

MAP 531 : Statistics refresher

21 septembre 2016

1 Decision theory

- Example
- Bias

2 Minimax and Bayes estimators

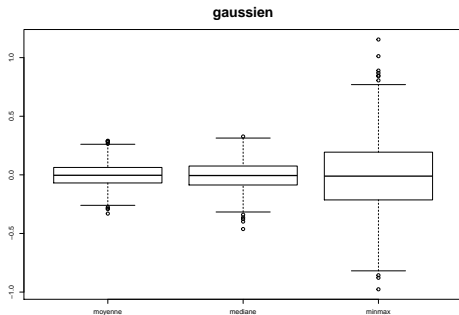
Estimation of the translation parameter

- Let (X_1, X_2, \dots, X_n) be an i.i.d. n -sample with probability density function given by $p_\theta(x) = q(x - \theta)$, $\theta \in \Theta = \mathbb{R}$ where q is symmetric : $q(x) = q(-x)$
- Estimator :
 - 1 $\hat{\theta}_n^{(1)}(X_1, \dots, X_n) = n^{-1} \sum_{i=1}^n X_i$ Empirical mean,
 - 2 $\hat{\theta}_n^{(2)}(X_1, \dots, X_n) = \text{median}(X_1, \dots, X_n)$, median.
 - 3 $\hat{\theta}_n^{(2)}(X_1, \dots, X_n) = 0.5 (\min(X_1, \dots, X_n) + \max(X_1, \dots, X_n))$.

Gaussian noise

$$q(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}x^2\right)$$

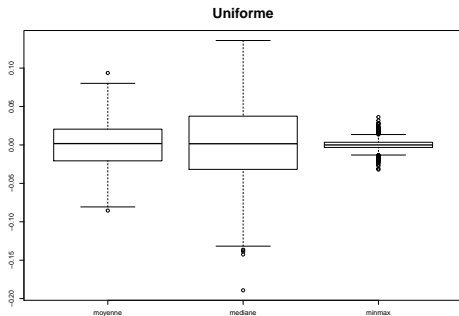
Sample : $n = 100$. Number
of sets : $m = 1000$.



Uniforme noise

$$q(x) = \mathbb{1}_{[-1/2, 1, 2]}(x)$$

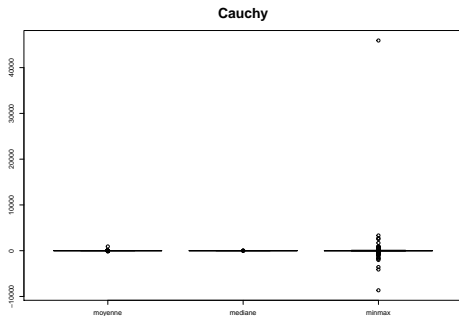
Sample : $n = 100$. Number
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Cauchy noise

$$q(x) = \frac{1}{1+x^2}$$

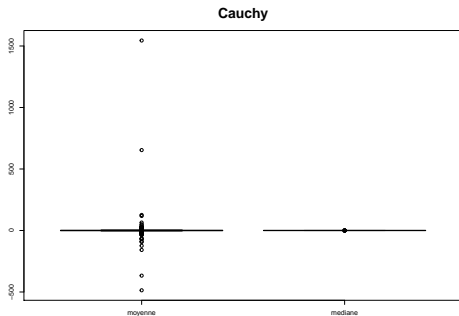
Sample : $n = 100$. Number
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Cauchy noise (zoom)

$$q(x) = \frac{1}{1+x^2}$$

Sample : $n = 100$. Number
of sets : $m = 1000$.



Estimator of the variance in a translation-scaling model

Let X_1, X_2, \dots, X_n be real random variables i.i.d. with p.d.f.

$$p_{\theta}(x) = \frac{1}{\sigma} q\left(\frac{x - \mu}{\sigma}\right), \quad \theta = (\mu, \sigma^2) \in \mathbb{R} \times \mathbb{R}_+^*$$

where $\int x^2 q(x) dx = 1$ and $\int x q(x) dx = 0$.

