NON-RESPONSE IN SAMPLE SURVEYS

By

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Missing Data Problems in Sample Surveys

- Missing data is a common problem and challenge for analysts.
- Non-response error arises from failure to obtain the responses from the respondents (sample units) during the survey.



Causes of Missing Data Problems in Sample Surveys

• There are many reasons why data could be missing, including:



- Unavailable (No Contact)
- Unwillingness (Refusal)
- Unable (Not able)



- A sensor failed.
- Someone purposefully turned off recording equipment.
- There was a power cut.
- The method of data capture was changed.



- An internet connection was lost.
- A network went down.
- A hard drive became corrupt.
- A data transfer was cut short.

***** Methods to Handel the Non-Response Problems

 To deal with the problem of non-response following techniques are frequently being used by survey practitioners

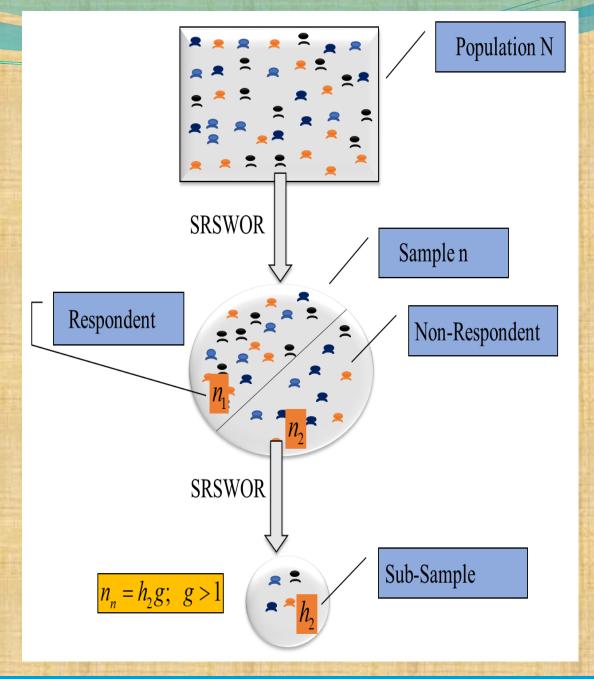
* Hansen and Hurwitz (1946) Technique

Hansen and Hurwitz (1946) suggested a sub-sampling of non-respondents technique to
deal with the problems of non-response which often occur in mail surveys. This technique
is applicable for human surveys.

The Hansen and Hurwitz (1946) technique may be summarized in the following steps:

STEP 1: Select a sample of respondents from the population and seek their responses through postal mail, email or online survey etc. by a fixed deadline.

STEP 2: Once the deadline is over, identify the non-respondents.



STEP 3: Select a sub-sample from the non-respondents and seek their responses through more engaging mode of surveying such as personal interview by trained interviewers.

STEP 4: Combine the data of both the rounds of survey to develop an estimate for the population parameter.

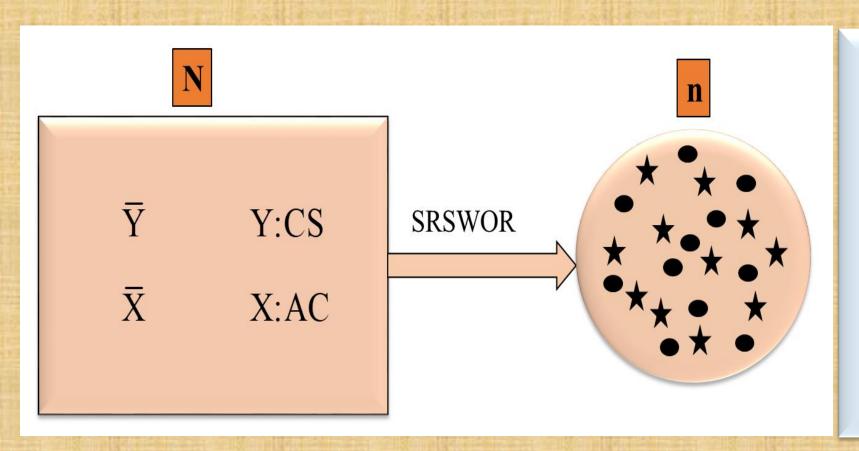
The unbiased estimator of population mean \overline{Y} is given as

