

AKSHATA BHAT  
abhat6@wisc.edu  
CS532 Homework 1

① Let's calculate the  $j^{\text{th}}$  component of  $\nabla_{\beta}(\beta^T A \beta)$

$$\frac{\partial}{\partial \beta_j} (\beta^T A \beta) = \frac{\partial}{\partial \beta_j} \sum_{k,l} (\beta_k a_{kl} \beta_l)$$

$$A = (a_{kl})$$

$$\beta^T A \beta = (\beta_1 \ \beta_2 \ \dots \ \beta_p) \begin{pmatrix} a_{11} & a_{12} & \dots & a_{1p} \\ a_{21} & \dots & \dots & \dots \\ \vdots & \dots & \dots & \dots \\ a_{pp} \end{pmatrix} \begin{pmatrix} \beta_1 \\ \beta_2 \\ \vdots \\ \beta_p \end{pmatrix}$$

$$\Rightarrow \frac{\partial}{\partial \beta_j} (\beta^T A \beta) = \frac{\partial}{\partial \beta_j} \left( \sum_{k \neq j} a_{kj} \beta_k \beta_j + \sum_{k \neq j} a_{jk} \beta_k \beta_j + a_{jj} \beta_j^2 \right)$$

$$= \sum_{k \neq j} a_{kj} \beta_k + \sum_{k \neq j} a_{jk} \beta_k + 2a_{jj} \beta_j$$

$$= \underbrace{\sum_{k \neq j} a_{kj} \beta_k + a_{jj} \beta_j}_{\downarrow} + \underbrace{\sum_{k \neq j} a_{jk} \beta_k + a_{jj} \beta_j}_{\downarrow}$$

$$(a_{1j} \ a_{2j} \ \dots \ a_{pj}) \begin{pmatrix} \beta_1 \\ \beta_2 \\ \vdots \\ \beta_p \end{pmatrix}$$

$$= A^T \beta$$

$$(a_{j1} \ a_{j2} \ \dots \ a_{jp}) \begin{pmatrix} \beta_1 \\ \beta_2 \\ \vdots \\ \beta_p \end{pmatrix}$$

$$= A \beta$$

Hence, when  $A$  is not symmetric,

$$\boxed{\nabla_{\beta}(\beta^T A \beta) = A \beta + A^T \beta}$$

$\therefore$  When  $A$  is symmetric  $A^T = A$ ,

$$\text{so } \nabla_{\beta}(\beta^T A \beta) = A \beta + A \beta = 2A \beta$$

$$\textcircled{2} \quad X = \begin{bmatrix} 0 & 0 & 0 \\ 1 & 1 & 2 \\ 2 & 3 & 5 \end{bmatrix} \quad y = \begin{bmatrix} 0 \\ 4 \\ 10 \end{bmatrix}$$

$$a) \quad X^T X = \begin{bmatrix} 0 & 1 & 2 \\ 0 & 1 & 3 \\ 0 & 2 & 5 \end{bmatrix} \begin{bmatrix} 0 & 0 & 0 \\ 1 & 1 & 2 \\ 2 & 3 & 5 \end{bmatrix} = \begin{bmatrix} 5 & 7 & 12 \\ 7 & 10 & 17 \\ 12 & 17 & 29 \end{bmatrix} = \begin{bmatrix} 5 & 7 & 12 \\ 7 & 10 & 17 \\ 12 & 17 & 29 \end{bmatrix}$$

Here, Column 3 is a linear combination of column 1 and 2.

$$C_3 = C_1 + C_2$$

Hence,  $\text{rank}(X^T X) < 3$ ,  $X^T X$  is not invertible

Also, to find  $\min_{\beta} \|y - X\beta\|_2^2$ , let's calculate

$$X^T X \beta = X^T y$$

$$\rightarrow \begin{bmatrix} 5 & 7 & 12 \\ 7 & 10 & 17 \\ 12 & 17 & 29 \end{bmatrix} \begin{bmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \end{bmatrix} = \begin{bmatrix} 24 \\ 34 \\ 58 \end{bmatrix}$$

Let's perform Gaussian Elimination on the augmented matrix

$$\left[ \begin{array}{ccc|c} 5 & 7 & 12 & 24 \\ 7 & 10 & 17 & 34 \\ 12 & 17 & 29 & 58 \end{array} \right]$$

