Physics-Informed Neural ODEs (PINODE)

Aleksei Sholokhov, Steven Brunton, J. Nathan Kutz, Hassan Mansour, Saleh Nabi

Tuesday 16th May, 2023

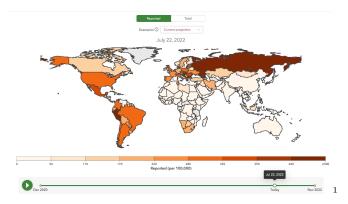
Plan of the Defense

Show topics and published papers. Mention covid

Feature Selection for Mixed-Effect Models

Mixed-effect models

- Used for analyzing combined data across a range of groups.
- ▶ Use covariates to separate the **population variability** from the **group variability**.
- **Borrow strength** across groups to estimate key statistics.

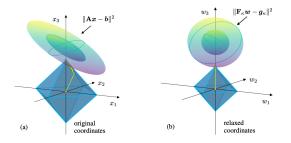


Feature Selection for Mixed-Effect Models

Practitioners:

- Often seek sparse models that select most informative covariates.
- ▶ Want to be **flexible but efficient** in using various sparsity-promoting terms.
- Want a library to be universal and compatible with e.g. sklearn.

Sparse Relaxed Regularized Regression (SR3) [9] showed great results for t linear models:



Goal: create a feature selection library that uses a relaxation approach for feature-selection in mixed-effect models.

Linear Mixed-Effect (LME) Models

Dataset: m groups (X_i, Z_i, y_i) , i = 1, ... m, each has n_i observations

- $X_i \in \mathbb{R}^{n_i \times p}$ group *i* design matrix for fixed features
- ▶ $Z_i \in \mathbb{R}^{n_i \times q}$ group *i* design matrix for random features
- $y_i \in \mathbb{R}^{n_i}$ group i observations

Model:

$$y_i = X_i \beta + Z_i u_i + \varepsilon_i$$

$$\varepsilon_i \sim \mathcal{N}(0, \Lambda_i)$$

$$u_i \sim \mathcal{N}(0, \Gamma)$$

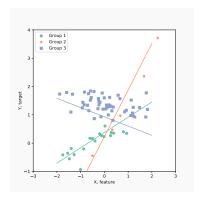
Equivalently:

$$y_i = X_i \beta + \omega_i$$

$$\omega_i \sim \mathcal{N}(0, Z_i \Gamma Z_i^T + \Lambda_i)$$

Simplifying assumption:

$$\Gamma = \operatorname{diag}\left(\left(\right)\gamma\right)$$



Notation

$$\begin{aligned} y_i &= X_i \beta + Z_i u_i + \varepsilon_i & i = 1 \dots m \\ \varepsilon_i &\sim \mathcal{N}(0, \Lambda_i) \\ u_i &\sim \mathcal{N}(0, \Gamma) \end{aligned} \tag{1}$$

- ▶ p number of fixed features, q number of random effects.
- $\beta \in \mathbb{R}^p$ fixed effects, or mean effects
- $\mathbf{v}_i \in \mathbb{R}^q$ random effects
- ▶ Γ ∈ $\mathbb{R}^{q \times q}$ covariance matrix of random effects, often Γ = diag ((γ))
- $ightharpoonup arepsilon_i \in \mathbb{R}^{n_i}$ observation noise
- ▶ $\Lambda_i \in R^{n_i \times n_i}$ covariance matrix for noise

Unknowns: β , u_i , γ , sometimes Λ_i .

Likelihood for Mixed Models

Optimization problem:

$$\mathcal{FS} - \mathcal{LME} \quad \min_{\beta \in \mathbb{R}^{p}, \ \gamma \in \mathbb{R}^{q}_{+}} \mathcal{L}(\beta, \gamma) + R(\beta, \gamma) \tag{2}$$

Where \mathcal{L} :

$$\mathcal{L}(\beta, \gamma) = \sum_{i=1}^{m} \frac{1}{2} (y_i - X_i \beta)^T (Z_i \Gamma Z_i^T + \Lambda_i)^{-1} (y_i - X_i \beta) +$$

$$+ \frac{1}{2} \log \det (Z_i \Gamma Z_i^T + \Lambda_i), \quad \Gamma = \operatorname{diag} ((\gamma))$$
(3)

- $ightharpoonup \mathcal{L}(eta,\gamma)$ is smooth on its domain, quadratic w.r.t. eta and $ar{\eta}$ -weakly-convex w.r.t. γ .
- $ightharpoonup R(\beta, \gamma)$ is closed, proper, with easily computed prox operator

