

Variable (Model) Selection

Chapter 5 of SPRM discusses some modeling strategies and metrics. We have discussed some of the latter already.

- Thus far we have mostly worked with example problems where predictor variables were identified in advance. Often in modeling, we may end up having several candidate models, all of which pass the usual hypothesis tests. How do we pick the best model?
- Variable selection is the process of choosing a subset of all available predictors. Depending on the modeling goal, we might choose differently, but in any case, we are interested in a model we can interpret or justify with respect to the problem at hand
- We have already considered model comparison (among nested models) via the F -tests, R^2 , adjusted R^2 , and other less well-defined, more subjective criteria

Variable (Model) Selection

- There are additional measures, such as likelihood-based criteria (AIC, BIC) for non-nested models, but no metric is universally “best”.
- When there are many variables available, however, comparing individual models by any means can become overwhelming. If there are p predictor variables, there are 2^p possible models (i.e, if there are 10 variables, we would have over 1000 candidate models, not including those with any interaction effects.). This also assumes just one predictor form (X , \sqrt{X} , etc) for each predictor.
- Thus, the critical questions then becomes: how to choose good models in an intelligent and efficient way. This applies perhaps more to exploratory modeling but formalizing model selection in all problems is useful, lest we appear to be ‘data-dredging’ or ‘fishing’ for results

Practical Model vs. the Ideal

- **To set criteria for models relative to one another, we need to consider the ideal world where there is a ‘correct’ model.**

What are consequences of including unnecessary variables or excluding necessary variables relative to this model?

For example, assume we have the most general model as follows (has q candidate predictors)

$$y_i = \beta_0 + \beta_1 x_{i1} + \dots + \beta_q x_{iq} + \epsilon_i$$

- **For this model, one of two conditions might hold (indicating the correct model)**
 1. All predictors have non-zero β - all are predictors
 2. Model should have $\beta_0, \beta_1, \beta_2, \dots, \beta_p$ non-zero but $\beta_{p+1}, \dots, \beta_q$ are zero, or these associated X s are not needed

Model vs. Ideal: Consequences

- We fit one model or the other, not knowing the 'true' state of nature. Then, what are consequences of
 - A. Instead of using all q predictors, we use only use $p < q$ predictors, and fit:

$$y_i = \beta_0 + \beta_1 x_{i1} + \dots + \beta_p x_{ip} + \epsilon_i$$

- If we fit a reduced model (omitting some X), we actually decrease variance on the β coefficients remaining in the model, but the estimates of β can be said to be biased if model (1) above is the correct model. Recall bias here means $E(\hat{\beta}) \neq \beta$ for the predictors in the model

- **The predictions \hat{y} are also biased.** In fact, for the model, the Mean Square Error (MSE) = $\text{bias}^2 + \text{variance}$. Bias is generally considered bad but we may have good reason to tolerate it a bit in some more advanced methods. Note that bias is unobservable practically.
- **Why are estimates said to be biased? A:** There is still ‘signal’ in the residuals that would have been explained by omitted X s

Model vs. Ideal: Consequences

- What are consequences if

B. We include all q predictors

- If we fit the full model, when we in fact should omit some X s, we increase variance for all β coefficients in the model, that is, estimating the extra $\beta_{p+1}, \dots, \beta_q$ adds (unnecessary) variance. We also increase variance in prediction.
- In practice bias is always unknown - we don't know the true state of nature, and cannot parse the MSE into bias vs. variance. We might tolerate bias (through omitting X s) to have reduced variance on β estimates and \hat{y} . In many situations, there will still be smaller MSE, due to the trade-offs that occur
- Note: these consequences don't hold if X are orthogonal (rarely met, but this is why low correlation among X s is a good property)

Purposes of Modeling

SPRM Section 5.3

- **An overarching consideration in model selection relates the the purpose and intended use. This might be:**
 - **Descriptive, Exploratory** (with respect to understanding relationships):

Here, we use the model to search for fundamental relationships. Typically start simply with only a few essential variables, and then choose variables and combinations of variables to build forward. Consideration of why variables should be in the model is required.
 - **Predictive:**

The focus is here on the high predictive ability of the model. We want predictions to be realistic and close to the sample data. Less consideration is given to which and how many variables are required, included, etc - 'black-box' modeling to some extent.

