanations for the resistance we found. First, it could be that our experiment simply was not long enough to induce responses. This may be true for community responses, which can take decades (Wilcox et al., 2016).

Second, it could be that our manipulations were not large enough to induce a response. That is, maybe a 50% decrease/increase in any given year is not abnormal given our sites historical range of variability (Knapp et al., 2017). We cannot definitively rule out this possibility, but we have reason to believe our manipulations *should* have been large enough. Using the methods described by Lemoine et al. (2016), we calculated the percent reduction and increase of mean growing season precipitation necessary to reach the 1% and 99% extremes of the historical precipitation regime at our site. The 1% quantile of precipitation at our site is 110 mm, a 47% reduction from the mean, and the 99% quantile is 414 mm, a 77% increase from mean growing season precipitation (Appendix 3). Thus, our drought treatment represented extreme precipitation amounts, especially in years where ambient precipitation was below average (Fig. 2A). The irrigation treatment may have been too small, however.

Third, the ecosystem and community resistance to drought and deluge we observed could be a real phenomenon in this system, at least over the five year time span we observed. The perennial plants in this cold desert ecosystem are tolerant to drought conditions (Bazzaz, 1979; Franks, 2011, A.R. Kleinhesselink, unpublished data), meaning a consistent ecosystem functional response at low soil moisture is perfectly reasonable. For example, many of the perennial grasses in our focal ecosystem avoid drought stress by growing early in the growing season (Blaisdell, 1958, A.R. Kleinhesselink, personal observation). Likewise, the dominant shrub in our focal ecosystem, *Artemisia tripartita*, has access to water deep in the soil profile thanks to a deep root system (Germino and Reinhardt, 2014). As for the irrigation treatments, the lack of a saturating response (e.g., a lower regression slope) suggests water availability remained a limiting factor at the levels of precipitation we were able to manipulate.

In conclusion, our results suggest that five years of $\pm 50\%$ ambient precipitation is not enough to induce a shift in ecosystem functional response in a sagebrush steppe. This is likely because component species are tolerant to drought conditions and water remains co-limiting at higher precipitation levels. Longer time series of chronic precipitation alteration may reveal plant community shifts that we did not observe (e.g., Wilcox et al., 2016). Our results suggest compositional shifts would have the largest impact at high rainfall because the current community maintained consistent ecosystem functional response at very low water availability.

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300 **7 AUTHOR CONTRIBUTIONS**

- 301 • Andrew T. Tredennick collected data, analyzed the data, wrote the paper, prepared figures and/or
302 tables, reviewed drafts of the paper.
- 303 • Andrew R. Kleinhesselink conceived and designed the experiments, performed the experiments,
304 collected data, reviewed drafts of the paper.
- 305 • Peter B. Adler conceived and designed the experiments, performed the experiments, collected data,
306 analyzed the data, reviewed drafts of the paper.

8 TABLES

Table 1. Results from statistical tests for clustering and dispersion of community composition among precipitation treatments. ‘adonis’ tests whether treatments form unique clusters in multidimensional space; ‘betadisper’ tests whether treatments have similar dispersion. For both tests, P values greater than 0.05 indicate there is no support that the treatments differ.

Year	Test	n	d.f.	F	P
2011	adonis	21	2	1.02	0.45
2011	betadisper	21	2	2.23	0.14
2012	adonis	22	2	1.10	0.29
2012	betadisper	22	2	0.21	0.81
2013	adonis	22	2	1.23	0.14
2013	betadisper	22	2	0.28	0.76
2014	adonis	22	2	0.95	0.58
2014	betadisper	22	2	0.35	0.71
2015	adonis	21	2	1.05	0.37
2015	betadisper	21	2	3.01	0.07
2016	adonis	21	2	1.07	0.35
2016	betadisper	21	2	0.50	0.62