Estimator of the variance in a translation-scaling model

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where $\int x^2 q(x) dx = 1$ and $\int x q(x) dx = 0$.

$$\begin{aligned} \mathbb{E}_\theta \left[\sum_{i=1}^n (X_i - \bar{X}_n)^2 \right] &= \mathbb{E}_\theta \left[\sum_{i=1}^n (X_i - \mu - (\bar{X}_n - \mu))^2 \right] \\ &= \mathbb{E}_\theta \left[\sum_{i=1}^n (X_i - \mu)^2 - n(\bar{X}_n - \mu)^2 \right] \\ &= n\sigma^2 - \sigma^2 = (n-1)\sigma^2. \end{aligned}$$

Estimator of the variance in a translation-scaling model

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$$p_\theta(x) = \frac{1}{\sigma} q\left(\frac{x - \mu}{\sigma}\right), \quad \theta = (\mu, \sigma^2) \in \mathbb{R} \times \mathbb{R}_+^*$$

where $\int x^2 q(x) dx = 1$ and $\int x q(x) dx = 0$.

$$\mathbb{E}_\theta \left[\sum_{i=1}^n (X_i - \bar{X}_n)^2 \right] = (n-1)\sigma^2$$

Therefore

$$S_n^2 = (n-1)^{-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2$$

is an unbiased estimator of σ^2

Unbiased estimation

In the scaling estimation issue, it is easy to build an unbiased estimator

$$\begin{aligned}\mathbb{E}_\theta[(X_1 - \bar{X}_n)^2] &= \mathbb{E}_\theta[(X_1 - \mu - (\bar{X}_n - \mu))^2] \\&= \mathbb{E}_\theta[(X_1 - \mu)^2] - 2\mathbb{E}_\theta[(X_1 - \mu)(\bar{X}_n - \mu)] + \mathbb{E}_\theta[(\bar{X}_n - \mu)^2] \\&= \mathbb{E}_\theta[(X_1 - \mu)^2] - \frac{2}{n}\mathbb{E}_\theta[(X_1 - \mu)^2] + \mathbb{E}_\theta[(\bar{X}_n - \mu)^2] \\&= \sigma^2 - 2\frac{\sigma^2}{n} + \frac{\sigma^2}{n} \\&= \frac{n-1}{n}\sigma^2.\end{aligned}$$

Unbiased estimation

$$\mathbb{E}_\theta[(X_1 - \bar{X}_n)^2] = \frac{n-1}{n} \sigma^2$$

Therefore

$$\frac{n}{n-1} (X_1 - \bar{X}_n)^2$$

is an **unbiased** estimator of σ^2

Unbiased estimation

Unbiased estimator

$$\frac{n}{n-1}(X_1 - \bar{X}_n)^2$$

Biased estimator

$$\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2$$

Which one is the best ?

Unbiased estimation

A biased estimator is not necessarily a “bad” estimator. Being unbiased means : the errors are symmetric around the true value.

Estimation of a uniform distribution support

Let X_1, X_2, \dots, X_n be independent random variables following $\text{Unif}([0, \theta])$, where $\theta \in \Theta = \mathbb{R}_+^*$. We denote $X_{n:n} = \max(X_1, \dots, X_n)$. For all $\theta \in \Theta$ and $x \in [0, \theta]$, we have

$$\mathbb{P}_\theta(X_{n:n} \leq x) = \mathbb{P}_\theta(\max(X_1, \dots, X_n) \leq x) = \prod_{i=1}^n \mathbb{P}_\theta(X_i \leq x) = (x/\theta)^n.$$

The p.d.f. of $X_{n:n}$ is therefore given by

$$n \frac{x^{n-1}}{\theta^n}$$

for $x \in [0, \theta]$. Thus :

$$\begin{aligned}\mathbb{E}_\theta[X_{n:n}] &= \int_0^\theta x \cdot n \frac{x^{n-1}}{\theta^n} dx = \frac{n}{n+1} \frac{\theta^{n+1}}{\theta^n} = \frac{n}{n+1} \theta, \\ \mathbb{E}_\theta[X_{n:n}^2] &= \int_0^\theta x^2 \cdot n \frac{x^{n-1}}{\theta^n} dx = \frac{n}{n+2} \frac{\theta^{n+2}}{\theta^n} = \frac{n}{n+2} \theta^2.\end{aligned}$$

Estimation of a uniform distribution support

Let X_1, X_2, \dots, X_n be independent random variables following $\text{Unif}([0, \theta])$, where $\theta \in \Theta = \mathbb{R}_+^*$. We denote $X_{n:n} = \max(X_1, \dots, X_n)$.

$$\mathbb{E}_\theta[X_{n:n}] = \frac{n}{n+1}\theta, \quad \mathbb{E}_\theta[X_{n:n}^2] = \frac{n}{n+2}\theta^2.$$

