Probability Review for Machine Learning

Motivation

Uncertainty arises through:

- Noisy measurements
- Variability between samples
- Finite size of data sets

Probability provides a consistent framework for the quantification and manipulation of uncertainty.

Sample Space

Sample space Ω is the set of all possible outcomes of an experiment.

Observations $\omega \in \Omega$ are points in the space also called sample outcomes, realizations, or elements.

Events $E \subset \Omega$ are subsets of the sample space.

In this experiment we flip a coin twice:

Sample space All outcomes $\Omega = \{HH, HT, TH, TT\}$

Observation $\omega = HT$ valid sample since $\omega \in \Omega$

Event Both flips same $E = \{HH, TT\}$ valid event since $E \subset \Omega$

Probability

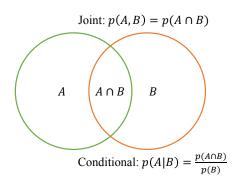
The probability of an event E, P(E), satisfies three axioms:

- 1: $P(E) \ge 0$ for every E
- **2**: $P(\Omega) = 1$
- 3: If E_1, E_2, \ldots are disjoint then

$$P(\bigcup_{i=1}^{\infty} E_i) = \sum_{i=1}^{\infty} P(E_i)$$

Joint and Conditional Probabilities

Joint Probability of A and B is denoted P(A, B). Conditional Probability of A given B is denoted P(A|B).



$$p(A,B) = p(A|B)p(B) = p(B|A)p(A)$$

Conditional Example

Probability of passing the final is 60% and probability of passing both the final and the midterm is 45%.

What is the probability of passing the final given the student passed the midterm?

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$$P(M|F) = P(M,F)/P(F)$$

= 0.45/0.60
= 0.75

Independence

Events A and B are independent if P(A, B) = P(A)P(B).

• Independent: A: first toss is HEAD; B: second toss is HEAD;

$$P(A, B) = 0.5 * 0.5 = P(A)P(B)$$

• Not Independent: A: first toss is HEAD; B: first toss is HEAD;

$$P(A,B) = 0.5 \neq P(A)P(B)$$

Independence

Events A and B are conditionally independent given C if

$$P(A, B|C) = P(B|C)P(A|C)$$

Consider two coins ²: A regular coin and a coin which always outputs HEAD or always outputs TAIL.

A=The first toss is HEAD; B=The second toss is HEAD; C=The regular coin is used. D=The other coin is used.

Then A and B are conditionally independent given C, but A and B are NOT conditionally independent given D.

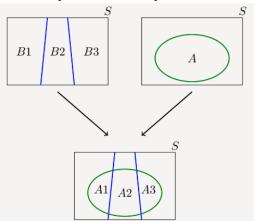
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 $^{^2 \}verb|www.probabilitycourse.com/chapter1/1_4_4_conditional_independence.| \\$ php

Marginalization and Law of Total Probability

Law of Total Probability ³

$$P(X) = \sum_{Y} P(X,Y) = \sum_{Y} P(X|Y)P(Y)$$



³www.probabilitycourse.com/chapter1/1_4_2_total_probability.php

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Bayes' Rule

Bayes' Rule:

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

$$P(\theta|x) = \frac{P(x|\theta)P(\theta)}{P(x)}$$

$$Posterior = \frac{\text{Likelihood} * Prior}{\text{Evidence}}$$

$$Posterior \propto Likelihood \times Prior$$

Bayes' Example

Suppose you have tested positive for a disease. What is the probability you actually have the disease?

This depends on the prior probability of the disease:

- P(T = 1|D = 1) = 0.95 (likelihood)
- P(T=1|D=0) = 0.10 (likelihood)
- P(D=1) = 0.1 (prior)

So
$$P(D = 1|T = 1) = ?$$

Bayes' Example

Suppose you have tested positive for a disease. What is the probability you actually have the disease?

$$P(T=1|D=1)=0.95$$
 (true positive)
 $P(T=1|D=0)=0.10$ (false positive)
 $P(D=1)=0.1$ (prior)

So
$$P(D = 1|T = 1) = ?$$

Use Bayes' Rule:

$$P(D=1|T=1) = \frac{P(T=1|D=1)P(D=1)}{P(T=1)} = \frac{0.95*0.1}{P(T=1)} = 0.51$$

$$P(T=1) = P(T=1|D=1)P(D=1) + P(T=1|D=0)P(D=0)$$

$$= 0.95*0.1 + 0.1*0.90 = 0.185$$

Random Variable

How do we connect sample spaces and events to data? A random variable is

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Random Variable

How do we connect sample spaces and events to data? A random variable is a mapping which assigns a real number $X(\omega)$ to each observed outcome $\omega \in \Omega$

For example, let's flip a coin 10 times. $X(\omega)$ counts the number of Heads we observe in our sequence. If $\omega = HHTHTHHTHT$ then $X(\omega) = 6$.

Discrete and Continuous Random Variables

Discrete Random Variables

- Takes countably many values, e.g., number of heads
- Distribution defined by probability mass function (PMF)
- Marginalization: $p(x) = \sum_{y} p(x, y)$

Continuous Random Variables

- Takes uncountably many values, e.g., time to complete task
- Distribution defined by probability density function (PDF)
- Marginalization: $p(x) = \int_{y} p(x, y) dy$

I.I.D.

Random variables are said to be independent and identically distributed (i.i.d.) if they are sampled from the same probability distribution and are mutually independent.

This is a common assumption for observations. For example, coin flips are assumed to be iid.