– **Explanatory:**

The goal here is to describe the process in a realistic and interpretable way. Lots of thinking required about which variables are important to have in the model. Parsimony is generally sought (smallest model that is complete). Confounding, effect modification must be thoroughly addressed

- **These purposes are not mutually exclusive, and ideally we'd like a little bit of all of these properties in our modeling strategy**

Variable Selection Procedures

- Given the large number of candidate models in many case, we may wish to specify a variable selection strategy *a priori* .
- Thinking through the problem carefully and specifying how to add/remove specific variables is itself a strategy (and one that should be employed!), but with many predictors this is not always clear
- Thus, we may want to use one of the following ‘objective’ approaches (SPRM Section 5.5)

Variable Selection Procedures

A. Forward Selection:

1. Begin with null or empty model (no predictors), add predictor with highest simple correlation with Y . A significance level for entry is established here and applied throughout
2. Add in the predictor that has the highest partial correlation with Y after adjusting for the X variable added above.
3. With new model, return to Step 2, looking at the remaining predictors via the same criterion. Enter those meeting the significance criterion.

Variable Selection Procedures

B. Backward Selection:

1. Begin with all predictors in the model. Remove weakest one (smallest t statistic) with significance level for removal established at this step and applied throughout
2. Re-assess remaining predictors and remove according to criterion above.
3. Stop when there are no more predictor variables to remove

Variable Selection Procedures

C. Stepwise Selection: (Foreword (FW) version)

1. Begin with null model, add predictor as in forward selection, with significance level specified
2. Add in the next predictor as in forward selection. At this stage consider omitting first predictor according to criteria for backward stepwise procedure. A separate significance level may be used for removing variables versus adding.
3. Proceed with adding and removing variables as above
4. Stop when no more adding/removing criteria are satisfied

Variable Selection Procedures

- **A Common (manual) approach (- typical univariable to multivariable analysis strategy)**
 1. Examine variables one at a time in relation to response, choose those meeting significance level specified
 2. Put all in multivariable model, assess significance at typical criterion
 3. Proceed with removing variables that are now nonsignificant
 4. Stop when no more nonsignificant predictors

This approach is thought of as a way of rigorously screening important predictors (meeting univariate and multivariable model significance criteria required). Perhaps not the best strategy . . . unless approached carefully

Variable Selection: Missing Values for Predictors

An important related issue is incompleteness of data for predictors, a common feature in observational data and sometimes even in carefully controlled experiments.

- To appropriately contrast models, one would need to work with the intersection of X variables that have non-missing values. This is naturally implied in backward selection at the outset.
- If the analysis cohort is not fixed at this or some value, then the number of observations/cases n will change over the modeling process. Predictor variable effects may change solely due to being based on different sets of observations (i.e. datasets).
- On the other hand, using the intersection of complete data discards information, decreases statistical power, and can lead to biased estimates
- *Missing data* techniques are needed - imputation methods, etc

Variable Selection: Example

From C&H Text Table 3.3, Supervisor performance data: The data consists of six candidate predictors in relation to an overall performance measure for supervisors.

X1 Handling of employee complaints
X2 Allowance of special privileges
X3 Opportunity to learn new things
X4 Raises based on performance
X5 Criticism of poor performance
X6 Rate of advancing to better jobs

```
. corr x1-x6  
(obs=30)
```

		x1	x2	x3	x4	x5	x6
x1		1.0000					
x2		0.5583	1.0000				
x3		0.5967	0.4933	1.0000			
x4		0.6692	0.4455	0.6403	1.0000		
x5		0.1877	0.1472	0.1160	0.3769	1.0000	
x6		0.2246	0.3433	0.5316	0.5742	0.2833	1.0000

- There is moderately high correlation among some predictors (it is a good idea to check this before applying any automated modeling procedures, due to problems introduced by multicollinearity)

```
. reg y x1 x2 x3 x4 x5 x6
```

Source	SS	df	MS	Number of obs = 30		
Model	3147.96634	6	524.661057	F(6, 23)	=	10.50
Residual	1149.00032	23	49.9565359	Prob > F	=	0.0000
Total	4296.96667	29	148.171264	R-squared	=	0.7326
				Adj R-squared	=	0.6628
				Root MSE	=	7.068

y	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
x1	.6131876	.1609831	3.81	0.001	.2801687	.9462066
x2	-.0730501	.1357247	-0.54	0.596	-.3538181	.2077178
x3	.3203321	.1685203	1.90	0.070	-.0282787	.668943
x4	.0817321	.2214777	0.37	0.715	-.3764293	.5398936
x5	.0383814	.1469954	0.26	0.796	-.2657018	.3424647
x6	-.2170567	.1782095	-1.22	0.236	-.5857111	.1515977
_cons	10.78708	11.58926	0.93	0.362	-13.18713	34.76128

```
.* check collinearity
. vif
```

Variable	VIF	1/VIF
x4	3.08	0.324862
x1	2.67	0.374945
x3	2.27	0.440326
x6	1.95	0.512403
x2	1.60	0.624652
x5	1.23	0.814260
Mean VIF	2.13	

- Some significant predictors, fairly high R^2 , collinearity not a problem (next page)
- Assuming there are not other model assumption problems, we can proceed to determine which sub-model might be best

Variable Selection: Example checking collinearity?