The estimator $(n+1)/n X_{n:n}$ is an **unbiased** estimator of θ .
The quadratic risk of $a_n X_{n:n}$ is

$$\begin{aligned} \mathbb{E}_\theta[(a_n X_{n:n} - \theta)^2] &= a_n^2 \mathbb{E}_\theta[X_{n:n}^2] - 2a_n \theta \mathbb{E}_\theta[X_{n:n}] + \theta^2 \\ &= \frac{na_n^2}{n+2}\theta^2 - \frac{2a_n n}{n+1}\theta^2 + \theta^2 = \theta^2 \left\{ \frac{na_n^2}{n+2} - \frac{2a_n n}{n+1} + 1 \right\} \end{aligned}$$

The minimum is reached for $a_n = (n+2)/(n+1)$ and equals

$$\mathbb{E}_\theta \left[\left(\frac{n+2}{n+1} X_{n:n} - \theta \right)^2 \right] = \frac{\theta^2}{(n+1)^2}$$

Estimation of a uniform distribution support

Let X_1, X_2, \dots, X_n be independent random variables following $\text{Unif}([0, \theta])$, where $\theta \in \Theta = \mathbb{R}_+^*$. We denote $X_{n:n} = \max(X_1, \dots, X_n)$. The estimator $(n+1)/n X_{n:n}$ is **unbiased** but for all $\theta \in \Theta$

$$\mathbb{E}_\theta \left[\left(\frac{n+2}{n+1} X_{n:n} - \theta \right)^2 \right] \leq \mathbb{E}_\theta \left[\left(\frac{n+1}{n} X_{n:n} - \theta \right)^2 \right]$$

We say that $(n+1)/n X_{n:n}$ is **inadmissible** for the **quadratic risk**.

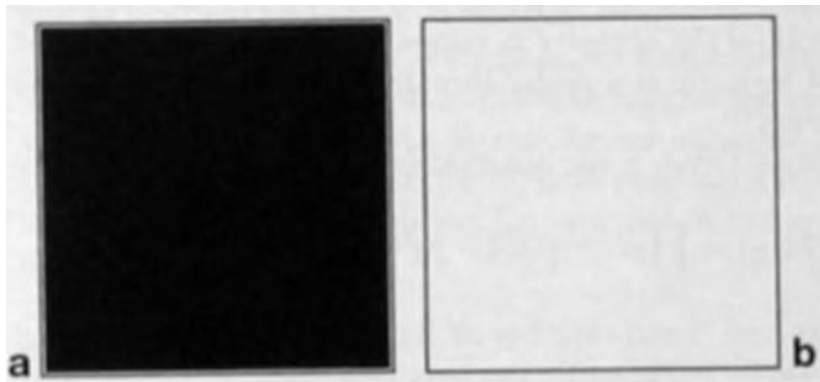
1 Decision theory

2 Minimax and Bayes estimators

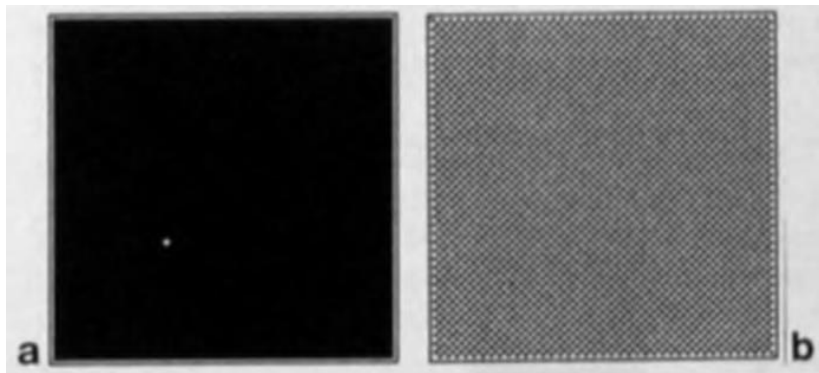
Observation



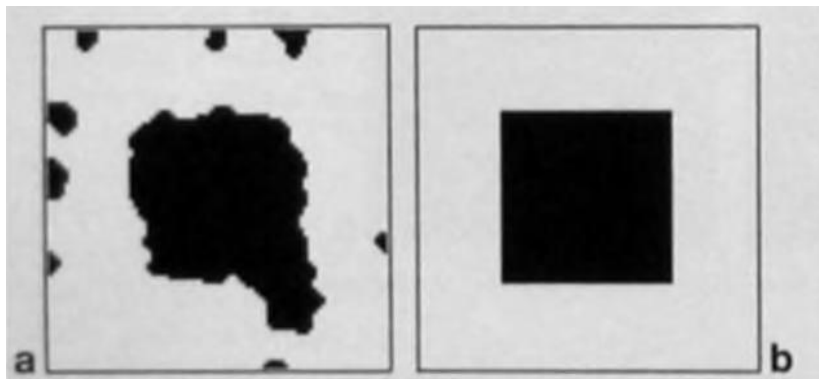
Solution triviale avec uniquement l'a priori



