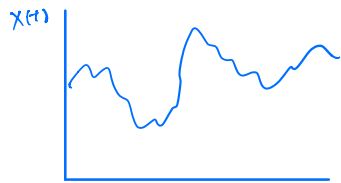


e.g. track position of a particle



$$\Omega = C[0, \infty)$$

infinite-dimensional sample space

$$P(A) \quad A \in C[0, \infty)$$

\Rightarrow NOT calculus-based probability

Random field

$$X(x, y)$$

$$X(x, y, z)$$

$$X(x, y, z, t)$$

Problem of Points

2 gamblers, \$1 pot

1st person wins 5 games get \$1

(4,3) \Rightarrow interrupted \Rightarrow how to split?

{ expectation
law of large numbers

Repeat the game n times

$$\Rightarrow \sum_{i=1}^n \mathbb{I}(X_i = 1) \geq n\alpha \quad \leftarrow \begin{array}{l} \text{A wins} \\ \text{A pays } \alpha \text{ proportion of the pot} \end{array}$$

$$\sum_{i=1}^n \mathbb{I}(X_i = 2) \geq n(1-\alpha)$$

$$\frac{1}{n} \sum_{i=1}^n \mathbb{I}(X_i = 1) \xrightarrow{LLN} E\mathbb{I}(X_i = 1) = P(X_i = 1) = p$$

$$\Rightarrow p = P(\text{A wins} \mid (4,3))$$

$$= \frac{1}{2} P(5,3) + \frac{1}{2} P(4,4)$$

$$= \frac{1}{2} + \frac{1}{2} \times \frac{1}{2} = \frac{3}{4}$$

Law of Large Numbers (LLN)

$$X_1, X_2, \dots \text{ i.i.d. } E|X_1| < \infty$$

$$\frac{1}{n} \sum_{i=1}^n X_i \rightarrow EX_1$$

Proof WLLN when $\text{var } X_1 < \infty$:

$$X_i : \Omega \rightarrow \mathbb{R}$$

$X_i(\omega)$ r.v.'s are functions

different ways of convergence for deterministic functions	{	$f_n(x) \rightarrow f_\infty(x)$	"pointwise"
		$\sup_{x \in \mathbb{R}} f_n(x) - f_\infty(x) \rightarrow 0$	"uniform"
		$\int_{-\infty}^{\infty} f_n(x) - f_\infty(x) ^p dx \rightarrow 0$	"L ^p "

Convergence in Probability:

$$X_n \xrightarrow{P} X_\infty$$

iff

$$\forall \varepsilon > 0,$$

$$P(|X_n - X_\infty| > \varepsilon) \rightarrow 0$$

as $n \rightarrow \infty$

WLLN:

$$\text{want to prove } P(|\bar{X}_n - EX_1| > \varepsilon) \rightarrow 0$$

To do this, let's first derive some inequalities.

Markov Inequality

$$\cdot W \geq 0 ; \quad EW < \infty$$

$$\cdot P(W > w) \leq \frac{EW}{w}$$

Proof:

$$P(W > w) = E \mathbb{I}(W > w)$$

$$\text{On } \{W > w\}, \quad \frac{W}{w} \geq 1$$

$$\begin{aligned} \Rightarrow E \mathbb{I}(W > w) &\leq E \frac{W}{w} \mathbb{I}(W > w) \\ &\leq \frac{EW}{w} \end{aligned}$$

Application 1. Chebyshev's inequality

$$\cdot X ; \quad \text{Var } X < \infty$$

$$\cdot P(|X - EX| > x) \leq \frac{\text{Var } X}{x^2}$$

Proof:

$$P(|X - EX| > x) = P(\underbrace{|X - EX|^2}_{w}, \text{ Markov's inequality} > x^2)$$

Application 2. Exponential inequality

"Light-tail r.v."

$$\cdot X ; \quad E \exp(\theta X) < \infty ; \quad \theta > 0$$

$$\cdot P(X > x) \leq \frac{E \exp(\theta X)}{e^{\theta x}} = e^{-\theta x} E \exp(\theta X)$$

↑
exponential decay in right tail

Proof:

$$P(X > x) = P(\underbrace{e^{\theta X}}_{w}, \text{ Markov's inequality} > e^{\theta x})$$

X_1, X_2, \dots i.i.d., $\text{var } X_i < \infty$

$$P(|\bar{X}_n - EX| > \varepsilon) \leq \frac{1}{\varepsilon^2} \frac{\text{var } X_1}{n} \rightarrow 0$$

↑ Chebyshev's

$$\Rightarrow \text{WLLN: } \bar{X}_n \xrightarrow{P} EX$$

X_1, X_2, \dots stationary, $\text{var } X_i < \infty$

$$(X_1, X_2, \dots) \stackrel{D}{=} (X_1, X_{1+2}, \dots)$$

$$X_i \stackrel{D}{=} X_j \quad \text{time-invariant}$$

$$(X_i, X_{i+j}) \stackrel{D}{=} (X_1, X_{1+j})$$

e.g. ① X_1, X_2, \dots i.i.d.

② $X_i = X_1 \quad \frac{1}{n} \sum_{i=1}^n X_i = X_1 \not\xrightarrow{P} EX_1 \quad (\text{perfect corr})$

$$\Rightarrow \frac{1}{n} \sum_{i=1}^n X_i \xrightarrow{P} EX_1 \quad \text{holds}$$

$$\text{if } \text{cov}(X_i, X_n) \rightarrow 0 \quad \text{as } n \rightarrow \infty$$

Proof:

$$\text{Var } \bar{X}_n \rightarrow 0 \quad ?$$

$$\frac{1}{n^2} \text{Var} \left(\sum_{i=1}^n X_i \right) = \frac{1}{n^2} \sum_{i=1}^n \sum_{j=1}^n \text{Cov}(X_i, X_j) \quad \leftarrow \text{cc}(i-j) ?$$

$$= \frac{1}{n^2} \left[\underbrace{n \text{Var } X_1}_{\rightarrow 0} + \underbrace{2 \sum_{i=1}^{n-1} (n-i) \text{cc}(i)}_{?} \right] \quad \leftarrow \text{Covariance lag } i \text{ } \text{Cov}(X_1, X_{1+i})$$

$$\Rightarrow \frac{1}{n^2} \sum_{i=1}^{n-1} (n-i) |\text{cc}(i)|$$

$$= \frac{1}{n} \sum_{i=1}^{n-1} \underbrace{\frac{n-i}{n}}_{\leq 1} |\text{cc}(i)|$$

$$\leq \frac{1}{n} \sum_{i=1}^{n-1} |\text{cc}(i)| \rightarrow 0 \quad \text{by assumption}$$