After series of row-wise operations, we get

$$\left[ \begin{array}{ccc|c} 1 & 0 & 1 & 2 \\ 0 & 1 & 1 & 2 \\ 0 & 0 & 0 & 0 \end{array} \right]$$

$$\Rightarrow \begin{aligned} \beta_1 + \beta_3 &= 2 \\ \beta_2 + \beta_3 &= 2 \end{aligned}$$

Only 2 pivotal points which also implies  $\text{rank}(X^T X) = 2$



Hence  $\beta_{OLS} = \begin{bmatrix} 2 - \beta_3 \\ 2 - \beta_3 \\ \beta_3 \end{bmatrix}$

$\beta_3$  is a free variable.

for any value of  $\beta_3$ , it will result in the same error.

Hence, multiple minimisers exist

Problem 3 →

$$\textcircled{3} \quad y_i = x_i^T \beta + \varepsilon_i \quad \varepsilon_i \sim N(0, \sigma^2)$$

$$\text{Given } \hat{\beta} = \operatorname{argmin}_{\beta} \left\{ \|y - X\beta\|_2^2 + \lambda \beta_2^2 + \frac{\lambda}{2} \beta_3^2 \right\}$$

Maximum Likelihood Estimate of  $\beta$  is given by (a posteriori)

$$L(X, y) = P(\beta) P(x, y)$$

$$= P(\beta) \prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(y_i - x_i^T \beta)^2}{2\sigma^2}\right)$$

$$\sigma = 1$$

$$\Rightarrow P(\beta) \left(\frac{1}{\sqrt{2\pi}}\right)^n \exp\left(-\frac{1}{2} \sum_{i=1}^n (y_i - x_i^T \beta)^2\right)$$

$$\Rightarrow \left(\frac{1}{\sqrt{2\pi}}\right)^n \exp\left(-\frac{1}{2} \sum_{i=1}^n (y_i - x_i^T \beta)^2 + \log P(\beta)\right)$$

maximising above equation, is similar to

$$\min_{\beta} \sum_{i=1}^n (y_i - x_i^T \beta)^2 - \log P(\beta)$$

$$= \min_{\beta} \|y - X\beta\|_2^2 - 2\log P(\beta) \quad - \textcircled{1}$$

We need to find the distribution of  $\beta$ , such that

$$-2\log P(\beta) \approx \lambda \beta_2^2 + \frac{\lambda}{2} \beta_3^2 \quad (\text{minimising is same})$$

$$\text{if we consider } \beta \sim \left( \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}, \begin{bmatrix} \infty & 0 & 0 \\ 0 & 1/\lambda & 0 \\ 0 & 0 & 2/\lambda \end{bmatrix} \right)$$

$$\begin{aligned} P(\beta) &= \frac{1}{\sqrt{\det(2\pi\Sigma)}} \exp\left(-\frac{1}{2} \beta^T \Sigma^{-1} \beta\right) \\ &= C \cdot \exp\left(-\frac{1}{2} [\beta_1 \ \beta_2 \ \beta_3] \begin{bmatrix} 0 & 0 & 0 \\ 0 & \lambda & 0 \\ 0 & 0 & \lambda/2 \end{bmatrix} \begin{bmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \end{bmatrix}\right) \\ &= C \cdot \exp\left(-\frac{1}{2} [0 + \lambda \beta_2^2 + \frac{\lambda}{2} \beta_3^2]\right) \end{aligned}$$



Substituting  $P(\beta)$  in ①

$$\Rightarrow \min_{\beta} \|y - X\beta\|_2^2 - 2 \left[ \log C - \frac{1}{2} \left[ \lambda \beta_2^2 + \frac{\lambda}{2} \beta_3^2 \right] \right]$$

$$= \min_{\beta} \|y - X\beta\|_2^2 + \lambda \beta_2^2 + \frac{\lambda}{2} \beta_3^2$$

Hence,  $\beta \sim \left( \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}, \begin{bmatrix} \infty & 0 & 0 \\ 0 & \frac{1}{\lambda} & 0 \\ 0 & 0 & \frac{2}{\lambda} \end{bmatrix} \right)$

Problem 4  $\rightarrow$

$$(4) \quad x_i^0 \sim N\left(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} 1 & 0 \\ 0 & 2 \end{bmatrix}\right) \quad A = \begin{bmatrix} 1 & 2 \\ 3 & 4 \end{bmatrix} \quad y_i^0 = Ax_i^0$$