- Collinearity relates to the degree of correlation between predictors. if two predictors are highly correlated, one will "stand in" for the other. This can cause confusion in model selection strategies, and even misleading results.
- The Variance Inflation Factor (VIF) is defined as follows:

$$VIF_j = \frac{1}{1 - R_j^2}, \quad j = 1, \dots, p$$

- Here, R_j^2 is the r-squared value for a model predicting covariate X_j with all the other predictors. The higher the value is, the larger the inflation factor becomes. Values over 10 may indicate too much correlation

Variable Selection: Example

- We can always fit all $2^6 = 64$ models, (not recommended).
- With some theory/context to guide us, we might follow a specific strategy. For example, starting with the above model, retaining specific factors by design, and sequentially removing others, then inspecting residuals, etc. This would be tractable here, but in models with larger number of predictors, still may be a lengthy process.
- Alternatively, we can proceed with the various automated variable selection procedures. These always should be used with CAUTION.
- The different variable selection procedures (forward, backward, stepwise) are typically implemented via a single module in computer packages, as forward and backward procedures can be thought of as variations on the stepwise approach

Variable Selection: Example

- The help file for Stata's stepwise procedure, which can be used with numerous modeling approaches (not just linear regression):

```
. help sw
```

Stepwise estimation

```
sw cmd regress [var1 ...] [weight] [if exp] [in range], { pr(#) | pe(#) | pr(#)
                  pe(#) } [ forward lr hier lockterm1 cmd_options ]
```

predict after sw behaves the same as predict after the particular estimation command;
see help for the particular estimation command for details.

sw performs stepwise estimation, the flavor of which is determined by the options:

pr(#)	backward selection
pr(#) hier	backward hierarchical selection
pr(#) pe(#)	backward stepwise
pe(#)	forward selection
pe(#) hier	forward hierarchical selection

`pr(#) pe(#) forward forward stepwise`

`pr(#)` specifies the significance level for removal from the model; terms with `p>=pr()` are eligible for removal.

`pe(#)` specifies the significance level for addition to the model; terms with `p<pe()` are eligible for addition.

`forward` specifies the forward-stepwise method when both `pr()` and `pe()` are also specified. Specifying both `pr()` and `pe()` without `forward` results in backward stepwise. Note that specifying only `pr()` results in backward selection and specifying only `pe()` results in forward selection.

- Our stepwise procedure would be called 'forward stepwise' here

Variable Selection: Example

Example – forward selection:

The probability to enter option, p_e , was set to .99 (*only* for illustrative purposes, we typically set at more conventional level) so that all predictor variables will enter based on statistical significance according to criteria described earlier

```
. sw regress y x1 x2 x3 x4 x5 x6, pe(.99)
              begin with empty model
p = 0.0000 <= 0.9900 adding x1
p = 0.1278 <= 0.9900 adding x3
p = 0.2082 <= 0.9900 adding x6
p = 0.5616 <= 0.9900 adding x2
p = 0.6426 <= 0.9900 adding x4
p = 0.7963 <= 0.9900 adding x5
```

Source	SS	df	MS
Model	3147.96634	6	524.661057
Residual	1149.00032	23	49.9565359

Number of obs	=	30
F(6, 23)	=	10.50
Prob > F	=	0.0000
R-squared	=	0.7326
Adj R-squared	=	0.6628

Total		4296.96667	29	148.171264		Root MSE	=	7.068

y		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]		

x1		.6131876	.1609831	3.809	0.001	.2801687		.9462065
x3		.3203321	.1685203	1.901	0.070	-.0282787		.6689429
x6		-.2170567	.1782095	-1.218	0.236	-.585711		.1515977
x2		-.0730501	.1357247	-0.538	0.596	-.353818		.2077178
x4		.0817321	.2214777	0.369	0.715	-.3764293		.5398936
x5		.0383814	.1469954	0.261	0.796	-.2657018		.3424647
_cons		10.78708	11.58926	0.931	0.362	-13.18713		34.76128

At stage 1, Stata fits all models with just one variable, and picks the model whose variable has the smallest p-value.

If that p-value is smaller than p_e (in this case yes) then it fits the 2-variable models, and chooses the model which has the smallest p-value for the second variable, as long as that p-value is $< p_e$.

