

生物统计学： 生物信息中的概率统计模型

2024年秋



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有关信息

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2024年生物统计学
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课程安排

- 生物背景和课程简介
- 传统生物统计学及其应用
- 生物统计学和生物大数据挖掘
 - Hidden Markov Model (HMM)及其应用
 - Markov Chain
 - HMM理论
 - HMM和基因识别 (Topic I)
 - HMM和序列比对 (Topic II)
 - 进化树的概率模型 (Topic III)
 - Motif finding中的概率模型 (Topic IV)
 - EM algorithm
 - Markov Chain Monte Carlo (MCMC)
 - 基因表达数据分析 (Topic V)
 - 聚类分析-Mixture model
 - Classification-Lasso Based variable selection
 - 基因网络推断 (Topic VI)
 - Bayesian网络
 - Gaussian Graphical Model
 - 基因网络分析 (Topic VII)
 - Network clustering
 - Network Motif
 - Markov random field (MRF)
 - Dimension reduction及其应用 (Topic VIII)
- 面向生物大数据挖掘的深度学习

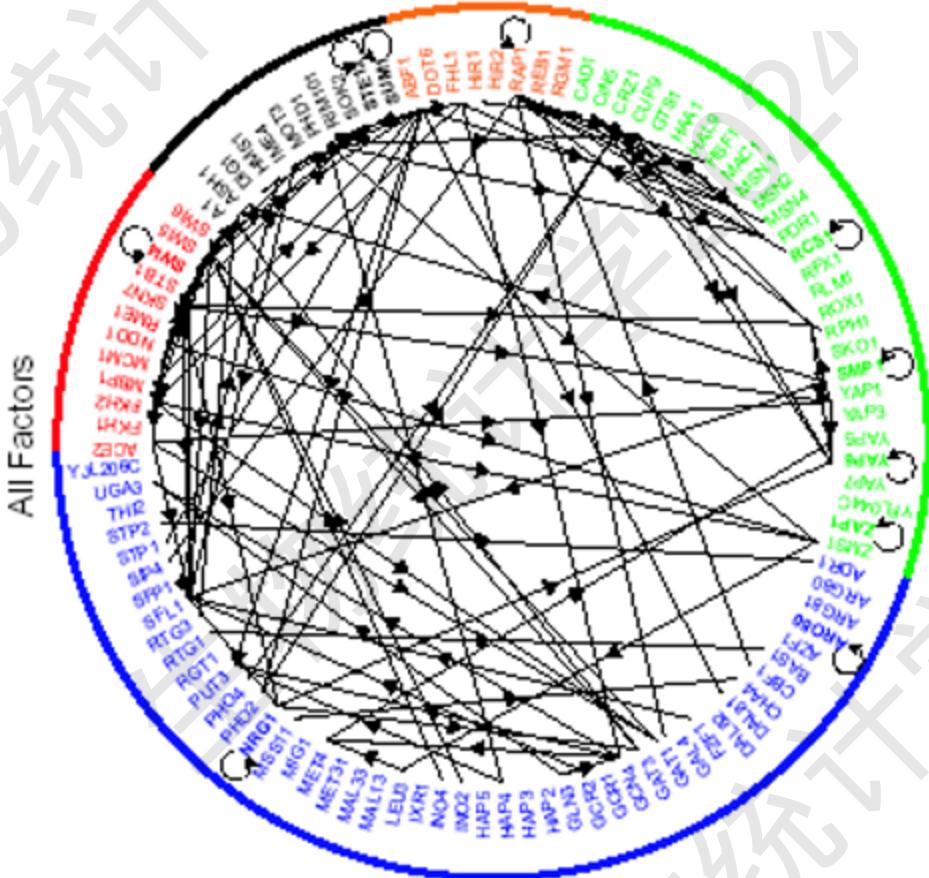
研究对象：
生物序列，
进化树，
生物网络，
基因表达
...

方法：
生物计算与生物统计

第7章: Regulatory Network

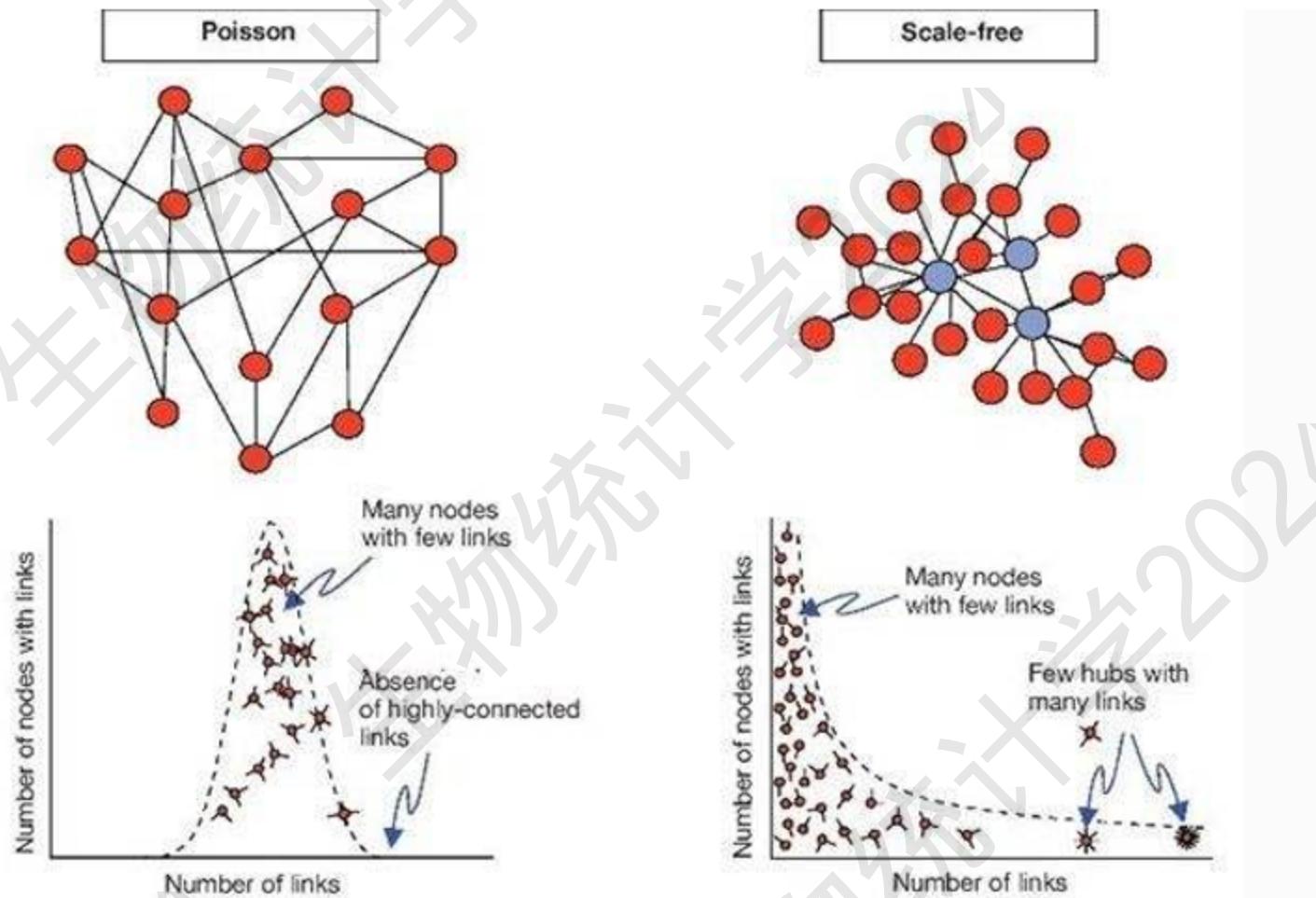
- Regulatory network
- Reverse engineering
- Bayesian network

Part I: Regulatory Network



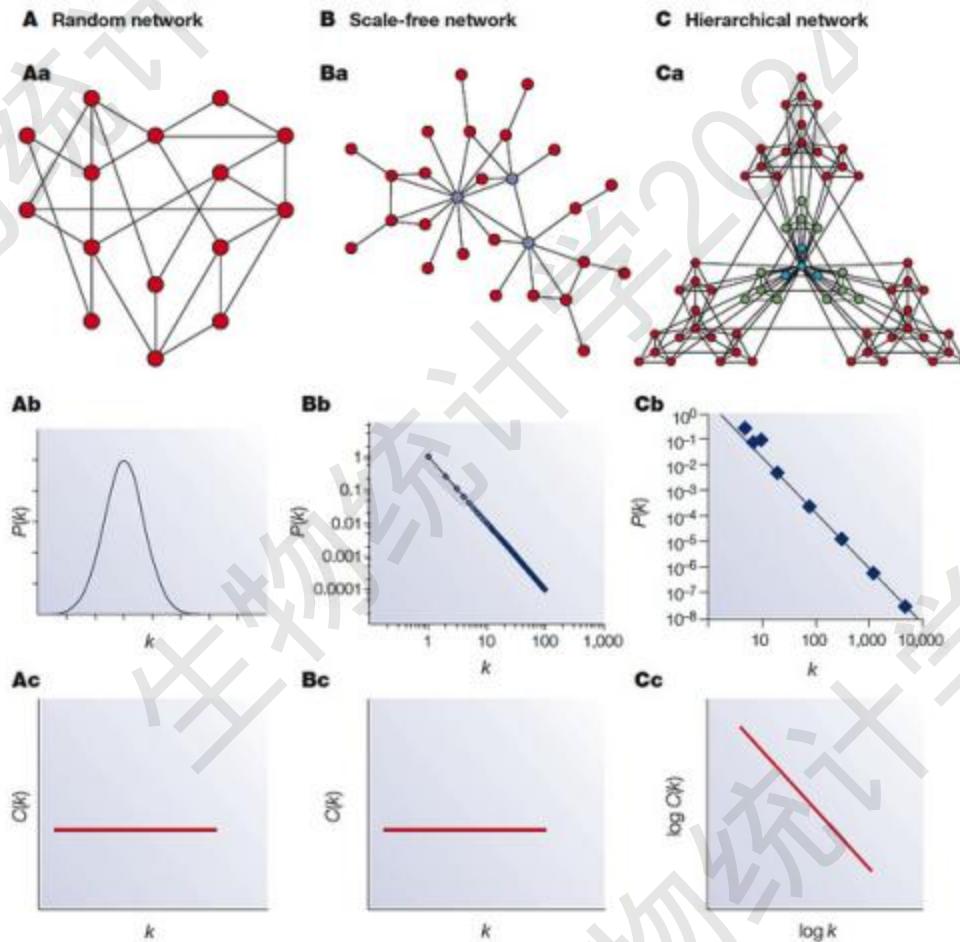
Lee et al. Science
2002.

Scale-free network



正态分布（左）与无标度网络（右）的比较

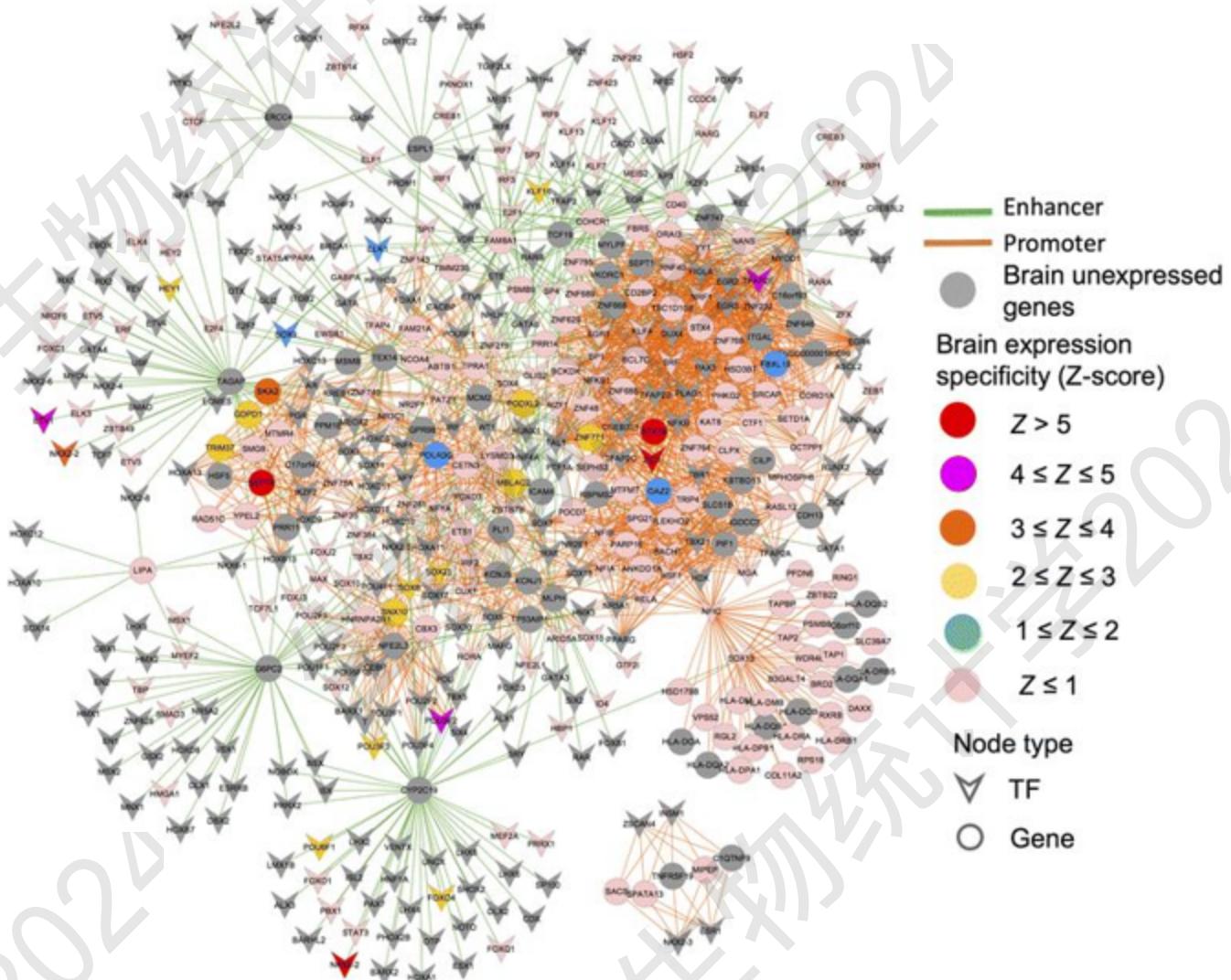
Scale-free network



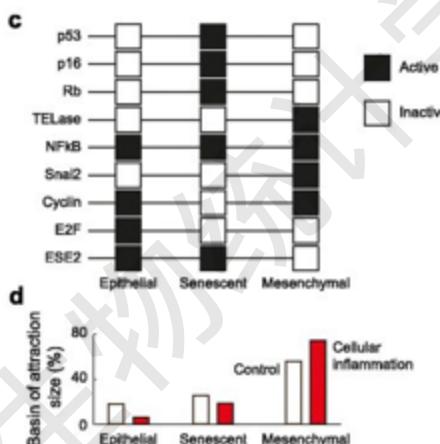
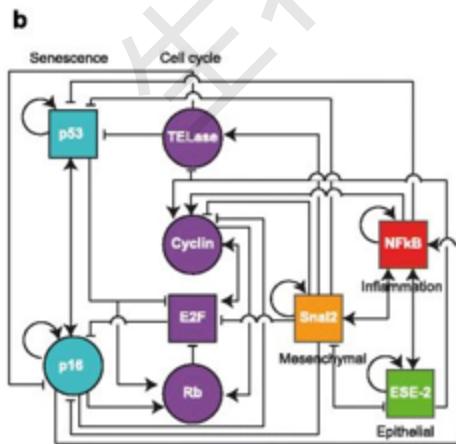
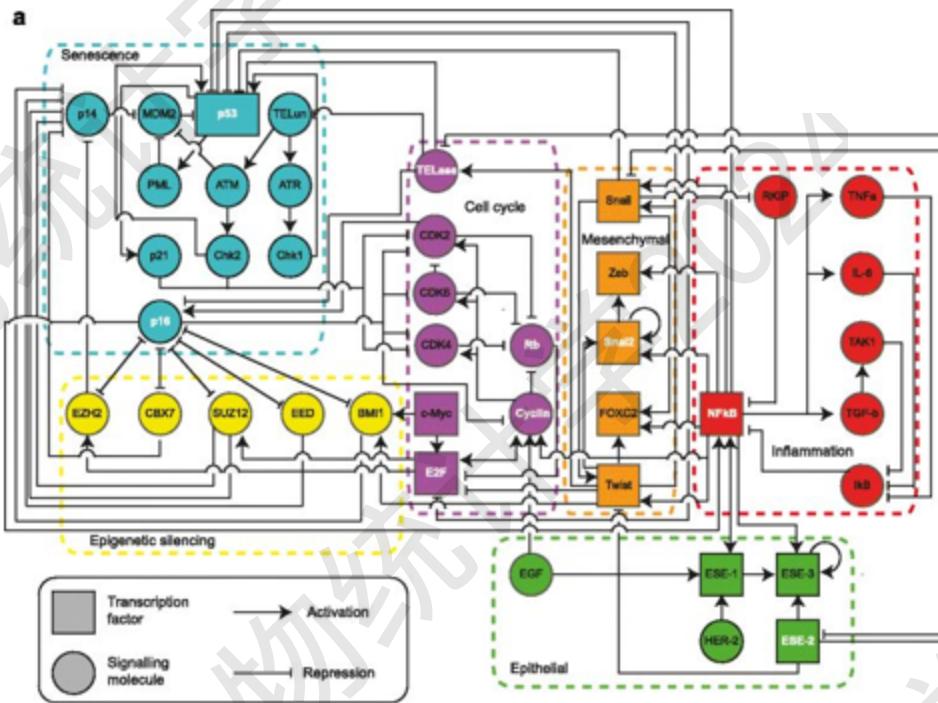
Scale-free network



Regulatory Network



Regulatory Network



Transcription Regulatory Code

- Each gene is regulated by a set of TFs.
- Each TF can regulate many genes.
- Which genes are regulated by which TFs on which conditions?
- How does regulator control the expression of its target gene?

How to Clarify Transcription Regulatory code ?

In silico.

- From sequence to gene regulatory network.
- Find all the potential TFBS upstream a gene.
- *Predict gene expression from gene sequence.*

Cell, 2004.



Too much noise!

Experimental methods

- Gel shift
- DNA footprinting.
- Reporter genes
-

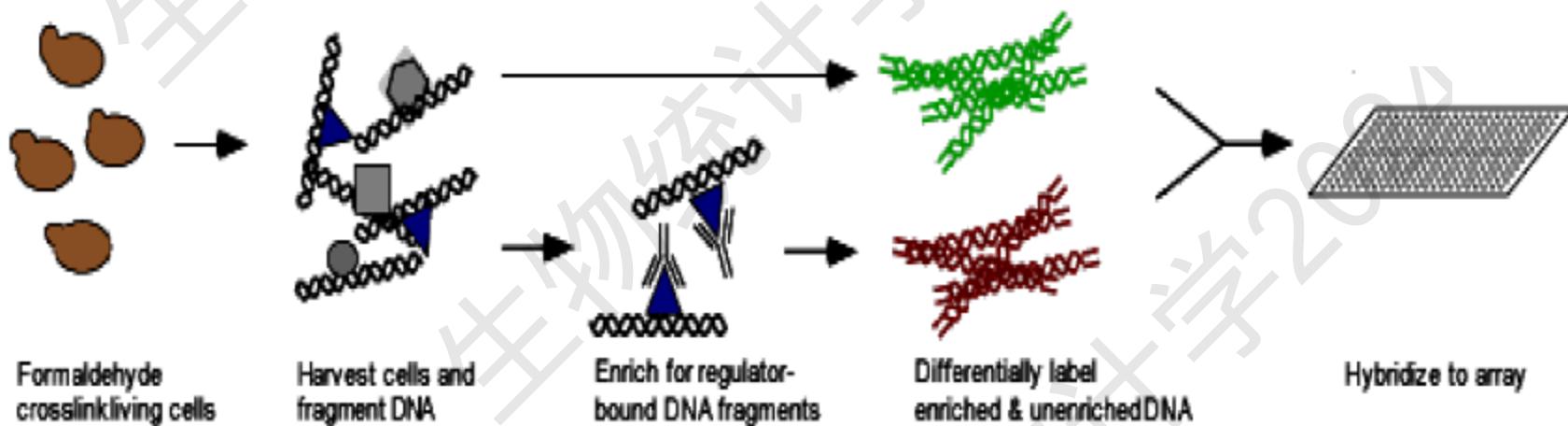


Not large scale

Not systematic

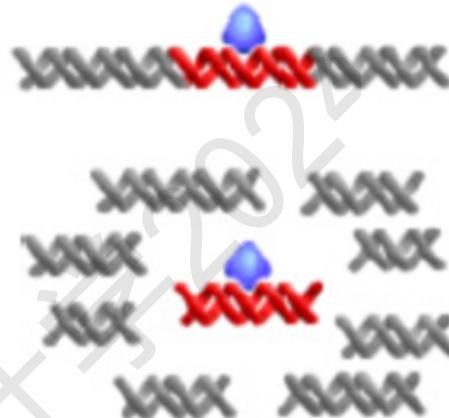
ChIP-chip Experiments

Identify all the target genes that can be directly or indirectly bind by a TF.