Regularization

 $ightharpoonup R(\beta, \gamma)$ is closed, proper, with easily computed *prox operator*

$$\begin{aligned} \operatorname{prox}_{\alpha R + \delta_{\mathcal{C}}}(\tilde{\beta}, \tilde{\gamma}) &:= \operatorname*{argmin}_{(\beta, \gamma) \in \mathcal{C}} R(\beta, \gamma) + \frac{1}{2\alpha} \| (\beta, \gamma) - (\tilde{\beta}, \tilde{\gamma}) \|_2^2, \\ & \text{where } \mathcal{C} := \mathbb{R}^p \times R_+^q \end{aligned} \tag{4}$$

Examples:

- $ightharpoonup R(x) = \lambda \sum_{i=1}^{p} w_i ||x_j||_1 \text{LASSO}$ and Adaptive LASSO penalties [1, 6]
- ► $R(x) = \lambda ||x||_0 \ell_0$ penalty [8, 5]
- ightharpoonup R(x) SCAD penalty ([2, 3])

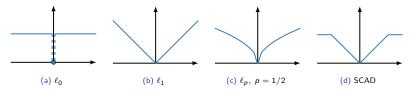


Figure: Four commonly-used regularizers which promote sparsity

SR3-Relaxation for Mixed-Effect Models ($\mathcal{MSR}3$)

Original problem $FS - \mathcal{LME}$:

$$\min_{\beta \in \mathbb{R}^p, \ \gamma \in \mathbb{R}^q_+} \mathcal{L}(\beta, \gamma) + R(\beta, \gamma) \tag{5}$$

Relaxed problem MSR3:

$$\min_{\beta,\tilde{\beta}\in\mathbb{R}^{p},\,\gamma,\tilde{\gamma}\in\mathbb{R}^{q}_{+}} \mathcal{L}(\beta,\gamma) + \phi_{\mu}(\gamma) + \kappa_{\eta}(\beta-\tilde{\beta},\gamma-\tilde{\gamma}) + R(\tilde{\beta},\tilde{\gamma})$$
(6)

where the $\emph{relaxation}$ κ_{η} decouples the likelihood and the regularizer

$$\kappa_{\eta}(\beta - \tilde{\beta}, \gamma - \tilde{\gamma}) := \frac{\eta}{2} \|\beta - \tilde{\beta}\|_{2}^{2} + \frac{\eta}{2} \|\gamma - \tilde{\gamma}\|_{2}^{2}, \quad \eta > \bar{\eta}$$
 (7)

and the perspective mapping ϕ_{μ} replaces $\gamma \geq$ 0 with a log-barrier

$$\phi_{\mu}(\gamma) := \begin{cases} -\mu \sum_{i=1}^{q} \ln(\gamma_i/\mu), & \mu > 0\\ \delta_{\mathbb{R}^q_+}(\gamma), & \mu = 0\\ +\infty, & \mu < 0 \end{cases}$$
(8)

Value Function Reformulation

 $\mathcal{MSR}3$ -relaxation replaces the original likelihood $\mathcal L$ with a value function $u_{\eta,\mu}$:

$$v_{\eta,\mu}(\tilde{\beta},\tilde{\gamma}) := \min_{(\beta,\gamma)} \mathcal{L}_{\eta,\mu}((\beta,\gamma),(\tilde{\beta},\tilde{\gamma}))
 := \min_{(\beta,\gamma)} \mathcal{L}(\beta,\gamma) + \phi_{\mu}(\gamma) + \kappa_{\eta}(\beta - \tilde{\beta},\gamma - \tilde{\gamma})$$
(9)

so MSR3-formulation (6) becomes

$$\min_{\beta \in \mathbb{R}^p, \ \gamma \in \mathbb{R}^q_+} \nu_{\eta,\mu}(\tilde{\beta},\tilde{\gamma}) + R(\tilde{\beta},\tilde{\gamma}) \tag{10}$$

When η is larger than the weak-convexity constant

- \triangleright $v_{\eta,\mu}$ is well-defined and continuously differentiable.
- As $\mu \to 0$ and $\eta \to \infty$, cluster points of solutions to $\mathcal{MSR}3$ are first-order stationary points for $\mathcal{FS} \mathcal{LME}$
- \triangleright $v_{\eta,\mu}$ don't need to be evaluated precisely.