The procedure is repeated until adding any other remaining variables would have the added variable's p-value being $> p_e$.

Variable Selection: Analysis Example

- Following C&H, one might use variable entry criteria of $t_{.05, n-p}$ or $|t| > 1$. Using the first one (corresponding to $|t|$ of about 1.5)

- **forward selection**

```
. sw regress y x1 x2 x3 x4 x5 x6, pe(.15)
                        begin with empty model
p = 0.0000 < 0.1500 adding x1
p = 0.1278 < 0.1500 adding x3
```

Source	SS	df	MS	Number of obs = 30		
Model	3042.3177	2	1521.15885	F(2, 27)	=	32.74
Residual	1254.64897	27	46.4684804	Prob > F	=	0.0000
Total	4296.96667	29	148.171264	R-squared	=	0.7080
				Adj R-squared	=	0.6864
				Root MSE	=	6.8168

y	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
x1	.6435176	.1184774	5.43	0.000	.400422	.8866132
x3	.2111918	.1344037	1.57	0.128	-.0645818	.4869655
_cons	9.87088	7.061224	1.40	0.174	-4.617554	24.35931

Variable Selection: Example

- **backward selection** Here, we set the prob(remove) option, pr, was set to .33 to correspond to a $|t|$ -statistic of 1.0.

```
. sw regress y x1 x2 x3 x4 x5 x6, pr(.33)
                        begin with full model
```

```
p = 0.7963 >= 0.3300 removing x5
```

```
p = 0.6426 >= 0.3300 removing x4
```

```
p = 0.5616 >= 0.3300 removing x2
```

Source	SS	df	MS	Number of obs = 30		
-----+-----				F(3, 26)	=	22.92
Model	3117.85753	3	1039.28584	Prob > F	=	0.0000
Residual	1179.10914	26	45.3503515	R-squared	=	0.7256
-----+-----				Adj R-squared	=	0.6939
Total	4296.96667	29	148.171264	Root MSE	=	6.7343
-----+-----						
y	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----						
x1	.6227297	.1181464	5.271	0.000	.3798763	.8655832
x6	-.1869508	.1448537	-1.291	0.208	-.4847019	.1108003
x3	.312387	.1541997	2.026	0.053	-.0045751	.6293491
_cons	13.57774	7.5439	1.800	0.084	-1.928967	29.08445

Variable Selection: Example

Comments:

- All models are about the same (forward, backward, and just choosing on our own). Note that X_1 and X_3 would always be included, but coefficients, p-values, etc a bit different in each.
- We should pay attention to β coefficients to note inconsistencies or nonsensical results. Here, these are not radically different model to model (should not be, as vif was low).
- **Caution:** sw should not be done mechanically, without careful consideration of results. We must always apply context and common sense to this approach. The most frequent criticism of automated procedures relates to the fact that they will arrive at answers, whether correct or not. To illustrate:

Variable Selection: A Simulated Data Example

- **A simple simulation experiment:** - Generate a moderately small dataset of completely random response Y and corresponding large number of predictors X . Run different stepwise procedures.

```
. set obs 40

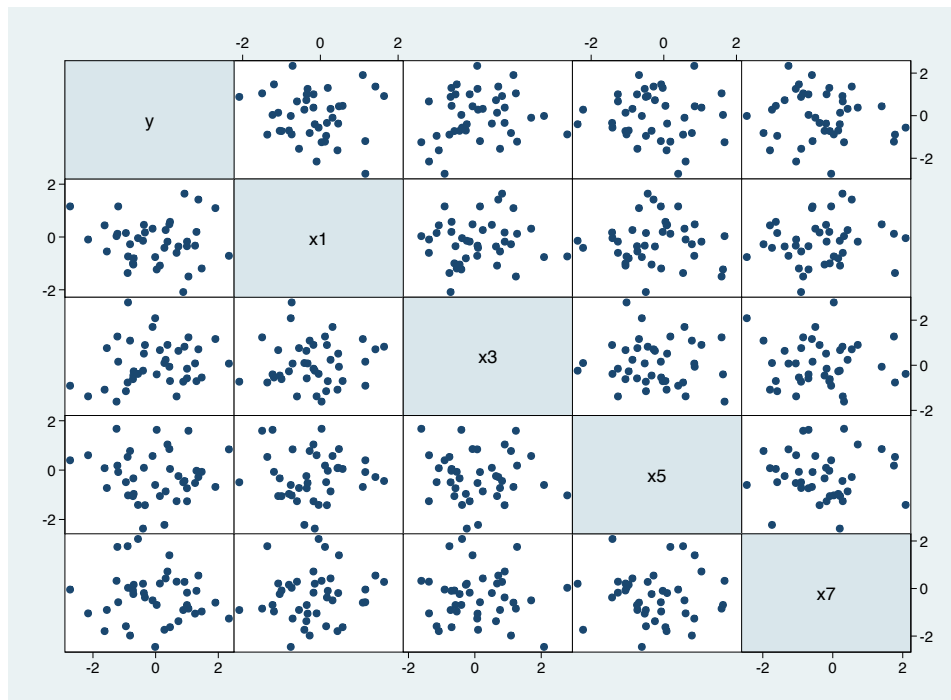
. set seed 44132254
. gen y=invnorm(uniform())

. gen x1=invnorm(uniform())

. for num 2/15: gen xX=invnorm(uniform())

-> gen x2=invnorm(uniform())
-> gen x3=invnorm(uniform())
-> gen x4=invnorm(uniform())
-> gen x5=invnorm(uniform())
-> gen x6=invnorm(uniform())
-> gen x7=invnorm(uniform())
. . .-
> gen x14=invnorm(uniform())
-> gen x15=invnorm(uniform())
```

```
. graph matrix y x1 x3 x5 x7
```