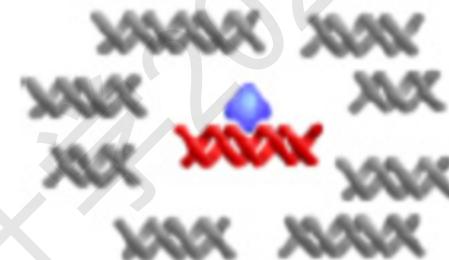


ChIP-chip Experiments

1. Cross-link
the proteins to DNA



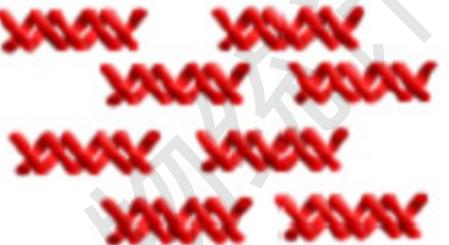
2. Fragment the DNA



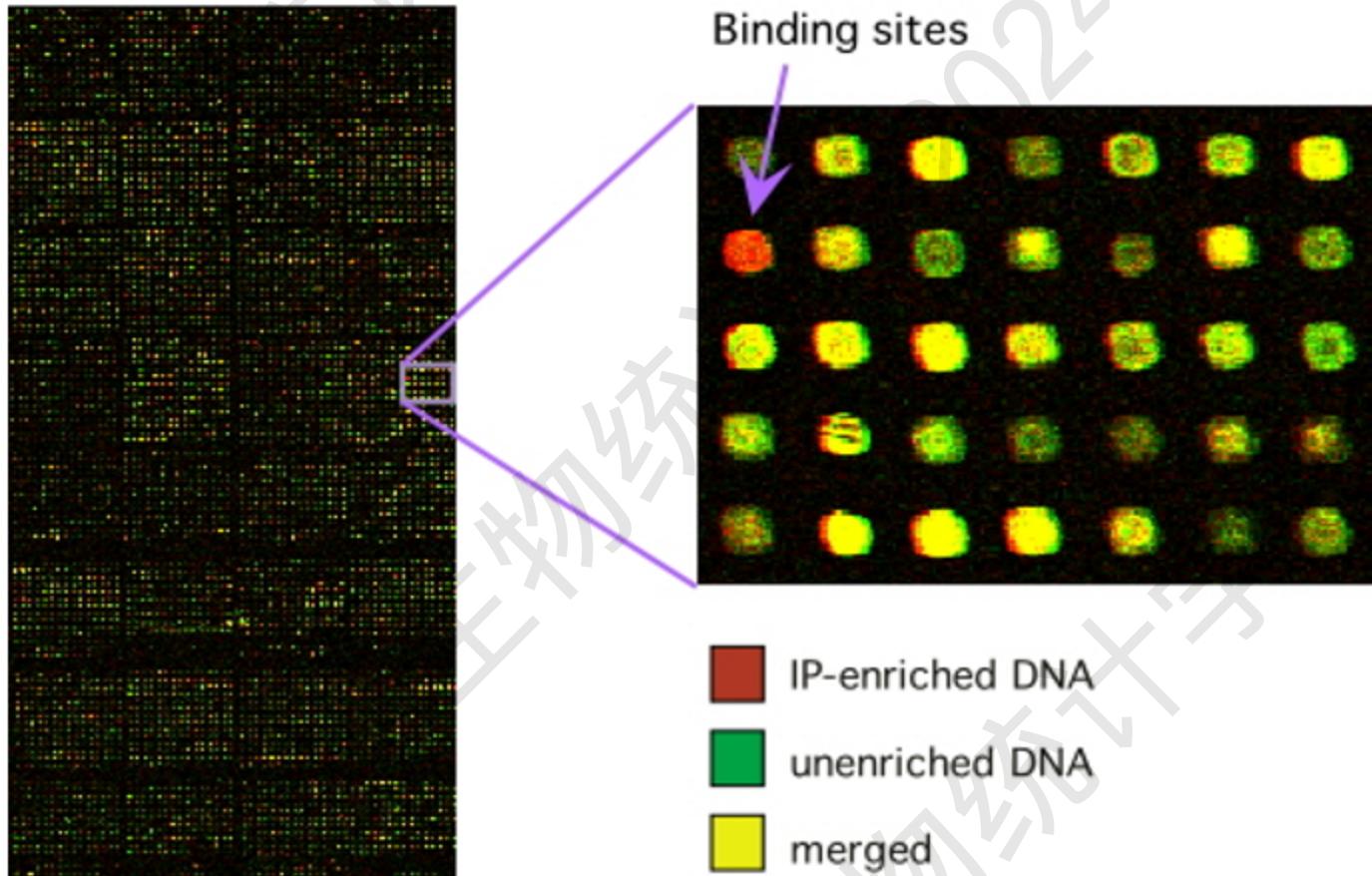
3. Immunoprecipitate
the protein of choice



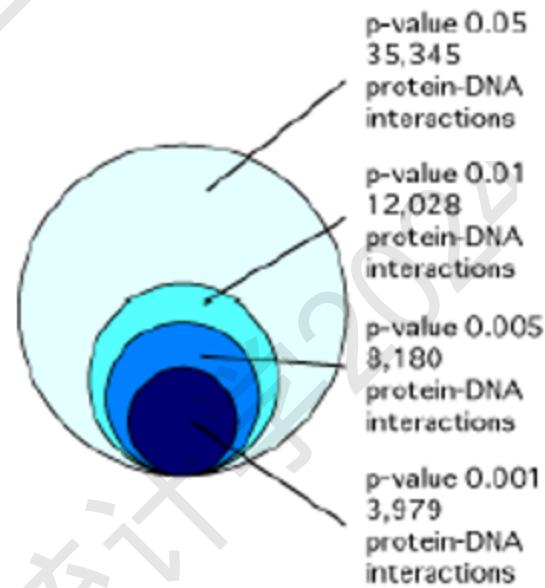
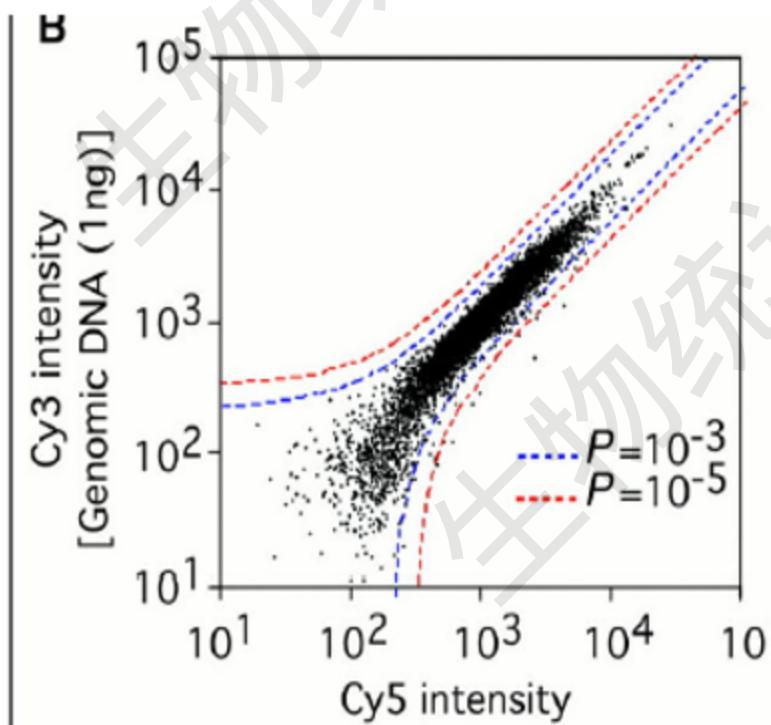
4. Amplify the DNA
attached to it



ChIP-chip Experiments

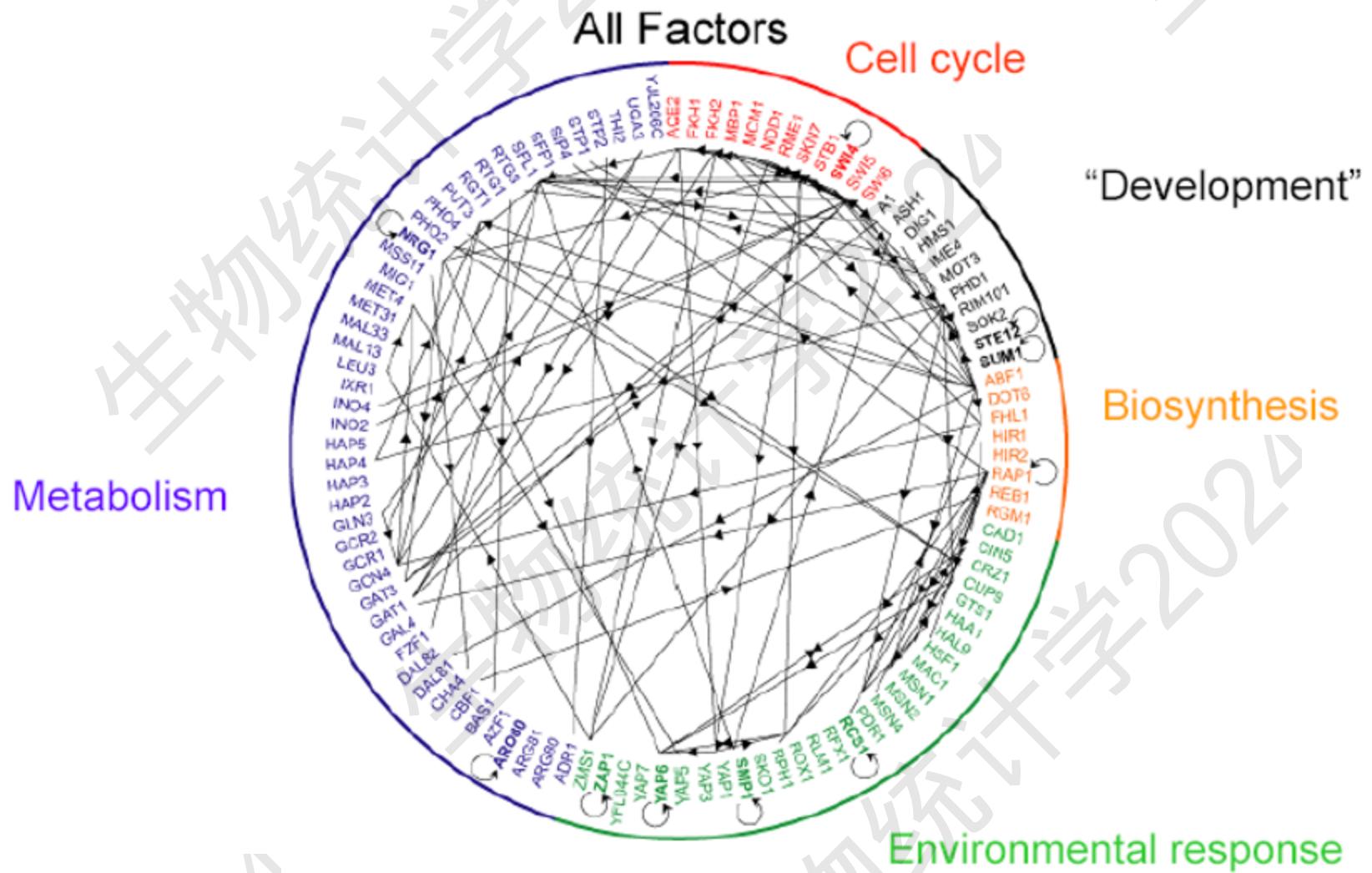


Protein-DNA Interactions

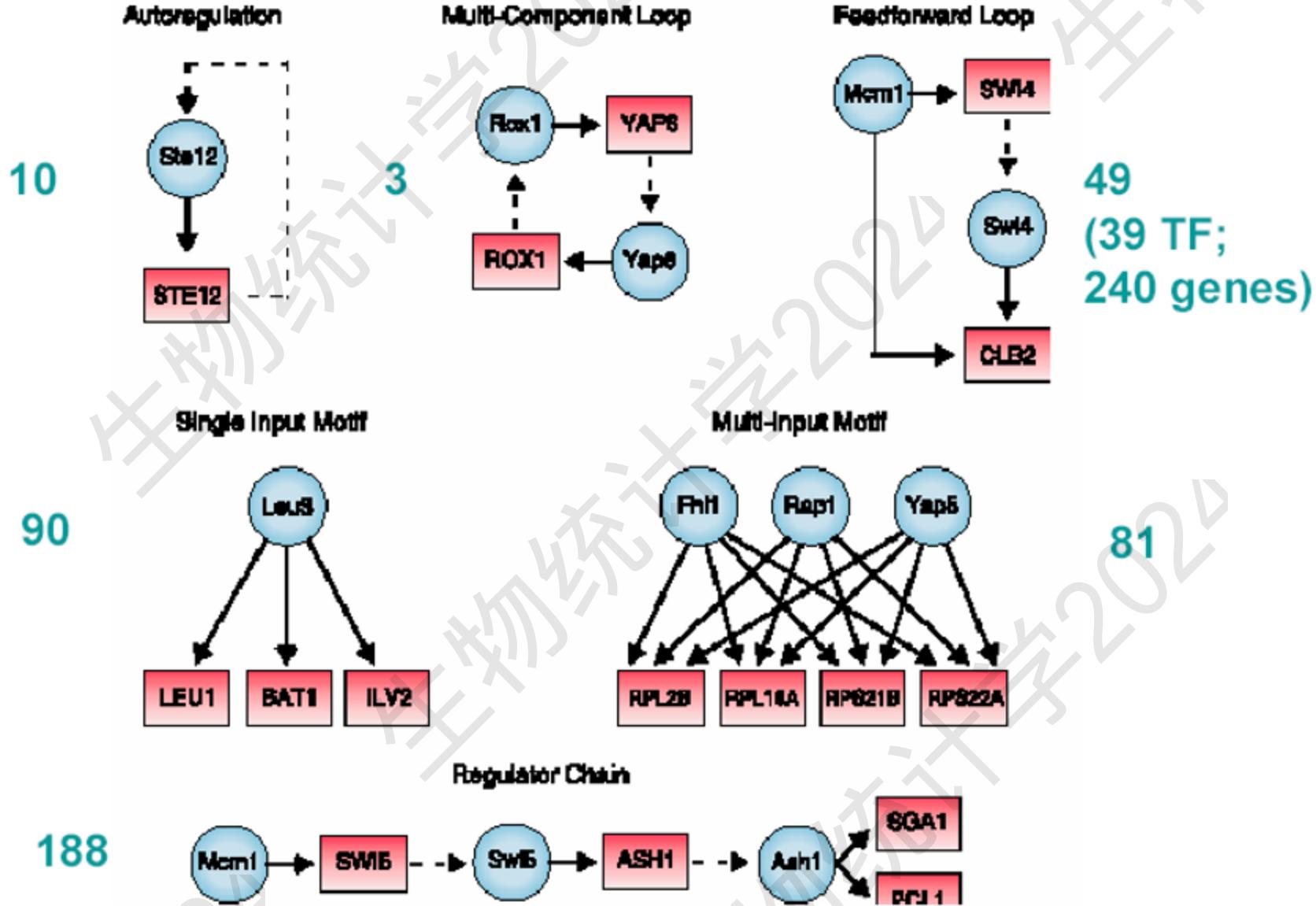


ChIP-chip Experiments

- 1 condition , 1 TF
Jason et.al. Nature (2001). Promoter-specific binding of Rap1 revealed by genome-wide maps of protein-DNA association.
- 1 condition, 106 TFs
Lee et.al. Science(2002). Transcription regulatory networks in *Saccharomyces cerevisiae*
- Multiple conditions , 203 TFs.
Harbison, et.al. (2004). Transcription regulatory code of a eukaryotic genome.



Lee, et al. Science, 2002.



Multiple Conditions

	Dat1 ⁺	Hap3 ⁺	Met18 ⁺	Pho4 ^{11,4+}	Sig1 ¹⁴⁺	Swi4 ⁺	YDR266C ⁺	+
A1 ⁺								
Abf1 ⁺	Dig1 ^{5,6+}	Hap4 ^{2,3+}	Met28 ³⁺	Pip2 ⁺	Sip3 ⁺	Swi5 ⁺	YDR520C ⁺	+
Abt1 ⁺	Dot6 ⁺	Hap5 ³⁺	Met31 ³⁺	Ppr1 ⁺	Sip4 ³⁺	Swi6 ⁺	YER051W ⁺	+
Aca1 ⁺	Ecm22 ⁺	Hir1 ⁺	Met32 ³⁺	Put3 ^{2,3+}	Skn7 ^{1,2,7+}	Tbs1 ⁺	YER130C ⁺	+
Ace2 ⁺	Eds1 ⁺	Hir2 ⁺	Met4 ³⁺	Rap1 ³⁺	Sko1 ⁺	Tec ^{15,6+}	YER184C ⁺	+
Adr1 ^{3,7+}	Fap7 ⁺	Hir3 ⁺	Mga1 ¹⁺	Rco1 ⁺	Smk1 ⁺	Thi2 ¹²⁺	YFL044C ⁺	+
Aft21,2 ⁺	Fhl1 ^{1,3,4+}	Hms1 ⁺	Mig1 ⁸⁺	Rcs1 ^{1,2,3+}	Smp1 ⁺	Tos8 ⁺	YFL052W ⁺	+
Arg80 ³⁺	Fkh1 ⁺	Hms2 ⁺	Mig2 ¹⁺	Rdr1 ⁺	Snf1 ⁺	Tye7 ⁺	YGR067C ⁺	+
Arg81 ³⁺	Fkh2 ^{1,2+}	Hog1 ⁺	Mig3 ⁺	Rds1 ¹⁺	Snt2 ⁺	Uga3 ^{3,4+}	Yhp1 ⁺	+
Ar080 ³⁺	Fzf1 ⁺	Hsf1 ^{1,2,7+}	Mot3 ^{1,2,3+}	Reb1 ^{1,2+}	Sok2 ⁵⁺	Ume6 ¹⁺	YJL206C1 ²⁺	+
Arr1 ¹⁺	Gal3 ⁺	Ifh1 ⁺	Msn1 ⁺	Rfx1 ⁺	Spt10 ⁺	Upc2 ⁺	YKL222C ⁺	+
Ash1 ⁵⁺	Gal4 ^{8,9+}	Ime1 ¹⁺	Msn2 ^{1,2,4,7,10}	Rgm1 ⁺	Spt2 ⁺	Usv1 ⁺	YKR064W ⁺	+
Ask10 ⁺	Gal80 ⁺	Ime4 ¹⁺	Msn4 ^{1,2,4,10}	Rgt1 ⁸⁺	Spt23 ⁺	War1 ⁺	YLR278C ⁺	+
Azf1 ⁺	Gat1 ^{3,4,7+}	Ino2 ⁺	Mss11 ⁵⁺	Rim101 ^{1,2+}	Srd1 ⁺	Wtm1 ⁺	YML081W ⁺	+
Bas1 ³⁺	Gat3 ⁺	Ino4 ⁺	Mth1 ⁸⁺	Rlm1 ⁵⁺	Stb1 ⁺	Wtm2 ⁺	YNR063W ⁺	+
Bye1 ⁺	Gcn4 ^{3,4+}	Ixr1 ⁺	Ndd1 ⁺	Rlr1 ⁺	Stb2 ⁺	Xbp1 ^{2,7+}	Yox1 ⁺	+
Cad11 ³⁺	Gcr1 ⁺	Kre33 ⁺	Ndt80 ⁺	Rme1 ⁺	Stb4 ⁺	Yap1 ^{1,2,7+}	YPR022C ⁺	+
Cbf1 ³⁺	Gcr2 ³⁺	Kss1 ^{5,6+}	Nnf2 ⁺	Rox1 ^{1,2+}	Stb5 ⁺	Yap3 ¹⁺	YPR196W ⁺	+
Cha4 ³⁺	Gln3 ^{3,4+}	Leu3 ³⁺	Nrg1 ^{1,2+}	Rph1 ^{1,2,3+}	Stb6 ⁺	Yap5 ¹⁺	Yrr1 ⁺	+
Cin5 ^{1,2+}	Gts1 ⁺	Mac1 ¹⁺	Oaf1 ⁺	Rpi1 ⁺	Ste12 ^{5,8+}	Yap6 ^{1,2+}	Zap1 ⁺	+
Crz1 ⁺	Gzf3 ^{1,4+}	Mal13 ⁺	Opi1 ⁺	Rpn4 ^{1,2+}	Stp1 ³⁺	Yap7 ^{1,2+}	Zms1 ⁺	+
Cst6 ⁺	Haa1 ⁺	Mal33 ^{1,2+}	Pdc2 ⁺	Rtg1 ^{3,4+}	Stp2 ⁺	YBL054W ⁺	+	+
Cup9 ⁺	Hac1 ⁺	Mbf1 ⁺	Pdr1 ²⁺	Rtg3 ^{1,2,3,4+}	Stp4 ⁺	YBR239C ⁺	+	+
Dal80 ⁴⁺	Hal9 ⁺	Mbp1 ^{1,2+}	Pdr3 ⁺	Rts2 ⁺	Sum1 ⁺	YBR267W ⁺	+	+
Dal81 ^{3,4+}	Hap1 ⁺	Mcm1 ^{5,6+}	Phd1 ⁵⁺	Sfl1 ⁺	Sut1 ⁺	YDR026C ⁺	+	+
Dal82 ^{3,4+}	Hap2 ⁴⁺	Mds3 ⁺	Pho2 ^{1,2,3,11+}	Sfp1 ^{1,2,3+}	Sut2 ⁺	YDR049W ⁺	+	+

¹ Highly hyperoxic

² Mildly hyperoxic

³ Amino acid starved

⁴ Nutrient deprived

⁵ Filamentation

⁶ Mating

⁷ Heat

⁸ Galactose

⁹ Raffinose

¹⁰ Acidic⁺

¹¹ Phosphate deprived⁺

¹² Vitamin deprived⁺

*All regulators
were profiled
in rich
medium

*A subset of these
were profiled
in at least
one of 12
other
environmental
conditions

Harbison et al.
Nature 2004.

