MEDIATION AND MODERATION IN REPEATED MEASURES DESIGNS: RECENT DEVELOPMENTS AND FUTURE DIRECTIONS

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Overview

- Personal History and Philosophy
- Introduction to Mediation
- Two-Instance Repeated Measures Designs
 - Mediation
 - Moderation
 - Conditional Process Analysis (AKA Moderated Mediation)
- New Designs & New Questions

Philosophical Approach to Statistics

Quantitative Psychologist (A personal definition)

Evaluate, develop, and improve statistical methods that:

- 1. are relevant to the types of research questions, and
- 2. use the types of data collected

in substantive psychology.

- Clinical
- Developmental
- Health
- Personality
- Social
- Cognitive
- Neuroscience
- and more...

Evaluating Methods

- Comparing the performance of multiple methods which are available for the same type of research question and data
- Explore how one type of analysis, traditionally used to answer a different research question, might be useful for a new research question

Examples:

 When assessing conditional moderated mediation, should we use regression based approaches or structural equation modeling approaches?

Hayes, A. F., **Montoya, A. K.**, & Rockwood, N. J. (2017). The analysis of mechanisms and their contingencies: PROCESS versus structural equation modeling. Australasian Marketing Journal, 25 (1), 76 - 81.

 Can measures of model fit be used to select number of factors in an exploratory factor analysis?

Montoya, A. K., & Edwards, M. C. (2017, Apr) Cautions on using model t to choose number of factors in EFA. Presented at the 2017 Annual Meeting of the National Council on Measurement in Education, San Antonio, TX.

 Can models like mediation and moderated-mediation be used to better understand and predict differential item functioning in item response theory?

Montoya, A. K. & Jeon, M. (in prep) MIMIC models for testing uniform and non-uniform differential item functioning as moderated mediation models.

Developing New Methods

Research Question

- Comparing means?
- Relationship between variables?
- Explaining a process?

Properties of the Data

- Independent vs. Paired Data
- Level of measurement
- Multilevel / Longitudinal

Identify combination where researchers **want** to be able to answer the research question with this specific type of data

Question: How does a treatment condition compare to a control?

Data: Multiple control conditions collected

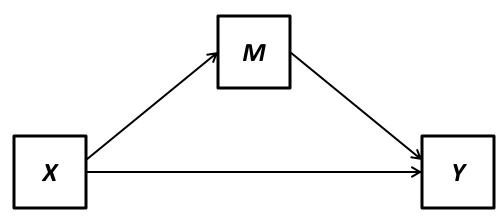
Montoya, A. K., Flaherty, B., & Cheryan, S. (2015, Feb) Collapsing conditions increasing type I error: Changing analysis of multiple control conditions. Presented at the Annual Meeting of the Society for Personality and Social Psychology, Long Beach, CA.

Question: For what levels of a moderating variable do group differ on some outcome?

Data: More than two groups

Hayes, A. F., & **Montoya, A. K.** (2017). A tutorial on testing, visualizing, and probing an interaction involving a multicategorical variable in linear regression. Communication Methods and Measures, 11 (1), 1-30.

Question: What Explains an Effect?



A simple mediation model connects an **assumed** causal variable (X) to an **assumed** outcome variable (Y), through some mechanism (M).

Many different kind of variables may act as mediators. Emotional variables, situational, individual level variables, cognitive variables, environmental variables, etc.

The goal of statistical mediation analysis is to determine if there is an effect of *X* on *Y* through *M*. This is typically done by estimating the *indirect effect* and testing if it is different than zero.

<u>Causality:</u> Causal order of the variables is an **assumption** of mediation. The term "effect" is used based on this assumption. The quality of causal inference is determined by study design and theory.

Examples of Mediation in Psychology

J Child Fam Stud (2015) 24:2038–2045 DOI 10.1007/s10826-014-0004-7



ORIGINAL PAPER

Eı of Di

The Influence of Mood on the Relation

Personality and Individual Differences 52 (2012) 828-832

Jon Yoı



Contents lists available at SciVerse ScienceDirect

Personality and Individual Differences

journal homepage: www.elsevier.com/locate/paid



Perfectionism, mattering, and depression: A mediational analysis

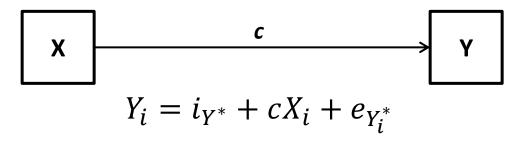
Gordon L. Flett^a, Ingrid Galfi-Pechenkov^{a,*}, Danielle S. Molnar^a, Paul L. Hewitt^b, Abby L. Goldstein^c

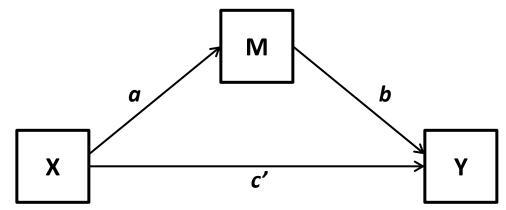
Caroline Davis^{1*}, Laura Mackew¹, Robert D. Levitan², Allan S. Kaplan², Jacqueline C. Carter³ and James L. Kennedy²

Mediation: Path Analysis

Consider *a*, *b*, *c*, and *c'* to be measures of the effect of the variables in the mediation model.

These could be measured using regression coefficients from OLS or path estimates in a structural equation model using maximum likelihood estimation.





Indirect effect of X on Y (through M) = $a \times b$

Direct effect of X on Y (not through M) = c'

Indirect effect = total effect - direct effect $a \times b = c - c'$

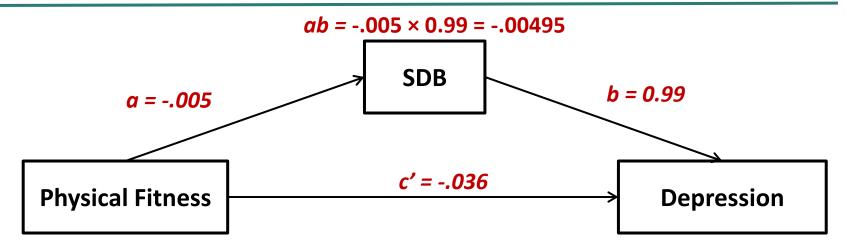
Total effect = direct effect + indirect effect

$$c = c' + a \times b$$

$$M_i = i_M + aX_i + e_{M_i}$$

$$Y_i = i_Y + c'X_i + bM_i + e_{Y_i}$$

Does Physical Fitness Impact Depression Through Sleep Disordered Breathing (SDB)?



Increase in physical fitness results in a *reduction* in sleep disordered breathing.

Controlling for physical fitness, increase in sleep disordered breathing results in an *increase* in depression

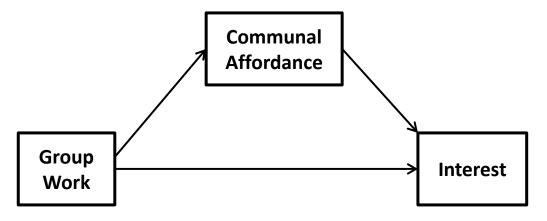
Indirect Effect: We estimate increase in physical fitness leads to a decrease in depression *through sleep disordered breathing*.

Direct Effect: Controlling for sleep disordered breathing, increase in physical fitness lead to a *decrease* in depression

What if I Use a Within-Subjects Design?

What if *X* is something we manipulate within-subjects instead of a between-subjects variable?

Study examining how group work in computer science classes might impact women's interest in computer science.



Study 1: Randomly assign individuals to read a computer science syllabus with group work or without group work

Study 2: Everyone reads two syllabi, one with group work and one without.

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Two-Instance Repeated Measures Designs

The causal variable of interest is the factor which differs by repeated measures.

X: varies between repeated measurements

M: measured in each of the two instances

Y: measured in each of the two instances

Examples:

- Participants read two scenarios. Interested in how scenario influences *Y* through *M*. Measure *M* and *Y* in each scenario.
- Pre-post test: Therapist measures certain symptoms and various outcomes before administering some intervention, and after administering the intervention.