$$\overline{y}_{HH} = \frac{n_1 \overline{y}_1 + n_2 \overline{y}_{h_2}}{n}$$

where \overline{y}_{h_2} denotes the mean of h_2 observations

from the sub-sample of non-responding units.

* Imputation Technique



A: Set of responding units

$$A = \{ \bigstar \star \star \star \star \star \}$$

 \overline{A} : Set of non-responding units

$$\overline{A} = \{ \bullet \bullet \bullet \bullet \bullet \}$$

$$\star\star\star\star\star$$
 r units

n Sample Size

General Imputation Method

$$\mathbf{y}_{.i} = \begin{cases} \mathbf{y}_{i} & if \quad i \in \mathbf{A} \\ \hat{\mathbf{y}}_{i} & if \quad i \in \overline{\mathbf{A}} \end{cases}$$

 \hat{y}_i denotes the imputed value for the ith non-responding unit

• The general point estimator of population mean takes the form

$$\overline{y}_s = \frac{1}{n} \left[\sum_{i=1}^r y_i + \sum_{i=1}^{n-r} \hat{y}_i \right]$$

where the value \hat{y}_i is the imputed value

Mean Method of Imputation

- In this method, no auxiliary information is used.
- The missing values are replaced with the mean of responding units of the study variable.
- Under this method the data after imputation becomes:

$$y_{.i} = \begin{cases} y_i & \text{if } i \in A \\ \\ \overline{y}_r & \text{if } i \in \overline{A} \end{cases}$$

 \overline{y}_r is the response mean

• The point estimator of the population mean is:

$$\overline{\mathbf{y}}_{m} = \frac{1}{n} \left[\sum_{i=1}^{r} \mathbf{y}_{i} + \sum_{i=1}^{n-r} \overline{\mathbf{y}}_{r} \right] = \overline{\mathbf{y}}_{r}$$

* Ratio Method of Imputation

- In the ratio method of imputation, we assume that imputation is carried out with the aid of an auxiliary variable x.
- The data after imputation becomes:

$$\mathbf{y}_{.i} = \begin{cases} \mathbf{y}_{i} & if \quad i \in \mathbf{A} \\ \hat{b} x_{i} & if \quad i \in \overline{\mathbf{A}} \end{cases}$$

where $\hat{\mathbf{b}} = \frac{\sum_{i=1}^{r} y_i}{\sum_{i=1}^{r} x_i} = \frac{\overline{y}_r}{\overline{x}_r}$

The point estimator of the population mean becomes

$$\overline{y}_{RAT} = \frac{1}{n} \left[\sum_{i=1}^{r} y_i + \sum_{i=1}^{n-r} \hat{b} x_i \right] = \frac{\overline{y}_r}{\overline{x}_r} \overline{x}_n \implies \overline{y}_{RAT} = \frac{\overline{y}_r}{\overline{x}_r} \overline{x}_n$$

$$\Rightarrow \overline{y}_{RAT} = \frac{\overline{y}_r}{\overline{x}_r} \overline{x}_n$$

* Hot Deck (HD) Method of Imputation

$$\mathbf{y}_{.i} = \begin{cases} \mathbf{y}_{i} & if \quad i \in \mathbf{A} \\ \mathbf{\overline{y}}_{g(i)} & if \quad i \in \mathbf{\overline{A}} \end{cases}$$

where $y_{g(i)}$ is the y value given by the donor unit $g(i) \in A$,

drawn at random (with replacement) from the r responding units.