Part II: Reverse Engineering

- Given: a (large) set of gene expression observations
- Goal: find the network fits that observation data.
- References:
 - Gardner, di Bernardo, Lorenz, and Collins. Inferring Genetic Networks and Identifying Compound Mode of Action via Expression Profiling. *Science* **301**, pp.102-105 (2003)
 - Michael Hecker, Sandro Lambeck, Susanne Toepfer, Eugene van Someren, Reinhard Guthke. Gene regulatory network inference: Data integration in dynamic models—A review. *BioSystems* **96** (2009) 86–103.

Reverse Engineering



Reverse Engineering



Bayesian Method

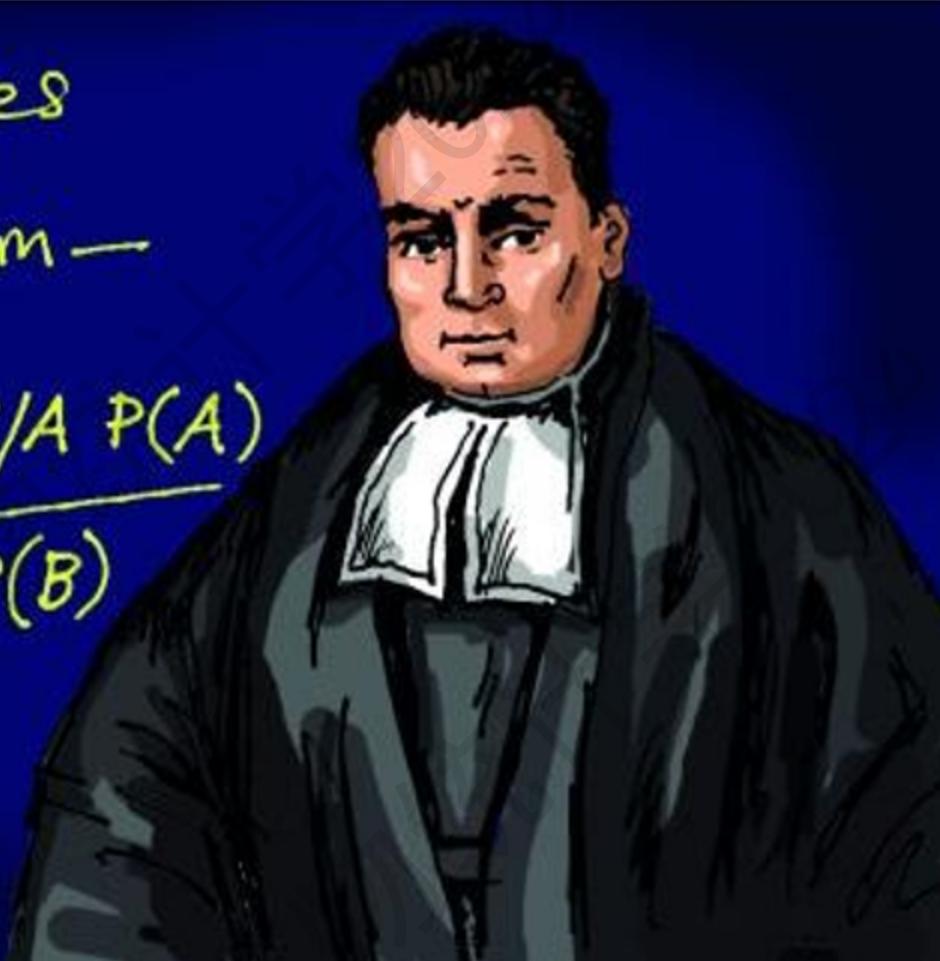
- 所谓的贝叶斯方法源于他生前为解决一个“逆概”问题写的一篇文章，而这篇文章是在他死后才由他的一位朋友发表出来的。在贝叶斯写这篇文章之前，人们已经能够计算“正向概率”，如“假设袋子里里面有N个白球，M个黑球，你伸手进去摸一把，摸出黑球的概率是多大”。而一个自然而然的问题是反过来：“如果我们事先并不知道袋子里里面黑白球的比例，而是闭着眼睛摸出一个（或好几个）球，观察这些取出来的球的颜色之后，那么我们可以就此对袋子里里面的黑白球的比例作出什么样的推测”。这个问题，就是所谓的逆概问题。

Bayesian Method

Thomas Bayes

Bayes' theorem —

$$P(A|B) = \frac{P(B|A) P(A)}{P(B)}$$



Bayesian Method

- Bayes' theorem:

$$P(A|B) = \frac{P(B|A) P(A)}{P(B)}$$

DREAM Project

- DREAM: Dialogue for Reverse Engineering Assessments and Methods.
- Objective: To catalyze the interaction between experiment and theory in the area of cellular network inference and quantitative model building in systems biology.
- <http://dreamchallenges.org/>
- <http://dreamchallenges.org/challenges/> (current DREAM)

Modeling Expression with Differential Equations

Assumes network behavior can be modeled as a system of linear differential equations of the form:

$$dx/dt = Ax + u$$

x is a vector representing the continuous-valued levels (concentrations) of each network component

A is the network model: an $N \times N$ matrix of coefficients describing how each x_i is controlled by upstream genes $x_j, x_k, \text{ etc.}$

u is a vector representing an external additive perturbation to the system

An example:

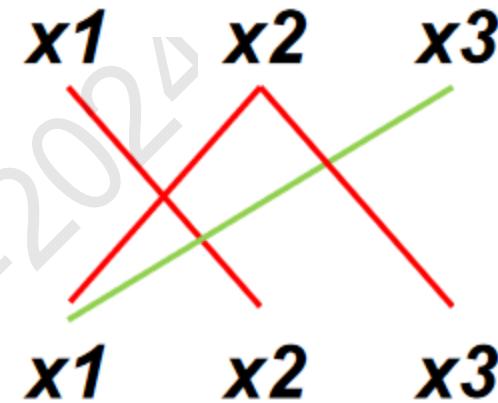
From discrete- to continuous-valued networks

Three genes: x_1, x_2, x_3

x_1 activates x_2

x_2 activates x_1 and x_3

x_3 inhibits x_1



$$dx/dt = Ax + u$$

$$dx_1/dt = a_{12}x_2 - a_{13}x_3$$

$$dx_2/dt = a_{21}x_1$$

$$dx_3/dt = a_{32}x_2$$

$$\frac{d}{dt} \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix} = \begin{bmatrix} 0 & a_{12} & -a_{13} \\ a_{21} & 0 & 0 \\ 0 & a_{32} & 0 \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix} + \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}$$

The steady state assumption

- Near a steady-state point, expression levels do not change over time.
- Under the steady-state assumption, the model reduces to $0 = \mathbf{Ax} + \mathbf{u} \rightarrow \mathbf{Ax} = -\mathbf{u}$
- A straightforward method to infer \mathbf{A} would be to apply N perturbations, \mathbf{u} , to the network, in each case measuring steady-state expression levels for the \mathbf{x} .
- However, in larger networks it may be impractical to apply so many perturbations
- As a simplifying assumption, ***consider that each gene has a maximum of k non-zero regulatory inputs.***

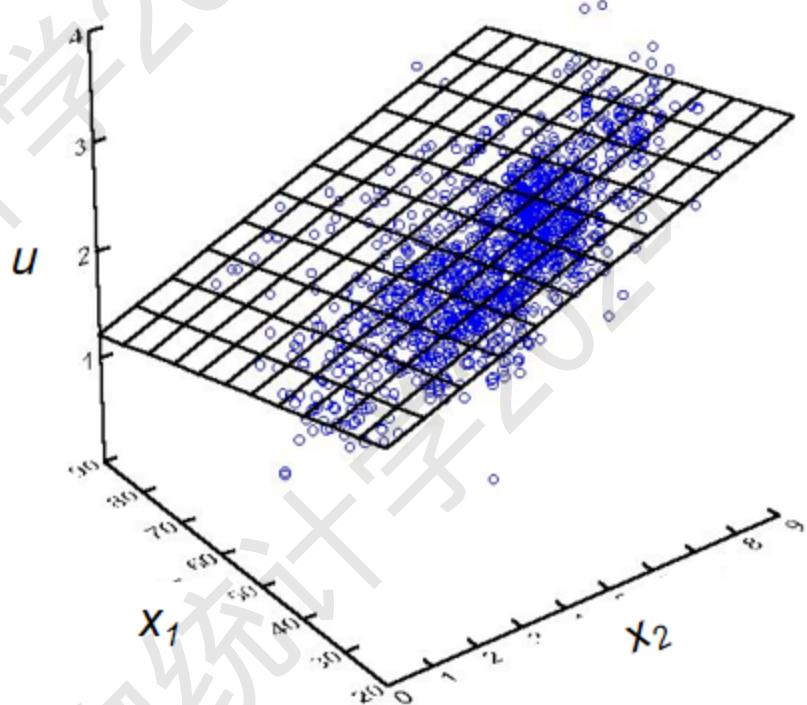
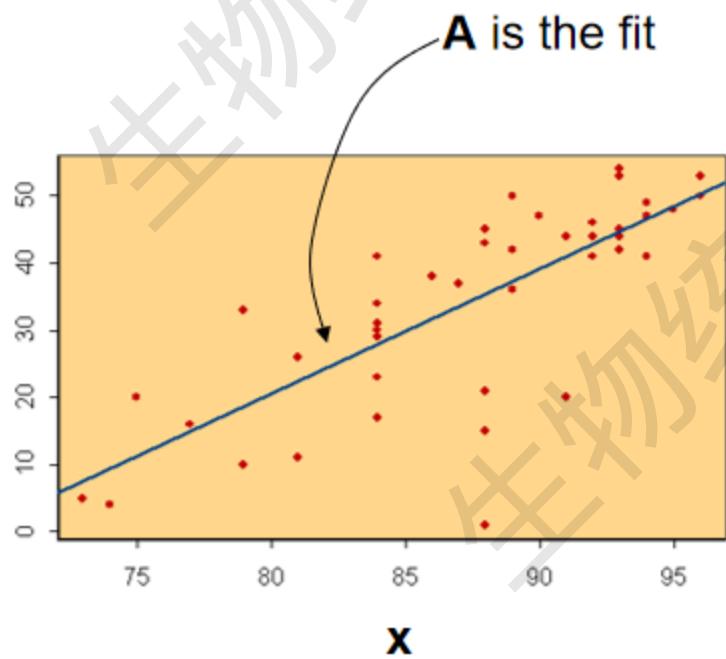
The Inference Procedure

$$Ax = -u$$

- Infer inputs to each gene separately
- For the given gene, consider all possible combinations of the k regulatory inputs
- For each combination, use multiple linear regression to determine optimal values of the k coefficients
- Choose the combination that fits the observed data with the least error

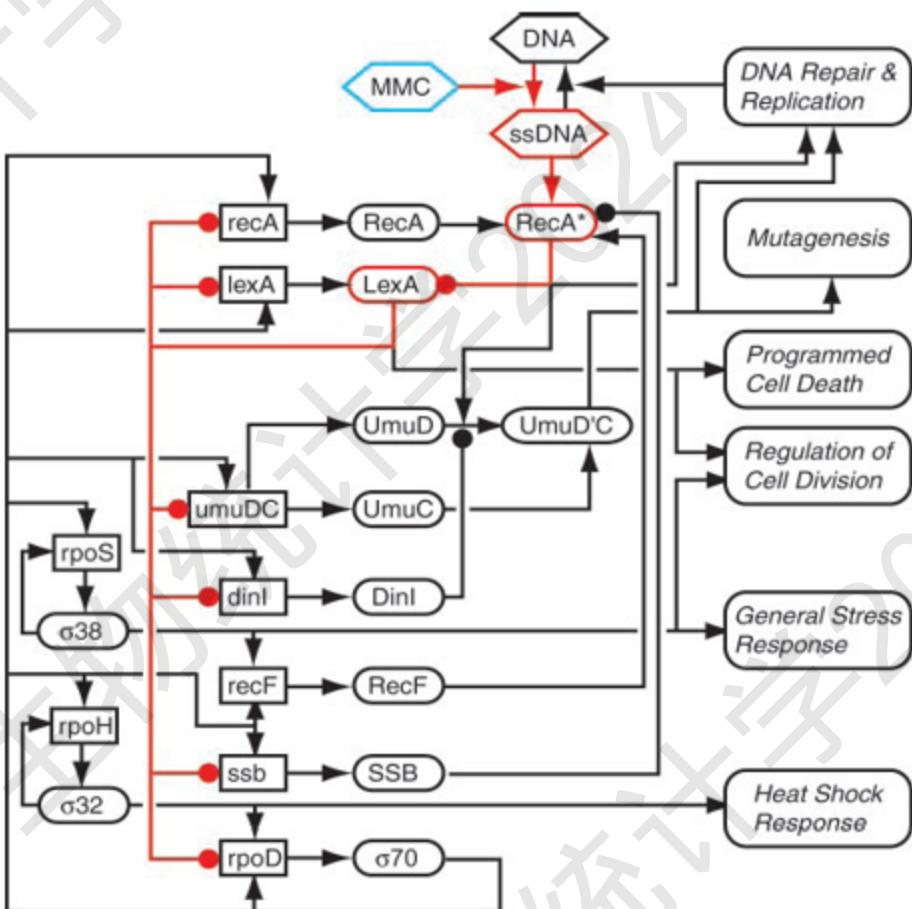
Multiple regression

$$u = -Ax$$



Application to SOS System

Fig. 1. Diagram of interactions in the SOS network. DNA lesions caused by mitomycin C (MMC) (blue hexagon) are converted to single-stranded DNA during chromosomal replication. Upon binding to ssDNA, the RecA protein is activated (RecA*) and serves as a coprotease for the LexA protein. The LexA protein is cleaved, thereby diminishing the repression of genes that mediate multiple protective responses. Boxes denote genes, ellipses denote proteins, hexagons indicate metabolites, arrows denote positive regulation, filled circles denote negative regulation. Red emphasis denotes the primary pathway by which the network is activated after DNA damage.

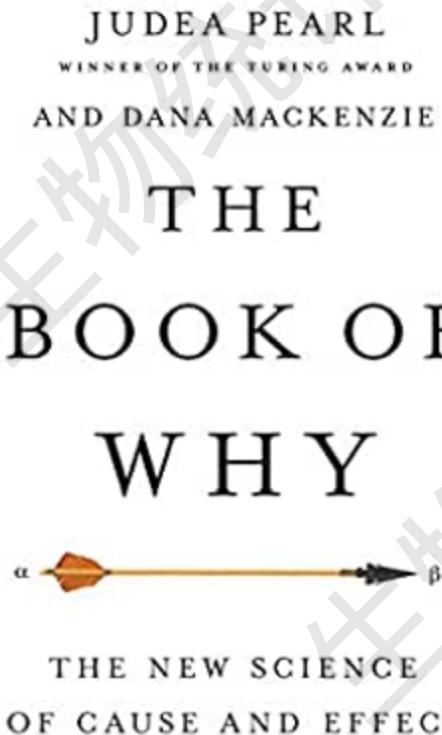


Part III: Bayesian Network

- 本部分Slides主要来自于N.Friedman and D.Heckman's slides.
- References:
- N.Friedman et al. Using Bayesian Networks to analyze expression data. *J. Comput. Biol.*, 7:601-620, 2000.

人工智能中的因果(Cause and Effect)推断

"JUST EXTRAORDINARY." —SCIENCE FRIDAY (NPR)



[美]朱迪·珀尔 JUDEA PEARL

—图灵奖获得者—

[英]达纳·麦肯齐 DANA MACKENZIE 著

江生 于华 译

THE BOOK OF WHY

为什么

关于因果关系的新科学



THE NEW SCIENCE
OF CAUSE AND EFFECT

中信出版集团

Motivation

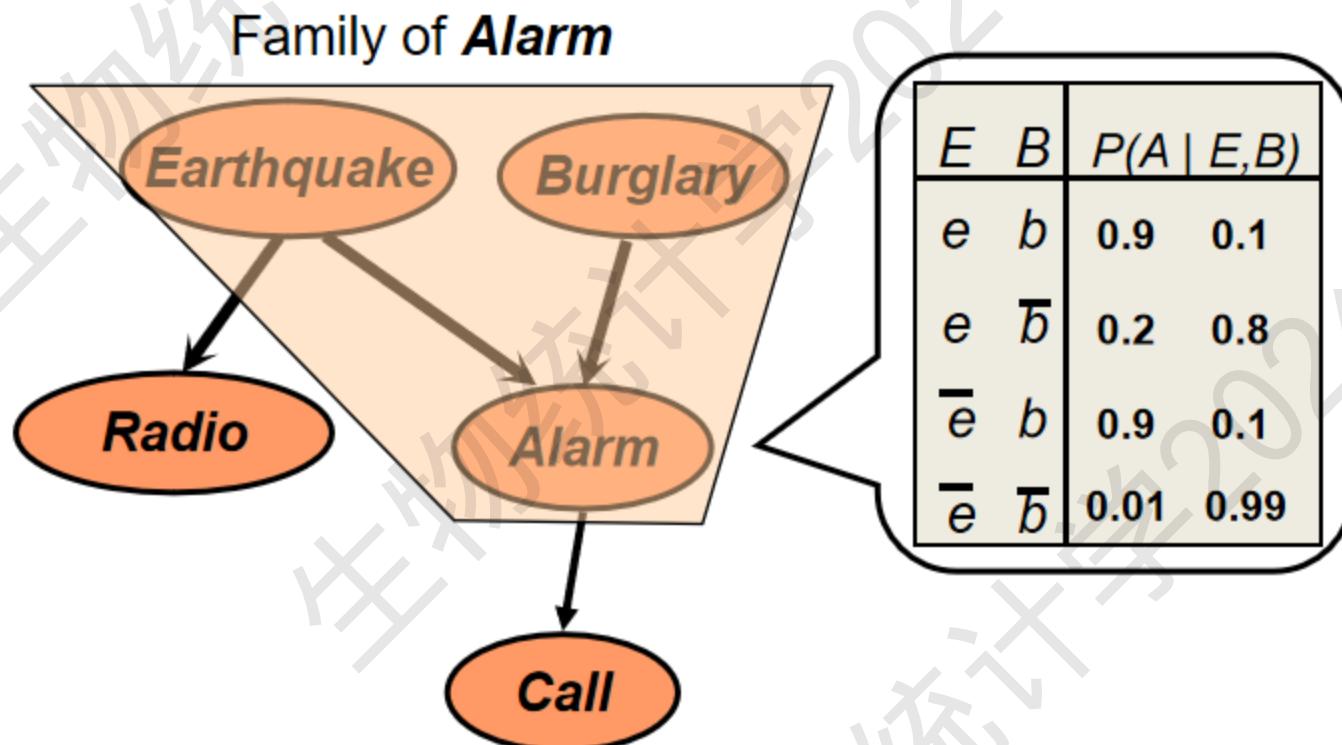
- Given gene expression data, what's the relationship between genes?
 - Who regulates who?
 - How does one gene regulate other gene?
- Exploring the relationship among features to construct a better classifier instead of treating them independently.

