Lecture 9: Probit and Complementary Log-log Models for Binary Response Data

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Part I: Probit Model Motivating Data for Probit Model: *Biological Assay*

- Biological assay or *Bioassay* is a general type of scientific experiment investigating effects of agents on biological systems
- In one common type of bioassay, different concentrations of a chemical compound, drug, etc. are applied to batches of biological experimental entities (animals, tissues, cells...). The number responding at each dose is then recorded as a response variable.
- This type of experiment typically produces proportions as response data. For example,
 - Number of experimental animals in batch (dose group) i: n_i
 - Number of animals respond (0/1) in batch i: y_i
 - Concentration of chemical compound in batch i: di

Tolerance – a latent (unobserved) variable

Consider the *Toxicity of cypermethrin to moths* example.
 Low dose → few die; high dose → many die.

Table 1: Mortality of tobacco budworm moths 72 hours after exposure to cypermethrin.

Sex of moth	Dose of	Number affected
Sex of filotif		
	cypermethrin	out of 20
Male	1.0	1
	2.0	4
	4.0	9
	8.0	13
	16.0	18
	32.0	20
Female	1.0	0
	2.0	2
	4.0	6
	8.0	10
	16.0	12
	32.0	16

Tolerance (continued)

- Which ones die? Those with low *tolerance* than dose d_i for cypermethrin will.
- Idea: There exists a latent (unobserved) tolerance U, which has a distribution f(U). Those with tolerance higher than d_i would survive. The binary outcome (death/not) of each moth i depends on whether its latent tolerance (u_i) exceeds the current dose (d_i) .

$$Y = \begin{cases} 1, & \text{if } U_i \leq d_i \\ 0, & \text{otherwise} \end{cases}$$
 (1)

Then the probability of death when exposed to dose d_i is

$$p_i = P(U \le d_i) = \int_{-\infty}^{d_i} f(u) \, du$$
 (2)

Now consider U as a random variable that we are interested in.

Tolerance: Logistic Distribution

If *U* follows a *logistic distribution*:

$$f(u) = \frac{\exp\{(u - \mu)/\tau\}}{\tau [1 + \exp\{(u - \mu)/\tau\}]^2}, -\infty < u < \infty$$
 (3)

where $-\infty < \mu < \infty$, and $\tau > 0$. $E(U) = \mu$, $Var(U) = \pi^2 \tau^2 / 3$.

$$p_i = \int_{-\infty}^{d_i} f(u) \, du = \int_{-\infty}^{d_i} \frac{\exp\{(u - \mu)/\tau\}}{\tau [1 + \exp\{(u - \mu)/\tau\}]^2} \, du = \frac{\exp\{(d_i - \mu)/\tau\}}{1 + \exp\{(d_i - \mu)/\tau\}}$$
 (4)

Let $\beta_0 = -\mu/\tau$ and $\beta_1 = 1/\tau$, we can rewrite Eqn. (3) as:

$$p_{i} = \frac{\exp(\beta_{0} + \beta_{1}d_{i})}{1 + \exp(\beta_{0} + \beta_{1}d_{i})}$$
 (5)

This yields the logistic regression model that we are familiar with.

$$logit(p_i) = \beta_0 + \beta_1 d_i.$$
 (6)

Note the logit function is the inverse function of logistic function (see Lecture 3, Slide 8).

Tolerance: Normal (Gaussian) Distribution

More naturally, we may consider the tolerance variable U as a continuous variable following a *normal distribution*:

$$f(u) = \frac{1}{\sigma\sqrt{2\pi}} \exp\{-\frac{1}{2}(\frac{u-\mu}{\sigma})^2\}, -\infty < u < \infty$$
 (7)

Then,

$$p_{i} = \int_{-\infty}^{d_{i}} f(u) du = \frac{1}{\sigma \sqrt{2\pi}} \int_{-\infty}^{d_{i}} \exp\{-\frac{1}{2} (\frac{u - \mu}{\sigma})^{2}\} du = \Phi\left(\frac{d_{i} - \mu}{\sigma}\right)$$
(8)

Let $\beta_0 = -\mu/\sigma$ and $\beta_1 = 1/\sigma$, we can rewrite the above Eqn. as

$$p_i = \Phi(\beta_0 + \beta_1 d_i). \tag{9}$$

This yields the probit regression model,

$$\Phi^{-1}(p_i) = \beta_0 + \beta_1 d_i. \tag{10}$$

The probability of response is linked with a linear combination of predictors via a probit function – the inverse of the cumulative distribution function of the standard normal distribution.