• Under the HD method of imputation the point estimator of population mean

$$\overline{y}_{HD} = \frac{1}{n} \left[\sum_{i=1}^{r} y_i + \sum_{i=1}^{n-r} y_{g(i)} \right]$$

Nearest Neighbour (NN) Method of Imputation

• Under the NN method the data after imputation becomes

$$\mathbf{y}_{.i} = \begin{cases} \mathbf{y}_{i} & if \quad i \in \mathbf{A} \\ \mathbf{\overline{y}}_{g(i)} & if \quad i \in \mathbf{\overline{A}} \end{cases}$$

where $y_{g(i)}$ is the y value given by the donor unit g(i) such that $\min_{g \in R_p} |x_g - x_i|$ occors for g = g(i).

If it results in more than one unit a donor is randomly selected from them.

• Under the NN method of imputation the point estimator of the population mean becomes $\begin{bmatrix} 1 \\ \hline \end{bmatrix}$

$$\overline{y}_{NN} = \frac{1}{n} \left[\sum_{i=1}^{r} y_i + \sum_{i=1}^{n-r} y_{g(i)} \right]$$

Regression Method of Imputation

$$\mathbf{y}_{.i} = \begin{cases} \mathbf{y}_{i} & if \quad i \in \mathbf{A} \\ \hat{\mathbf{y}}_{i} & if \quad i \in \overline{\mathbf{A}} \end{cases}$$

$$where \hat{\mathbf{y}}_{i} = \hat{a} + \hat{b}x_{i}$$

where
$$\hat{y}_i = \hat{a} + \hat{b}x_i$$

$$\hat{b} = \frac{s_{yx}(r)}{s_x^2(r)}; \ \hat{a} = \overline{y}_r - \hat{b}\overline{x}_r$$

Under the Regression method of Imputation the point estimator of the population mean

$$\overline{y}_{\text{Re }g} = \frac{1}{n} \left[\sum_{i=1}^{r} y_i + \sum_{i=1}^{n-r} (\hat{a} + \hat{b}x_i) \right] \Rightarrow \overline{y}_{reg} = \overline{y}_r + \hat{b}(\overline{x}_n - \overline{x}_r)$$

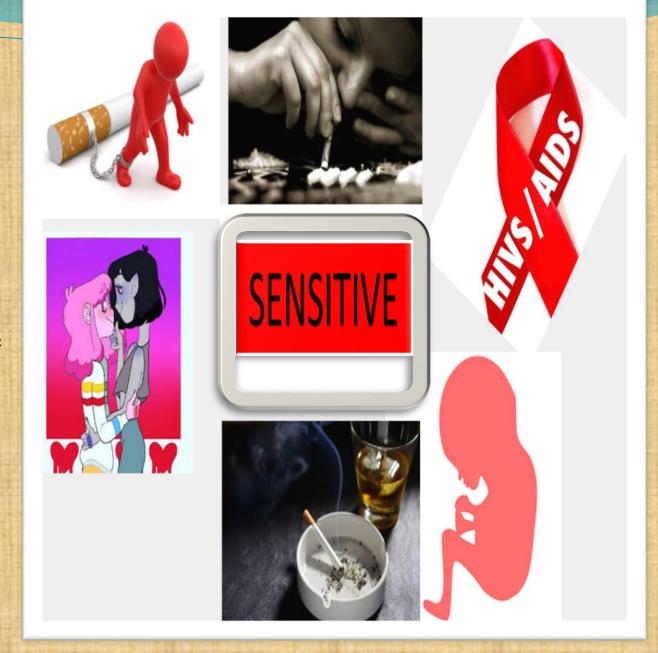
$$\Rightarrow \overline{y}_{reg} = \overline{y}_r + \hat{b}(\overline{x}_n - \overline{x}_r)$$

Dealing with non-response arise due to sensitive issues

- Economists, psychologists, sociologists, managers, and policy makers have many reasons for asking personal questions.
- Biometric sample surveys need get readymade information for future planning and policy implementations related to the subject matters of highly sensitive issues.