Non-Examples:

- Does calorie consumption impact body image through weight gain over time?
- Any instance where repeated-measure factor is a "nuisance" (e.g. studying schools, but not interested in comparing schools directly).

Judd, Kenny, and McClelland (2001)

Judd, C. M., Kenny, D. A., & McClelland, G. H. (2001). Estimating and testing mediation and moderation in within-subject designs. *Psychological Methods*, *6*, 115-134.

A "causal steps", Baron and Kenny type logic to determining whether M is functioning as a mediator of X's effect on Y when both M and Y are measured twice in different instances but on the same people.

1. On average, does Y differ by instance?

Paired t-test on Y's

2. On average, does *M* differ by instance?

Paired t-test on M's

3. Does difference in *M* predict a difference in *Y*?

$$\widehat{Y_{1i} - Y_{2i}} = \widehat{c'} + \widehat{b}(M_{1i} - M_{2i}) + \widehat{d}(M_{1i} + M_{2i} - \frac{1}{n} \sum_{i=1}^{n} (M_{1i} + M_{2i}))$$

4. Does the difference in *M* account for all the difference in *Y*?

$$\widehat{Y_{1i} - Y_{2i}} = \widehat{c'} + \widehat{b}(M_{1i} - M_{2i}) + \widehat{d}(M_{1i} + M_{2i} - \frac{1}{n} \sum_{i=1}^{n} (M_{1i} + M_{2i}))$$

Montoya & Hayes, 2017

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Two-Condition Within-Participant Statistical Mediation Analysis: A Path-Analytic Framework

Amanda K. Montoya and Andrew F. Hayes The Ohio State University

Researchers interested in testing mediation often use designs where participants are measured on a dependent variable Y and a mediator M in both of 2 different circumstances. The dominant approach to assessing mediation in such a design, proposed by Judd, Kenny, and McClelland (2001), relies on a series of hypothesis tests about components of the mediation model and in so thased on an estimate of or formal inference about the indirect effect of the rectard to the class approach in the path-analytic framework that is now commonly used in between-participant mediation analysis. By so doing, it is apparent how to estimate the indirect effect of a within-participant manipulation on some outcome through a mediator as the product of paths of influence. This path-analytic approach eliminates the need for discrete hypothesis tests about components of the model to support a claim of mediation, as Judd et al.'s method requires, because it relies only on an inference about the product of paths—the indirect effect. We generalize methods of inference for the indirect effect widely used in between-participant designs to this within-participant version of mediation analysis, including bootstrap confidence intervals and Monte Carlo confidence intervals. Using this path-analytic approach, we extend the method to models with multiple mediators operating in parallel and serially and discuss the comparison of indirect effects in these more complex models. We offer macros and code for SPSS, SAS, and Mplus that conduct these analyses.

Keywords: mediation, indirect effect, path analysis, within-participant design, resampling methods

Statistical mediation analysis allows an investigator to answer questions about the process by which some presumed causal variable X Departes to affect an outcome variable Y. Using simple principles of linear modeling (though other analytical approaches are possible; Imai, Keele, & Tingley, 2010; Pearl, 2010, 2012), mediation analysis is used to quantify and test the pathways of influence from X to Y. In a mediation process, one of those pathways consists of a sequence of causal steps in which X affects a mediator variable M, which in turn causally influences Y. This indirect effect of X—the conjunction of the effect of X on M and the effect of X on M and the effect of X on W and the effect of X or M and X or M an

Mediation analysis is commonplace in the social sciences, business, medical research, and many other areas. For example, White,

term interaction with a member of the same religion or a different religion (X) influenced intergroup bias (Y) through five different emotions (e.g., anger and sadness; M). Littleton (2015) found that pregnant women who had a history of sexual victimization (X) had higher rates of depression (M), which predicted increased somatic complaints (e.g., back pair, Y). Schuldt, Guillory, and Gay (2016) examined how the weight of a person recommending a recipe (X) influenced the perceived healthiness of the recipe (Y) through the perceived health of the recommender (M).

Abu-Rayya, Bliuc, and Faulkner (2015) investigated how long-

Discussions of mediation analysis and its application are most typically couched in terms of or conducted using data from research designs that are cross-sectional or "between-participant" in nature. Typically in these designs, participants are measured once on a proposed mediator M and dependent variable P, as in the examples above. This may occur following random assignment of participants into one of two conditions (V) that vary via some manipulation (e.g., a "teatment" vs. a "control" group) that is presumed to cause differences in M and Y. Alternatively, measurement of M and Y may occur contemporaneously with the observation of X (rather than random assignment). For expositional convenience, we refer to designs of this sort (i.e., with or without random assignment to X) throughout this article as "between-participant" designs.

Less attention in the methodology literature has been dedicated to mediation analysis when the data come from repealed measurement of the same people on variables in the mediation process, even though such designs are common. In this article we address mediation analysis in a specific category of repeated measures designs. Researchers sometimes measure a dependent variable Y and a mediator M in two different situations or circumstances CO, Criticisms of approach outlined by Judd, Kenny, and McClelland

- Path-analytic approach to mediation in two-instance repeated-measures designs
- Generalization to multiple mediator models
 - Parallel Mediation
 - Serial Mediation

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Judd Criticisms and Misuses

All criticisms of the causal steps approach apply to this approach:

- There is no explicit quantification of the indirect effect
 - Inference about an indirect effect should be the result of a test on a quantification of the indirect effect
- Requiring that there must be a total effect is too restrictive
 - The direct and indirect effect could be of opposite sign
 - There is greater power to detect the indirect effect than total effect (Judd, Kenny, 2014, Psych Science)

This method has been used by a variety of researchers:

- Approximately 300 citing papers, with around 140 using this method
- Many researchers do not report or estimate the partial regression coefficient for the sum of the mediators
- Because the estimate of the indirect effect is not made explicit, researchers often misinterpret the coefficients
 - b path is often interpreted as indirect effect
- Extensions to more complicated models have been poorly implemented

Advantages of the Path-Analytic Approach

Model based approach, rather than piecewise hypotheses

Provides an estimate of the indirect, total, and direct effects

 Allows us to conduct inferential tests directly on an estimate of the indirect effect

Connects researchers understanding of between-subjects mediation to within-subjects mediation

Reduce misinterpretation of regression coefficients

Using a path analytic framework will help extend the simple mediation model to more complicated questions

- Multiple mediators
- Moderated mediation
- Integration of between and within-subjects designs

Path-Analytic Method for Two-Instance Repeated-Measures Design

Total Effect c:

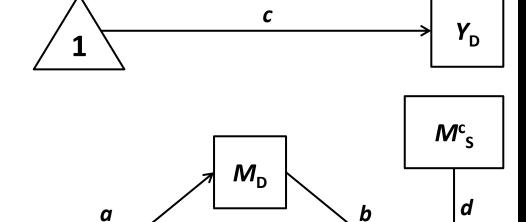
$$Y_{2i} - Y_{1i} = c + \epsilon_{Y^*i}$$

a path:

$$M_{2i} - M_{1i} = a + \epsilon_{Mi}$$

b path and c' path:

$$Y_D = c' + bM_D + dM_S^c + \epsilon_{Yi}$$



^c Indicates mean centered

Indirect effect of *instance* on Y (through M) = $a \times b$

Direct effect of *instance* on Y (not through M) = c'

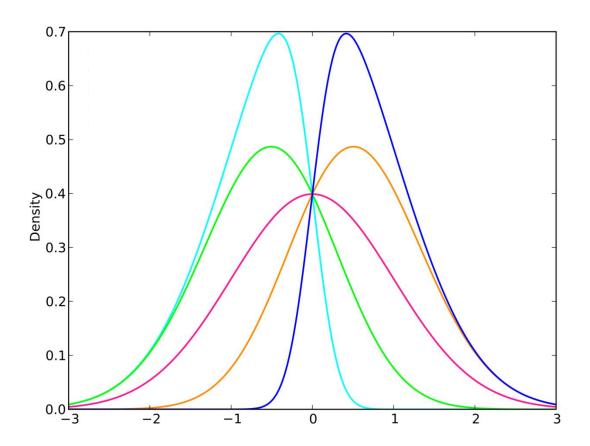
Total effect = direct effect + indirect effect

$$c = c' + a \times b$$

Inference about the Indirect Effect

Why is this so hard?