Value Function Reformulation

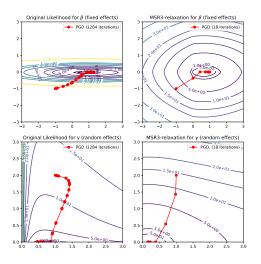


Figure: Comparison of the level-sets for the original likelihood (left) and $\mathcal{MSR}3$ -likelihood (right), for fixed (top) and random (bottom) effects.

Designing an Algorithm

 $G_{\nu,\eta}$ encodes both gradient of a Lagrangian (lines 1-2) and the complementarity condition (line 3):

$$G_{\nu,\eta}((\beta,\gamma,\nu),(\tilde{\beta},\tilde{\gamma})) := \begin{bmatrix} \nabla_{\beta} \mathcal{L}(\beta,\gamma) + \eta(\beta-\tilde{\beta}) \\ \nabla_{\gamma} \mathcal{L}(\beta,\gamma) + \eta(\gamma-\tilde{\gamma}) - \nu \\ \nu \bigodot \gamma - \mu \mathbf{1} \end{bmatrix}$$
(11)

We apply Newton method to G while geometrically decreasing μ . Lemma: For every $(\mu, \eta) \in \mathbb{R}_+ \times \mathbb{R}_{++}$,

$$(\hat{\beta}, \hat{\gamma}) = \underset{(\beta, \gamma)}{\operatorname{argmin}} \mathcal{L}_{\eta, \mu}((\beta, \gamma), (\tilde{\beta}, \tilde{\gamma}))$$

$$\iff \exists \hat{v} \in \mathbb{R}^{q}_{+} \text{ s.t. } G_{\nu, \eta}((\beta, \gamma, \hat{v}), (\tilde{\beta}, \tilde{\gamma})) = 0$$

$$(12)$$

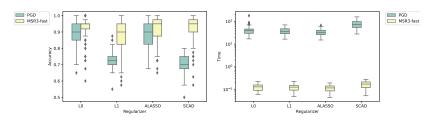
If $\mu > 0$, then $\hat{\mathbf{v}} = -\nabla \phi_{\mu}(\hat{\gamma})$, and if $\mu = 0$, then $\hat{\mathbf{v}}$ is the unique KKT multiplier associated with the constraint $0 \le \gamma$.

MSR3-fast Algorithm

```
1 progress ← True; iter = 0;
      2 \beta^+, \tilde{\beta}^+ \leftarrow \beta_0; \quad \gamma^+, \tilde{\gamma}^+ \leftarrow \gamma_0; \quad v^+ \leftarrow 1 \in \mathbb{R}^q; \quad \mu \leftarrow \frac{v^{+'} \gamma^+}{10\pi}
        3 while iter < max_iter and ||G_{\mu}(\beta^+, \gamma^+, \nu^+)|| > \text{tol} and progress
                               do
      \mathbf{4} \mid \ \mid \quad \beta \leftarrow \beta^+; \quad \gamma \leftarrow \gamma^+; \quad \tilde{\beta} \leftarrow \tilde{\beta}^+; \quad \tilde{\gamma} \leftarrow \tilde{\gamma}^+
        5 \mid [dv, d\beta, d\gamma] \leftarrow \nabla G_{\mu}((\beta, \gamma, v), (\tilde{\beta}, \tilde{\gamma}))^{-1} G_{\mu}((\beta, \gamma, v), (\tilde{\beta}, \tilde{\gamma}))
                                                      \alpha \leftarrow 0.99 \times \min \left(1, -\frac{\gamma_i}{d\alpha_i}, \forall i: d\gamma_i < 0\right)
      6 \beta^+ \leftarrow \beta + \alpha d\beta; \gamma^+ = \gamma + \alpha d\gamma; v^+ \leftarrow v + \alpha dv
                                      if \|\gamma^+ \odot v^+ - q^{-1}\gamma^+^T v^+ \mathbf{1}\| > 0.5 q^{-1} v^{+T} \gamma^+ then continue:
                                                 else
        8
                                                      \tilde{\beta}^+ = \operatorname{prox}_{\alpha R}(\beta^+); \ \tilde{\gamma}^+ = \operatorname{prox}_{\alpha R + \delta_{\mathbb{D}_+}}(\gamma^+); \ \mu = \frac{1}{10} \frac{v^{+} \gamma^+}{\sigma}
                                                 end
 10
                                                   progress = (\|\beta^+ - \beta\| \ge \text{tol or } \|\gamma^+ - \gamma\| \ge \text{tol or } \|\tilde{\beta}^+ - \tilde{\beta}\| \ge \text{tol or } \|\tilde{\beta}^+ - \tilde{\beta}
 11
                                                           \|\tilde{\gamma}^+ - \tilde{\gamma}\| > \mathsf{tol})
 12
                                                   iter += 1
13 end
14 return \tilde{\beta}^+, \tilde{\gamma}^+
```