Bayesian Network

- Directed acyclic graph (DAG).
 - Nodes: random variables.
 - Edges: direct influence.
- Set of conditional probability distributions.
- Joint distribution.

$$p(\mathbf{X}) = \prod_{i=1}^n p(X_i \mid \text{parents}(X_i)).$$

Bayesian Networks: Example



$$P(B, E, A, C, R) = P(B)P(E)P(A | B, E)P(R | E)P(C | A)$$

隐马氏模型的数学问题

- 识别问题—已知若干个隐马氏模型及其参数，对一个观测样本，决定它来自哪一个模型。
- 解码问题—由观测样本得到隐状态；
- 学习问题—由观测样本得到参数组 λ ；

Bayesian Network

- 初级：参数学习
- 中级：图分解
- 高级：近似算法
- 特级：EM算法

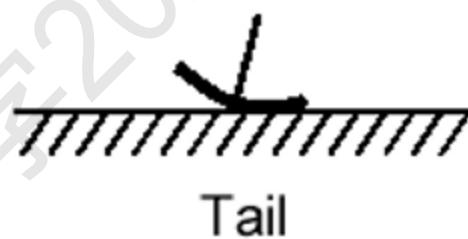
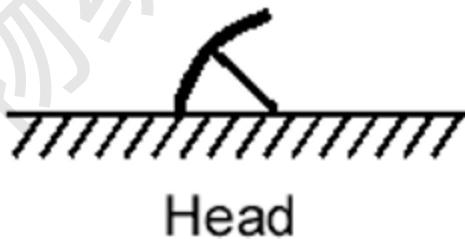
Learning Problems

- Estimation of the parameters.
- Construct the structure.

Let's start from the basic parameter estimation problem.

A: Learning Parameters

Simple Case: Binomial Experiment



- ◆ When tossed, it can land in one of two positions: Head or Tail
- ◆ We denote by θ the (unknown) probability $P(H)$.

Estimation task:

- ◆ Given a sequence of toss samples $x[1], x[2], \dots, x[M]$ we want to estimate the probabilities $P(H) = \theta$ and $P(T) = 1 - \theta$

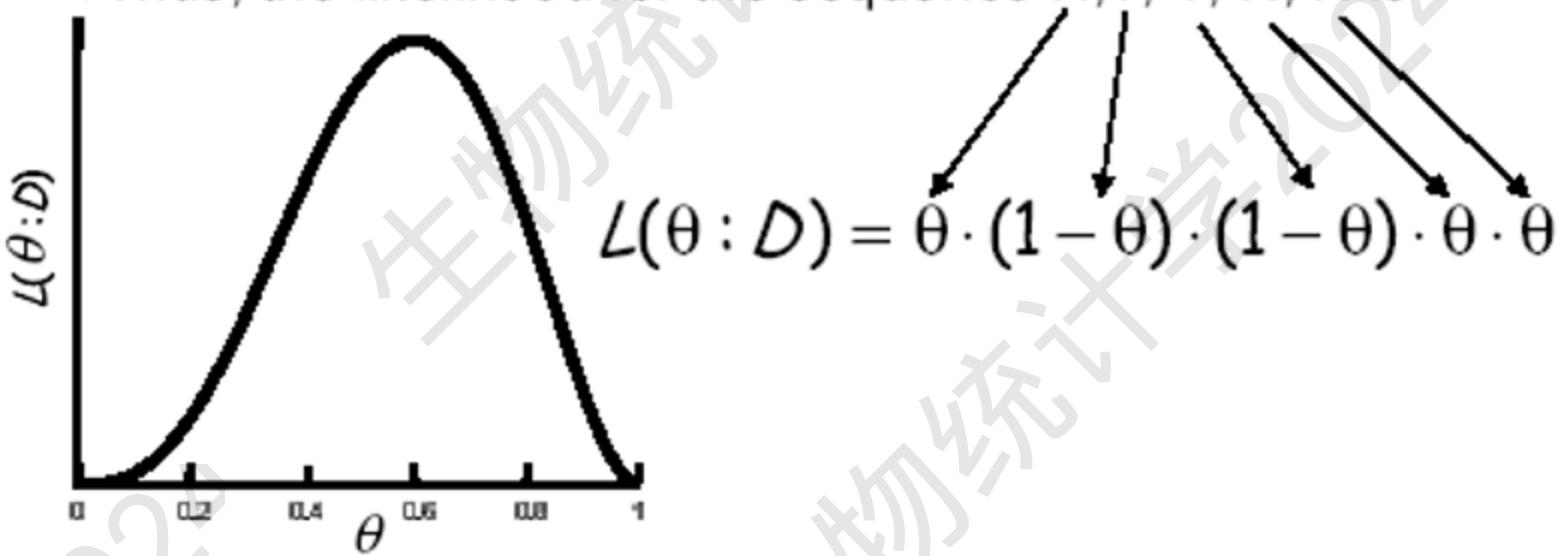
Likelihood Function

- ◆ How good is a particular θ ?

It depends on how likely it is to generate the observed data

$$L(\theta : D) = P(D | \theta) = \prod_m P(x[m] | \theta)$$

- ◆ Thus, the likelihood for the sequence H, T, T, H, H is



Sufficient Statistics

- ◆ To compute the likelihood in the thumbtack example we only require N_H and N_T
(the number of heads and the number of tails)

$$L(\theta : D) = \theta^{N_H} \cdot (1 - \theta)^{N_T}$$

N_H and N_T are **sufficient statistics** for the binomial distribution

- ◆ A **sufficient statistic** is a function that summarizes, from the data, the relevant information for the likelihood
 - If $s(D) = s(D')$, then $L(\theta | D) = L(\theta | D')$

Maximum Likelihood Estimation (MLE)

- MLE principle: Learn parameters that maximize the likelihood function.
- This is one of the most commonly used estimation in statistics (Classical approach) and intuitively appealing.

MLE In Binomial Case

- Applying the MLE principle we get

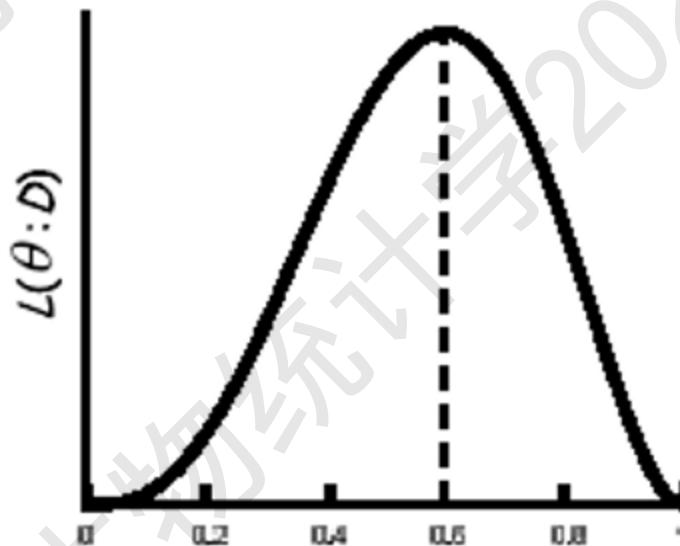
$$\hat{\theta} = \frac{N_H}{N_H + N_T}$$

(Which coincides with what one would expect)

Example:

$$(N_H, N_T) = (3, 2)$$

MLE estimate is $3/5 = 0.6$

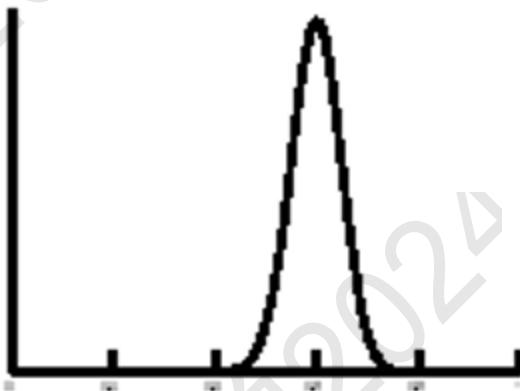


MLE is Not Enough

- ◆ MLE commits to a specific value of the unknown parameter(s)



Coin



Thumbtack

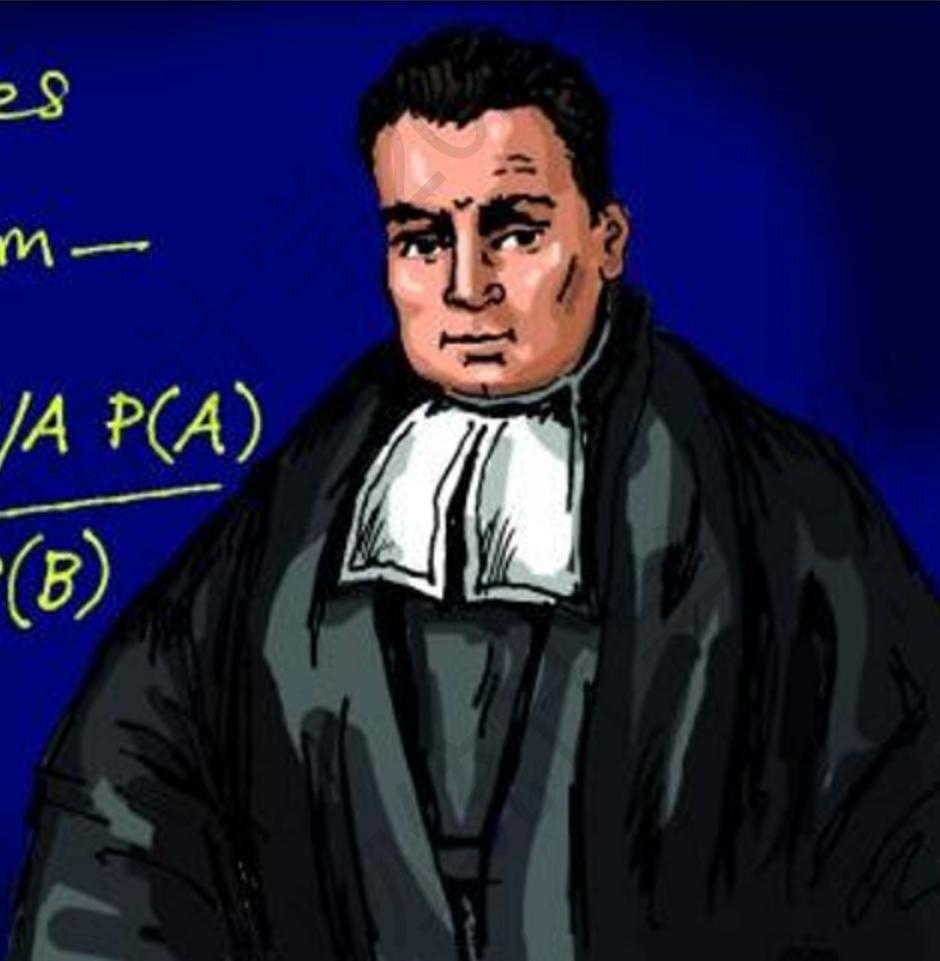
- ◆ MLE is the same in both cases
- ◆ Confidence in prediction is clearly different

Bayesian Inference

Thomas Bayes

Bayes' theorem —

$$P(A|B) = \frac{P(B|A) P(A)}{P(B)}$$



Bayesian Inference

- Representing uncertainty about parameters using a probability distribution over parameters, data.
- Using Bayes' rule to learn.
 - Data (D) and the $p(x|\xi)$ probability distribution
 - Prior distribution $p(\theta|\xi)$

$$p(\theta|D, \xi) = \frac{p(\theta|\xi) p(D|\theta, \xi)}{p(D|\xi)}$$

$$p(D|\xi) = \int p(D|\theta, \xi) p(\theta|\xi) d\theta$$

Binomial Experiment Revised

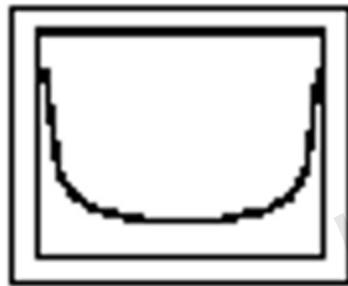
Prior: Beta distribution

$$\begin{aligned} p(\theta) &= \text{Beta}(\alpha_H, \alpha_T) \\ &= \frac{\Gamma(\alpha_H + \alpha_T)}{\Gamma(\alpha_H) + \Gamma(\alpha_T)} \theta^{\alpha_H - 1} (1 - \theta)^{\alpha_T - 1} \end{aligned}$$

Posterior:

$$\begin{aligned} p(\theta|D) &= \text{Beta}(\alpha_H + N_H, \alpha_T + N_T) \\ &= \frac{\Gamma(\alpha_H + \alpha_T + N_H + N_T)}{\Gamma(\alpha_H + N_H) + \Gamma(\alpha_T + N_T)} \theta^{N_H + \alpha_H - 1} (1 - \theta)^{N_T + \alpha_T - 1} \end{aligned}$$

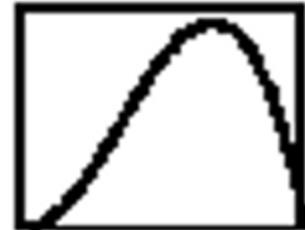
Beta Distribution



$\text{Beta}(0.5, 0.5)$



$\text{Beta}(1, 1)$



$\text{Beta}(3, 2)$



$\text{Beta}(19, 39)$

MAP (Maximum A-Posterior Probability)

- Using MAP, we can obtain an estimation of the parameter

$$\tilde{\theta} = \frac{\alpha_H + N_H}{\alpha_H + \alpha_T + N_H + N_T}$$

- Recall that the MLE is

$$\hat{\theta} = \frac{N_H}{N_H + N_T}$$

Intuition

- The hyperparameters a_H and a_T can be thought of imaginary counts (pseudo-counts) from our experience.
- Equivalent sample size= $a_H + a_T$.
- The larger the equivalent sample size, the more confident we are about the true probability.

Bayesian Inference vs. MLE

Frequentist Approach:

- ◆ Assumes there is an unknown but fixed parameter θ
- ◆ Estimates θ with some confidence
- ◆ Prediction by using the estimated parameter value

Bayesian Approach:

- ◆ Represents uncertainty about the unknown parameter
- ◆ Uses probability to quantify this uncertainty:
 - Unknown parameters as random variables
- ◆ Prediction follows from the rules of probability:
 - Expectation over the unknown parameters

Bayesian Inference vs. MLE (Cont.)

- In our example, MLE and Bayesian prediction differ.
- However, If prior is well-behaved (does not assign 0 density to any feasible parameter value), then both MLE and Bayesian prediction converge to the same value, the "true" distribution.

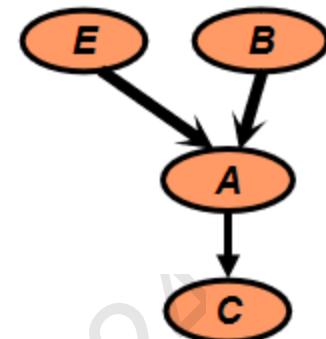
Bayesian Network

- 初级：参数学习
- 中级：图分解
- 高级：近似算法
- 特级：EM算法

Learning Parameters

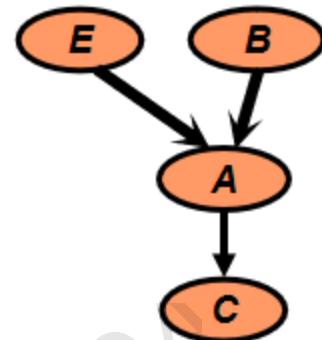
- Training data has the form:

$$D = \begin{bmatrix} E[1] & B[1] & A[1] & C[1] \\ \vdots & \vdots & \vdots & \vdots \\ E[M] & B[M] & A[M] & C[M] \end{bmatrix}$$



Likelihood Function

- Assume i.i.d. samples
- Likelihood function is



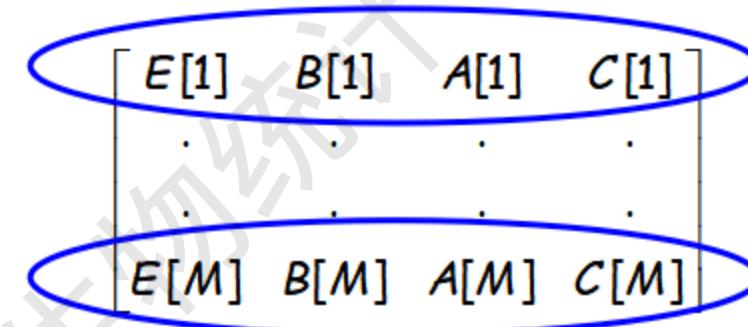
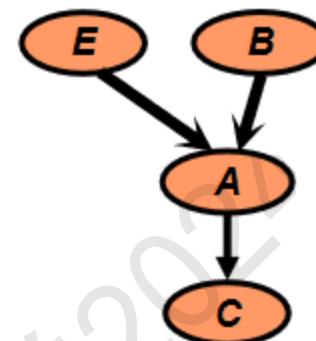
$$L(\Theta : D) = \prod_m P(E[m], B[m], A[m], C[m] : \Theta)$$

Likelihood Function

- By definition of network, we get

$$L(\Theta : D) = \prod_m P(E[m], B[m], A[m], C[m] : \Theta)$$

$$= \prod_m \left\{ \begin{array}{l} P(E[m] : \Theta) \\ P(B[m] : \Theta) \\ P(A[m] | B[m], E[m] : \Theta) \\ P(C[m] | A[m] : \Theta) \end{array} \right\}$$

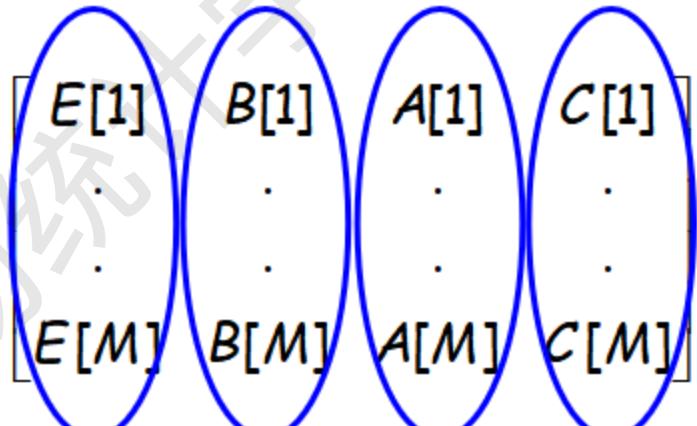
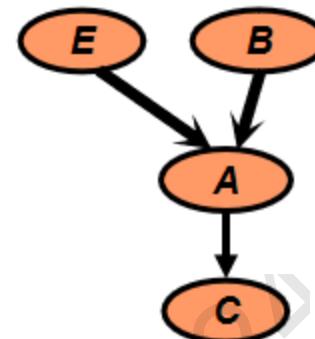


Likelihood Function

- Rewriting terms, we get

$$\begin{aligned} L(\Theta : D) &= \prod_m P(E[m], B[m], A[m], C[m] : \Theta) \\ &= \prod_m P(E[m] : \Theta) \\ &\quad \prod_m P(B[m] : \Theta) \\ &\quad \prod_m P(A[m] | B[m], E[m] : \Theta) \\ &\quad \prod_m P(C[m] | A[m] : \Theta) \end{aligned}$$

4 Subnetworks



General Bayesian Networks

Generalizing for any Bayesian network:

$$\begin{aligned}L(\Theta : D) &= \prod_m P(x_1[m], \dots, x_n[m] : \Theta) \\&= \prod_i \prod_m P(x_i[m] | Pa_i[m] : \Theta_i) \\&= \prod_i L(\Theta_i : D)\end{aligned}$$

The likelihood **decomposes** to small ones according to the structure of the network.