- The product of two normal distributions is not necessarily normal. The shape of the distribution of the indirect effect depends on the true indirect effect.
- There are many instances where the null hypothesis (ab = 0) could be true



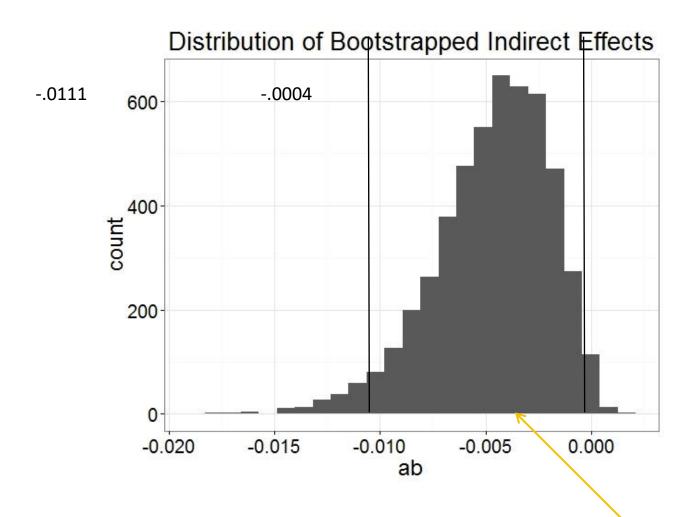
Bootstrap Confidence Intervals (Percentile)

Empirically estimate sampling distribution of the indirect effect. From this distribution compute confidence intervals which can be used for estimation and hypothesis testing.

Method

- 1. Randomly sample *n* cases from your dataset with replacement.
- 2. Estimate the indirect effect using resampled dataset, call this $ab^{(1)}$
- 3. Repeat steps 1 and 2 a total of K times where K is many (10,000 recommended), each time calculated $ab^{(k)}$.
- 4. The sampling distribution of the $ab^{(i)}$'s can be used as an estimate of the sampling distribution of the indirect effect.
- 5. For a 95% confidence interval the lower and upper bounds will be the 2.5^{th} and 97.5^{th} percentiles of the K estimates of the indirect effect.

Bootstrap Confidence Intervals



Zero is not contained in the confidence interval [-.0111, -.0004] so we conclude the indirect effect is different from zero with 95% confidence.

ab = -.0047

Bootstrap Confidence Intervals (Percentile)

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Appeal

- No assumptions about the sampling distribution of the indirect effect
- Provides point estimate of indirect effect
- Can calculate confidence intervals
- Good method balance Type I Error and Power

Evaluating Methods of Inference

RESEARCH QUESTION: Which methods of inference for the indirect effect work best for within-subjects across a variety of situations?

Methods:

- Causal steps (JKM)
- Joint significance
- Sobel Test
- Bootstrapping (Percentile)
- Monte Carlo Confidence Intervals

Population Characteristics:

Sample size: 20, 50, 100, 200

a path: 0, 0.14, 0.39, 0.59

b path: 0, 0.14, 0.39, 0.59

c' path: 0, 0.14, 0.39, 0.59

d path: 0, 0.14, 0.39, 0.59

 $\rho_{\rm m}$: 0, .3, .6, .9

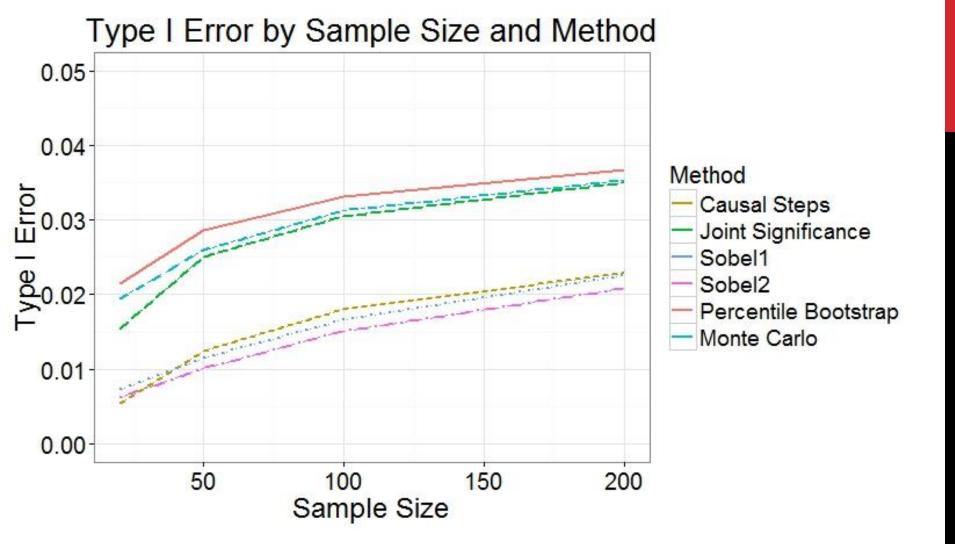
 ρ_{v} : 0, .3, .6, .9

Some combination were impossible, which left 6848 conditions

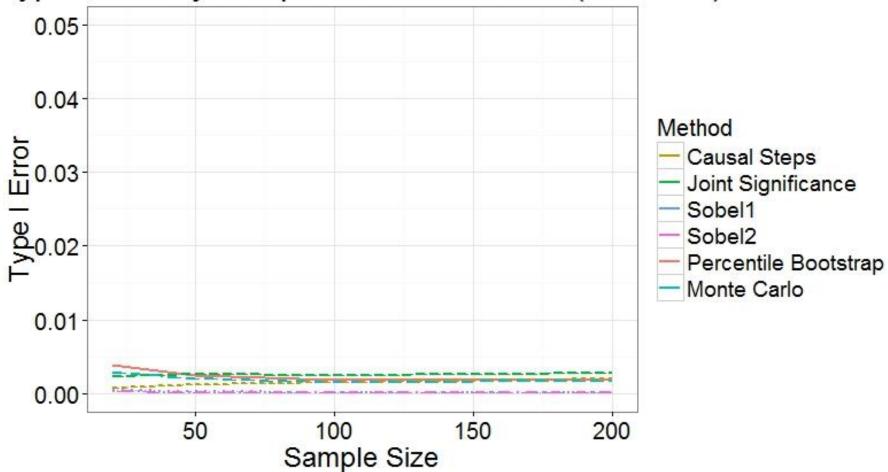
Generated 1000 data sets per condition

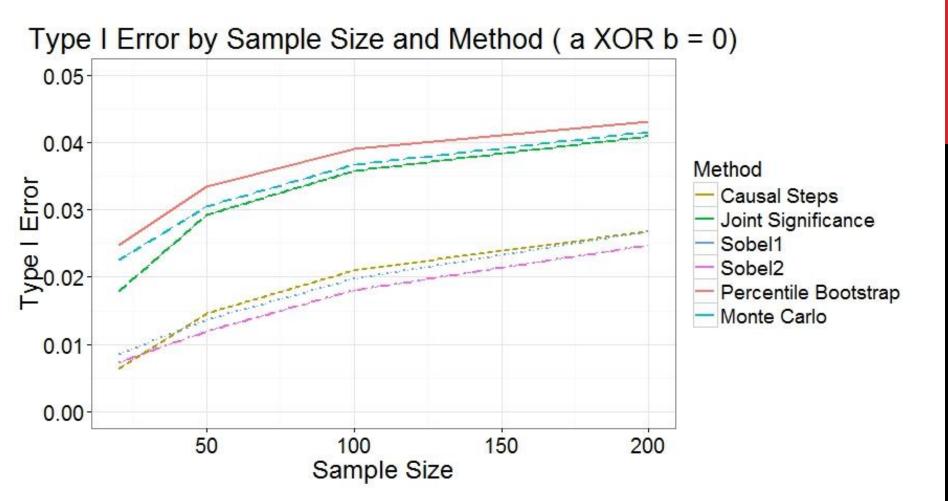
Used each method of inference on each dataset