Nothing going on here in terms of Y vs. X relationship. From full model and selection procedures, we obtain the following:

- ordinary approach, test all initially

```
. regress y x1-x15
```

Source	SS	df	MS	Number of obs =	40
Model	23.8400721	15	1.58933814	F(15, 24) =	1.47
Residual	26.018029	24	1.08408454	Prob > F =	0.1958
Total	49.8581011	39	1.27841285	R-squared =	0.4782
				Adj R-squared =	0.1520
				Root MSE =	1.0412

y	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
x1	-.169245	.2573697	-0.66	0.517	-.70043	.3619399
x2	.0157974	.2589846	0.06	0.952	-.5187206	.5503154
x3	-.1640166	.2333674	-0.70	0.489	-.6456633	.3176301
x4	.5187105	.2476059	2.09	0.047	.0076771	1.029744
x5	-.0623123	.2048912	-0.30	0.764	-.485187	.3605624
x6	-.0997396	.1762514	-0.57	0.577	-.4635046	.2640253
x7	-.4994487	.350675	-1.42	0.167	-1.223206	.2243088
x8	.0897644	.2014728	0.45	0.660	-.3260551	.5055839
x9	-.1801775	.1904875	-0.95	0.354	-.5733245	.2129695
x10	.4787773	.2071367	2.31	0.030	.0512682	.9062863
x11	-.4467917	.2575459	-1.73	0.096	-.9783404	.084757

x12	-.2609168	.2570423	-1.02	0.320	-.7914259	.2695924
x13	-.2458519	.1920375	-1.28	0.213	-.6421979	.1504942
x14	-.3334932	.3279154	-1.02	0.319	-1.010277	.3432909
x15	-.5789872	.2251485	-2.57	0.017	-1.043671	-.1143034
_cons	-.1613593	.2109941	-0.76	0.452	-.5968297	.2741112

- NS model overall as expected, some β s are 'significant', but considering number of parameters/tests, none would pass (for example, if we used $.05/15 = .0033$ for significance). **Now run backward stepwise and forward stepwise procedures**

```
.* BACKWARD procedure

. sw reg y x1-x15, pr(.2)
      begin with full model
p = 0.9519 >= 0.2000 removing x2
p = 0.7241 >= 0.2000 removing x5
p = 0.6650 >= 0.2000 removing x8
p = 0.5513 >= 0.2000 removing x6
p = 0.6059 >= 0.2000 removing x1
p = 0.3503 >= 0.2000 removing x3
p = 0.3605 >= 0.2000 removing x12
```

p = 0.3166 >= 0.2000 removing x14

p = 0.3473 >= 0.2000 removing x13

Source	SS	df	MS	Number of obs =	40
Model	19.4624581	6	3.24374302	F(6, 33) =	3.52
Residual	30.395643	33	.921080091	Prob > F =	0.0084
Total	49.8581011	39	1.27841285	R-squared =	0.3904
				Adj R-squared =	0.2795
				Root MSE =	.95973

y	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
x11	-.3043337	.2139114	-1.42	0.164	-.7395397	.1308723
x15	-.5838627	.1856534	-3.14	0.004	-.9615774	-.206148
x10	.3292684	.149957	2.20	0.035	.0241785	.6343583
x4	.4684389	.1832929	2.56	0.015	.0955268	.8413511
x7	-.2236951	.1704186	-1.31	0.198	-.5704144	.1230243
x9	-.2142553	.1600246	-1.34	0.190	-.5398277	.1113171
_cons	-.2397916	.1655045	-1.45	0.157	-.576513	.0969297

```
.* FORWARD procedure
```

```
. sw reg y x1-x15, pe(.2)
```

```
begin with empty model
```

```
p = 0.0476 < 0.2000 adding x10
```

```
p = 0.1050 < 0.2000 adding x15
```

```
p = 0.0083 < 0.2000 adding x4
```