General Bayesian Networks (Cont.)

- **Decomposition \Rightarrow Independent estimation problems**
- If the parameters for each family are not related, they can be estimated independently of each other.

From Binomial to Multinomial

- ◆ For example, suppose X can have the values $1, 2, \dots, K$
- ◆ We want to learn the parameters $\theta_1, \theta_2, \dots, \theta_K$

Sufficient statistics:

- ◆ N_1, N_2, \dots, N_K - the number of times each outcome is observed

Likelihood function:

$$L(\theta: D) = \prod_{k=1}^K \theta_k^{N_k}$$

MLE:

$$\hat{\theta}_k = \frac{N_k}{\sum_{\ell} N_{\ell}}$$

From Beta to Dirichlet Distribution

Prior: Dirichlet distribution

$$p(\theta) = \text{Dir}(\theta | \alpha_1, \dots, \alpha_K)$$

$$= \frac{\Gamma(\alpha_1 + \dots + \alpha_K)}{\prod_{k=1}^K \Gamma(\alpha_k)} \prod_{k=1}^K \theta_k^{\alpha_k}$$

Posterior:

$$p(\theta | D) = \text{Dir}(\theta | \alpha_1, \dots, \alpha_K)$$

$$= \frac{\Gamma(\alpha_1 + \dots + \alpha_K + N_1 + \dots + N_K)}{\prod_{k=1}^K \Gamma(\alpha_k + N_k)} \prod_{k=1}^K \theta_k^{\alpha_k + N_k}$$

From Beta to Dirichlet Distribution (Cont.)

The MAP is

$$\theta_k = \frac{\alpha_k + N_k}{\sum_{l=1}^K (\alpha_l + N_l)}$$

The marginal likelihood is

$$\begin{aligned} P(D|G) &= \int P(D|\theta, G)P(\theta|G)d\theta \\ &= \frac{\Gamma(\sum_{k=1}^K \alpha_k)}{\prod_{k=1}^K \Gamma(\alpha_k)} \int_0^1 \prod_{k=1}^K \theta_k^{N_k + \alpha_k - 1} d\theta_k \\ &\stackrel{?}{=} \frac{\Gamma(\sum_{k=1}^K \alpha_k)}{\Gamma(\sum_{k=1}^K \alpha_k + \sum_{k=1}^K N_k)} \prod_{k=1}^K \frac{\Gamma(\alpha_k + N_k)}{\Gamma(\alpha_k)} \end{aligned}$$

Likelihood for Multinomial Network

- When we assume that $P(X_i | Pa_i)$ is multinomial, we get further decomposition:

$$\begin{aligned}L_i(\Theta_i; D) &= \prod_m P(x_i[m] | Pa_i[m]; \Theta_i) \\&= \prod_{pa_i[m]} \prod_{Pa_i[m]=pa_i} P(x_i[m] | pa_i; \Theta_i) \\&= \prod_{pa_i} \prod_{x_i} P(x_i | pa_i; \Theta_i)^{N(x_i, pa_i)} = \prod_{pa_i} \prod_{x_i} \theta_{x_i | pa_i}^{N(x_i, pa_i)}\end{aligned}$$

- For each value pa_i of the parents of X_i , we get an independent multinomial problem
- The MLE is $\hat{\theta}_{x_i | pa_i} = \frac{N(x_i, pa_i)}{N(pa_i)}$

Bayesian Inference for Multinomial Network

- Given data, we can compute the posterior for each multinomial independently. The posteriors are also Dirichlet with parameters

$$\alpha(X_j=1|pa_j) + N(X_j=1|pa_j), \dots, \alpha(X_j=k|pa_j) + N(X_j=k|pa_j)$$

- The predictive distribution is then represented by parameters

$$\tilde{\theta}_{x_j|pa_j} = \frac{\alpha(x_j, pa_j) + N(x_j, pa_j)}{\alpha(pa_j) + N(pa_j)}$$

More Generalizations

- Likelihood from exponential family.
 - Binomial distribution
 - Multinomial distribution
 - Poisson distribution
 - Gamma distribution
 - Normal distribution
- Conjugated distributions.

Learning Parameters: Summary

- Estimation relies on **sufficient statistics**

- For multinomials: counts $N(x_i, pa_i)$
 - Parameter estimation

$$\hat{\theta}_{x_i | pa_i} = \frac{N(x_i, pa_i)}{N(pa_i)}$$

MLE

$$\tilde{\theta}_{x_i | pa_i} = \frac{\alpha(x_i, pa_i) + N(x_i, pa_i)}{\alpha(pa_i) + N(pa_i)}$$

Bayesian (Dirichlet)

- Both are asymptotically equivalent.

B. Learning Structure From Data

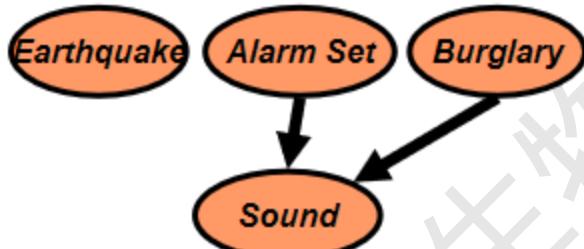
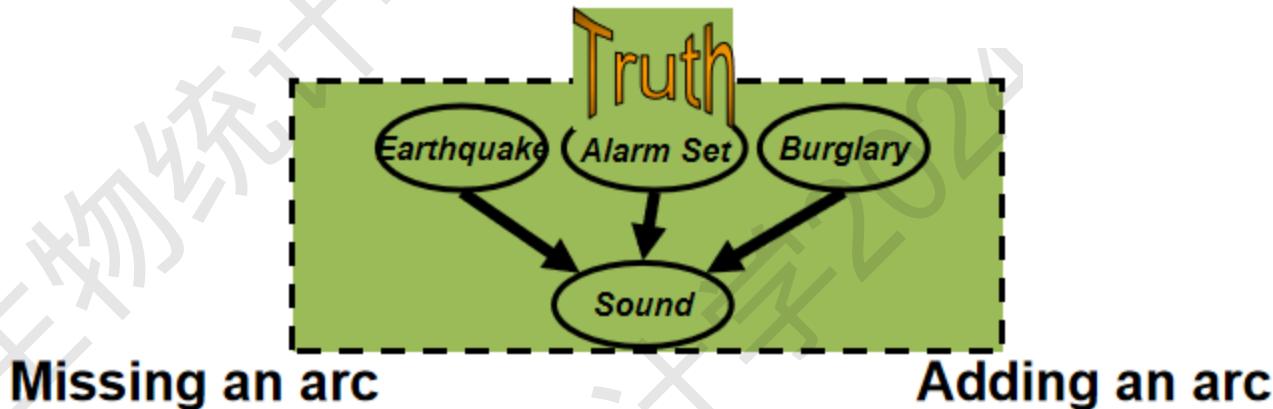
Bayesian Network

- 初级：参数学习
- 中级：图分解
- 高级：近似算法
- 特级：EM算法

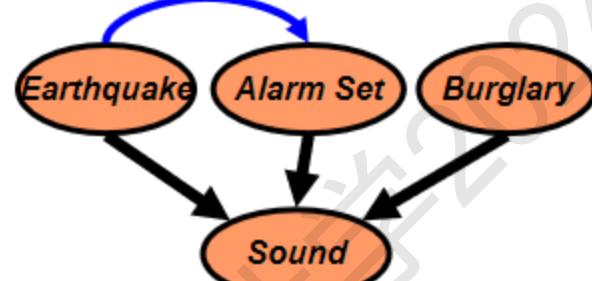
近似算法

- 从所有的网络结构空间进行搜索最优网络结构是一个NP问题，难以快速求解。
- 有两种常用的方法快速求解：
 - 贪心算法：假设现有结构为最优，每次调整一条边（增加、删除、改变方向）直到评分函数值最低为止
 - 直接通过网络结构增加约束来减少搜索空间，例如将网络结构限定为树形结构等

Why Struggle for Accurate Structure?



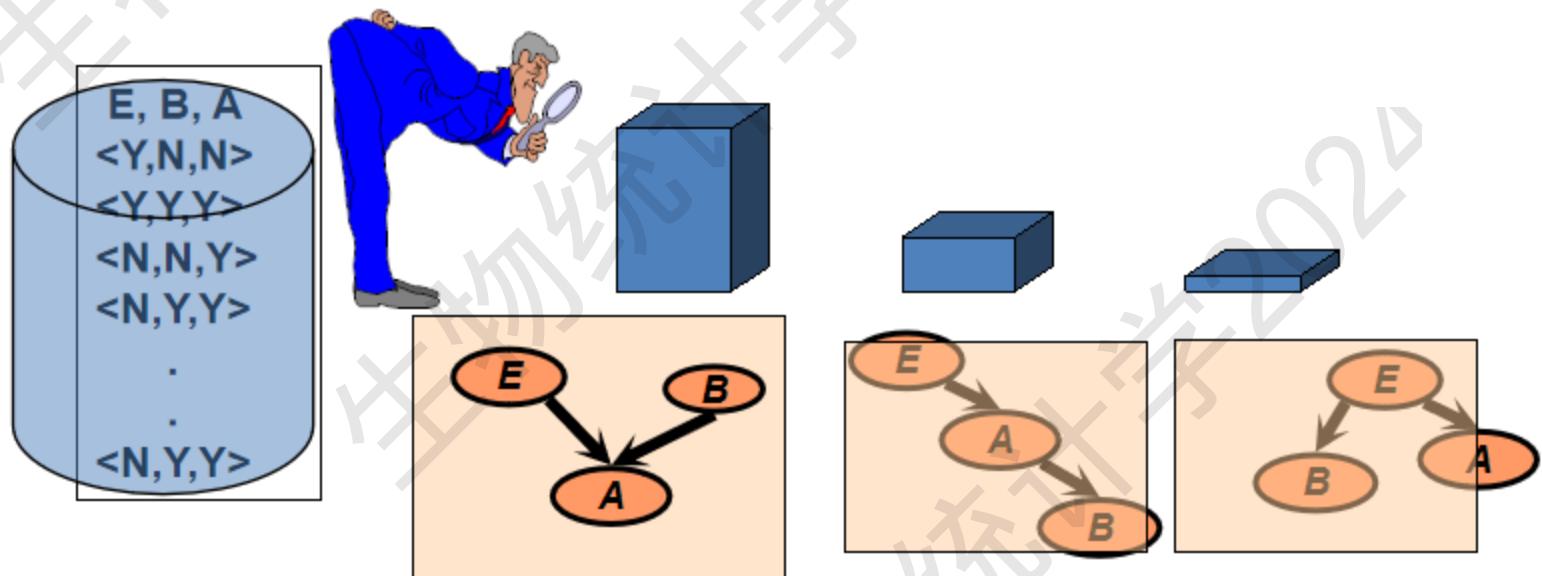
- Cannot be compensated for by fitting parameters
- Wrong assumptions about domain structure



- Increases the number of parameters to be estimated
- Wrong assumptions about domain structure

Score-based Learning

Define scoring function that evaluates how well a structure matches the data



Search for a structure that maximizes the score

Score Function I

Which structure is good?

- BDe scores (Heckman)

$$\text{BDe}(G : D) = \log \int P(D | G, \Theta)P(\Theta | G)d\Theta + \log P(G)$$

Marginal likelihood

Structure Prior

Marginal Likelihood (Multinomial Case)

- If data are complete, we can obtain the close form.

$$P(D|G) = \prod_{i=1}^n \prod_{j=1}^{q_i} \frac{\Gamma(\sum_{k=1}^{r_i} \alpha_{ijk})}{\Gamma(\sum_{k=1}^{r_i} \alpha_{ijk} + \sum_{k=1}^{r_i} N_{ijk})} \prod_{k=1}^{r_i} \frac{\Gamma(\alpha_{ijk} + N_{ijk})}{\Gamma(\alpha_{ijk})}$$

N_{ijk} : Number of cases where $X_i = k$, $P_{\alpha X_i} = j$

r_i : number of states of X_i

q_i : number of instance of parents of X_i .

Practical Consideration

Super exponential number (in the number of variables) of possible structures.

- How do we find the best graphs?
- How do we assign structure and parameter priors to all possible graphs?

Structure Prior Choice

- All possible structures are equally likely.
- Fix (or forbid) some arcs.
- Choosing a prior proportions to the similarity to a prior network structure.

Model Selection

- Theorem: finding the best BN structure among those structures with at most k parents in NP-hard problem ($k>1$).
- Heuristic searching
 - Greedy
 - MCMC

Score Function II

Which structure is good?

- BIC/MDL scores
 - BIC: Bayesian Information Criterion.
 - MDL: Minimum Description Length.

$$\text{BIC}(G, \Theta : D) = \log P(D | G, \Theta) - \frac{\log N}{2} \# \text{ param in } G$$

Fitness to data

Complexity regularization

Minimum Description Length Principle

- Universal coding.
 - Description length of the compressed form (model) of data.
 - Description length of the model itself used in the compression.

Minimum Description Length Principle (Cont.)

- Bayesian network case.
 - Modeling of data (Probability distribution).
 - Network coding (number of parameters).

See: N.Friedman. Learning Bayesian networks with local structure.

Decomposability

- Key property of the Bayesian network with complete data.

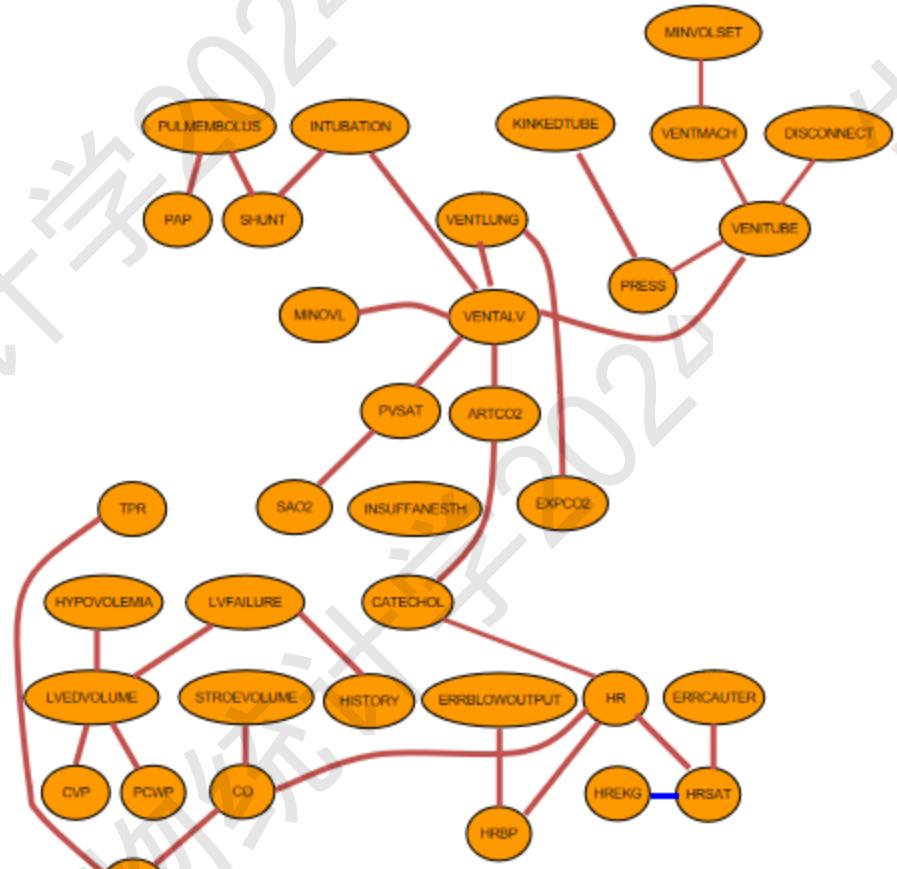
$$\text{score}(G) = \sum \text{score} (\text{ family of } X \text{ in } G)$$

Tree-structured Networks

Trees: At most one parent per variable.

Why trees?

- Elegant math=>we can solve the optimization problem
- Sparse parameterization to avoid over-fitting



Learning Trees

- Let $p(i)$ denote parent of X_i
- The Bayesian score can be written as sum of edge scores.

$$\begin{aligned}Score(G : D) &= \sum_i Score(X_i : Pa_i) \\&= \sum_i (Score(X_i : X_{p(i)}) - Score(X_i)) + \sum_i Score(X_i)\end{aligned}$$

Improvement over
“empty” network

Score of “empty”
network

Learning Tree

- Set edge weight as: $\text{Score}(X_j \rightarrow X_i) - \text{Score}(X_i)$.
- Well studied Problem in graph theory: Find the tree with maximum weight. It can be solved by maximum spanning tree algorithm (MST) in an efficient way.

Kruskal's Algorithm on MST

```
begin Kruskal;  
    sort the arcs in A in decreasing order of their weights;  
    LIST = Ø;  
    while |LIST| <  $n - 1$  do  
        begin  
            if the next arc does not create a cycle then add  
                it to LIST  
            else discard it  
        end;  
end;
```

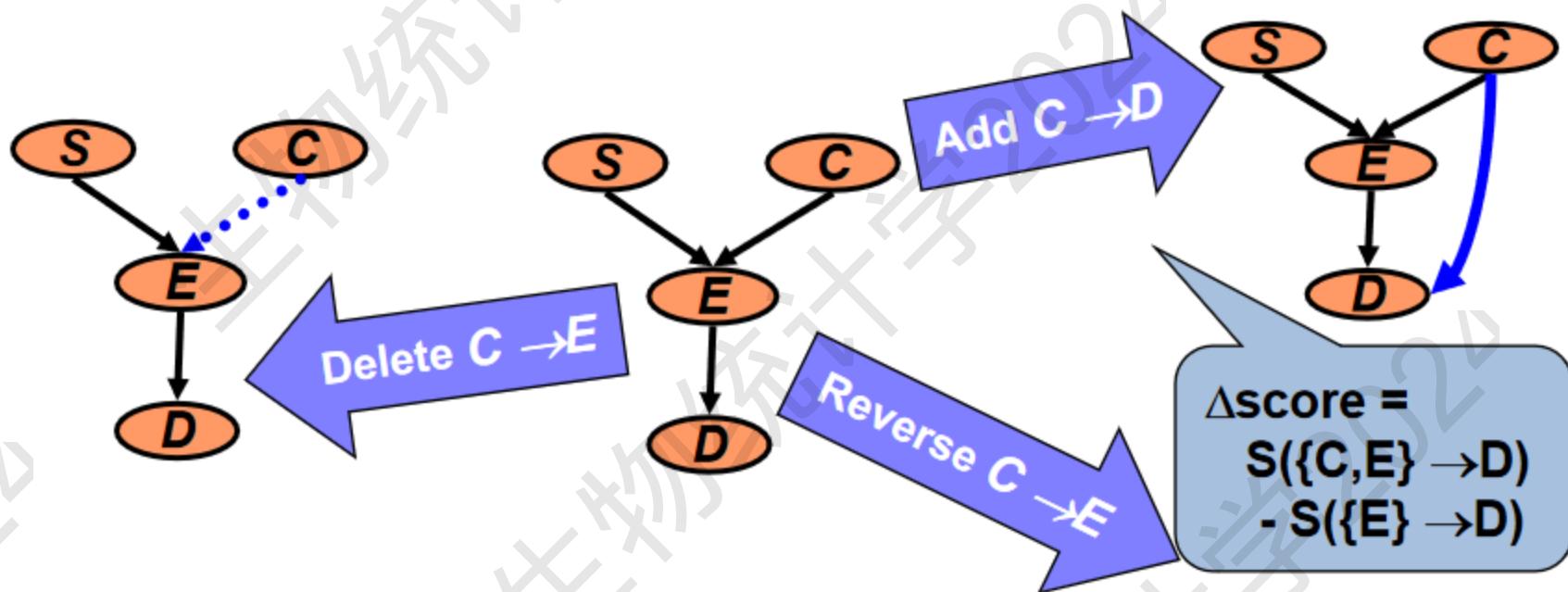
Heuristic Search: Beyond Trees

- Define a search space:
 - search states are possible structures
 - operators make small changes to structure
- Search techniques:
 - Greedy hill-climbing
 - Best first search
 - Simulated Annealing
 - ...