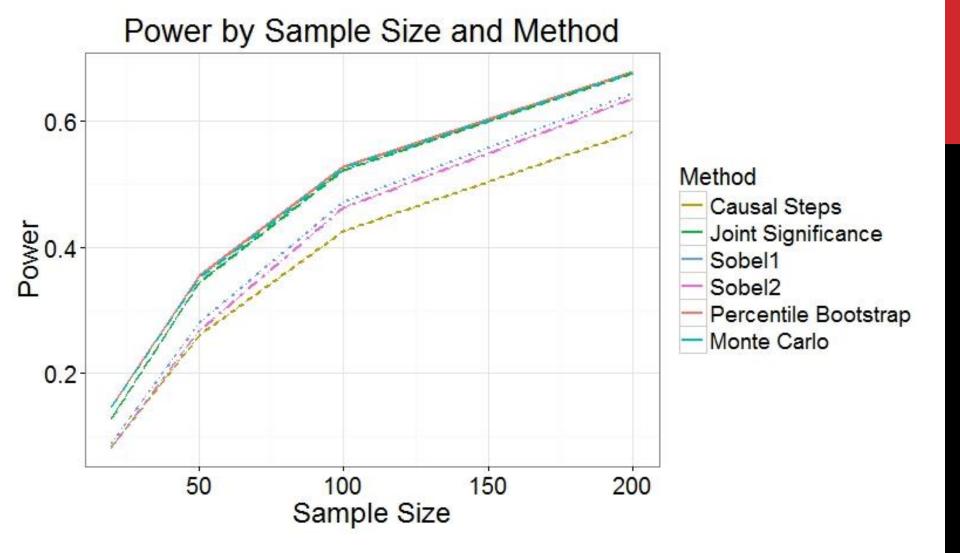
Recorded rejection rate for each method in each condition ($\alpha = .05$)



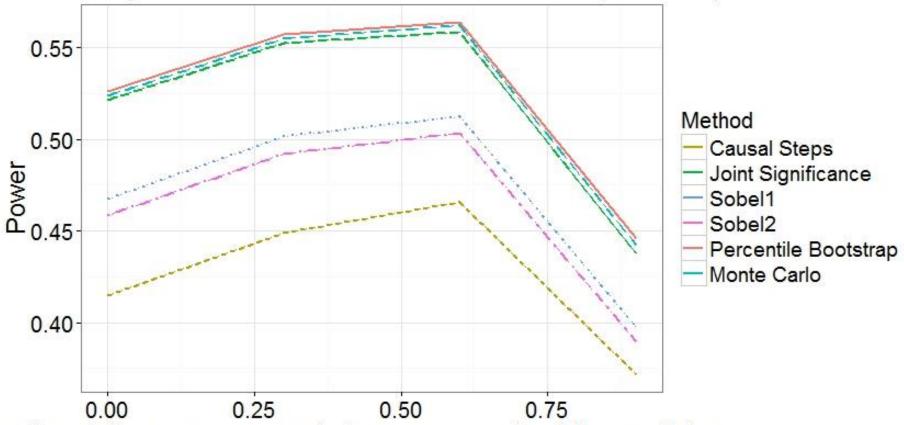
Type I Error by Sample Size and Method (a & b = 0)







Power by Mediator Correlation and Method (N = 100)



Correlation among repeated measurements of the mediator

Implications

- Best methods (based on Type I Error and Power)
 - bootstrap confidence interval
 - Monte Carlo method
 - joint significance test

- provide point estimates and confidence intervals for the indirect effect
- Within-subjects designs buys additional power in mediation analysis when:
 - repeated observations are not independent
 - M's are not too correlated.
- *Proximity:* Mediators that are *too correlated* may result in a loss of power due to increases in standard error of *b* path. (Judd & Kenny, 2014; Hoyle & Kenny, 1999)

So if it's better, how do we do it?

MEMORE is a macro available for SPSS and SAS for conducting (MEdiation and MOderation in REpeated measures designs). Documentation and download at akmontoya.com. Mediation functions described in Two-Condition Within-Participant Statistical Mediation Analysis: A Path-Analytic Framework

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Mediation analysis is commonplace in the social sciences, business, medical research, and many other areas. For example, White, Abu-Rayva, Bliuc, and Faulkner (2015) investigated how longterm interaction with a member of the same religion or a different religion (X) influenced intergroup bias (Y) through five different emotions (e.g., anger and sadness; M). Littleton (2015) found that pregnant women who had a history of sexual victimization (X) had higher rates of depression (M), which predicted increased somatic complaints (e.g., back pain; Y), Schuldt, Guillory, and Gay (2016) examined how the weight of a person recommending a recipe (X) influenced the perceived healthiness of the recipe (Y) through the perceived health of the recommender (M).

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SPSS Syntax:

MEMORE Y = Y1 Y2
$$/M$$
 = M1 M2 $/model$ = 1.

SAS Syntax:

Coming in 2019: Statistical mediation analysis: Within-subjects designs. By Hayes, Montoya, Preacher, and Page-Gould. Guilford Press

This article was published Online First June 30, 2016. Amanda K. Montoya and Andrew F. Hayes, Department of Psychology, The Ohio State University.

The Onio State University.