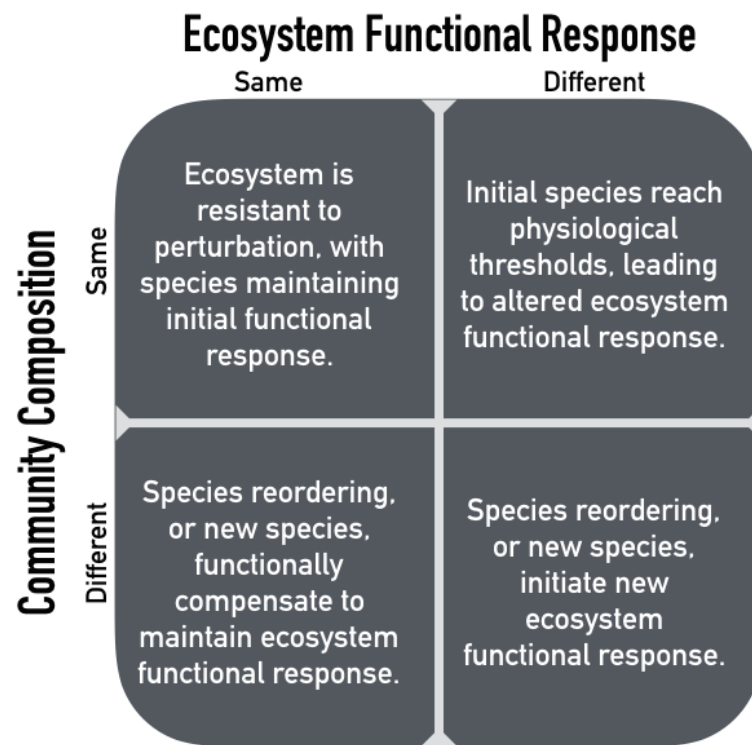


Figure 1. Possible outcomes of chronic resource alteration.

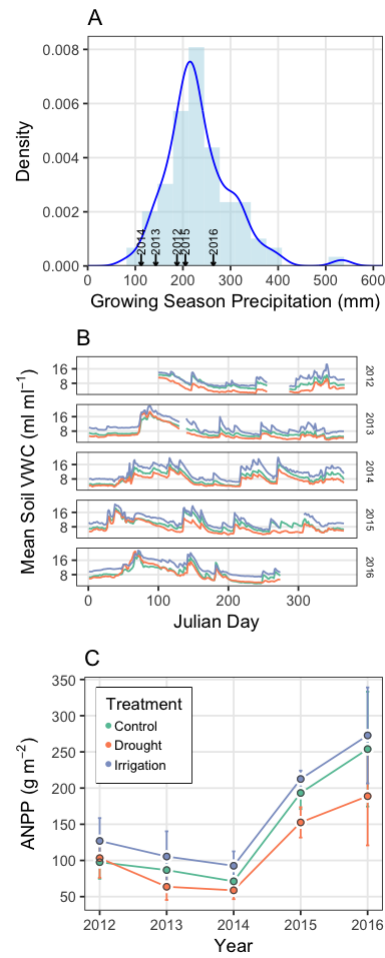


Figure 2. (A) Probability density of historical precipitation from 1926-2016, with the years of the experiment shown with arrows on the x -axis. (B) Observed soil volumetric water content (VWC) over the course of the experiment. (C) Mean (points) ANPP and its standard deviation (error bars) for each year of the experiment.