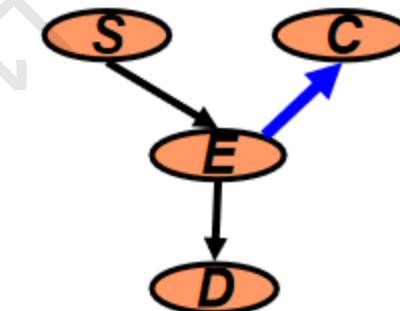
Local Search

- Start with a given network
 - empty network
 - best tree
 - a random network
- At each iteration
 - Evaluate all possible changes
 - Apply change based on score
- Stop when no modification improves score

Typical Operations In Heuristic Search



To update score after local change,
only re-score families that changed



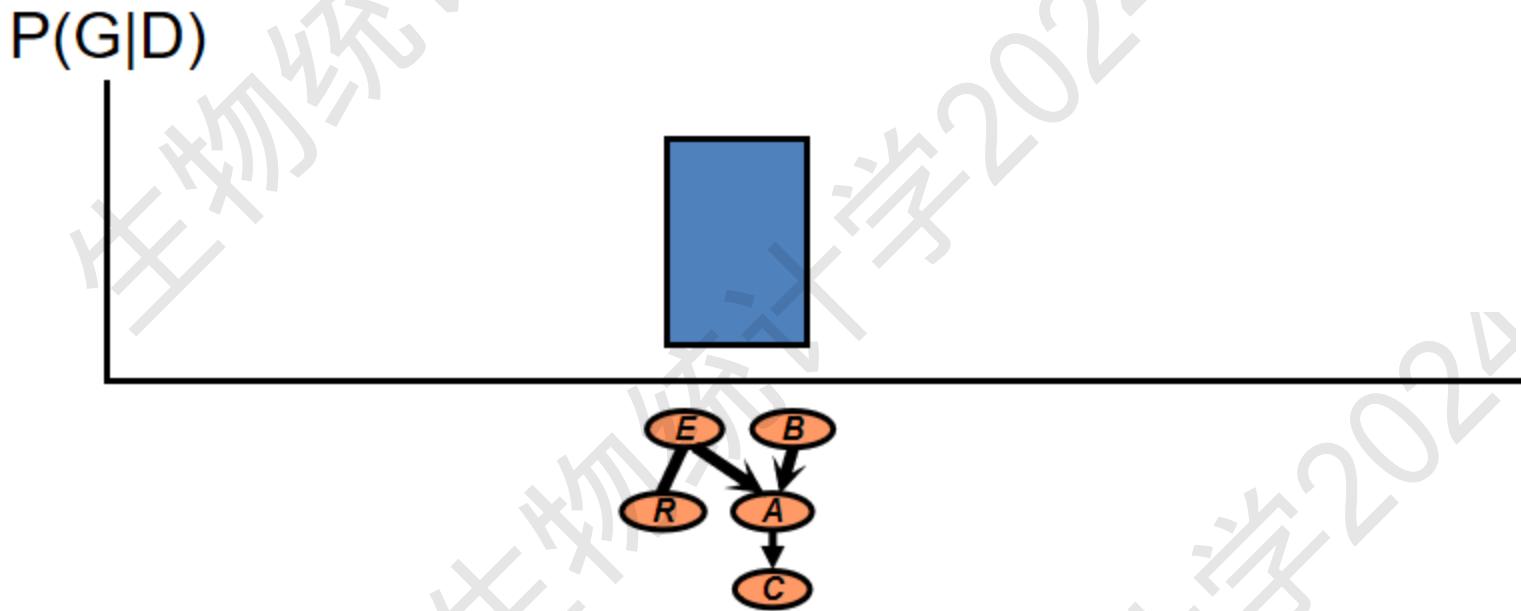
Local Search: Possible Pitfalls

- Local search can get stuck in:
 - **Local Maxima:**
 - All one-edge changes reduce the score
 - **Plateaus:**
 - Some one-edge changes leave the score unchanged

Escape From Traps

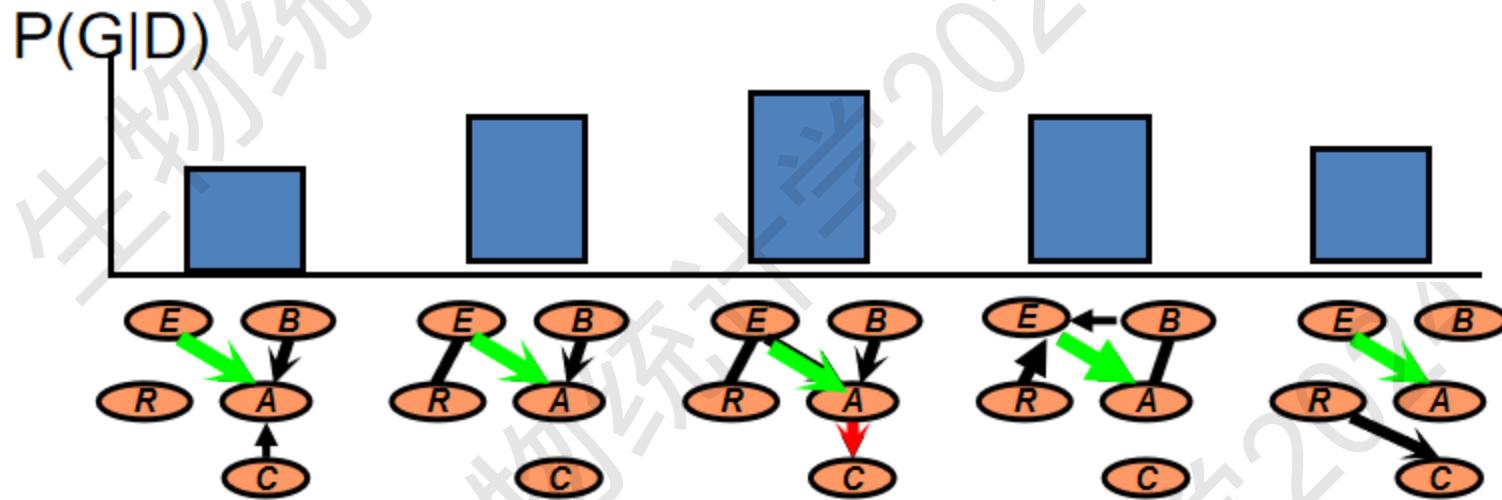
- Random restarts.
- Simulated annealing
 - Take the bad score with probability proportion to $\exp(\Delta\text{score}/t)$.
 - Cool down slowly.

Discovering Structure



- Current practice: model selection
 - Pick a single high-scoring model
 - Use that model to infer domain structure

Discovering Structure



Problem

- Small sample size \Rightarrow many high scoring models
- Answer based on one model often useless.
- We want features common to many models.

Bayesian Approach

- Posterior distribution over structures
- Estimate probability of **features**

- Edge $X \rightarrow Y$
- Path $X \rightarrow \dots \rightarrow Y$
- ...

$$P(f | D) = \sum_G f(G) P(G | D)$$

Feature of G ,
e.g., $X \rightarrow Y$

Indicator function
for feature f

Bayesian score
for G

Practical Implementation

- Bootstrap method.
 - Randomly generate m “perturbed” sample sets.
 - For each sample set, choose a best model G_i .
 - Average the feature among these m structures.

$$P(f(G) | D) \approx \frac{1}{n} \sum_{i=1}^n f(G_i)$$

C: Dealing With Missing Data

1. Structure known, how to learn the parameters?
2. Structure unknown, how to learn the structure and parameters?

Bayesian Network

- 初级：参数学习
- 中级：图分解
- 高级：近似算法
- 特级：EM算法

Incomplete Data

Data is often **incomplete**

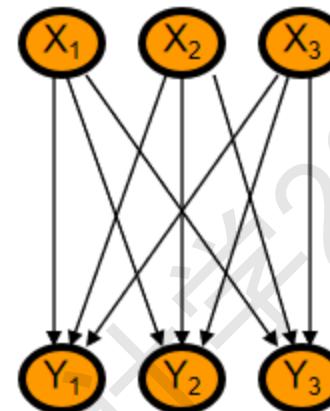
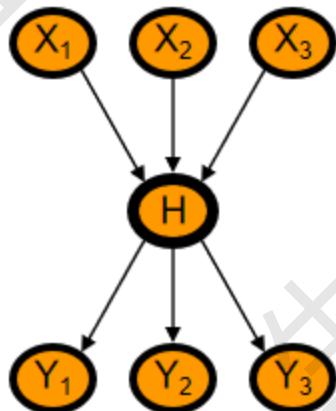
- Some variables of interest are not assigned values.

This phenomenon happens when we have

- **Missing values:**
 - Some variables unobserved in some instances
- **Hidden variables:**
 - Some variables are never observed
 - We might not even know they exist

Hidden (Latent) Variables

- Why should we care about unobserved variables?



17 parameters

$$17 = 1 + 1 + 1 + 8 + 2 + 2 + 2$$

27 parameters

$$27 = 1 + 1 + 1 + 8 + 8 + 8$$

More Computation

- The likelihood of the data does **not** decompose.
- Complete data.

$$\log L(\Theta : D = (x_1, \dots, x_n)) = \sum_i \log P(x_i | \text{Pa}(x_i))$$

- Incomplete data.

$$\log L(\Theta : D = (x_1, \dots, x_k)) = \log \sum_{x_{k+1}, \dots, x_n} \prod_i P(x_i | \text{Pa}(x_i))$$

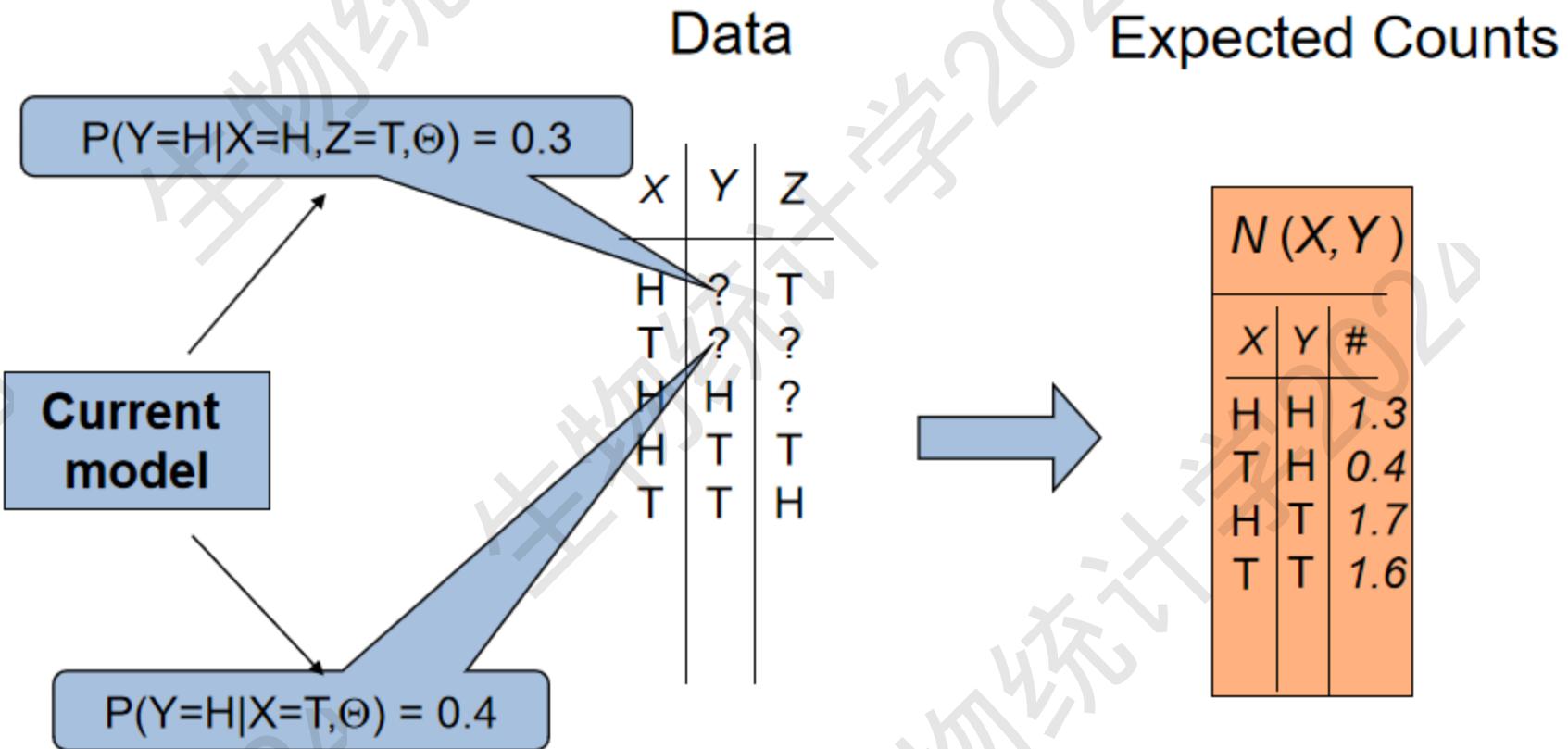
Learning Parameters With Incomplete Data

- Expectation maximization (EM) iteration algorithm is the general purpose method for learning from incomplete data.
 - E-Step.
 - M-Step.

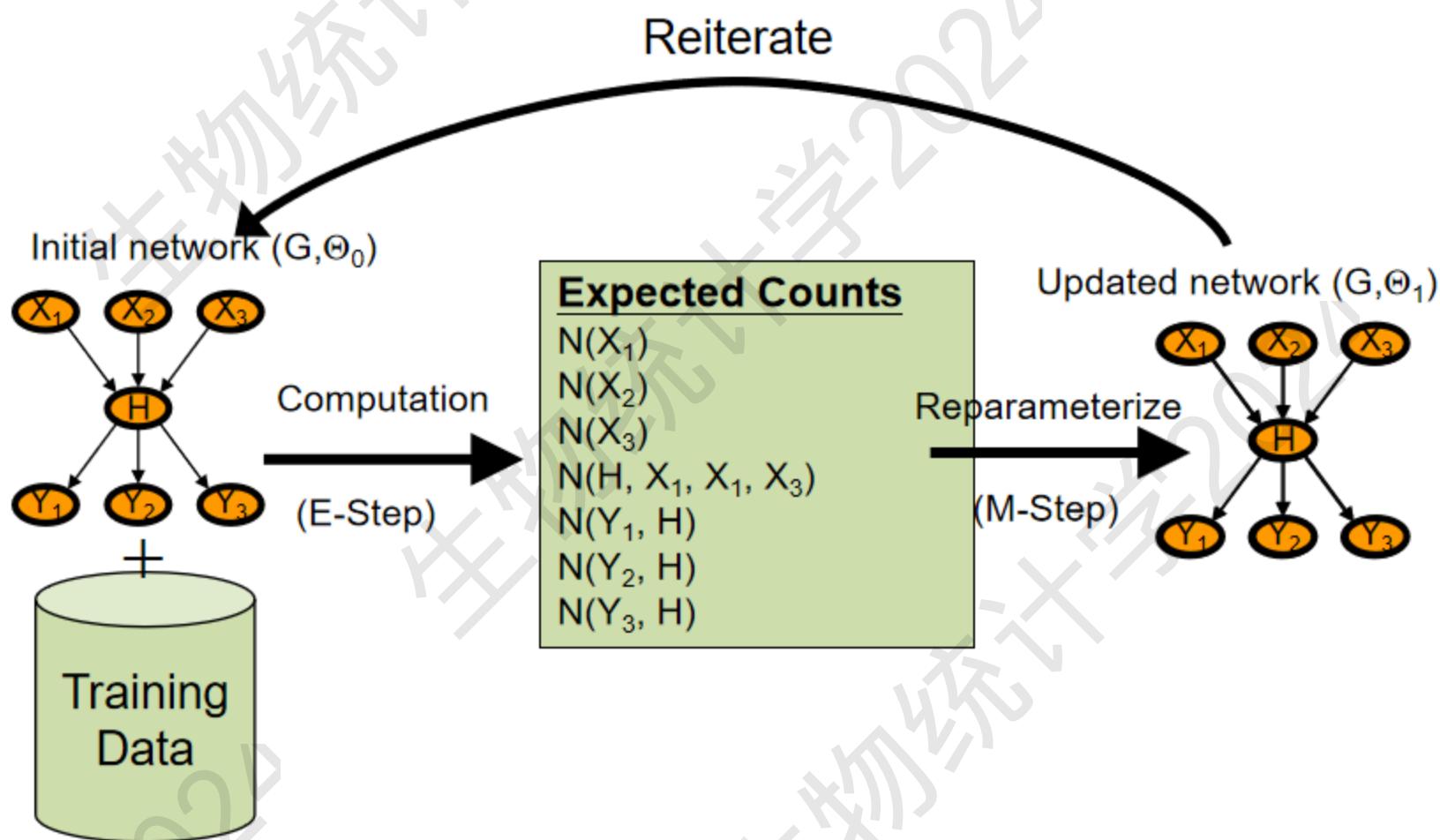
EM Intuition

- If we had true counts, we could estimate parameters.
- But with missing values, counts are unknown.
- We “complete” counts using probabilistic inference based on current parameter assignment.
- We use completed counts as if real to re-estimate parameters.

EM Algorithm



EM Algorithm (Cont.)



EM Algorithm (Cont.)

Formal Guarantees:

- $L(\Theta_1:D) \geq L(\Theta_0:D)$
 - Each iteration improves the likelihood
- If $\Theta_1 = \Theta_0$, then Θ_0 is a **stationary point** of $L(\Theta:D)$
 - Usually, this means a local maximum

Computational Bottleneck

Computation of expected counts in E-Step

- Need to compute posterior for each unobserved variable in each instance of training set.
- All posteriors for an instance can be derived from one pass of standard BN inference.

Summary: Parameter Learning With Incomplete Data

- Incomplete data makes parameter estimation hard
- Likelihood function
 - Does not have closed form
 - Is multimodal
- Finding maximum likelihood parameters:
 - EM
 - Gradient ascent
- Both exploit inference procedures for Bayesian networks to compute expected sufficient statistics

Incomplete Data: Structure Scores

Recall, Bayesian score:

$$\begin{aligned} P(G | D) &\propto P(G)P(D | G) \\ &= P(G) \boxed{\int P(D | G, \Theta)P(\Theta | G)d\theta} \end{aligned}$$

With incomplete data:

- Cannot evaluate **marginal likelihood** in closed form.
- We have to resort to **approximations**:
 - Evaluate score around MAP parameters
 - Need to find MAP parameters (e.g., EM)

Naïve Approach

- Perform EM for each candidate graph.
- Computationally expensive:
 - Parameter optimization via EM — non-trivial
 - Need to perform EM for all candidate structures
 - Spend time even on poor candidates
- In practice, considers only a few candidates.

Structural EM

Recall, in complete data we had

- Decomposition \Rightarrow efficient search.

Idea:

- Instead of optimizing the real score...
- Find **decomposable** alternative score.
- Such that maximizing new score \Rightarrow improvement in real score.

Structural EM (Cont.)

Idea:

- Use current model to help evaluate new structures

Outline:

- Perform search in (Structure, Parameters) space.
- At each iteration, use current model for finding either:
 - Better scoring parameters: “parametric” EM step.
 - Better scoring structure: “structural” EM step.

Structural EM Steps

Assume $B_0 = (G_0, \Theta_0)$ is “current” hypothesis.

Goal: Maximize **expected score**, given B_0

$$E[Score(B : D^+) | D, B_0] = \sum Score(B : D^+) P(D^+ | D, B_0)$$

where D^+ denotes **completed** data sets.

Theorem:(progress)

If $E[Score(B : D^+) | D, B_0] > E[Score(B_0 : D^+) | D, B_0]$

$$\Rightarrow Score(B : D) > Score(B_0 : D).$$

- This implies that by improving the expected score, we find networks that have higher objective score.

