

## RANKING AGGREGATION

In the simplest formulation, a full ranking on a set of items  $\llbracket n \rrbracket$  is seen as the permutation  $\sigma \in \mathfrak{S}_n$  that maps an item  $i$  to its rank  $\sigma(i)$ . Given a collection of  $N \geq 1$  permutations  $\sigma_1, \dots, \sigma_N$ , the goal of ranking aggregation is to find  $\sigma^* \in \mathfrak{S}_n$  that best summarizes it. A popular approach, called **Kemeny’s rule**, consists in solving the **NP-hard** following optimization problem:

$$\min_{\sigma \in \mathfrak{S}_n} \sum_{t=1}^N d(\sigma_t, \sigma),$$

where  $d(., .)$  is the Kendall’s tau distance, i.e.:

$$d(\sigma, \sigma') = \sum_{1 \leq i < j \leq n} \mathbb{I}\{(\sigma(j) - \sigma(i))(\sigma'(j) - \sigma'(i)) < 0\}$$

**Previous work:** Numerous results, e.g. bounds on the cost of approximation procedures, consistency relationships between Kemeny aggregation and other voting rules...

**Our contribution:** In a general statistical framework for Kemeny aggregation, we describe optimal elements and provide statistical guarantees for the generalization properties of an empirical median ranking in the form of rate bounds.

## STATISTICAL FRAMEWORK

Suppose that the dataset is composed of  $N$  i.i.d copies  $\Sigma_1, \dots, \Sigma_N$  of a generic random variable  $\Sigma \sim P$ . A true median for  $P$  w.r.t  $d$  is any solution of the minimization problem:

$$\min_{\sigma \in \mathfrak{S}_n} L(\sigma),$$

where  $L(\sigma) = \mathbb{E}_{\Sigma \sim P}[d(\Sigma, \sigma)]$  is **the risk** of  $\sigma$ .

What is the performance of **Kemeny empirical medians**, i.e. solutions  $\hat{\sigma}_N$  of

$$\min_{\sigma \in \mathfrak{S}_n} \hat{L}_N(\sigma), \quad (1)$$

where  $\hat{L}_N(\sigma) = \frac{1}{N} \sum_{t=1}^N d(\Sigma_t, \sigma)$  ?

## OPTIMALITY OF A CONSENSUS

The risk of a permutation candidate  $\sigma \in \mathfrak{S}_n$  can be written as

$$L(\sigma) = \sum_{i < j} p_{i,j} \mathbb{I}\{\sigma(i) > \sigma(j)\} + \sum_{i < j} (1 - p_{i,j}) \mathbb{I}\{\sigma(i) < \sigma(j)\}.$$

So if  $\exists$  a permutation  $\sigma$  with the property that  $\forall i < j$  s.t.  $p_{i,j} \neq 1/2$ ,

$$(\sigma(j) - \sigma(i)) \cdot (p_{i,j} - 1/2) > 0, \quad (2)$$

it would be necessarily a median for  $P$ .

**Definition 1.** The probability distribution  $P$  on  $\mathfrak{S}_n$  is said to be **stochastically transitive** if it fulfills the condition:  $\forall (i, j, k) \in \llbracket n \rrbracket^3$ ,

$$p_{i,j} \geq 1/2 \text{ and } p_{j,k} \geq 1/2 \Rightarrow p_{i,k} \geq 1/2.$$

In addition, if  $p_{i,j} \neq 1/2$  for all  $i < j$ ,  $P$  is said to be **strictly stochastically transitive**.

**Theorem 1.** If the distribution  $P$  is stochastically transitive, there exists  $\sigma^* \in \mathfrak{S}_n$  such that (2) holds true. In this case, we have

$$L^* = \sum_{i < j} \left\{ \frac{1}{2} - \left| p_{i,j} - \frac{1}{2} \right| \right\},$$

the excess of risk of any  $\sigma \in \mathfrak{S}_n$  is given by

$$L(\sigma) - L^* = 2 \sum_{i < j} |p_{i,j} - 1/2| \cdot \mathbb{I}\{(\sigma(j) - \sigma(i))(p_{i,j} - 1/2) < 0\}$$

and the set of medians of  $P$  is the class of equivalence of  $\sigma^*$  w.r.t. the equivalence relationship:

$$\sigma \mathcal{R}_P \sigma' \Leftrightarrow (\sigma(j) - \sigma(i))(\sigma'(j) - \sigma'(i)) > 0 \text{ for all } i < j \text{ such that } p_{i,j} \neq 1/2.$$

In addition, the mapping  $s^*$  (equivalent of the **Copeland score**) defined by

$$s^*(i) = 1 + \sum_{k \neq i} \mathbb{I}\{p_{i,k} < \frac{1}{2}\}$$

belongs to  $\mathfrak{S}_n$  and is the unique median of  $P$  iff  $P$  is strictly stochastically positive.