Application to Synthetic Problems

- ▶ The number of fixed effects *p* and random effects *q* is 20.
- $\beta = \gamma = \frac{1}{2}[1, 2, 3, \dots, 10, 0 \dots, 0]$
- ▶ 9 groups with sizes [10, 15, 4, 8, 3, 5, 18, 9, 6]
- \triangleright $X_i \sim \mathcal{N}(0, I)^p$, $Z_i = X_i$, $\varepsilon_i \sim \mathcal{N}(0, 0.3^2 I)$
- ► Each experiment is repeated 100 times.
- ▶ Grid-search for $\eta \in [10^{-4}, 10^2]$, golden search for $\lambda \in [0, 10^5]$
- Final model is chosen to maximize BIC



- + $\mathcal{MSR}3$ -relaxation improves feature selection performance of the original likelihood.
- + $\mathcal{MSR}3$ -fast optimization accelerates the compute time by $\sim 10^2$.
- Initialization of η is problem-specific

Comparison to Other Libraries

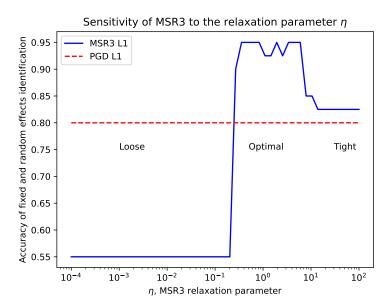
Algorithm	MSR3-Fast (ℓ_1)	glmmLasso ² [4]	$lmmLasso^3[7]$	PGD (ℓ_1)
Accuracy, %	88	48	66	73
FE Accuracy, %	86	52	47	56
RE Accuracy, %	91	45	84	91
Time, sec	0.19	1.37	11.51	38.39
Iterations, num	34	50	-	7693



 $^{^2 {\}sf https://rdrr.io/cran/glmmLasso/man/glmmLasso.html}$

³https://rdrr.io/cran/lmmlasso/

Choice of η



ℓ_0 -based Covariate Selection for Bullying Study from GBD

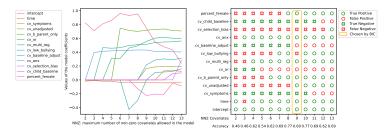


Figure: Fixed and random covariate selection for Bullying dataset⁴. The model selected 9 covariates, 7 of which were historically significant, and did not select 4 covariates, 1 of which was historically significant.

⁴Institute for Health Metrics and Evaluation (IHME). Bullying Victimization Relative Risk Bundle GBD 2020. Seattle, United States of America (USA), 2021.

Software

The code is available on GitHub: https://github.com/aksholokhov/pysr3

- All estimators are fully compatible to sklearn library.
- Implements SR3 for linear, generalized-linear, and linear mixed-effect models.
- Has tutorials, tests, and documentation.

Data-Driven Modeling of Physical Systems

- 1) People used to model physical systems with first-principle knowledge 2) Data-Driven modelling of dynamical systems became a big thing 3) However, it requires a lot of data
- 4) Incorporating prior knowledge is a big recent trend, so history does a spiral

Incorporating Knowledge into Models

1) There are multiple ways of incorporating knowledge into system 4) The overall umbrella term for it is physics-informed machine learning 2) Some use the equations that model phenomena 3) Some take aspects of it, e.g. symmetries and preservation laws, and forces A network to respect those 5) Our work falls into the first category of approaches

$$x \in \mathbb{R}^n$$

$$\frac{dx}{dt} = f(x)$$

$$x_0$$



$$x\in \mathbb{R}^n$$

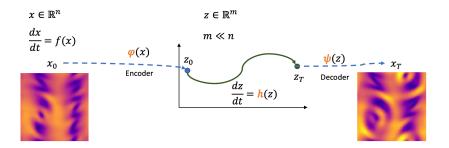
$$\frac{dx}{dt} = f(x)$$

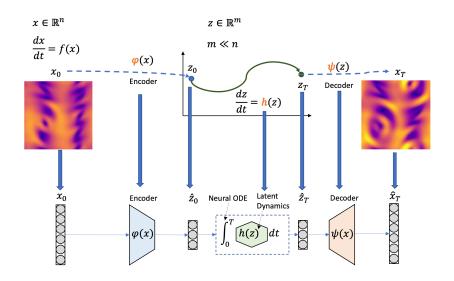
$$x_T = x_0 + \int_0^T f(x)dt$$











Physics-Informed Loss

1) We introduce physics to this system by adding a term which regularises latent gradient field. 2) In particular, it forces it to be equal to what a true physics should be under such projection. 3) We can not evaluate physics everywhere but we can at particular carefully-selected points. We call these points collocation points. 4) We feed a lot of collocation points and ask a network to do interpolation