Source	SS	df	MS	Number of obs =	40
Model	15.519939	3	5.17331299	F(3, 36) =	5.42
Residual	34.3381621	36	.953837837	Prob > F =	0.0035
Total	49.8581011	39	1.27841285	R-squared =	0.3113
				Adj R-squared =	0.2539
				Root MSE =	.97665

y	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
x10	.3516723	.1512455	2.33	0.026	.0449321	.6584125
x15	-.5013255	.1781587	-2.81	0.008	-.8626481	-.1400028
x4	.4964532	.1776622	2.79	0.008	.1361375	.8567688
_cons	-.1433977	.1570761	-0.91	0.367	-.4619629	.1751674

- Using stepwise procedures with the same criteria, two models contrary to each other, with overall F-test significance, arise from this data. ???
- Models are still reasonably consistent (all identify X_4 , X_{10} , X_{15}), and, adjusting for multiple comparisons, there would effectively be no model.
- But, in real-life modeling, we have expectations of relationships, do not always rigorously adjust for multiple comparisons and number of models considered, and so we might accept findings as real
- Also, illustrates why external validation (checking model with data not used to build it) is highly valuable

Modeling with Collinearity Present

Here is a correlation matrix for blood pressure (Y) in relation to six predictors: age, body surface area, weight, duration of hypertension, pulse, and a stress measure

```
. corr bp weight bsa dur pulse stress age  
(obs=20)
```

	bp	weight	bsa	dur	pulse	stress	age
bp	1.0000						
weight	0.9501	1.0000					
bsa	0.8659	0.8753	1.0000				
dur	0.2928	0.2006	0.1305	1.0000			
pulse	0.7214	0.6593	0.4648	0.4015	1.0000		
stress	0.1639	0.0344	0.0184	0.3116	0.5063	1.0000	
age	0.6591	0.4073	0.3785	0.3438	0.6188	0.3682	1.0000

The correlation between weight and BSA is very high (0.875)

Modeling with Collinearity Present

Checking the VIF:

```
. quietly reg bp weight bsa dur pulse stress age  
. vif
```

Variable	VIF	1/VIF
-----+-----		
weight	8.42	0.118807
bsa	5.33	0.187661
pulse	4.41	0.226574
stress	1.83	0.545005
age	1.76	0.567277
dur	1.24	0.808205
-----+-----		
Mean VIF	3.83	

Weight and BSA have fairly high VIF - are linear functions of the other variables (with each other being the strongest predictors most likely)

Modeling with Collinearity Present

Run the stepwise model:

```
. sw reg bp weight bsa dur pulse stress age, pe(.05)
```

```
p = 0.0000 < 0.0500 adding weight
```

```
p = 0.0000 < 0.0500 adding age
```

```
p = 0.0078 < 0.0500 adding bsa
```

Source	SS	df	MS	Number of obs	=	20
Model	556.943853	3	185.647951	F(3, 16)	=	971.93
Residual	3.05614729	16	.191009206	Prob > F	=	0.0000
				R-squared	=	0.9945
				Adj R-squared	=	0.9935
Total	560	19	29.4736842	Root MSE	=	.43705

bp	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
weight	.9058219	.0489895	18.49	0.000	.8019688	1.009675
age	.7016201	.0439595	15.96	0.000	.60843	.7948101
bsa	4.627399	1.521068	3.04	0.008	1.40288	7.851918
_cons	-13.66724	2.646638	-5.16	0.000	-19.27786	-8.056613

Both weight and BSA are retained, model fit is very good

Modeling with Collinearity Present

What about this model?

reg bp weight		age						
Source		SS		df	MS	Number of obs	=	20
-----+-----						F(2, 17)	=	978.25
Model		555.176061		2	277.58803	Prob > F	=	0.0000
Residual		4.82393934		17	.283761138	R-squared	=	0.9914
-----+-----						Adj R-squared	=	0.9904
Total		560		19	29.4736842	Root MSE	=	.53269

bp		Coefficient		Std. err.	t	P> t	[95% conf. interval]	
-----+-----								
weight		1.032961		.0311563	33.15	0.000	.9672269	1.098695
age		.7082517		.0535141	13.23	0.000	.595347	.8211565
_cons		-16.57936		3.007463	-5.51	0.000	-22.92456	-10.23417