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ship Program under Grant DGH:1345012. Correspondence concerning this article should be addressed to Amanda K. Montoya or Andrew F. Hayes, Department of Psychology, The Ohio State University, 1827 Neil Ave Mall, Columbus, OH 43210. E-mail: montoya.29@soa.edu or hayes.338@soa.edu

```
Variables:
Y = buy2
            buyl
M = hazard2 hazard1
                                                                                         Model Information
Computed Variables:
Ydiff =
                                 buy1
Mdiff =
                hazard2
                                hazardl
Mavg = (
                hazard2
                                 hazardl )
                                                          Centered
Sample Size:
Outcome: Ydiff = buy2
                                                                                         Total Effect Model
Model
       Effect
                     SE
                                          df
                                                             LLCI
                                                                       ULCI
       -.5636
                   .1932
                           -2.9168
                                      21.0000
                                                  .0082
                                                           -.9655
                                                                      -.1618
Outcome: Mdiff = hazard2
                                                                                         Model for M_2-M_1
Model
                                                             LLCI
       Effect
                                                                        ULCI
·X.
        .8000
                            3.1024
                                     21.0000
                                                  .0054
                                                            .2637
                                                                      1.3363
Outcome: Ydiff - buy2
                                  buyl
Model Summary
                 R-sq
                            MSE
                                                 dfl
                                                           df2
     .7721
                .5961
                          .3667
                                   14.0213
                                              2.0000
                                                        19.0000
                                                                    .0002
                                                                                           Model for Y_2-Y_1
Mode1
          coeff
                                                                         ULCI
                       SE
                                                               LLCI
"X"
                             -.5399
         -.0851
                                                             -.4152
                     .1577
                                       19.0000
                                                    .5955
                                                                         .2449
Mdiff
         -.5981
                    .1131
                             -5.2869
                                       19.0000
                                                    .0000
                                                             -.8349
                                                                        -.3613
Mavg
         -.181B
                    .1683
                             -1.0803
                                       19.0000
                                                    .2935
                                                             -.5341
                                                                         .1705
        ******* TOTAL, DIRECT, AND INDIRECT EFFECTS
Total effect of X on Y
    Effect
                   SE
                                                          LLCI
                                                                     ULCI
                .1932
                                               .0082
    -.5636
                        -2,9168
                                   21,0000
                                                         -.9655
                                                                   -.1618
Direct effect of X on Y
    Effect
                   SE
                                                          LLCI
                                                                     ULCI
                                                                                         Total, Direct, and Indirect
                                       df
    -.0851
                .1577
                         -.5399
                                   19.0000
                                               .5955
                                                         -.4152
                                                                    .2449
                                                                                         Effects
Indirect Effect of X on Y through M
         Effect
                    BootSE
                            BootLLCI
                                      BootULCI
Indl
         -.4785
                     .1363
                              -.7423
                                        -.2063
Indirect Key
                      Midiff
                                       Ydiff
Ind1 X
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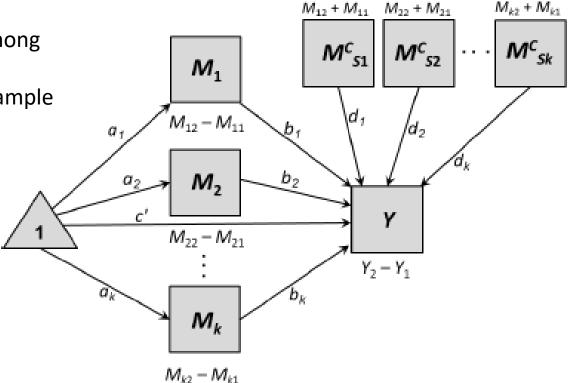
Multiple Mediation Models: Parallel

Simple or Multiple Parallel Mediation

- Up to 5 mediators
- Three methods of inference
 - Sobel test
 - Monte Carlo Cl
 - Bootstrapping (Percentile or BC)

 Pairwise contrasts among indirect effects

Can save bootstrap sample



Multiple Mediation Models: Serial

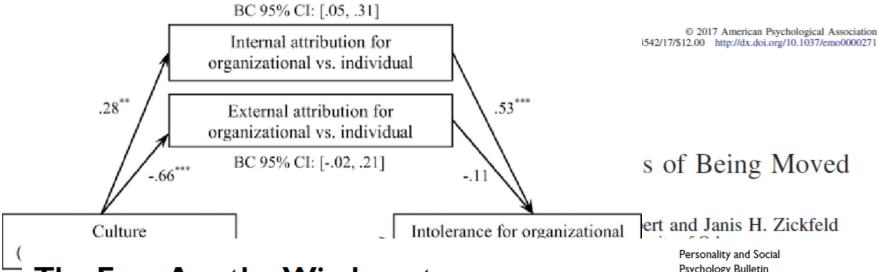
Serial Mediation

- Add serial = 1
- Current version: 2 mediators
- New version: Up to 5 mediators
- Serial argument will lead to easier implementation of $M_{22} + M_{21}$ $M_{12} - M_{11}$ moderated serial mediation M^{c}_{s2} M_1 g_2 d_1 c'V $Y_{2} - Y_{1}$ g_1 M_2 M^{c}_{S1}

 $M_{22} - M_{21}$

 $M_{12} + M_{11}$

Examples of Using MEMORE



The Eyes Are the Windows to the Mind: Direct Eye Gaze Triggers the Ascription of Others' Minds

Personality and Social
Psychology Bulletin
2016, Vol. 42(12) 1666–1677
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and Social Psychology, Inc
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DOI: 10.1177/0146167216669124