Tolerance - Expressed as Effective Dose

More specifically, we may be interested in

- The dose that is expected to result in a $(100 \times p)\%$ response, for example for p = 0.50, 50% of subjects respond
- This quantity ED can be derived from the model. Which dose level can give a $p_i = 50\%$? This can be computed based on Eqn. (9).
- Based on the probit model, the effective dose ED estimate for a response probability of p is given by

$$\widehat{ED} = \frac{1}{\hat{\beta}_1} (\Phi^{-1}(p) - \hat{\beta}_0) \tag{11}$$

The effective dose estimate for logit model is

$$\widehat{ED} = \frac{1}{\hat{\beta}_1} (\log \frac{p}{1-p} - \hat{\beta}_0) \tag{12}$$

• ED_{50} , ED_{90} are often of interest

Example: Anti-pneumococcus serum

Table 2: Number of deaths from pneumonia amongst batches of 40 mice exposed to different doses of a serum

Dose of serum	Number of deaths		
	out of 40		
0.0028	35		
0.0056	21		
0.0112	9		
0.0225	6		
0.0450	1		

Anti-pneumococcus serum: $logit(p_i) = \overline{\beta_0 + \beta_1 log(d_i)}$

First, we fit the logistic model with which we are familiar.

```
. gen ldose=log(dose)
. glm y ldose, family (binomial n) nolog
Generalized linear models
                                          Number of obs =
                                          Residual df =
Optimization : ML
                                          Scale parameter =
Deviance = 2.80889589
                                          (1/df) Deviance = .9362986
Pearson = 2.917404799
                                          (1/df) Pearson = .9724683
Variance function: V(u) = u*(1-u/n)
                                          [Binomial]
Link function : g(u) = \ln(u/(n-u))
                                          [Logit]
                                                    = 4.715478
                                          AIC
                                                        = -2.019418
                                          BIC
Log likelihood = -9.788693777
                   OIM
            Coef. Std. Err. z P>|z| [95% Conf. Interval]
     Idose | -1.829621 .2545469 -7.19 0.000 -2.328524 -1.330719
            -9.189392 1.25511 -7.32 0.000 -11.64936 -6.729422
     cons
```

. predict yhat_logitmod
(option mu assumed; predicted mean count)

Anti-pneumococcus serum: probit $(p_i) = \beta_0 + \beta_1 \log(d_i)$

Now we fit the probit model by specifying the link function "link (probit)". Deviance is still good for evaluating goodness-of-fit.

```
glm y ldose, family (binomial n) link (probit) nolog
                                 default: [ink ( logit )
Generalized linear models
                                                 Number of obs =
Optimization : ML
                                                 Residual df
                                                 Scale parameter =
Deviance = 3.193092195 Grouped Binary Data
Pearson = 3.197015026 Grouped Binary Data
                                                 (1/df) Deviance = 1.064364
                                                 (1/df) Pearson = 1.065672
Variance function: V(u) = u*(1-u/n)
                                                 [Binomial]
Link function : g(u) = invnorm(u/n)
                                                 [Probit]
                                                 AIC
                                                                = 4.792317
                                                 BIC
Log likelihood = -9.98079193
                                                                   -1.635222
                  OIM
             Coef. Std. Err. z P>|z| [95% Conf. Interval]
      ldose | -1.054488 .1328879 -7.94
                                               0.000 \quad -1.314943 \quad -.7940325
               -5.276716 .6441273 -8.19
                                               0.000
                                                        -6.539183 -4.01425
      cons
```

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predict yhat probitmod (option mu assumed; predicted mean count)