We know that,  $E[A\varepsilon] = AE[\varepsilon]$   
and  $\text{Cov}(A\varepsilon) = A\text{Cov}(\varepsilon)A^T$

a)

$$E(y_i) = E[Ax_i] = AE[x_i] = A \begin{bmatrix} 0 \\ 0 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}$$

$$\begin{aligned} \text{Cov}(y_i) &= \text{Cov}(Ax_i) = A\text{Cov}(x_i)A^T \\ &= \begin{bmatrix} 1 & 2 \\ 3 & 4 \end{bmatrix} \begin{bmatrix} 1 & 0 \\ 0 & 2 \end{bmatrix} \begin{bmatrix} 1 & 3 \\ 2 & 4 \end{bmatrix} \\ \text{Cov}(y_i) &= \begin{bmatrix} 9 & 19 \\ 19 & 41 \end{bmatrix} \end{aligned}$$

b) Now,  $x_i \sim N\left(\begin{bmatrix} 1 \\ 1 \end{bmatrix}, \begin{bmatrix} 1 & 0 \\ 0 & 2 \end{bmatrix}\right)$

$$E(y_i) = E[Ax_i] = AE[x_i] = \begin{bmatrix} 1 & 2 \\ 3 & 4 \end{bmatrix} \begin{bmatrix} 1 \\ 1 \end{bmatrix} = \begin{bmatrix} 3 \\ 7 \end{bmatrix}$$

$$\text{Cov}(y_i) = A\text{Cov}(x_i)A^T$$

Since the covariance matrix of  $x_i$  is same as

(a),  $\text{Cov}(y_i) = \begin{bmatrix} 9 & 19 \\ 19 & 41 \end{bmatrix}$

```
% Solution to HW1 - Problem 5
```

```
X = mvnrnd([0;0;0], [2,0,0;0,4,0;0,0,10],100); % Generating X for 100 samples
```

```
E = normrnd(0, sqrt(12), [100,1]); % Generating error for 100 samples
```

```
B = [0;1;-1]; % Regression coefficients
```

```
Y = X*B+E;
```

```
% (5a)
```

```
B_ols = inv(X'*X)*X'*Y; % Calculating regression parameters using ordinary least squares
```

```
disp("5a")
```

```
disp("B_ols : ")
```

```
disp(B_ols)
```

```
z = 1.645; % For constructing 90% confidence interval, z = 1.645
```

```
sd = sqrt(12);
```

```
%sqrt(sum((Y - X*B_ols).^2)/100); % Calculating standard deviation from the sample
```

```
v = diag(inv(X'*X)) ;
```

```
disp("90% Confidence Interval for B (B_1, B_2 and B_3): ")
```

```
CI = [B_ols - (z * (sd * sqrt(v))), B_ols + (z * (sd * sqrt(v)))];
```

```
disp(CI)
```

```
disp("Yes, the interval covers B_1");
```

```
disp(" ")
```

```
% (5b)
```

```
disp("5b")
```

```
count = 0;
```

```
for i=1:500
```

```
    E = normrnd(0, sqrt(12), [100,1]);
```

```
    Y = X*B+E;
```

```
    B_ols = inv(X'*X)*X'*Y;
```

```
    sd = sqrt(12); % sqrt(sum((Y - X*B_ols).^2)/100); % Calculating standard deviation from  
the sample
```

```
    v = diag(inv(X'*X));
```

```
    CI = [B_ols - (z * (sd * sqrt(v))), B_ols + (z * (sd * sqrt(v)))];
```

```
    if B(1) > CI(1,1) && B(1) < CI(1,2)
```

```
        count = count + 1;
```

```
    end
```

```
end
```

```
disp("Percentage of time confidence interval covers B_1 is")
```

```
disp((count*100)/500)
```

```
% (5c)
```

```
disp("5c")
```

```
E = -6 + (6+6)*rand(100,1);
```

```
B_ols = inv(X'*X)*X'*Y; % Calculating regression parameters using ordinary least squares
```

```
z = 1.645; % For constructing 90% confidence interval, z = 1.645
```

```
sd = sqrt(12); % sqrt(sum((Y - X*B_ols).^2)/100); % Calculating standard deviation from the  
sample
```

```
v = diag(inv(X'*X));
```

```
CI = [B_ols - (z * (sd * sqrt(v))), B_ols + (z * (sd * sqrt(v)))];
```

```
count = 0;
```

```
for i=1:500
```

```
    E = -6 + (6+6)*rand(100,1);
```

```
    Y = X*B+E;
```

```
    B_ols = inv(X'*X)*X'*Y;
```

```
    sd = sqrt(12); % sqrt(sum((Y - X*B_ols).^2)/100); % Calculating standard deviation from  
the sample
```

```
    v = diag(inv(X'*X)) ;
```

```
    CI = [B_ols - (z * (sd * sqrt(v))), B_ols + (z * (sd * sqrt(v)))];
```

```
    if B(1) > CI(1,1) && B(1) < CI(1,2)
```

```
        count = count + 1;
```

```
    end
```



```
end
disp("Percentage of time confidence interval covers B_1 is")
disp((count*100)/500)
```