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Saara Khalid¹, Jason C. Deska¹, and Kurt Hugenberg¹

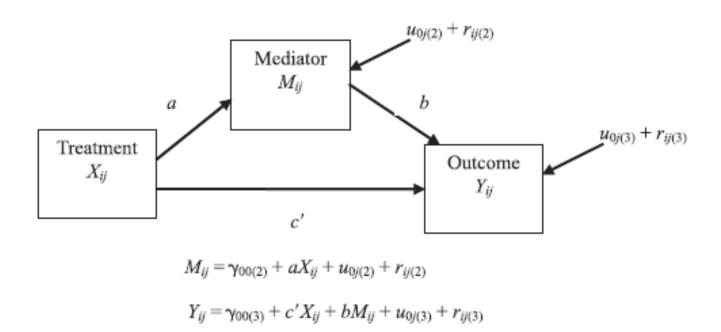
Intermediary Influence of the Brain

Siobhán Harty*, Francesco Sella and Roi Cohen Kadosh

Department of Experimental Psychology, University of Oxford, Oxford, UK

Evaluating Methods

- Need to compare and evaluate when certain methods are better than others, and provide reasonable recommendations to substantive researchers about when to use which method.
- Alternative methods may be useful for answering these questions
 - Structural Equation Modeling
 - Multilevel Modeling



Overview

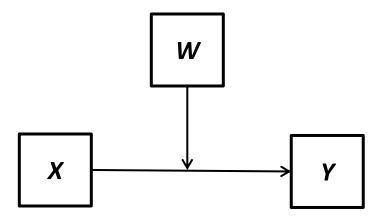
- Personal History and Philosophy
- Introduction to Mediation
- Two-Instance Repeated Measures Designs
 - Mediation
 - Moderation
 - Conditional Process Analysis (AKA Moderated Mediation)
- New Designs & New Questions

Moderation: A New Question

What does this relationship depend on?

When does an effect exist or not exist?

When is it positive and when is it negative?



The relationship between the focal predictor (X) and an outcome (Y) is said to be moderated when the size or direction of this relationship depends on W.

Moderation helps us understand *boundary conditions* of effects: for whom or when is the effect large or small, present or absent, positive or negative.

X and W are frequently described as "interacting" in their prediction of Y.

Examples of Moderation in Psychology

Self-Conscious Emotions and Self-Regulation in the Promotion of Condom Use

ion 522

Michaela Hynie

York University

Development and Psychopathology 27 (2015), 97–109 © Cambridge University Press 2015 doi:10.1017/S0954579414001321

SPECIAL SECTION ARTICLE

Religion priming and an oxytocin receptor gene (*OXTR*) polymorphism interact to affect self-control in a social context

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Moderation in Two-Instance Repeated-Measures Designs

Judd, C. M., McClelland, G. H., and Smith, E. R. (1996). Testing Treatment by Covariate Interactions When Treatment Varies Within Subjects. *Psychological Methods*, 1(4), 366-378.

Psychological Methods 1996, Vol. 1, No. 4, 166-178 Copyrigh: 1996 by the American Psychological Association, Inc., 1002/1893/06451.00

Testing Treatment by Covariate Interactions When Treatment Varies Within Subjects

Charles M. Judd and Gary H. McClelland University of Colorado at Boulder

Eliot R. Smith

In contrast to the situation when an independent or treatment variable varies between subjects, procedures for testing treatment by covariante interactions are not commonly understood when the treatment varies within subjects. The purpose of this article to to identify analytic approaches that test such interactions. Two design scenarios are discussed, one in which the covariate is measured only a single time for each subject and hence varies only between subjects, and the other in which the covariate is measured at each level of the treatment variable and thence varies both within and between subjects. In each case, alternative analyses are identified and their assumptions and relative efficiencies conspared.

An issue that arises with some frequency in data analysis in psychological research concerns the relationship between some measured variable and the dependent variable and whether that relationship depends on or varies across levels of a manipulated or experimental independent variable. For instance, in a clinical intervention study, we might randomly assign patients to one of two conditions, either a treatment intervention or a placebo intervention control condition. Prior to treatment, we measure a characteristic of the patients, probably focusing on the prior course and severity of their illness. Following the treatment, we assess the outcome variable of symptom severity. The primary question of interest, of course, is whether the outcome variable is affected by the manipulated treatment: Did the treatment make a difference on subsequent symptom severity? Additionally, however, we may well want to know whether the relationship between the treatment and posttreatment symptom severity depends on the patient's pretreatment course of

Charles M. Judd and Gary H. McClelland, Department of Psychology, University of Colorado at Boulder, Eliot R. Smith, Department of Psychological Sciences, Purdue Uni-

This work was partially supported by National Institute of

Mental Health Grant R01 MH45049.

Correspondence concerning this article should be addressed to Charles M. Judd. Department of Psychology, University of Colonado, Boulder, Colonado 80309.

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illness. It may be, for instance, that the treatment's effect is greater for patients whose pretreatment symptoms were relatively severe. Equivalently, it may be that posttreatment symptom severity is less well predicted by pretreatment correct of illness in the case of patients in the intervention condition than in the case of patients in the control condition.

The pretreatment measure of Illness coarse is typically called a covariate. The analysis that is of interests is an analysis of cowariance (ANCOVA), including the treatment by covariate interaction (Judd & Mc-Clelland, 1989). The two questions of interest are (a) Is there an overall treatment main effect? and (b) Is there a Treatment × Covariate interaction? If eninteraction is significant, it indicates that the covariate: outcome variable relationship depends on the treatment variable. Equivalently, it suggests that the effect of the treatment on the outcome variable depends on the level of the covariance.

The analysis is readily conducted using multiple regression, making the standard assumption that represents on the represent of the record of

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 Z_i + \varepsilon_i$$

and

 $Y_i = \beta_0 + \beta_1 X_i + \beta_2 Z_i + \beta_3 X_i Z_i + \varepsilon_i$

In the first equation, β_1 represents the magnitude of

366

Does the degree to which W predicts Y depend on which instance a person is in?

Or

Does effect of <u>instance</u> on Y depend on an individual's W?

Data should be a two-instance repeated-measures design with a **person level covariate**.

Moderation in Two-Instance Repeated-Measures Designs

1. Setup two regression equations, one for each instance

$$Y_{1i} = b_{10} + b_{11}W_i + \epsilon_{1i}$$

 $Y_{2i} = b_{20} + b_{21}W_i + \epsilon_{2i}$ Is b_{11} different from b_{21} ?

2. Take the difference between those two regression equations

$$Y_{2i} - Y_{1i} = (b_{10} - b_{20}) + (b_{11} - b_{21})W_i + (\epsilon_{1i} - \epsilon_{2i}) = b_0 + b_1W_i + \epsilon_i$$