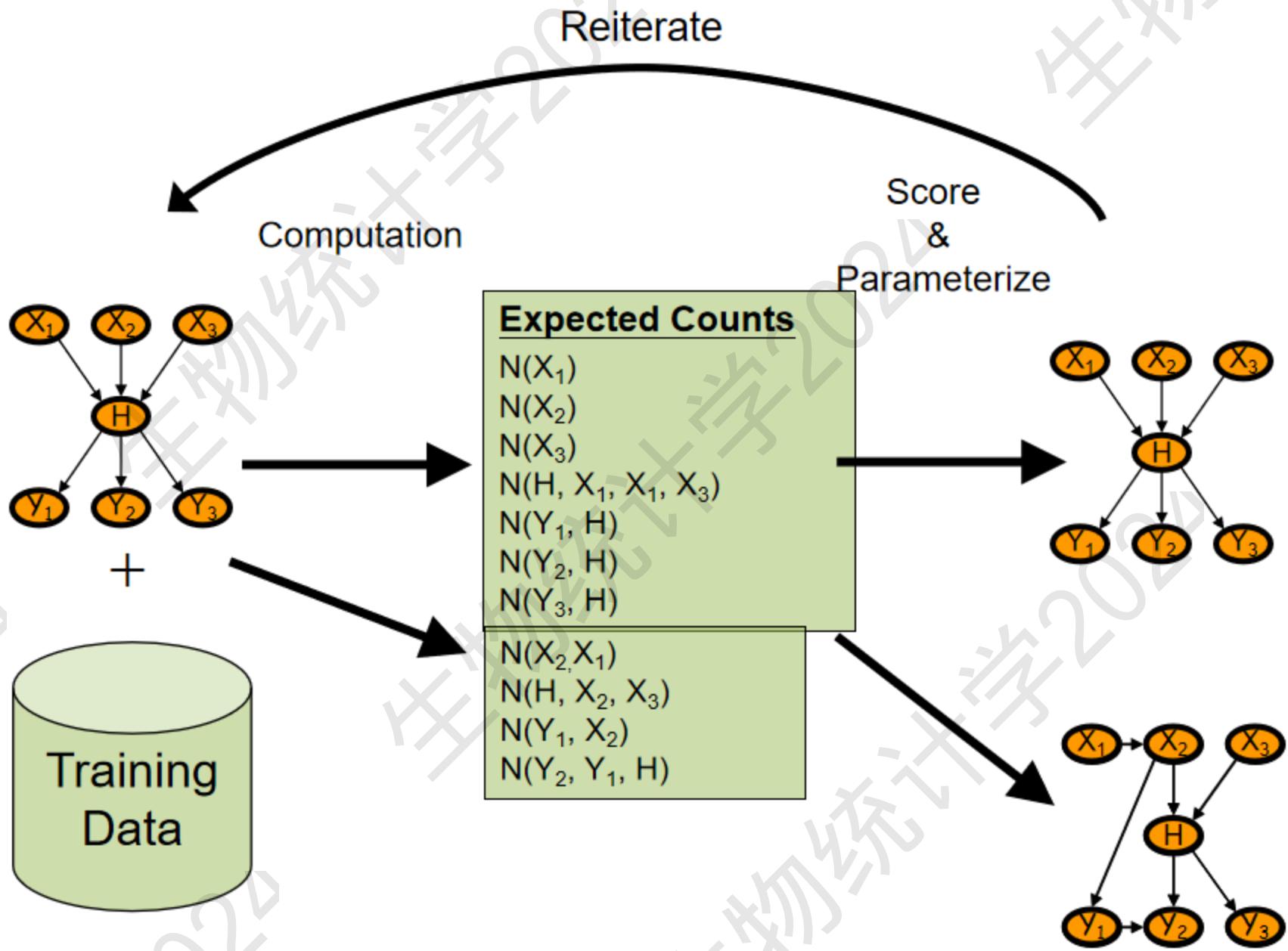
Structural EM for BIC/MDL

For the BIC/MDL score, we get that

$$\begin{aligned} \text{E}[\text{BIC}(B : D^+) | D, B_0] \\ &= \text{E}[\log P(D^+ | B) | D, B_0] - \text{Penalty}(B) \\ &= E\left[\sum_i N(X_i, Pa_i) \log P(X_i | Pa_i) | D, B_0\right] - \text{Penalty}(B) \\ &= \sum_i E[N(X_i, Pa_i) | D, B_0] \log P(X_i | Pa_i) - \text{Penalty}(B) \end{aligned}$$

Consequence:

- We can use complete-data methods, where we use expected counts, instead of actual counts.



The Structural EM Procedure

Input: $B_0 = (G_0, \Theta_0)$

loop for $n = 0, 1, \dots$ until convergence

Improve parameters:

$\Theta^*_{n+1} = \text{Parametric-EM}(G_n, \Theta_n)$

let $B^*_{n+1} = (G_{n+1}, \Theta^*_{n+1})$

Improve structure:

Search for a network $B_{n+1} = (G_{n+1}, \Theta_{n+1})$ s.t.

$E[\text{Score}(B_{n+1}; D) | B^*_{n+1}] > E[\text{Score}(B^*_{n+1}; D) | B^*_{n+1}]$

- Parametric-EM() can be replaced by Gradient Ascent, Newton-Raphson methods, or accelerated EM.
- Early stopping parameter optimization stage avoids “entrenchment” in current structure.

Applications

准备工作阶段

确定特征属性

获取训练样本

对每个类别计算 $p(y_i)$

对每个特征属性计算所有划分的条件概率

分类器训练阶段

以 $p(x|y_i)p(y_i)$
最大项作为 x 所属类别

对每个类别计算
 $p(x|y_i)p(y_i)$

应用阶段

App1: Expression Data Analysis

Reference:

- N.Friedman et al. Using Bayesian Networks to analyze expression data. *J. Comput. Biol.*, 7:601-620, 2000.
- A.Hartemink et al. Combining location and expression data for principled discovery of genetic regulatory network models. PSB 2002.

Motivation

- Extract meaningful information from gene expression data.
 - Infer regulatory mechanism.
 - Reveal function of proteins.
 -

Case 1: Cell-cycle Data

- Yeast cell-cycle data (P.Spellman, *Mol. Biol. of the cell*, 1998).
- 7 time series under different cell cycle synchronization methods (alpha, beta factor, CDC15, CDC24, CDC28,cln2,3).
- 6177 ORFs, 77 time points.
- 800 genes are identified related to cell cycle process (big variation).

Bayesian Network Model

- Random Variables
 - Individual genes
 - Experimental condition
 - Cell phase.
- Discretization: 3 levels, -1,0,1, depending on whether the expression level is significantly lower than, similar to, great than the respective control. However, this may not be necessary (For continuous variable, a linear Gaussian conditional model can be used).

Learning Bayesian Network (Cont.)

- Sparse candidate algorithm: identify small number of candidate parents for each gene based on simple local statistics (such as mutual information).
- Bootstrap confidence estimation:
 - Use re-sampling to generate perturbations of training data.
 - Use the number of times of feature is repeated among networks from these datasets to estimate confidence of Bayesian network features.

Sparse Candidate Algorithm

Input:

- A data set $D = \{\mathbf{x}^1, \dots, \mathbf{x}^N\}$,
- An initial network B_0 ,
- A decomposable score
 $\text{Score}(B \mid D) = \sum_i \text{Score}(X_i \mid \text{Pa}^B(X_i), D)$,
- A parameter k .

Output: A network B .

Loop for $n = 1, 2, \dots$ until convergence

Restrict

Based on D and B_{n-1} , select for each variable X_i a set C_i^n ($|C_i^n| \leq k$) of candidate parents.

This defines a directed graph $H_n = (\mathcal{X}, E)$, where
 $E = \{X_j \rightarrow X_i \mid \forall i, j, X_j \in C_i^n\}$.

(Note that H_n is usually cyclic.)

Maximize

Find network $B_n = (G_n, \Theta_n)$ maximizing
 $\text{Score}(B_n \mid D)$ among networks that satisfy $G_n \subseteq H_n$ (i.e., $\forall X_i, \text{Pa}^{G_n}(X_i) \subseteq C_i^n$).

Return B_n

Figure 1: Outline of the *Sparse Candidate* algorithm

Estimate Feature Significance Bootstrap Method

- For $i = 1 \dots m$ (in our experiments, we set $m = 200$).
 - Re-sample with replacement N instances from D . Denote by D_i the resulting dataset.
 - Apply the learning procedure on D_i to induce a network structure G_i .
- For each feature f of interest calculate

$$\text{conf}(f) = \frac{1}{m} \sum_{i=1}^m f(G_i)$$

where $f(G)$ is 1 if f is a feature in G , and 0 otherwise.

Markov Relation

- Pairs with 80% confidence were evaluated against original clustering.
 - 70% of these were intra-cluster.
 - The rest show interesting inter-cluster relations.
- Most pairs are functionally related.

Markov Relation (Cont.)

Table 2: List of top Markov relations

Confidence	Gene 1	Gene 2	notes
1.0	YKL163W-PIR3	YKL164C-PIR1	Close locality on chromosome
0.985	PRY2	YKR012C	No homolog found
0.985	MCD1	MSH6	Both bind to DNA during mitosis
0.98	PHO11	PHO12	Both nearly identical acid phosphatases
0.975	HHT1	HTB1	Both are Histones
0.97	HTB2	HTA1	Both are Histones
0.94	YNL057W	YNL058C	Close locality on chromosome
0.94	YHR143W	CTS1	Homolog to EGT2 cell wall control, both do cytokinesis
0.92	YOR263C	YOR264W	Close locality on chromosome
0.91	YGR086	SIC1	
0.9	FARI	ASH1	Both part of a mating type switch, expression uncorelated
0.89	CLN2	SVS1	Function of SVS1 unknown, possible regulation mediated through SWI6
0.88	YDR033W	NCE2	Homolog to transmembrane proteins, suggesting both involved in protein secretion
0.86	STE2	MFA2	A mating factor and receptor
0.85	HHF1	HHF2	Both are Histones
0.85	MET10	ECM17	Both are sulfite reductases
0.85	CDC9	RAD27	Both participate in Okazaki fragment processing

Order Relation

- Dominant gene: genes are indicative or potential source of the cell-cycle process.
- Dominance score: describing how strong that one gene can be the ancestor of other genes in the network.

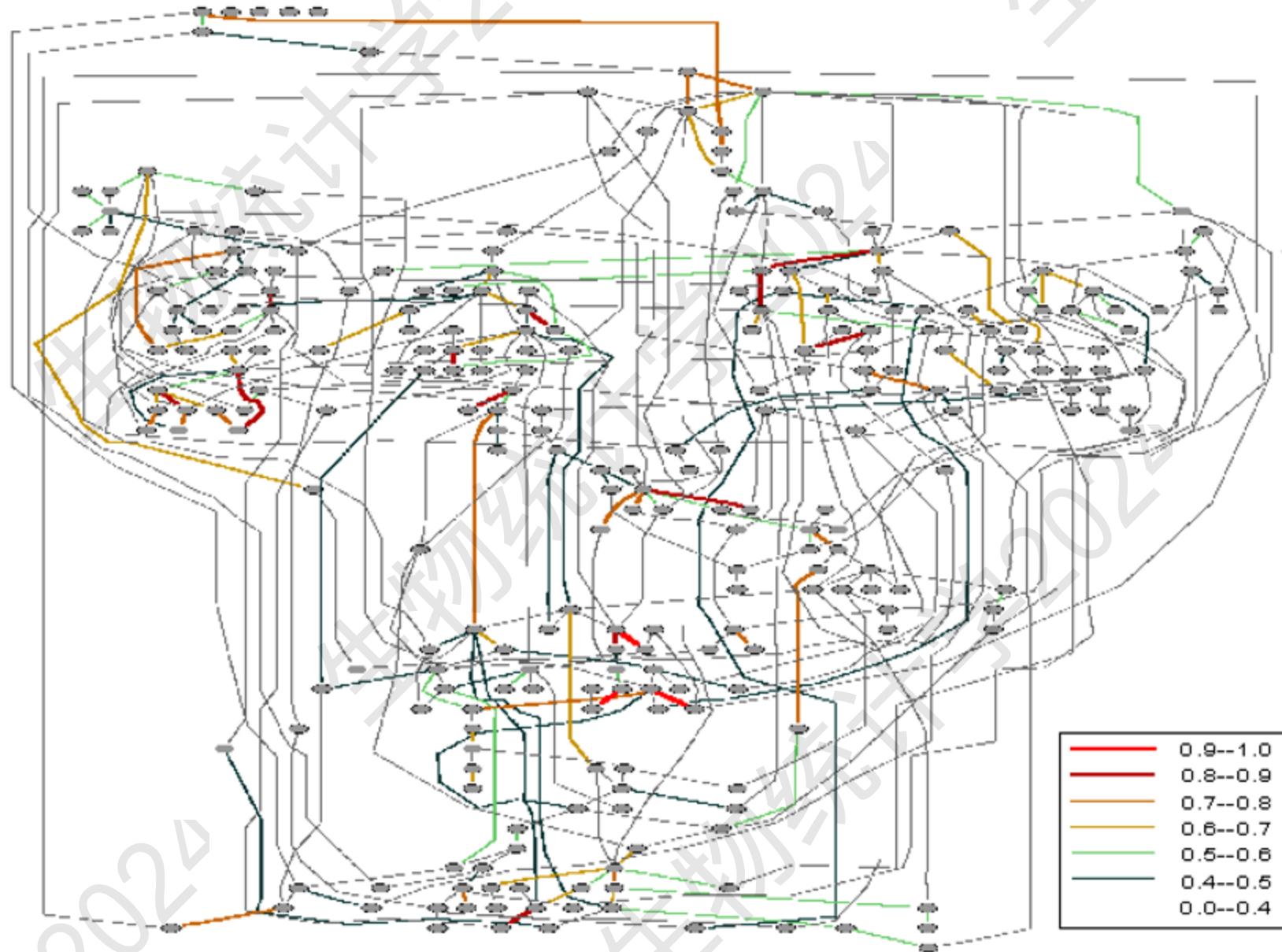
Dominant Genes

Table 1: List of dominant genes in the ordering relations (top 14 out of 30)

Gene/ORF	Dominance Score	# of descendant genes		notes
		> .8	> .7	
YLR183C	551	609	708	Contains forkheaded assosiated domain, thus possibly nuclear
MCD1	550	599	710	Mitotic chromosome determinant, null mutant is inviable
CLN2	497	495	654	Role in cell cycle START, null mutant exhibits G1 arrest
SRO4	463	405	639	Involved in cellular polarization during budding
RFA2	456	429	617	Involved in nucleotide excision repair, null mutant is inviable
YOL007C	444	367	624	
GAS1	433	382	586	Glycophospholipid surface protein, Null mutant is slow growing
YOX1	400	243	556	Homeodomain protein that binds leu-tRNA gene
YLR013W	398	309	531	
POL30	376	173	520	Required for DNA replication and repair, Null mutant is inviable
RSR1	352	140	461	GTP-binding protein of the ras family involved in bud site selection
CLN1	324	74	404	Role in cell cycle START, null mutant exhibits G1 arrest
YBR089W	298	29	333	
MSH6	284	7	325	Required for mismatch repair in mitosis and meiosis

Cell cycle control and initiation: CLN1, CLN2, CDC5.

Network Learned



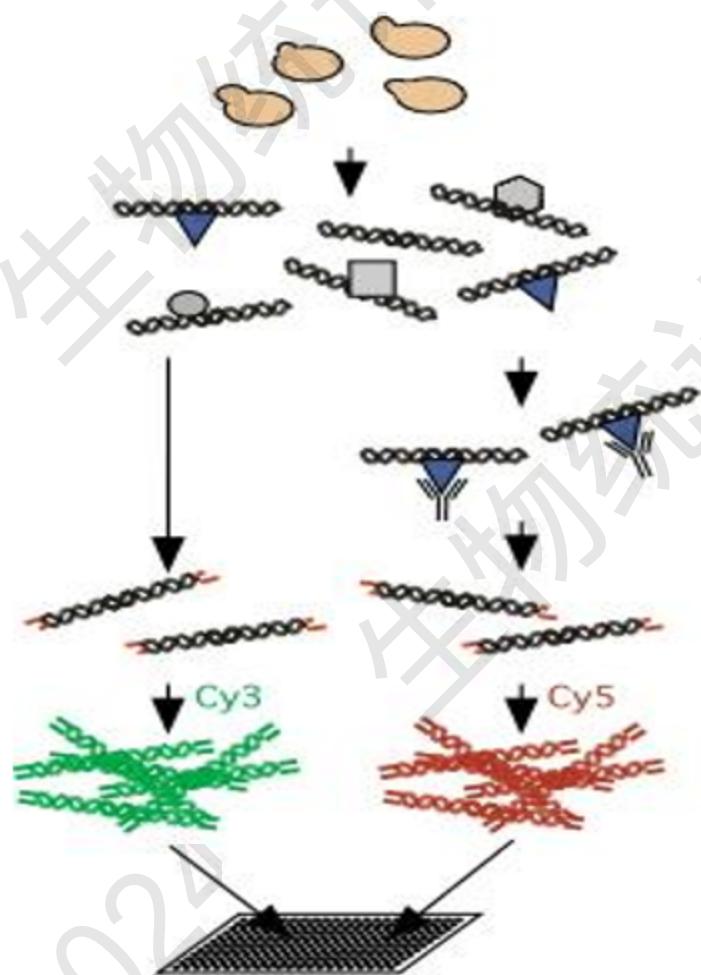
Case 2: Pheromone and Mating Response

- 6135 genes, 320 samples under different conditions.
- 32 genes are selected.
 - Pheromone response signaling pathway.
 - Mating response.
- Location data (transcription factor and DNA binding experiment, chip-chip data) are included as prior constraints.

Genes Selected

Gene	Color Mnemonic	Function of Corresponding Protein
STE2	magenta	transmembrane receptor peptide (present only in MAT α strains)
STE3	red	transmembrane receptor peptide (present only in MAT α strains)
GPA1	green	component of the heterotrimeric G-protein (G α)
STE4	green	component of the heterotrimeric G-protein (G β)
STE18	green	component of the heterotrimeric G-protein (G γ)
FUS3	blue	mitogen-activated protein kinase (MAPK)
STE7	yellow	MAPK kinase (MAPKK)
STE11	yellow	MAPKK kinase (MAPKKK)
STE25	yellow	scaffolding peptide holding together Fus3, Ste7, and Ste11 in a large complex
STE12	blue	transcriptional activator
KSS1	orange	alternative MAPK for pheromone response (in some dispute)
STE20	orange	p21-activated protein kinase (PAK)
STE20	orange	unknown function but necessary for proper function of Ste11
MFA1	magenta	a-factor mating pheromone (present only in MAT α strains)
MFA2	magenta	a-factor mating pheromone (present only in MAT α strains)
MFALPHA1	red	a-factor mating pheromone (present only in MAT α strains)
MFALPHA2	red	a-factor mating pheromone (present only in MAT α strains)
STE6	magenta	responsible for the export of a-factor from MAT α cells (present only in MAT α strains)
FAR1	blue	substrate of Fus3 that leads to G1 arrest; known to bind to STE4 as part of complex of proteins necessary for establishing cell polarity required for shmoo formation after mating signal has been received
FUS1	blue	required for cell fusion during mating
AGA1	blue	anchor subunit of a-agglutinin complex; mediates attachment of Aga2 to cell surface
AGA2	magenta	binding subunit of a-agglutinin complex; involved in cell-cell adhesion during mating by binding Sag1 (present only in MAT α strains)
SAG1	red	binding subunit of a-agglutinin complex; involved in cell-cell adhesion during mating by binding Aga2 (present only in MAT α strains; also known as Aga1)
BAR1	magenta	protease degrading a-factor (present only in MAT α strains)
SST2		involved in desensitization to mating pheromone exposure
KAR3		essential for nuclear migration step of karyogamy
TEC1		transcriptional activator believed to bind cooperatively with Ste12 (more active during induction of filamentous or invasive growth response)
MCM1		transcription factor believed to bind cooperatively with Ste12 (more active during induction of pheromone response)
SIN3		implicated in induction or repression of numerous genes in pheromone response pathway
TUP1		implicated in repression of numerous genes in pheromone response pathway
SNF2	aqua	implicated in induction of numerous genes in pheromone response pathway (component of SWI-SNF global transcription activator complex)
SWI1	aqua	implicated in induction of numerous genes in pheromone response pathway (component of SWI-SNF global transcription activator complex)

Location Analysis (Chip-chip)



- Crosslink protein to DNA in vivo with formaldehyde
- Break open cells and shear DNA
- Immunoprecipitate
- Reverse-crosslinks, blunt DNA and ligate to unidirectional linkers
- LM-PCR
- Hybridize to array

<http://inside.wi.mit.edu/young/pub/locationanalysis.html>

Bayesian Network Model

- Random variables
 - 32 genes.
 - Mating type (Mata, Mata).
- Discrimination: to 4 levels while preserving over 98% of the original total mutual information between pairs of genes.
- Location data: set the constraints specifying which edges are required to be present and which are required to be absent.