## CONNECTION TO OTHER VOTING RULES

Extension of voting rules to a distribution  $P$ :

- Copeland score of item  $i$ :  
 $s(i) = \sum_{k \neq i} \mathbb{I}\{p_{i,k} \leq 1/2\} - \mathbb{I}\{p_{i,k} > 1/2\}$
- Borda score of item  $i$ :  $s(i) = \mathbb{E}_P[\Sigma(i)]$

**Proposition 1.** (BORDA CONSENSUS) The probability distribution  $P$  on  $\mathfrak{S}_n$  is said to be **strongly stochastically transitive** if  $\forall (i, j, k) \in \llbracket n \rrbracket^3$ :

$$p_{i,j} \geq 1/2 \text{ and } p_{j,k} \geq 1/2 \Rightarrow p_{i,k} \geq \max(p_{i,j}, p_{j,k}).$$

Then under this condition, and for  $i < j$ ,  $p_{i,j} \neq \frac{1}{2}$ , there exists a unique  $\sigma^* \in \mathfrak{S}_n$  such that (2) holds true, corresponding to the Kemeny and Borda consensus both at the same time.

## UNIVERSAL RATES

The performance of a Kemeny empirical median  $\hat{\sigma}_N$  is mesured by its excess risk:

$$L(\hat{\sigma}_N) - L^* \leq 2 \max_{\sigma \in \mathfrak{S}_n} |\hat{L}_N(\sigma) - L(\sigma)|.$$

We establish the following result.

**Proposition 2.** The excess risk of  $\hat{\sigma}_N$  is upper bounded:

(i) In expectation by

$$\mathbb{E}[L(\hat{\sigma}_N) - L^*] \leq \frac{n(n-1)}{2\sqrt{N}}$$

(ii) With probability higher than  $1 - \delta$  for any  $\delta \in (0, 1)$  by

$$L(\hat{\sigma}_N) - L^* \leq \frac{n(n-1)}{2} \sqrt{\frac{2 \log(n(n-1)/\delta)}{N}}.$$

We then establish the tightness of the upper bound by providing a lower bound for the **minimax risk** :

$$\mathcal{R}_N \stackrel{\text{def}}{=} \inf_{\sigma_N} \sup_P \mathbb{E}_P[L_P(\sigma_N) - L_P^*], \quad (3)$$

where the sup. is taken over all distr. on  $\mathfrak{S}_n$ .

**Proposition 3.** The minimax risk for Kemeny aggregation is lower bounded as follows:

$$\mathcal{R}_N \geq \frac{1}{16e\sqrt{N}}.$$

## FAST RATES

For  $h > 0$ , we define the **low noise** condition:

$$\mathbf{NA}(h): \min_{i < j} |p_{i,j} - 1/2| \geq h.$$

Let  $\hat{p}_{i,j} = (1/N) \sum_{m=1}^N \mathbb{I}\{\Sigma_m(i) < \Sigma_m(j)\}$ . We establish exponential rates of convergence.

**Proposition 4.** Assume that  $P$  is stochastically transitive and fulfills condition **NA**( $h$ ) for some  $h > 0$ . The following assertions hold true.

(i) For any empirical Kemeny median  $\hat{\sigma}_N$ , we have:

$$\mathbb{E}[L(\hat{\sigma}_N) - L^*] \leq \frac{n^2(n-1)^2}{8} e^{-\frac{N}{2} \log\left(\frac{1}{1-4h^2}\right)}.$$

(ii) With probability at least  $1 - (n(n-1)/4)e^{-\frac{N}{2} \log\left(\frac{1}{1-4h^2}\right)}$ , the mapping

$$\hat{s}_N(i) = 1 + \sum_{k \neq i} \mathbb{I}\{\hat{p}_{i,k} < \frac{1}{2}\}$$

for  $1 \leq i \leq n$  belongs to  $\mathfrak{S}_n$  and is the unique solution of (1).

**Proposition 5.** Let  $h > 0$  and define

$$\tilde{\mathcal{R}}_N(h) = \inf_{\sigma_N} \sup_P \mathbb{E}_P[L_P(\sigma_N) - L_P^*],$$

where the sup. is taken over all stochastically transitive distr. on  $\mathfrak{S}_n$  satisfying **NA**( $h$ ). We have:

$$\tilde{\mathcal{R}}_N(h) \geq \frac{h}{4} e^{-N 2h \log\left(\frac{1+2h}{1-2h}\right)}. \quad (4)$$

Let  $\alpha_h = \frac{1}{2} \log(1/(1-4h^2))$  and  $\beta_h = 2h \log((1+2h)/(1-2h))$ . We have  $\alpha_h \sim \frac{1}{2} \beta_h$  when  $h \rightarrow \frac{1}{2}$ .

## COMPUTATIONAL BENEFIT

Under the low-noise condition, the Copeland method (complexity  $O(N \binom{n}{2})$ ) outputs the exact NP-hard Kemeny consensus (Proposition 4 (ii)).

## REFERENCES

- [1] J.Y. Audibert and A.B. Tsybakov. *Fast Learning Rates For Plug-in Classifiers*. Annals of Statistics, 2007.
- [2] V. Koltchinskii and O. Beznosova. *Exponential Convergence Rates in Classification*. COLT, 2005.