3. Regression weight for person level covariate in Step 2 tests moderation.

Estimate equation above and test if b_1 is significantly different from zero

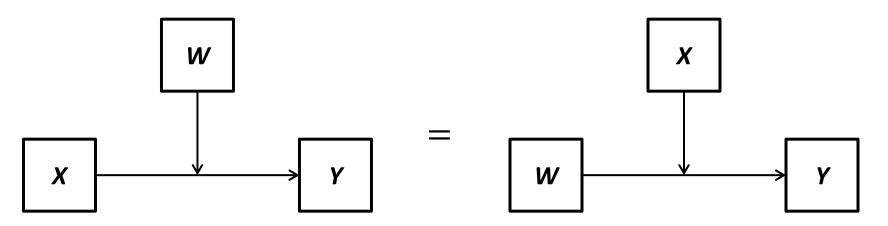
Symmetry in Within-Subjects Moderation

Does the effect of instance depend on *W*?

$$Y_{2i} - Y_{1i} = (b_{10} - b_{20}) + (b_{11} - b_{21})W_i + (\epsilon_{1i} - \epsilon_{2i}) = b_0 + b_1W_i + \epsilon_i$$

 $Y_{2i} - Y_{1i}$ is a quantification of the effect of instance, which means that if W predicts $Y_{2i} - Y_{1i}$ then the effect of instance depends on W.

b_1 is a test of exactly that!



What's Missing?

- Definitions of Conditional Effects
- Probing Conditional Effects
 - Pick-a-point approach
 - Johnson-Neyman
- How to deal with multiple moderators
 - Three-way interacts etc.
 - Multiple two-way interactions

Issues all addressed in Montoya (under review) Moderation analysis in two-instance repeated-measures designs: Probing methods and multiple moderator models

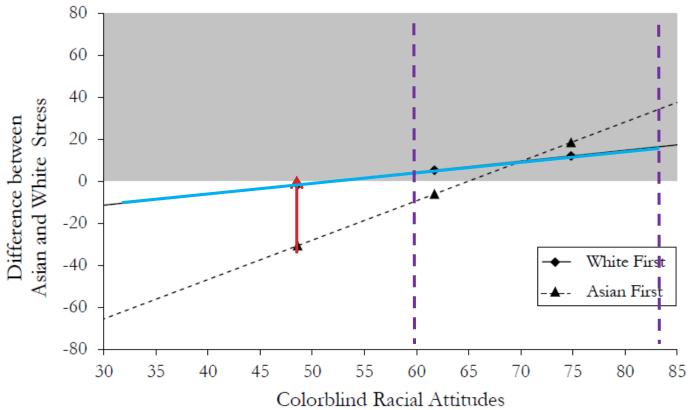
Does the effect of avatar race on stress depend on colorblind racial attitudes?

Sample of 28 white college students successively operate Black, Asian, and White racialized avatars in predominantly White environment in Second Life

X: Avatar Race (White, Black, Asian)

W: Colorblind Racial Attitudes

Y: Stress as measured by heart rate variability



Tawa, J. & **Montoya, A. K.** (under review) White students physiological stress while operating non-White avatars and the moderating role of awareness of racial privilege.

Conditional Effects in Within-Subjects Moderation

$$Y_{2i} - Y_{1i} = (b_{20} - b_{10}) + (b_{21} - b_{11})W_i + (\epsilon_{1i} - \epsilon_{2i}) = b_0 + b_1W_i + \epsilon_i$$

Given a value of W what is the effect of instance on the outcome?

 $Y_{2i} - Y_{1i}$ is a quantification of the effect of instance, which means that the conditional effect of instance $\theta_{X\to Y}(W) = b_0 + b_1 W$

Given a specific instance what is the effect of W on the outcome?

$$Y_{1i} = b_{10} + b_{11}W_i + \epsilon_{1i}$$

$$Y_{2i} = b_{20} + b_{21}W_i + \epsilon_{2i}$$

$$\theta_{W\to Y}(X)=b_{X1}$$

Probing an Effect of Instance on Outcome: The "Pick-a-Point" Approach

Select a value of the moderator (W)

$$\theta_{X \to Y}(W) = b_0 + b_1 W$$

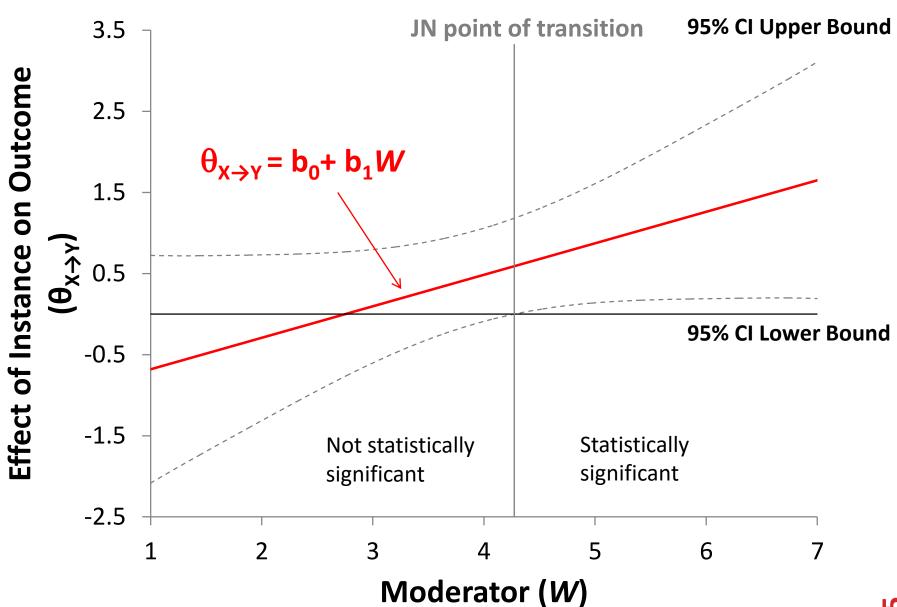
The estimated standard error of $\theta_{X\to Y}(W)$ is

$$s_{\theta_{X\to Y}(W)} = \sqrt{(s_{b_0}^2 + 2Ws_{b_0b_1} + W^2s_{b_1}^2)}$$
 Squared standard error of b_0 Squared standard error of b_1 Covariance of b_0 and b_1

The ratio of the effect to standard error is t-distributed as $t(df_{residual})$ under the null hypothesis that the effect of instance is zero at that moderator value.

$$\frac{\theta_{X \to Y}(W)}{s_{\theta_{X \to Y}(W)}} \sim t_{df}$$

A Plot of the "Region of Significance"



The Johnson-Neyman Technique

Find the value or values of the moderator (W) the conditional effect of instance is exactly significant.

Do not need to select values of W in advance.

What value of W produces a ratio of $\theta_{X\to Y}(W)$ to its standard error exactly equal to the critical t value (t_{crit}) required to reject the null hypothesis that $\theta_{X\to Y}(W)$ is equal to zero at that value of W?

$$t_{crit} = \frac{b_0 + b_1 W}{\sqrt{s_{b_0}^2 + 2W s_{b_0 b_1} + W^2 s_{b_1}^2}}$$

Isolating W yields to the solution in the form of a quadratic equation which always has two roots, though not always two that are interpretable.

MEMORE

MEMORE will estimate and probe moderation models when the focal predictor is a repeated-measures variable.

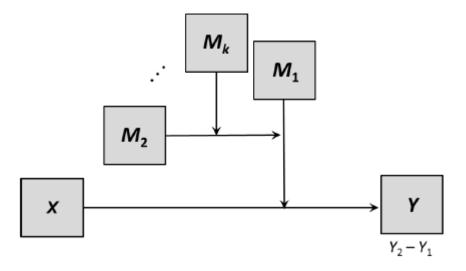
Multiple moderator models are also included!

Can have up to 5 moderators

Model 2: Additive Moderation

 M_1 M_2 \cdots M_k Y Y_2-Y_1

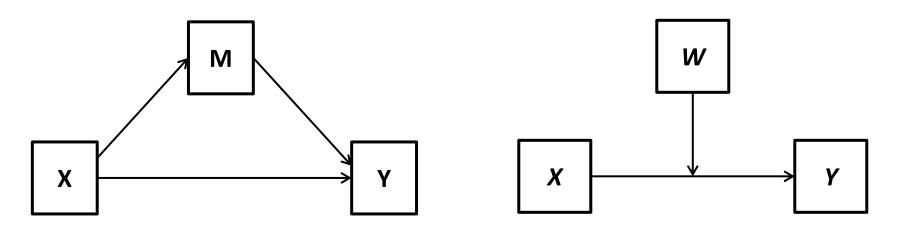
Model 3: Multiplicative Moderation



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Combining Mediation and Moderation: Conditional Process Analysis



Research questions:

- Does the process through which X affects Y through M depend on W?
- Are there certain groups where X affects Y through M and certain groups where this process does not occur?

Conditional process analysis allows a mediated process to be moderated. Now the indirect effect can be defined as a *function of the moderator*.

Examples of CPA in Psychology

Personality and Individual Differences 76 (2015) 104-110



Contents lists available at ScienceDirect

Personality and Individual Differences

journal homepage: www.elsevier.com/locate/paid



The existential model of perfectionism and depressive symptoms: Tests of incremental validity, gender differences, and moderated mediation



Dayna L. Sherry a,*, Simon B. Sherry b, Paul L. Hewitt c, Aislin Mushquash d, Gordon L. Flett e

Cultural Demands and Family Relationships



Evelina Lou¹, Richard N. Lalonde¹, and Benjamin Giguère²

^a Queen Elizabeth II Health Sciences Centre, 1276 South Park Street, Halifax, Nova Scotia B3H2Y9, Canada

b Department of Psychology and Neuroscience, Dalhousie University, 1355 Oxford Street, PO Box 15000, Halifax, Nova Scotia B3H4R2, Canada

Department of Psychology, University of British Columbia, 2136 West Mall, D.T. Kenny Building, Vancouver, British Columbia V6T1Z4, Canada

^d Mental Health Outpatient Programs, St. Joseph's Care Group, 710 Victoria Avenue East, Thunder Bay, Ontario P7C5P7, Canada

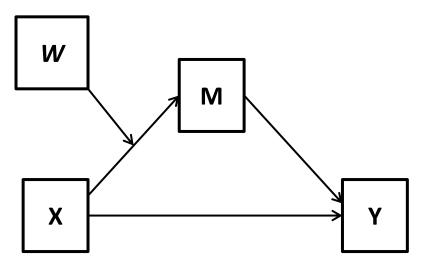
Compariment of Psychology, York University, Behavioural Sciences Building, 4700 Keele Street, Toronto, Ontario M3J1P3, Canada

CPA in Two-Instance Repeated-Measures Designs

Using the path analytic approach outlined in Montoya & Hayes (2017) we can now allow for moderation of a mediated pathway.

No pre-existing work addresses this type of question with this type of data.

First stage moderated mediation allows *W* to moderate the path between the within-subjects factor and the mediator.



Between subjects version

CPA in Two-Instance Repeated-Measures Designs

First stage moderated mediation allows *W* to moderate the path between the within-subjects factor and the mediator.

$$M_{1i} = a_{01} + a_{11}W_i + \epsilon_{M1i}$$

$$M_{2i} = a_{02} + a_{12}W_i + \epsilon_{M2i}$$

Allowing the effect of W to vary by instance, which is the same as allowing the effect of instance to depend on W.

$$M_{2i} - M_{1i} = (a_{20} - a_{10}) + (a_{21} - a_{11})W_i + (\epsilon_{M2i} - \epsilon_{M1i}) = a_0 + a_1W_i + \epsilon_{Mi}$$

How do we quantify the effect of instance on *M*?

$$\theta_{X\to M}(W) = a_0 + a_1 W_i$$

The model for the relationship between *M* and *Y* can remain the same.