Anti-pneumococcus serum: Effective dose

Based on Slide 7, we calculate the effective dose for 50 and 90% response: Model

```
. di (invnormal(0.5) – (-5.276716))/(-1.054488) –5.004055 &D \beta_0 \beta_1 . di exp(-5.004055) .00671068 . di (invnormal(0.9) – (-5.276716))/(-1.054488) –6.2193857 . di exp(-6.2193857) .00199047
```

- . gen phat_probit=yhat_probitmod/n
- . list y n dose phat_probit

_				
	y 	n	dose	phat_p~t
1.	35	40	.0028	.8216594
2.	21	40	.0056	.5756557
3.	9	40	.0112	.294556
4.	6	40	.0225	.1010245
5.	1	40	.045	.0223934
-	+			

The effective doses are $ED_{50} = 0.0067$ and $ED_{90} = 0.0020$ based on the log dose model with probit link.

Anti-pneumococcus serum: probit(p_i) = $\beta_0 + \beta_1 \log(d_i)$

This model still has a binomial outcome even though the link function is a function of the normal distribution

- The link function in a generalized linear model (GLM) answers "How is what is predicted by the linear function related to the mean of y (i.e., response variable)?"
 - Ordinary linear regression has the *identity* link, as the linear predictor directly predicts the mean of y for a given x
 - For binomial outcomes, the default link is the logit or log(p/1-p). Note that indeed p is the mean of y (mean of a sum of 0/1 responses is the proportion responding)
 - Probit regression is another GLM regression for **binary/binomial outcome**. The *probit* link indicates that the linear predictor predicts the probit transform of the mean p. The probit for p is the value from the (standard) normal distribution that is associated with the probability p.

Anti-pneumococcus serum: probit $(p_i) = \overline{\beta_0 + \beta_1 \log(d_i)}$

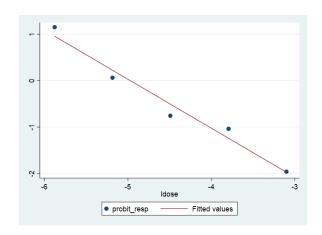
Can we just do the transform ourself and fit the model via ordinary regression. In this instance, we can.

```
. gen probdie=y/n
. gen probit_resp=invnormal(probdie)
```

list dose y n probdie ldose probit_resp

- I	+ dose	у	n	probdie	ldose	probit_~p
1. 2. 3. 4. 5.	.0028 .0056 .0112 .0225 .045	35 21 9 6 1	40 40 40 40 40	.875 .525 .225 .15 .025	-5.878136 -5.184988 -4.491841 -3.79424 -3.101093	1.150349 .0627067 755415 -1.036433 -1.959964
-	+					+

. twoway scatter probit_resp Idose



Anti-pneumococcus serum: probit $(p_i) = \beta_0 + \beta_1 \log(d_i)$

Here we will run a linear regression (OLS)

. regress probit_resp Idose

Source	SS	df	MS	Numb	er of obs	=	5
+				- F(1,	3)	=	95.29
Model	5.3558733	1	5.3558733	3 Prob	> F	=	0.0023
Residual	.168610761	3	.056203587	⁷ R-sq	uared	=	0.9695
+				- Adj I	R-squared	=	0.9593
Total	5.52448406	4	1.38112101	l Root	MSE	=	.23707
probit_resp	Coef.	Std. Err.	t	P> t	[95% C	onf.	Interval]
	-1.053787 -5.239319	.1079493 .4961589	-9.76 -10.56	0.002 0.002	-1.3973 -6.81831	_	7102445 -3.66032

Is it surprising? Estimates are very similar (not identical - least squares (LS) estimation used here). This method was widely used before availability of modern computing, because LS estimation can be carried out even without (good) computers.