- Highly sensitive issues such as sexual behavior, domestic violence, tax offender, HIV infection status, drug addiction, extra marital affair etc.
- Actual answers of these questions are hidden or misguided by people.
- Data obtained are definitely open to error if surveys are conducted through classical methods.



* Randomized Response Technique

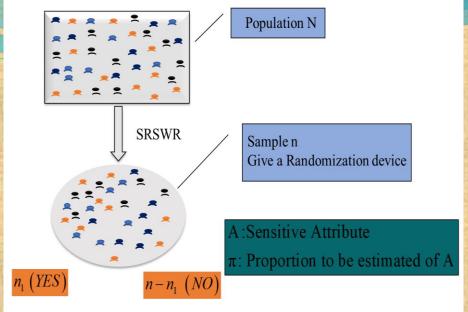
- Introduced by Warner (1965)
- When study under characteristics is sensitive in nature.
- Use randomization device to acquire the truthful response from respondents.
- To estimate the Proportion of the population possessing sensitive characteristic.

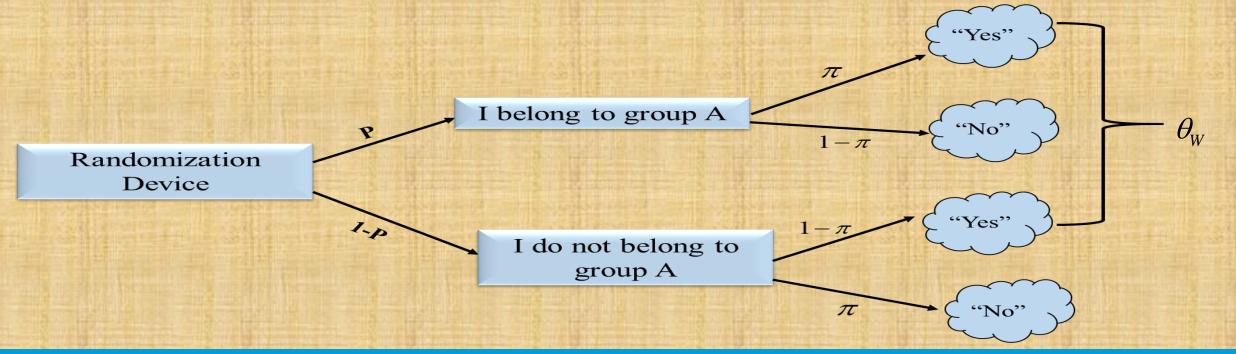


Gradual Development: RRT

* Warner (1965) Technique

• Each respondent is provided an identical randomization device as:





• With the help of a randomized device, the respondent replies only "Yes" or "No" answers in a random sample of n respondents. The probability of "Yes" answer:

$$\theta_{W} = P\pi + (1 - P)(1 - \pi)$$

The unbiased estimator of population proportion

$$\hat{\pi}_W = \frac{\hat{\theta}_W - (1 - P)}{2P - 1} \qquad P \neq 0.5$$

$$\hat{\theta}_{W} = \frac{n_{1}}{n}$$

where $\hat{\theta}_W$ is the observed proportion of "Yes" answer in the sample of n units drawn by the SRSWR sampling.

