$s_1$	$s_2$	$ar{X}$	S
0	0	0	0
0	1	0.5	0.7071
0	2	1	1.4142
1	0	0.5	0.7071
1	1	1	0
1	2	1.5	0.7071
2	0	1	1.4142
2	1	1.5	0.7071
2	2	2	0

# 1.1)

$\bar{X}$	Probability				
0	1/9				
0.5	2/9				
1	3/9				
1.5	2/9				
2	1/9				

## 1.2)

S	Probability	
0	3/9	
$\sqrt{0.5}=0.7071$	4/9	
$\sqrt{2} = 1.4142$	2/9	

calculations were performed using computer <sup>1</sup>

```
const means = {}
const stdevs = {}
let samples = []
for (let i of [0, 1, 2]) {
 for (let j of [0, 1, 2]) {
   let sample = [i, j]
   sampleMean = mean(sample)
   sampleStdev = stdev(sample)
   means[sampleMean] = 1 + (means[sampleMean] ?? 0)
   stdevs[sampleStdev] = 1 + (stdevs[sampleStdev] ?? 0)
    samples = [...samples, {i, j, sampleMean, sampleStdev}]
 }
}
console.table(samples)
console.table(
  Object.keys(means).sort()
    .reduce((acc, x) => [
      ...acc, { x, probabilty: '${means[x]}/${samples.length}' }], []))
console.table(
  Object.keys(stdevs).sort()
    .reduce((acc, x) => [
      ...acc, { x, probabilty: '${stdevs[x]}/${samples.length}' }], []))
```

 $<sup>^1 {\</sup>rm JavaScript}$  with ECMAScript 2020 (ES11) features.

Exercise 14a on page 305.

The approximate probability that the professor has finished grading before the 11:00pm news is 0.6026.

The time the professor spent grading is 250 minutes.

$$6:50 + 10mins = 7:00$$
 $11:00 - 7:00 = 4hours$ 
 $4hx \cdot 60 \frac{mins}{hour} = 240mins$ 
 $10mins + 240mins = 250mins$ 

According to Central Limit Theorum

$$\lim_{n \to \infty} P\left(\frac{T_0 - n\mu}{\sqrt{n}\sigma} \leqslant z\right) = \Phi(z)$$

n is large enough, according to rule of thumb

$$n = 40 > 30$$

The values from the problem into the formula

$$\frac{T_0-n\mu}{\sqrt{n}\mu}=\frac{250mins-40\cdot 6mins}{\sqrt{40}\cdot 6mins}=0.2635$$

The value from table A.3 on page 789 the textbook is used

$$\Phi(0.2635) = 0.6026$$

Exercise 18ac on page 305.

 $\mathbf{a}$ 

The approximate probability that the amount purchased is at least 12 gallons is 0.8106

$$\lim_{n\to\infty} P\left(\frac{\bar{X}-\mu}{\sigma/\sqrt{n}}\leqslant z\right) = \Phi(z)$$
 
$$P\left(\frac{\bar{X}-\mu}{\sigma/\sqrt{n}}\leqslant z\right) = P\left(\frac{12-11.5}{4/\sqrt{50}}\leqslant z\right)P(0.8838\leqslant z) = \Phi(z)$$
 
$$\Phi(0.8838) = 0.8106$$

**c**)

The 95th percentile for the total amount purchased by 50 randomly selected customers is approximately 622 gallons.

$$\frac{T_0 - n\mu}{\sqrt{n}\sigma} \leqslant z$$

$$T_0 \leqslant z\sqrt{n}\sigma + n\mu$$

According to table A.3 in the textbook:  $\Phi(.95) \approx 1.65$ 

$$T_0 \le 1.65\sqrt{50} \cdot 4 + 50 \cdot 11.5 = 621.6690$$

Exercise 1ab on page 346, but use this subset of data:

 $\mathbf{a}$ 

For this point estimate,  $\bar{X}$  is used:

$$\hat{\mu} = \bar{X} = \frac{X_1 + \dots + X_n}{n} = 113.7$$

The estimated standard error is the sample standard deviation divided by  $\sqrt{n}$ 

$$\hat{\sigma}_{\bar{X}} = \frac{s}{\sqrt{n}} = 13.0$$

```
const IQ = [
   82, 96, 102, 103, 106, 107, 108, 108, 108, 109, 110, 110, 111,
   113, 113, 113, 115, 115, 118, 119, 121, 122, 127, 136, 140, 146,
]
const xBar = mean(IQ)
const stdErr = stdev(IQ) / Math.sqrt(IQ.length)
console.table([{ xBar, stdErr }])
```

b)

The estimate of the value that separates the lowest 50% from the highest 50%, is the sample median.

$$\hat{m} = \tilde{X} = \frac{113 - 111}{2} = 112$$

The estimated standard error determined by bootstrapping method was 2.5

```
const boostrapSample = sample => {
   return sample.map(() => {
      const r = Math.floor(Math.random() * sample.length)
      return sample[r]
   })
}

const B = 200;
const bootstrapMedians = []
for (let i = 0; i < B; i++) {
   bootstrapMedians.push(median(boostrapSample(IQ)))
}

const SB = Math.sqrt((1/(B - 1) * variance(bootstrapMedians)))
console.log(SB)</pre>
```

Exercise 8a on page 348.

 $\mathbf{a}$ 

The estimate of p, the proportion which are not defective

$$\hat{p} = 1 - \frac{X}{n} = 1 - \frac{12}{80} = 0.85$$

The estimate standard error

$$\hat{\sigma}_{\hat{p}} = \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} = \sqrt{\frac{0.85 \cdot 0.15}{80}} = 0.0399$$

Exercise 12 on page 348.

For an unbiased estimator

$$E(\hat{\theta}) - \theta = 0 \implies \theta = E(\hat{\theta})$$

For the estimator in the problem

$$E(\hat{\sigma^2}) = E\left(\frac{(n_1 - 1)S_1^2 + (n_2 - 1)S_2^2}{n_1 + n_2 - 2}\right)$$

$$= \frac{1}{n_1 + n_2 - 2} E\left((n_1 - 1)S_1^2 + (n_2 - 1)S_2^2\right)$$

$$= \frac{1}{n_1 + n_2 - 2} \left(E((n_1 - 1)S_1^2) + E((n_2 - 1)S_2^2)\right)$$

$$= \frac{(n_1 - 1)E(S_1^2) + (n_2 - 1)E(S_2^2)}{n_1 + n_2 - 2}$$

In example 7.6 of the textbook, it is shown that

$$E(S^2) = \sigma^2$$

In the problem it states both types of fertilizer have the same variance

$$E(S_1^2) = E(S_2^2) = \sigma^2$$

Substitute into above to show the estimate is unbiased

$$E(\hat{\sigma^2}) = \frac{(n_1 - 1)\sigma^2 + (n_2 - 1)\sigma^2}{n_1 + n_2 - 2} = \frac{(n_1 - 1 + n_2 - 1)\sigma^2}{n_1 + n_2 - 2} = \frac{(n_1 + n_2 - 2)\sigma^2}{n_1 + n_2 - 2} = \sigma^2$$