Results: Extrapolation to Unknown Regimes

1) We show that network can indeed interpolate between collocations. Moreover, it can fill the whole unknown regimes of behaviour. (Duffing example with explanation)

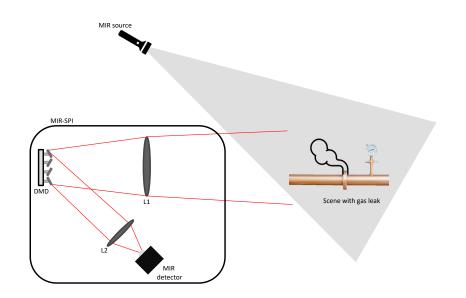
Results: Stable Long-Term Predictions

Fig 3.3.7

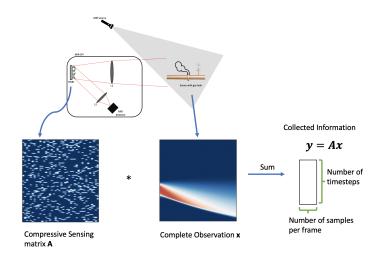
Results: Learning From Collocations

1) Finally we show that collocations can be even more useful than the data itself. 2) The difference is especially prominent in low-data regime. 3) It shows that collocations are powerful source of information and that the network can indeed interpolate between them.

Single-Pixel Imaging

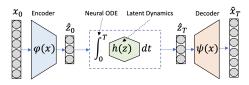


Single-Pixel Imaging

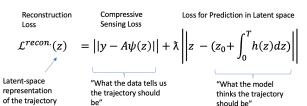


Compressive Sensing with Reduced-Order Models

Offline Step: Train a Data-Driven Reduced-Order Model



Online Step: Reconstruct Complete Observations by Optimizing in Latent Space



Results: Burger's Equation

Results on Burgers Maybe results on a harder problem

Results: Kolmogorov Flow OR Real Example

Conclusion

Results on Burgers Maybe results on a harder problem

References

References:

- Howard D. Bondell, Arun Krishna, and Sujit K. Ghosh. Joint Variable Selection for Fixed and Random Effects in Linear Mixed-Effects Models. <u>Biometrics</u>, 66(4):1069–1077, dec 2010.
- [2] Jianqing Fan and Runze Li. Variable Selection via Nonconcave Penalized Likelihood and its Oracle Properties. <u>Journal of the American Statistical Association</u>, 96(456):1348–1360, dec 2001.
- [3] Yingying Fan and Runze Li. Variable selection in linear mixed effects models. <u>The Annals of Statistics</u>, 40(4):2043–2068, aug 2012.
- [4] Andreas Groll and Gerhard Tutz. Variable selection for generalized linear mixed models by I 1-penalized estimation. <u>Statistics and Computing</u>, 24(2):137–154, 2014.
- [5] Richard H. Jones. Bayesian information criterion for longitudinal and clustered data. <u>Statistics in Medicine</u>, 30(25):3050–3056, nov 2011.
- [6] Bingqing Lin, Zhen Pang, and Jiming Jiang. Fixed and random effects selection by REML and pathwise coordinate optimization. <u>Journal of Computational and</u> <u>Graphical Statistics</u>, 22(2):341–355, 2013.
- [7] Jürg Schelldorfer, Peter Bühlmann, and SARA VAN DE GEER. Estimation for high-dimensional linear mixed-effects models using I1-penalization. <u>Scandinavian</u> <u>Journal of Statistics</u>, 38(2):197–214, 2011.
- [8] Florin Vaida and Suzette Blanchard. Conditional Akaike information for mixed-effects models. Biometrika, 92(2):351–370, jun 2005.
- [9] Peng Zheng and Aleksandr Aravkin. Relax-and-split method for nonconvex inverse problems. Inverse Problems, 36(9), 2020.