Dropping BSA variable results in very small loss in R^2 .
 Model dropping weight and keeping BSA variable is similar,
 not quite as good ($R^2 = 0.88$). It is not clear that both
 should be in the model, depends on questions of interest

Some Additional Criteria for Evaluating Models - non-nested models

SPRM 5.4

- We have already discussed the Mean Squared Error as the ‘variance’ of the whole model, $MSE = \frac{SSE}{n-p-1}$. Among two models, materially smaller MSE would be preferred in general
- The R^2 and adjusted version R_{adj}^2 : Latter is useful for *non-nested* models
- Two measures that can be used *whether or not models are nested* relate to the ‘information’ in the model. The Akaike Information Criterion (AIC) and Bayes Information Criterion (BIC) provide measures that balance information extracted from the data (fit) and number of parameters

Some Additional Criteria for Evaluating Models

AIC:

$$AIC = n \log_e(SSE_p/n) + 2p$$

where p is the *total* number of parameters (intercept included)

BIC:

$$BIC = n \log_e(SSE_p/n) + p \log_e(n)$$

- penalizes more heavily for having lots of parameters p relative to observations n
- These quantities for a given model are not particularly interpretable, but rather in contrasting two models, are useful. Among two candidates, the model with smaller AIC or BIC would be preferred. Models need not be nested, and these quantities take into account number of parameters

Example: Structured vs. Automated Approach

Surgical Unit data: The variables are

- x_1 : blood clotting measure
 - x_2 : prognostic index
 - x_3 : enzyme measure
 - x_4 : liver function measure
 - x_5 : age
 - x_6 : gender
 - y : survival - outcome
 - $\ln y$: log survival - outcome
-
- Predict $\ln(\text{survival time})$. It is suspected that there may be an interaction effect between x_2 and x_3 and thus we generate an interaction term of $x_2 \times x_3$ for consideration.

- **A systematic (manual) modeling approach**, followed by stepwise procedure

```
. gen x2x3=x2*x3
. reg lny x1 x2 x3 x4 x5 x6 x2x3
```

Source	SS	df	MS	Number of obs = 54		
				F(7, 46)	=	22.46
Model	9.90882875	7	1.41554696	Prob > F	=	0.0000
Residual	2.89889608	46	.06301948	R-squared	=	0.7737
				Adj R-squared	=	0.7392
Total	12.8077248	53	.241655186	Root MSE	=	.25104
lny	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
x1	.0950729	.0296256	3.21	0.002	.0354397	.1547061
x2	.0150846	.0069173	2.18	0.034	.0011608	.0290084
x3	.0177798	.0052646	3.38	0.001	.0071828	.0283768
x4	-.0012851	.0560579	-0.02	0.982	-.1141239	.1115537
x5	-.0049206	.0032525	-1.51	0.137	-.0114675	.0016263
x6	.0627443	.0735302	0.85	0.398	-.0852642	.2107528
x2x3	-.0000244	.0000769	-0.32	0.752	-.0001792	.0001303
_cons	3.903096	.5111992	7.64	0.000	2.874105	4.932087

```
.* store some results on this model for later
```

```
. est store A
```

The above is our full model, stored it as Model A. Now we look at some other models. Drop least significant predictors other than interaction

```
. reg lny x1 x2 x3 x5 x2x3
```

Source	SS	df	MS	Number of obs = 54		
Model	9.85956218	5	1.97191244	F(5, 48) = 32.11		
Residual	2.94816266	48	.061420055	Prob > F = 0.0000		
Total	12.8077248	53	.241655186	R-squared = 0.7698		
				Adj R-squared = 0.7458		
				Root MSE = .24783		

lny	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
x1	.0959735	.0217194	4.42	0.000	.0523037	.1396434
x2	.016032	.0067313	2.38	0.021	.0024979	.0295661
x3	.0185256	.0050542	3.67	0.001	.0083634	.0286878
x5	-.0048902	.0030783	-1.59	0.119	-.0110795	.001299
x2x3	-.0000331	.0000749	-0.44	0.660	-.0001838	.0001175

_cons		3.846634	.4977812	7.73	0.000	2.845777	4.84749
-------	--	----------	----------	------	-------	----------	---------

```
. set store B
```

```
.* drop interaction and other ~null variable (x5)
```

```
. reg lny x1 x2 x3
```