Part II: The complementary log-log model – Tolerance follows a Gumbel distribution

If the tolerance variable U follows a *Gumbel distribution*:

$$f(u) = \frac{1}{\kappa} e^{(u-\alpha)/\kappa} \exp\{-e^{(u-\alpha)/\kappa}\}, \quad -\infty < u < \infty$$
 (13)

Where $-\infty < \alpha < \infty$ and $\kappa > 0$ are unknown parameters.

$$p_{i} = \int_{-\infty}^{d_{i}} f(u) du = \int_{-\infty}^{d_{i}} \frac{1}{\kappa} e^{(u-\alpha)/\kappa} \exp\{-e^{(u-\alpha)/\kappa}\} du = 1 - \exp\{-e^{(d_{i}-\alpha)/\kappa}\}$$
(14)

Let $\beta_0 = -\alpha/\kappa$ and $\beta_1 = 1/\kappa$, and we can rewrite Eqn (12) as

$$p_i = 1 - \exp\{-\exp(\beta_0 + \beta_1 d_i)\},$$
 or (15)

$$\log\{-\log(1-p_i)\} = \operatorname{cloglog}(p_i) = \beta_0 + \beta_1 d_i \tag{16}$$

This yields the complementary log-log regression model. The cloglog function links the probability with the linear combination of predictors. The effective dose estimate is $\widehat{ED} = \frac{1}{\hat{\beta}_1} \{ \log[-\log(1-p)] - \hat{\beta}_0 \}$. (17)

Anti-pneumococcus serum: $cloglog(p_i) = \beta_0 + \beta_1 log(d_i)$

Run a cloglog model by specifying the link function "link (cloglog)"

```
. glm y ldose, family (binomial n) link(cloglog) nolog
Generalized linear models
                                          Number of obs
Optimization : ML
                                          Residual df
                                          Scale parameter =
Deviance = 1.309885583 Crouped Sinary Data
Pearson = 1.361478279 CoF \checkmark
                                          (1/df) Deviance = .4366285
                                          (1/df) Pearson = .4538261
Variance function: V(u) = u*(1-u/n)
                                          [Binomial]
Link function : q(u) = \ln(-\ln(1-u/n))
                                          [Complementary log-log]
                                          AIC
                                                  = 4.415675
                                                     = -3.518428
Log likelihood = -9.039188624
                                          BIC
                   OIM
         y | Coef. Std. Err. z P>|z| [95% Conf. Interval]
     -7.511787 .9324851 -8.06
                                        0.000
                                                -9.339424 -5.68415
     cons
```

```
. predict yhat_cloglogmod
(option mu assumed; predicted mean y)
```

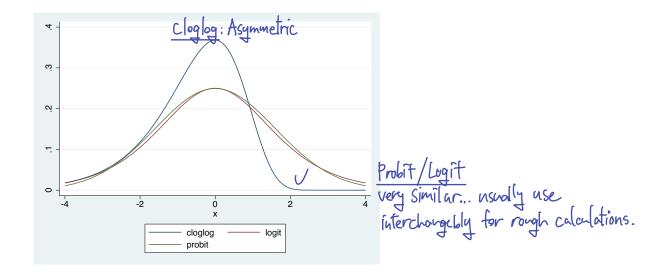
Deviance can still be used to evaluate the goodness-of-fit of the model.

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Tolerance distributions comparison

This figure shows the tolerance density function for the complementary log-log distribution (0,1), the logistic (0,1), and $N(0,1.6^2)$. The normal distribution is chosen to have the same mean and variance as the standard logistic distribution.

```
. twoway function cloglog = \exp(x) \cdot \exp(-\exp(x)), range(-4 4) 
 || function logit = \exp(x)/(1 + \exp(x))^2, range(-4 4) 
 || function probit = 0.39894 \cdot \exp(-0.5 \cdot (x/1.6)^2)/1.6, range(-4 4)
```