$$Y_{Di} = c' + bM_{Di} + dM_{Si}^c + \epsilon_{Yi}$$

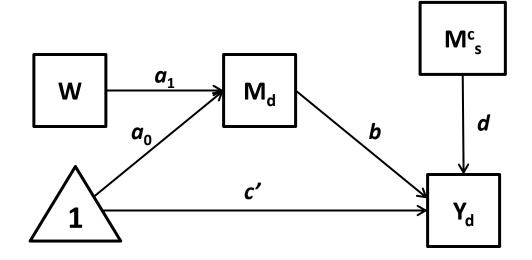
Translating to Path Model

First stage moderated mediation allows *W* to moderate the path between the within-subjects factor and the mediator.

$$M_{2i} - M_{1i} = a_0 + a_1 W_i + \epsilon_{Mi}$$

 $\theta_{X \to M}(W) = a_0 + a_1 W_i$

$$Y_{Di} = c' + bM_{Di} + dM_{Si}^c + \epsilon_{Yi}$$



What is the indirect effect?

$$\theta_{X \to M}(W) \times b = (a_0 + a_1 W)b = a_0 b + a_1 b W$$

Indirect effect is a *function* of the moderator

Making Inference

$$\theta_{X \to M}(W) \times b = (a_0 + a_1 W)b = a_0 b + a_1 b W$$

Conditional Indirect Effects:

Select a value of *W*, plug that into the equation for the indirect effect, and use bootstrapping to estimate the indirect effect at that value

Does the indirect effect *depend* on the moderator?

If $a_1b = 0$ then the indirect effect *does not* depend on W

$$\theta_{X \to M}(W) \times b = a_0 b + 0 * W = a_0 b$$

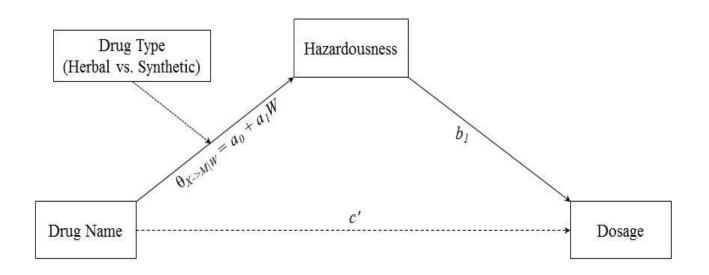
 a_1b can be called the **index of moderated mediation**

A test on the index will indicate if the indirect effect depends on *W.* We can do this formal test using bootstrapping.

Do simple names increase drug dosing through perceptions of safety? Does it depend on type of drug?

The Dark Side of Fluency: Fluent Names Increase Drug Dosing

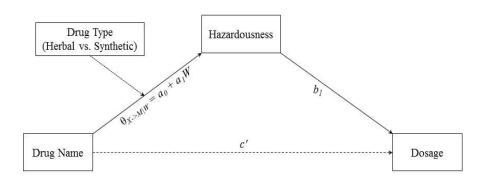
S. Dohle & A. K. Montoya (in press) Journal of Experimental Psychology: Applied



Hypothesis: more complex drug names would decrease dosing through an increase in perceived hazardousness, and a subsequent negative effect of hazardousness on dosing.

Hypothesis: Drug names would have more of an effect on perceived hazardousness for synthetic compared to herbal drugs.

Do simple names increase drug dosing through perceptions of safety? Does it depend on type of drug?



We found that in **both** the herbal and synthetic condition, <u>more complex names</u> were perceived as more hazardous. There was no significant difference based on drug type.

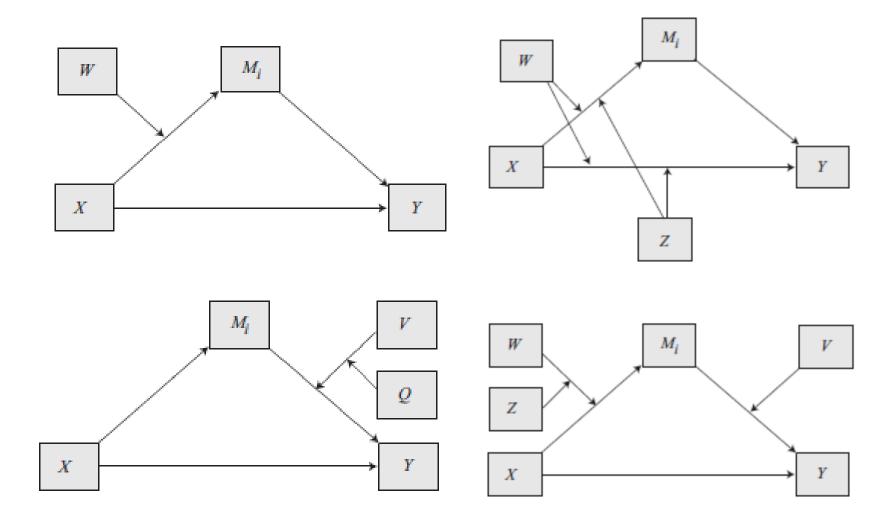
Higher perceived <u>hazardousness</u> was significantly related to reduced drug dosage.

The indirect effect in the in both conditions such that <u>more complex names led to lower doses through hazardousness</u> [Herbal: $a_0b = -6.78$ [-11.18, -3.40], Synthetic: $(a_0+a_1)b = -7.70$ [-11.98, -3.91]].

The *index of moderated mediation* was not significantly different from zero ($a_1b = -.92$ [-3.89, 1.58]) . Therefore we concluded there was <u>no strong evidence</u> that drug type influenced the indirect effect of drug name complexity on dosage through hazardousness.

Coming Soon: MEMORE Models 4 - ...

MEMORE will be expanded to include a variety of moderated mediation models.



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Future Directions

New Designs:

- More than two conditions / instances
- New types of data
 - Counts
 - Survival Times

Generalized linear mixed models

New Questions:

- Meta-analysis
 - Effect size in mediation
 - Meta-analysis of mediation
 - Meta-mediation analysis

Summary

- Mediation, moderation, and conditional process analysis are very important to answering research questions in psychology
- These methods for repeated-measures designs are underdeveloped
- I've made headway in developing these methods for two-instance repeated measures designs
 - Path analytic approach
 - Easy to use tool
- There is much work to do in expanding these analyses in repeatedmeasures designs
- Linear regression based methods need to be compared to structural equation modeling approaches and multilevel approaches
- Additional related questions:
 - Meta-analysis

Resources

I am available for questions now and forever via email at montoya.29@osu.edu

Things to look forward to:

Hayes, A. F., **Montoya, A. K.**, Preacher, K. J., & Page-Gould, E. (under contract). *Statistical mediation analysis: Within-participant designs*. New York: The Guilford Press.

MEMORE can be downloaded from akmontoya.com

Slides available at github.com/akmontoya/York2017

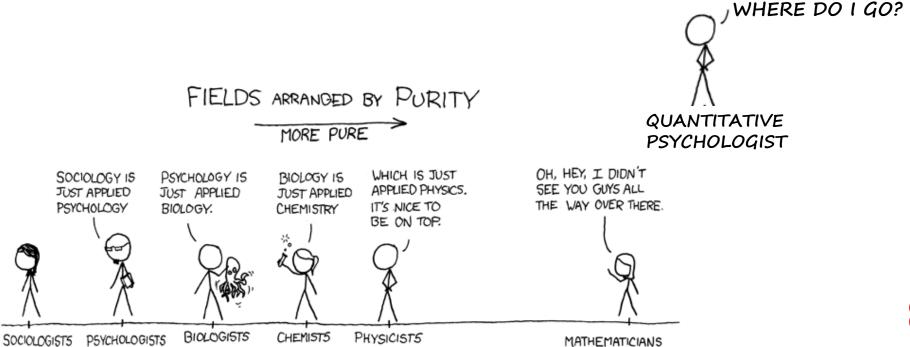
Thank you!



The Mechanisms and Contingencies (MAC) Lab

The Ohio State University

Thank you, Dr. Hayes and the Mechanisms and Contingencies Lab! Thanks to the National Science Foundation Graduate Research Fellowship, the Counsel of Graduate Students, and The Ohio State University Distinguished Dean's University Fellowship for supporting my research. And thanks to all of you for attending!



Two-Condition Repeated Measures Mediation

We can quantify paths of influence in the two condition repeated measures case.

3. Does difference in *M* predict a difference in *Y*?

In this model there are two measures of the effect of M on $Y: g_{11}$ and g_{21}

$$E(Y_{1i}) = g_{10} + g_{11}M_{1i}$$

$$E(Y_{2i}) = g_{20} + g_{21}M_{2i}$$

Look at how the difference in *M* affects the difference in *Y*, subtract these two equations.

$$E(Y_{1i} - Y_{2i}) = g_{10} - g_{20} + g_{11}M_{1i} - g_{21}M_{2i}$$

Then apply a rotation to get:

$$\boldsymbol{E}(Y_{1i} - Y_{2i}) = g_{10} - g_{20} + \frac{g_{21} + g_{11}}{2}(M_{1i} - M_{2i}) + \frac{g_{11} - g_{21}}{2}(M_{1i} + M_{2i})$$

This means if we regress the difference in Y's onto the difference and sum of the M's, the regression coefficient for the difference in M's will estimate the average effect of M on Y.

$$\widehat{Y_{1i} - Y_{2i}} = \hat{h} + \hat{b}(M_{1i} - M_{2i}) + \hat{d}(M_{1i} + M_{2i})$$

Two-Condition Repeated Measures Mediation

We can quantify paths of influence in the two condition repeated measures case.

4. Does the difference in M account for all the difference in Y?

As noted before, differences between *Y*'s reflect an effect of *X*, so expected differences in *Y*'s when the differences in *M*'s is zero reflect the direct effect.

$$E(Y_{1i} - Y_{2i}) = h + b(M_{1i} - M_{2i}) + d(M_{1i} + M_{2i})$$

The quantity h reflects the expected difference in Y's when the difference in M's is zero and the sum of M's is zero. Grand mean centering the sum term renders the intercept of this equation interpretable as the average difference between conditions on Y after accounting for differences in M.

$$E(Y_{1i} - Y_{2i}) = h + d \frac{1}{n} \sum_{i=1}^{n} (M_{1i} + M_{2i}) + b(M_{1i} - M_{2i}) + d(M_{1i} + M_{2i}) - \frac{1}{n} \sum_{i=1}^{n} (M_{1i} + M_{2i})$$

Thus $h + d \frac{1}{n} \sum_{i=1}^{n} (M_{1i} + M_{2i})$ is a quantification of the direct effect. An estimate of this value could be the intercept from the following equation:

$$\widehat{Y_{1i} - Y_{2i}} = \widehat{c'} + \widehat{b}(M_{1i} - M_{2i}) + \widehat{d}(M_{1i} + M_{2i} - \frac{1}{n} \sum_{i=1}^{n} (M_{1i} + M_{2i}))$$