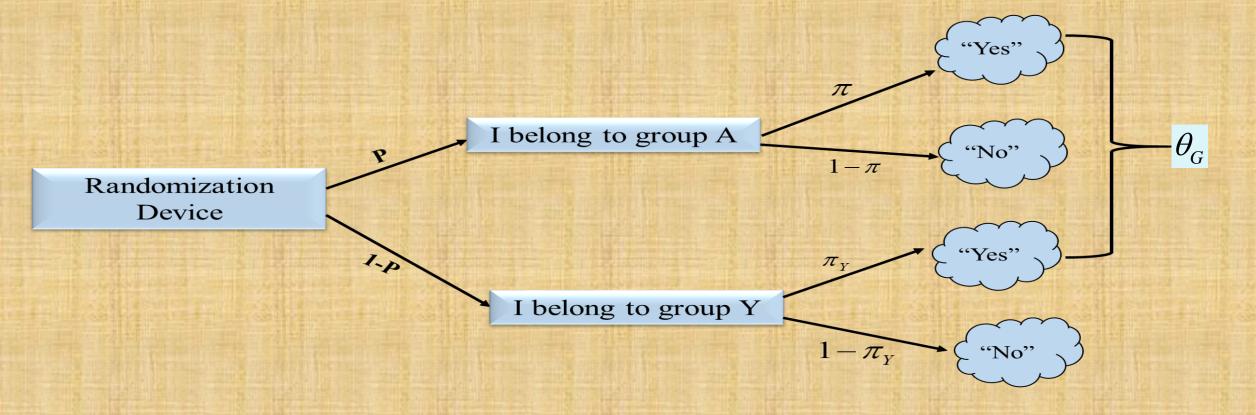
• The Variance is:

$$V(\hat{\pi}_W) = \frac{\pi(1-\pi)}{n} + \frac{P(1-P)}{n(2P-1)^2}$$

Greenberg et al. (1969) Technique



• Each respondent is provided an identical randomized device as:



- For example, in estimating the proportion of persons having the extra marital relations in a certain community the two questions may be:
- i. Are you having extra marital relations?
- ii. Did you born in the month of March?

* When the Proportion of unrelated character is known

• With the help of a randomized device, the respondent replies only "Yes" or "No" answers in a random sample of n respondents. The probability of "Yes" answer in the population: $\theta_G = P\pi + (1-P)\pi_{_Y}$

Let n_1 be the number of observed "Yes" answer in the sample of n units. So that $\hat{\theta}_G = \frac{n_1}{n}$

• The unbiased estimator of π is:

$$\hat{\pi}_G = \frac{\hat{\theta}_G - (1 - P)\pi_Y}{P}$$

• The variance of the estimator $\hat{\pi}_G$ is:

$$V(\hat{\pi}_G) = \frac{\theta_G - (1 - \theta_G)}{n P^2}$$

Population: $(N < \infty)$; A:Sensitive Attribute

Y: Non-Sensitive Unrelated Attribute

 π : Proportion to be estimated of A

 π_{Y} : Proportion of Non-Sensitive unrelated attribute Y

When the Proportion of unrelated character is unknown

Here π_Y the proportion of unrelated character Y in the population is unknown

• In this case the probability of "Yes" answer:

Probability of "yes" answer using device 1

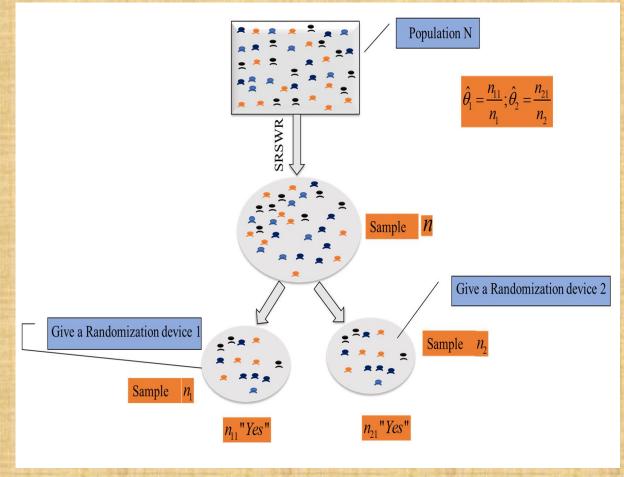
$$\theta_1 = P_1 \pi + \left(1 - P_1\right) \pi_Y$$

Probability of "yes" answer using device 2

$$\theta_2 = P_2 \pi + \left(1 - P_2\right) \pi_Y$$

• The unbiased estimator of π is:

$$\hat{\pi}_{G} = \frac{(1 - P_{2})\hat{\theta}_{1} - (1 - P_{1})\hat{\theta}_{2}}{P_{1} - P_{2}}$$



where $\hat{\theta}_1$ and $\hat{\theta}_2$ are the observed proportion of "Yes" answer in the first and second sample respectively.