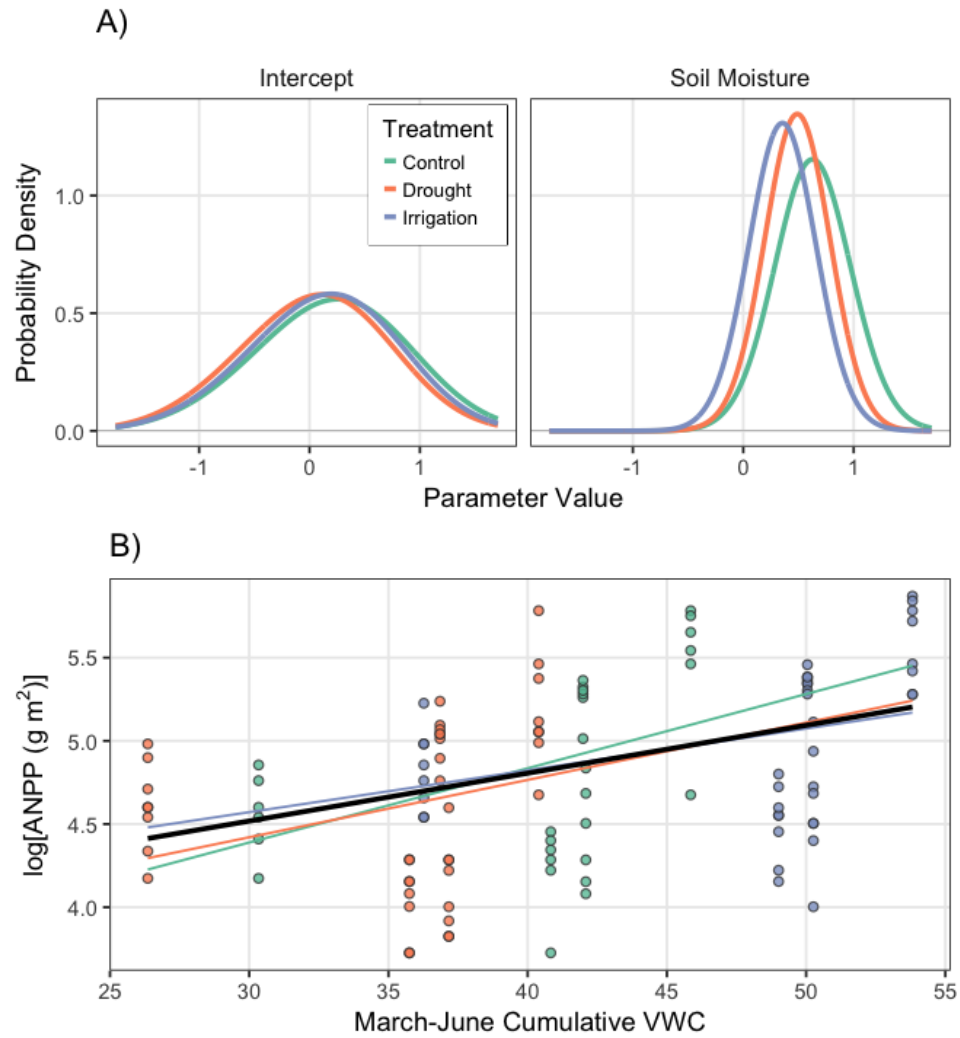


Figure 3. Results from the generalized linear mixed effects model. (A) Posterior distributions of treatment-level parameters ('Intercept' and the effect of 'Soil Moisture'). Kernel bandwidths of posterior densities were adjusted by a factor of 5 for visual clarity. (B) Scatterplot of the data and model estimates shown a solid lines. Model estimates come from treatment level coefficients (colored lines) and the overall mean coefficients (heavy black line). Note that we show log(ANPP) on the y-axis of panel B; this same plot can be seen on the arithmetic scale in supporting material Fig. S2-1.

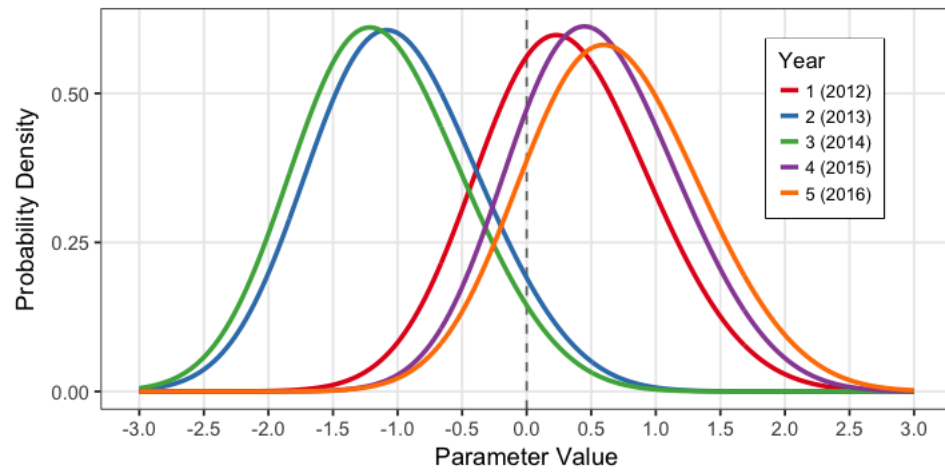


Figure 4. Posterior distributions of random year effects (intercept offsets). Kernel bandwidths of posterior densities were adjusted by a factor of 5 for visual clarity.

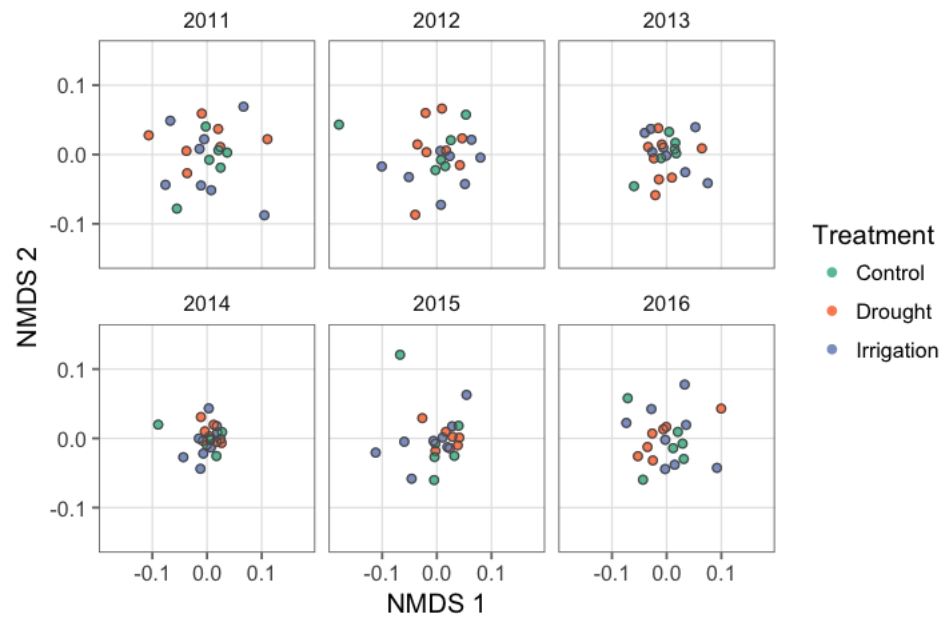


Figure 5. Nonmetric multidimensional scaling scores representing plant communities in each plot, colored by treatment.

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