Probability Distribution Statistics

Mean: First Moment, μ

$$E[x] = \sum_{i=1}^{\infty} x_i p(x_i)$$
 (univariate discrete r.v.)
$$E[x] = \int_{-\infty}^{\infty} x p(x) dx$$
 (univariate continuous r.v.)

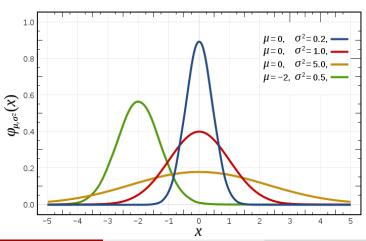
Variance: Second (central) Moment, σ^2

$$Var[x] = \int_{-\infty}^{\infty} (x - \mu)^2 p(x) dx$$
$$= E[(x - \mu)^2]$$
$$= E[x^2] - E[x]^2$$

Univariate Gaussian Distribution

Also known as the Normal Distribution, $\mathcal{N}(\mu, \sigma^2)$

$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\{-\frac{1}{2\sigma^2}(x-\mu)^2\}$$



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Multivariate Gaussian Distribution

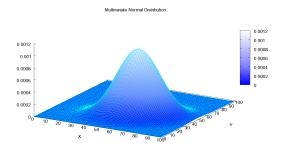
Multidimensional generalization of the Gaussian.

 \mathbf{x} is a D-dimensional vector

 μ is a D-dimensional mean vector

 Σ is a $D\times D$ covariance matrix with determinant $|\Sigma|$

$$\mathcal{N}(\mathbf{x}|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma|^{1/2}} \exp\{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)\}$$



Covariance Matrix

Recall that \mathbf{x} and μ are D-dimensional vectors Covariance matrix Σ is a matrix whose (i, j) entry is the covariance

$$\Sigma_{ij} = Cov(\mathbf{x}_i, \mathbf{x}_j)$$

$$= E[(\mathbf{x}_i - \mu_i)(\mathbf{x}_j - \mu_j)]$$

$$= E[(\mathbf{x}_i \mathbf{x}_j)] - \mu_i \mu_j$$

so notice that the diagonal entries are the variance of each elements. The covariant matrix has the property that it is symmetric and positive-semidefinite (this is useful for whitening).

Inferring Parameters

We have data X and we assume it is sampled from some distribution. How do we figure out the parameters that 'best' fit that distribution? Maximum Likelihood Estimation (MLE)

$$\hat{\theta}_{MLE} = \underset{\theta}{\operatorname{argmax}} P(X|\theta)$$

Maximum A posteriori Probability (MAP)

$$\hat{\theta}_{MAP} = \underset{\theta}{\operatorname{argmax}} P(\theta|X)$$

We are trying to infer the parameters for a Univariate Gaussian Distribution, mean (μ) and variance (σ^2) .

$$\mathcal{N}(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\{-\frac{1}{2\sigma^2}(x-\mu)^2\}$$

The likelihood that our observations x_1, \ldots, x_N were generated by a univariate Gaussian with parameters μ and σ^2 is

Likelihood =
$$p(x_1 ... x_N | \mu, \sigma^2) = \prod_{i=1}^{N} \frac{1}{\sqrt{2\pi\sigma^2}} \exp\{-\frac{1}{2\sigma^2} (x_i - \mu)^2\}$$

For MLE we want to maximize this likelihood, which is difficult because it is represented by a product of terms

Likelihood =
$$p(x_1 ... x_N | \mu, \sigma^2) = \prod_{i=1}^{N} \frac{1}{\sqrt{2\pi\sigma^2}} \exp\{-\frac{1}{2\sigma^2} (x_i - \mu)^2\}$$

So we take the log of the likelihood so the product becomes a sum

Log Likelihood =
$$\log p(x_1 \dots x_N | \mu, \sigma^2)$$

= $\sum_{i=1}^N \log \frac{1}{\sqrt{2\pi\sigma^2}} \exp\{-\frac{1}{2\sigma^2}(x_i - \mu)^2\}$

Since log is monotonically increasing $\max L(\theta) = \max \log L(\theta)$

The log Likelihood simplifies to

$$\mathcal{L}(\mu, \sigma) = \sum_{i=1}^{N} \log \frac{1}{\sqrt{2\pi\sigma^2}} \exp\{-\frac{1}{2\sigma^2} (x_i - \mu)^2\}$$
$$= -\frac{1}{2} N \log(2\pi\sigma^2) - \sum_{i=1}^{N} \frac{(x_i - \mu)^2}{2\sigma^2}$$

Which we want to maximize. How?

To maximize we take the derivatives, set equal to 0, and solve:

$$\mathcal{L}(\mu, \sigma) = -\frac{1}{2}N\log(2\pi\sigma^2) - \sum_{i=1}^{N} \frac{(x_i - \mu)^2}{2\sigma^2}$$

Derivative w.r.t. μ , set equal to 0, and solve for $\hat{\mu}$

$$\frac{\partial \mathcal{L}(\mu, \sigma)}{\partial \mu} = 0 \implies \hat{\mu} = \frac{1}{N} \sum_{i=1}^{N} x_i$$

Therefore the $\hat{\mu}$ that maximizes the likelihood is the average of the data points.

Derivative w.r.t. σ^2 , set equal to 0, and solve for $\hat{\sigma}^2$

$$\frac{\partial \mathcal{L}(\mu, \sigma)}{\partial \sigma^2} = 0 \implies \hat{\sigma}^2 = \frac{1}{N} \sum_{i=1}^{N} (x_i - \hat{\mu})^2$$