Source		SS	df	MS		Number of obs =	54
<hr/>							
Model		9.69918607	3	3.23306202		F(3, 50) =	52.00
Residual		3.10853876	50	.062170775		Prob > F =	0.0000
<hr/>							
Total		12.8077248	53	.241655186		R-squared =	0.7573
<hr/>							
						Adj R-squared =	0.7427
						Root MSE =	.24934

lny		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
<hr/>							
x1		.0954583	.0216921	4.40	0.000	.0518884	.1390283
x2		.01334	.0020347	6.56	0.000	.0092532	.0174268
x3		.0164517	.0016299	10.09	0.000	.0131779	.0197254
_cons		3.766176	.2267583	16.61	0.000	3.310718	4.221633

```
. est store C
```

```
.
```

```
.
.
.* FOR ILLUSTRATIVE PURPOSES, go too far, drop an important variable
```

```
. reg lny x1 x2
```

Source		SS	df	MS	Number of obs = 54		
-----+-----					F(2, 51)	=	9.09
Model		3.36505402	2	1.68252701	Prob > F	=	0.0004
Residual		9.44267081	51	.185150408	R-squared	=	0.2627
-----+-----					Adj R-squared	=	0.2338
Total		12.8077248	53	.241655186	Root MSE	=	.43029

lny		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----							
x1		.0630202	.0370214	1.70	0.095	-.0113034	.1373437
x2		.013129	.0035111	3.74	0.000	.0060801	.0201778
_cons		5.23573	.3000092	17.45	0.000	4.633437	5.838024

```
.
. * store results
. est store D
.
.
. *
```

```
.* fit some model not nested in models B through D
```

```
.
```

```
. reg lny x2 x3 x4
```

Source	SS	df	MS	Number of obs =	54
Model	9.19359439	3	3.06453146	F(3, 50) =	42.40
Residual	3.61413045	50	.072282609	Prob > F =	0.0000
Total	12.8077248	53	.241655186	R-squared =	0.7178
				Adj R-squared =	0.7009
				Root MSE =	.26885

lny	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
x2	.0110093	.0024043	4.58	0.000	.0061801	.0158385
x3	.0126091	.0019547	6.45	0.000	.0086831	.0165352
x4	.1297686	.0417491	3.11	0.003	.0459131	.2136241
_cons	4.40582	.1989817	22.14	0.000	4.006154	4.805487

```
. est store E
```

- Now contrast all models

```
. display information-based measures
. est stats A B C D E
```

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
A	54	-37.77142	2.342978	8	11.31404	27.22592
B	54	-37.77142	1.88797	6	8.224059	20.15796
C	54	-37.77142	.4577638	4	7.084472	15.04041
D	54	-37.77142	-29.54156	3	65.08312	71.05007
E	54	-37.77142	-3.611096	4	15.22219	23.17813

Note: N=Obs used in calculating BIC; see [R] BIC note

We use minimum of *AIC* or *BIC* as our model selection criteria, in addition to other considerations. So, the best is model C.

$$\ln y_i = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3$$

This model had best R^2 among 'smaller' models, best *AIC*, *BIC*

- **How about using stepwise regression `sw`?**

We specify both `pr` and `pe`. Typically, `pr` should be greater than `pe`.

```
. sw reg lny x1 x2 x3 x4 x5 x6 x2x3, pe(0.05) pr(0.1) forward
                        begin with empty model
p = 0.0000 < 0.0500 adding x2x3
p = 0.0026 < 0.0500 adding x1
p = 0.0017 < 0.0500 adding x3
p = 0.0300 < 0.0500 adding x2
p = 0.7722 >= 0.1000 removing x2x3
```

Source	SS	df	MS	Number of obs =	54
Model	9.69918607	3	3.23306202	F(3, 50) =	52.00
Residual	3.10853876	50	.062170775	Prob > F =	0.0000
Total	12.8077248	53	.241655186	R-squared =	0.7573
				Adj R-squared =	0.7427
				Root MSE =	.24934

lny	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
x2	.01334	.0020347	6.56	0.000	.0092532 .0174268

x1	.0954583	.0216921	4.40	0.000	.0518884	.1390283
x3	.0164517	.0016299	10.09	0.000	.0131779	.0197254
_cons	3.766176	.2267583	16.61	0.000	3.310718	4.221633

- **In this case, arrives at same 'best model'.** Note: interaction term entered before main effects (a generally undesirable property). Stepwise procedures provide means to constrain the form of model to some extent, preventing this anomaly

Summary: Variable Selection

- Earlier measures and new quantities introduced here (AIC, BIC) are useful, while context and 'domain knowledge' should guide modeling most
- Automated procedures can be useful but must be used cautiously. There is not universal agreement about approaches, although backward stepping procedures seem to be preferred
- Model selection is a large and evolving area of statistics, many more tools available