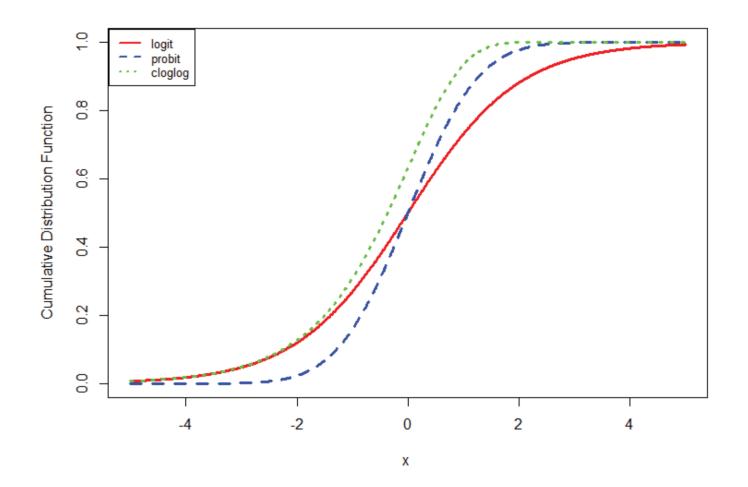


Generally, logit and probit models perform similarly. The normal distribution and the logistic distribution (for the latent tolerance variable) have similar shapes, but the logistic distribution has slightly longer/heavier tails.

Unlike logit/probit, the cloglog function is asymmetric, and may be used when the probability of response is very small or large. Depending on the scentific problems, different tolerance distributions might serve better as the theoretical model.

Comparing Transforms of p

Recall from Lecture 3 and Eqn (4), (8) and (14) in this lecture.



The cloglog function (the green dotted line) approaches 1 sharply but zero slowly and can be used to model rare event.

Examining the Three Models

Table 3: Model Comparison using Deviance.

	Model	Deviance	Deviance d.f.	P-value	\hat{ED}_{50}	
	logit	2.809	1	0.42	.006588 >	
Harder <	probit	3.193	1	0.36	.006711	no huge difference
to interest	cloglog	1.310 ل	1	0.73	.006029	

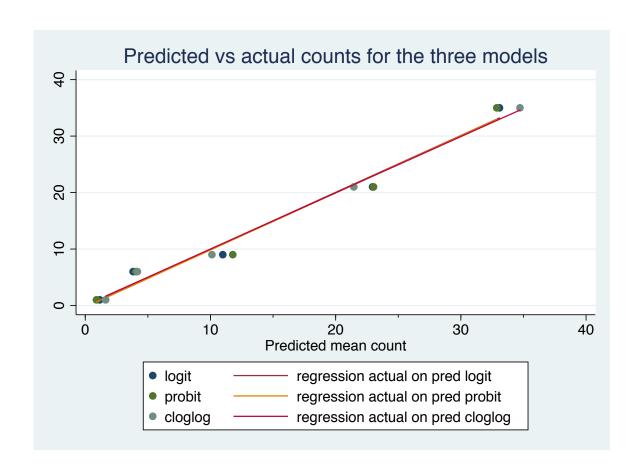
Predicted counts: Simply a better fit of cloglog.

May Still choose logif for better interpretation....

List Idose y n yhat_logitmod yhat_probitmod yhat_cloglogmod

-	 Idose	у	n	yhat_l~d	yhat_p~d	yhat_c~d
1. 2. 3. 4. 5.	-5.878136 -5.184988 -4.491841 -3.79424 -3.101093	35 21 9 6 1	40 40 40 40 40	33.08491 22.95006 10.98707 3.823065 1.154902	32.86637 23.02623 11.78224 4.040979 .8957366	34.72205 21.45233 10.11815 4.16571 1.634881
-	+					+

Examining the Three Models



There is no clear winner - complementary log-log may be a bit better overall. ED_{50} estimate appears closer to the empirical data (50% affected just exceeded for dose .0056)

Summary

Logit, Probit, and Cloglog Models In far Binary

- For binary outcome data, probit and cloglog are useful alternative models to the logistic regression.
- The probit model has a long history and may have performance similar to logit. It is often used when the binary response (0/1) is whether a latent variable exceed a threshold, and is also known as the latent variable model. For example, the tolerance in Bioassay data.
- The Complimentary Log-Log (cloglog) function is asymmetric. It is often used when the probability of an event is very small (or large).
- Deviance can still be used for assessing goodness-of-fit and for model comparisons (formally for nested models of the same link and informally for non-nested models).
- Logit is the most commonly used model for binary outcome data and logit function is the default link option in statistical software.