5a)

```
B_ols :
    -0.1862
     0.8574
    -1.2067
```

90% Confidence Interval for B (B\_1, B\_2 and B\_3):

```
    -0.5682    0.1957
     0.5353    1.1796
    -1.3993   -1.0141
```

Yes, the interval covers B\_1

5b)

```
Percentage of time confidence interval covers B_1 is
    89.4000
```

5c)

```
Percentage of time confidence interval covers B_1 is
    88.4000
```

```

% Solution to Problem 6
data = readtable('bikeshare.csv');
Y = data.cnt;
X = [ones(size(Y)), data.holiday, data.temp, data.hum, data.windspeed];

%Ordinary Least Squares
disp('6a') %Ordinary Least Squares
B_ols = inv(X'*X)*X'*Y;
disp('B_olsf for [intercept, holiday, temperature, humidity, windspeed]')
disp(B_ols)
disp("Coefficients B_2, B_4 and B_5 for attributes holiday, humidity and windspeed are negative, which indicate that")
disp("usage of bike is low when there is a holiday, when humidity and windspeed is high. We also see that magnitude ")
disp("of B_3 is high, indicating that temperature has a high positive correlation with bike usage - probably people ")
disp("like to bike during summers. Hence coefficients seem convincing.")
disp(" ")

disp("6b")
%Ridge Regression
s = size(X);
I = eye(s(2));

disp('Ridge Regression with lambda = 0.1')
lambda = 0.1;
B_ridge = inv((X'*X) + (lambda * I))*X'*Y

disp('Ridge Regression with lambda = 1')
lambda = 1;
B_ridge = inv((X'*X) + (lambda * I))*X'*Y

disp('Ridge Regression with lambda = 10')
lambda = 10;
B_ridge = inv((X'*X) + (lambda * I))*X'*Y

disp('Ridge Regression with lambda = 100')
lambda = 100;
B_ridge = inv((X'*X) + (lambda * I))*X'*Y

disp('Ridge Regression with lambda = 1000')
lambda = 1000;
B_ridge = inv((X'*X) + (lambda * I))*X'*Y
disp(" ")
disp(" - - - - - ")

```

```

6a)
B_olsf for [intercept, holiday, temperature, humidity, windspeed]
    1.0e+03 *

    4.1156
   -0.6136
    6.6103
   -3.1093
   -4.8085

```

Coefficients B\_2, B\_4 and B\_5 for attributes holiday, humidity and windspeed are negative,



which indicate that  
usage of bike is low when there is a holiday, when humidity and windspeed is high. We also  
see that maginitude  
of B\_3 is high, indicating that temperature has a high positive correlation with bike usage  
- probably people  
like to bike during summers. Hence coefficients seem convincing.

6b)

Ridge Regression with  $\lambda = 0.1$

B\_ridge =

```
1.0e+03 *  
  
4.0568  
-0.6101  
6.5955  
-3.0493  
-4.6627
```

Ridge Regression with  $\lambda = 1$

B\_ridge =

```
1.0e+03 *  
  
3.6454  
-0.5814  
6.4484  
-2.5952  
-3.6448
```

Ridge Regression with  $\lambda = 10$

B\_ridge =

```
1.0e+03 *  
  
2.6060  
-0.4050  
5.0866  
-0.7355  
-0.9642
```

Ridge Regression with  $\lambda = 100$

B\_ridge =

```
1.0e+03 *  
  
2.3612  
-0.0637  
2.2269  
1.0719  
0.2401
```

Ridge Regression with  $\lambda = 1000$

B\_ridge =

```
1.0e+03 *
```

1.4655  
0.0262  
0.8656  
0.8852  
0.2555

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