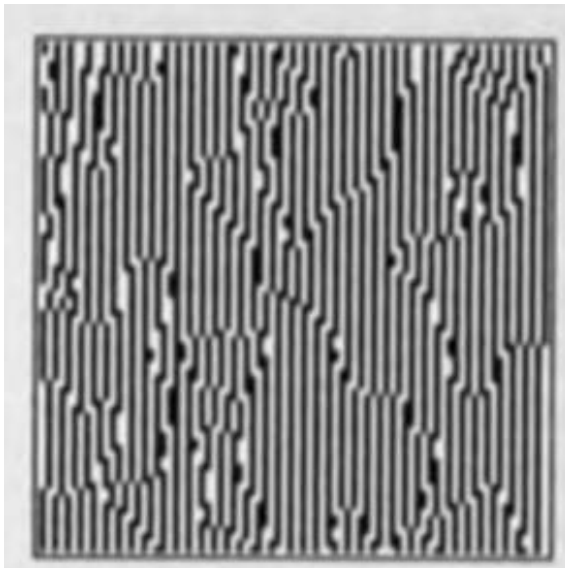
Images ne respectant pas les contraintes



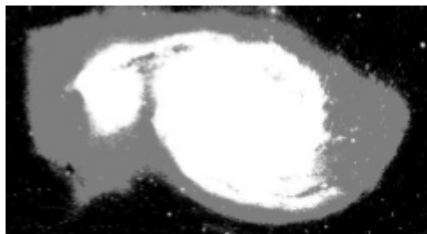
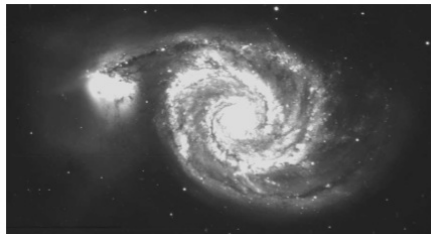
Une solution et l'image réelle



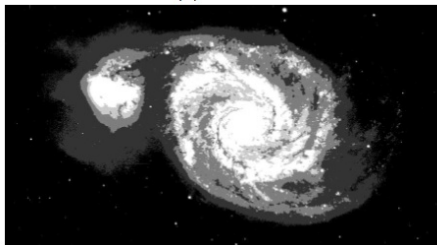
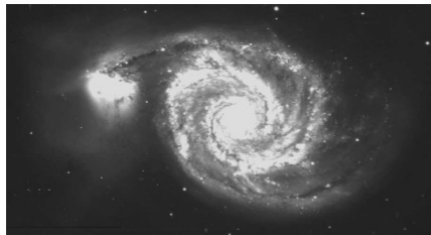
En cas de mauvais apriori



Autre exemples en imagerie : Segmentation utilisant des HMM



Segmentation utilisant des HMM



Segmentation



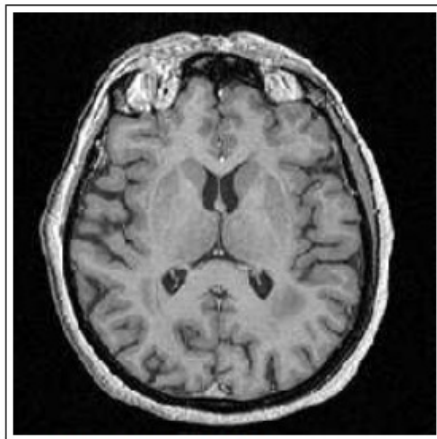
a. Image radar originale ©ERS-1

Segmentation

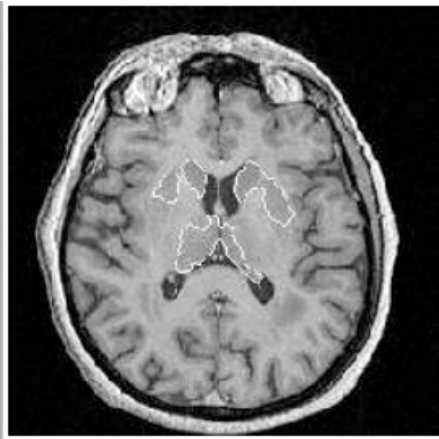


h Image segmentée en régions

Segmentation



a. Image cérébrale IRM.



Résultat de la segmentation