• The variance of the estimator $\hat{\pi}_G$ is:

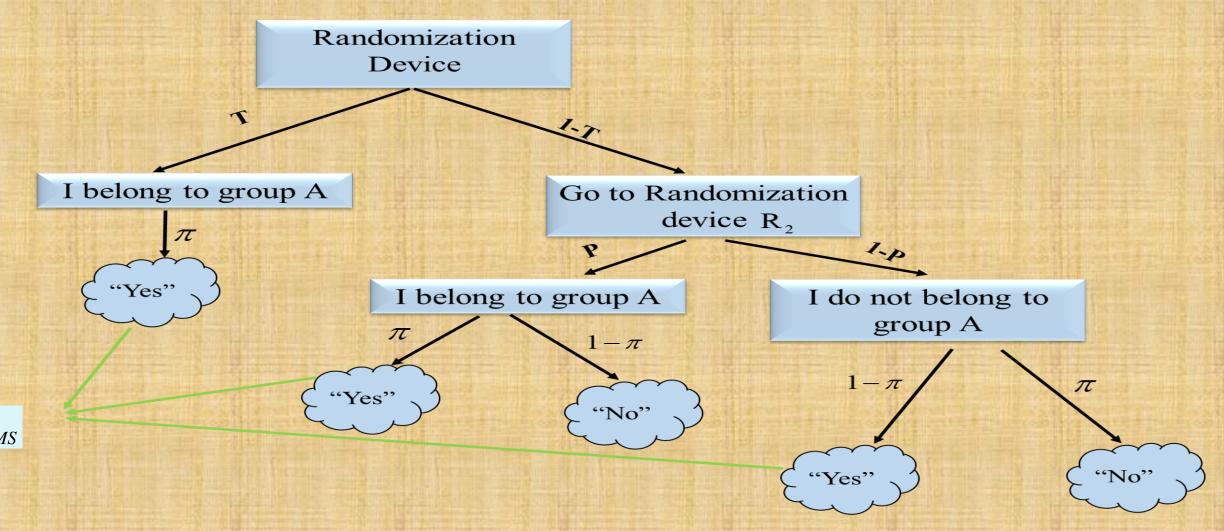
$$V(\hat{\pi}_G) = \frac{1}{(P_1 - P_2)^2} \left[\frac{(1 - P_2)^2 \theta_1 (1 - \theta_1)}{n_1} + \frac{(1 - P_1)^2 \theta_2 (1 - \theta_2)}{n_2} \right]$$

For the best choice of n_1 and n_2 , it should follow the relation

$$\frac{n_1}{n_2} = \sqrt{\frac{\theta_1 (1 - \theta_1)}{\theta_2 (1 - \theta_2)}} \frac{(1 - P_2)}{(1 - P_1)}$$

Mangat and Singh (1990) Technique

• Each respondent is provided an identical randomized device



• With the help of a randomized device, the respondent replies only "Yes" or "No" answers in a random sample of n respondents. The probability of "Yes" answer:

$$\theta_{MS} = T\pi + (1-T)\{P\pi + (1-P)(1-\pi)\}$$

$$\hat{\theta}_{MS} = \frac{n_1}{n}$$

• The unbiased estimator of π is:

$$\hat{\pi}_{MS} = \frac{\hat{\theta}_{MS} - (1 - T)(1 - P)}{2P - 1 + 2T(1 - P)}$$



$$V(\hat{\pi}_{MS}) = \frac{\pi(1-\pi)}{n} + \frac{(1-T)(1-P)\{1-(1-T)(1-P)\}}{n\{2P-1+2T(1-P)\}^2}$$

Population: $(N < \infty)$

A:Sensitive Attribute

 π : Proportion to be estimated of A