Learning Bayesian Network

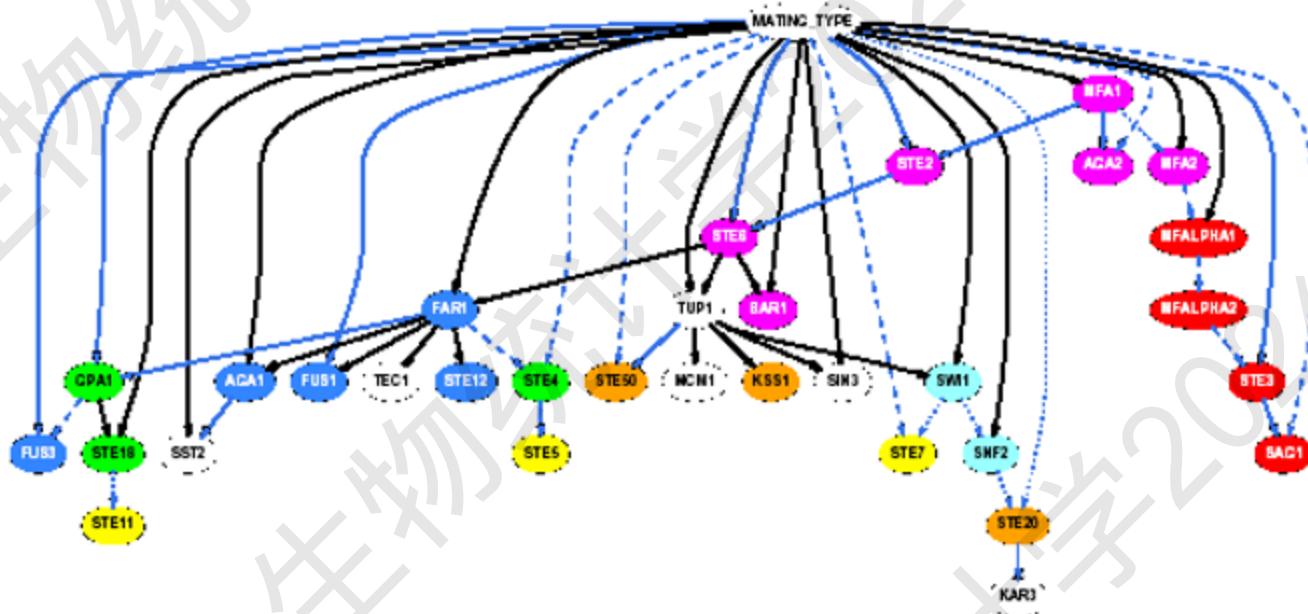
- Score: Bayesian score metric (BSM).
- Local heuristic searching algorithm: simulated annealing.
- Caching: keeping the top 500 structures recorded.
- Feature induction: Average features within top 500 structures.

Learning Bayesian Network (Cont.)

$$\begin{aligned} p(E_{XY}|D) &= \sum_S p(E_{XY}|D, S) \cdot p(S|D) \\ &= \sum_S 1_{XY}(S) \cdot e^{\text{BSM}(S)} \end{aligned}$$

Approximation:

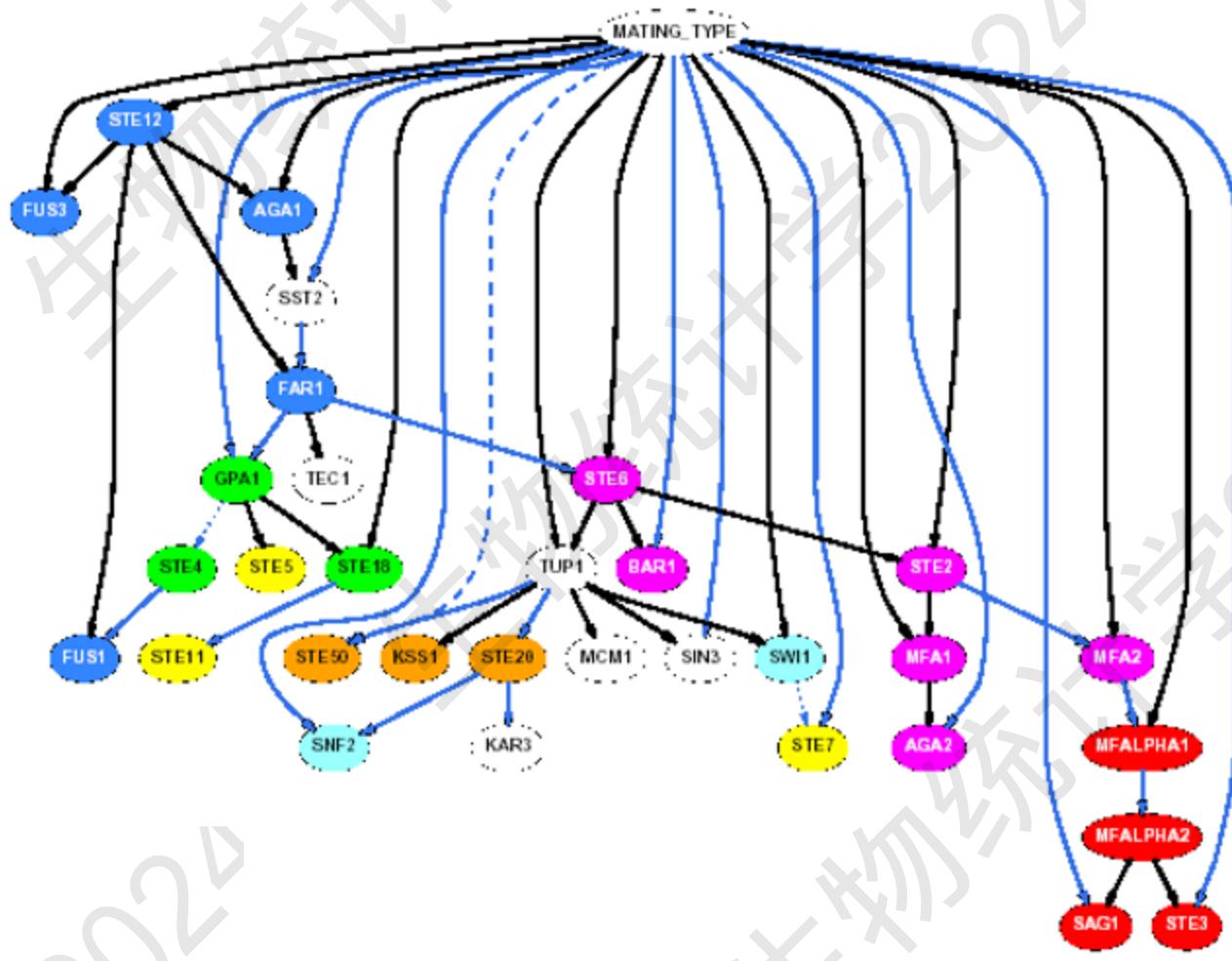
Learned Network Without Constraint



Node color: Different function.

Edge color: Solid black (0.99-1.0), dash blue (0.75-0.99), dot blue (0.5-0.75).

Learned Network With Constraints



Constraints included:

STE1
2

FUS
1

FUS
3

AGA
1

App2. Bayesian Classifier

- Reference:
 - N.Friedman. Building classifier using Bayesian networks. Proc. NCAI 1277-1284, 1996.
 - O.D.King et al. Predicting Gene Function From Patterns of Annotation. *Genome Research* 13: 896-904, 2003.

Basic Problem

- Given a dataset
 - Here X_i stands for the training data, c stands for the class label, assuming we have m classes,
 - We estimate the probability.

$$P(C_i | X), i=1,2,\dots,m$$

- The classifier is then denoted by:

$$\arg \max_i P(C_i | X)$$

How can we estimate the posterior probability?

Naïve Bayesian Network

- Assumption: all the variables are independent, given the class label.
- Joint distribution. $P((v_1, v_2 \dots v_{m-1}, v_m) | C) = \prod_i P(v_i | C)$

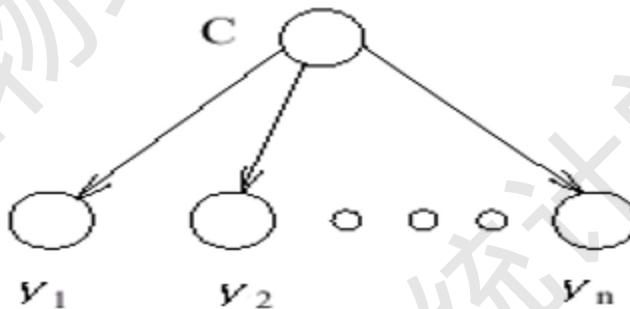


Figure 1. The structure of the naive Bayes network.

Tree Argumented Naive Bayes (TAN) Model

- Bayesian network with the class as the root, will each attribute's parent set contain class and at most one other attribute.

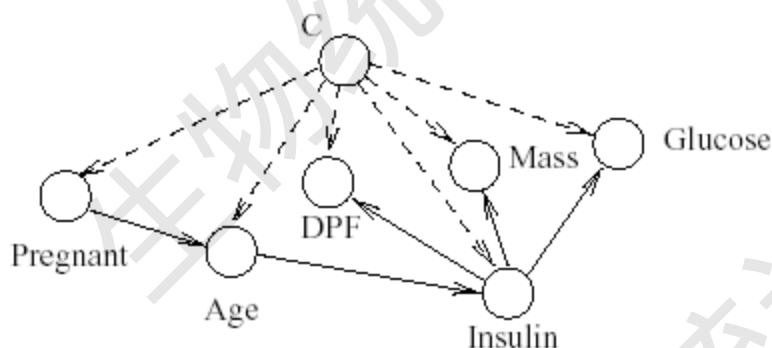


Figure 3. A TAN model learned for the data set "pima." The dashed lines are those edges required by the naive Bayesian classifier. The solid lines are correlation edges between attributes.

GO Function Prediction

- Motivation: GO is the controlled vocabulary of gene functions. Predict gene function by the pattern of annotation.
- Idea: If the annotation of two attribute tend to occur together in the database, then a gene holding one attribute is likely to hold the other as well.

Gene Ontology Structure

AmiGO

Search GO:
 Terms Gene Products

[Top Docs](#) [Gene Ontology](#) [GO Links](#) [GO Summary](#)

GO:0003673 : Gene Ontology (6448)

- ⓘ [GO:0008150 : biological process \(6447\)](#)
 - ⓘ [GO:0007610 : behavior \(1\)](#)
 - ⓘ [GO:0000004 : biological process unknown \(1893\)](#)
 - ⓘ [GO:0009987 : cellular process \(2257\)](#)
 - ⓘ [GO:0007154 : cell communication \(175\)](#)
 - ⓘ [GO:0008219 : cell death \(24\)](#)
 - ⓘ [GO:0030154 : cell differentiation \(0\)](#)
 - ⓘ [GO:0008151 : cell growth and/or maintenance \(2168\)](#)
 - ⓘ [GO:0006928 : cell motility \(0\)](#)
 - ⓘ [GO:0006944 : membrane fusion \(33\)](#)
 - ⓘ [GO:0007275 : development \(341\)](#)
 - ⓘ [GO:0008371 : obsolete biological process \(0\)](#)
 - ⓘ [GO:0007582 : physiological processes \(4520\)](#)
 - ⓘ [GO:0016032 : viral life cycle \(1\)](#)
 - ⓘ [GO:0005575 : cellular component \(6436\)](#)
 - ⓘ [GO:0003674 : molecular function \(6435\)](#)

Formalization

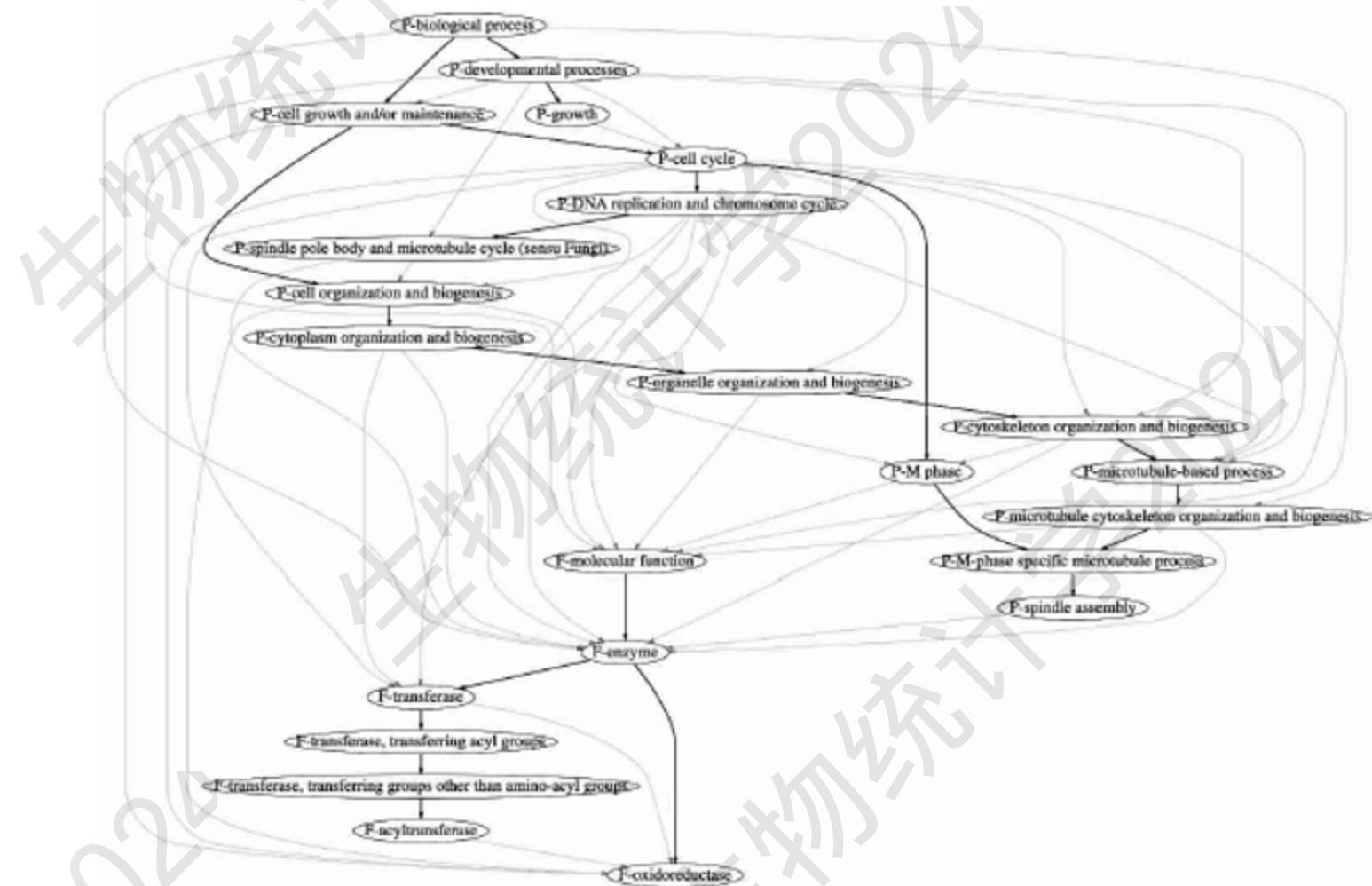
- GO attributes j . X_j indicate function. $X_j(i) = 1$ if gene is annotated with j .
- Attribution set $\text{nad}(X_j)$: neither ancestor nor descendant attribute of one attribute j in the GO DAG.
- The task is to estimate the probability

$$q(i,j) = \Pr(X_j = 1 \mid \mathbf{nad}(X_j) = \mathbf{nad}(X_j)(i))$$

Bayesian Network Model

- Nodes: GO attribute covers more than 10 genes, and no descendant covers more than 10 genes.
 - SGD, 170.
 - FlyBase, 218.
- Constraints: just considering those structures logically consist with GO DAG.

Fragment of Learned Bayesian Network



Further Reading

- N.Friedman et al. A structural EM algorithm for phylogenetic inference. *RECOMB2001*.
- E.Segal et al. From promoter sequence to gene expression data. *RECOMB2002*.
- E.Segal. Regulatory module. *Nature Genetics* 34: 2003.

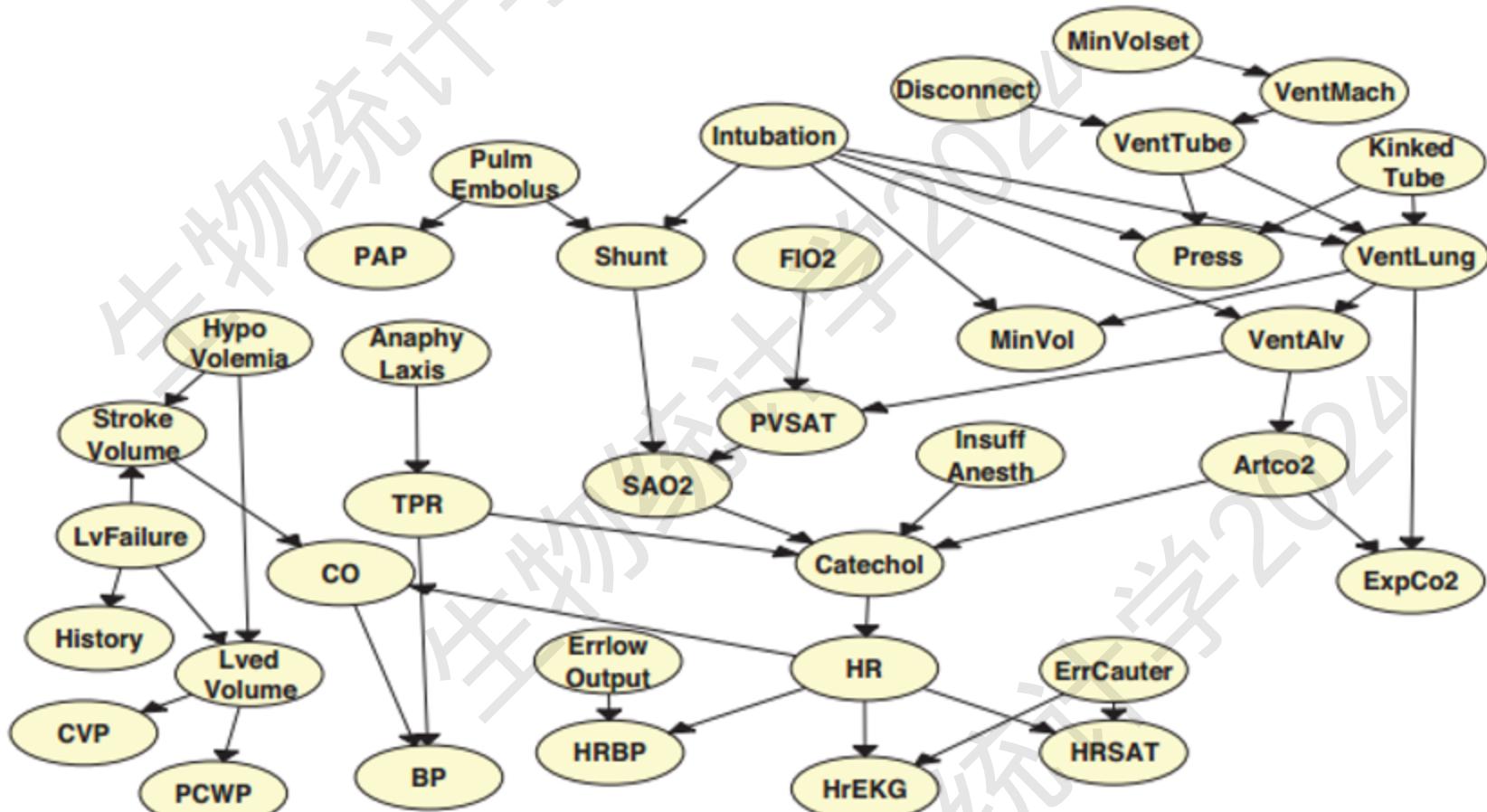
Bayesian Network Sources

- Peoples
 - N.Friedman <http://www.cs.huji.ac.il/~nir/>
 - D.Heckman
<http://www.research.microsoft.com/~heckerman/>
 - J. PEARL http://bayes.cs.ucla.edu/jp_home.html
 - F.V.Jensen <http://www.cs.auc.dk/~fvj/>
 -

Bayesian Network Sources

- Bayesian Network Repository
<http://www.cs.huji.ac.il/labs/compbio/Repository/.>
- Systems
 - Bayesian Networks Software Package listing
<http://www.cs.berkeley.edu/~zuwhan/bn.html.>
 - Microsoft Belief Network Tools
<http://www.research.microsoft.com/research/dtg/msbn/>
 - Hugin <http://hugin.dk/>
 -

Case 3: ICU predictions



(a)

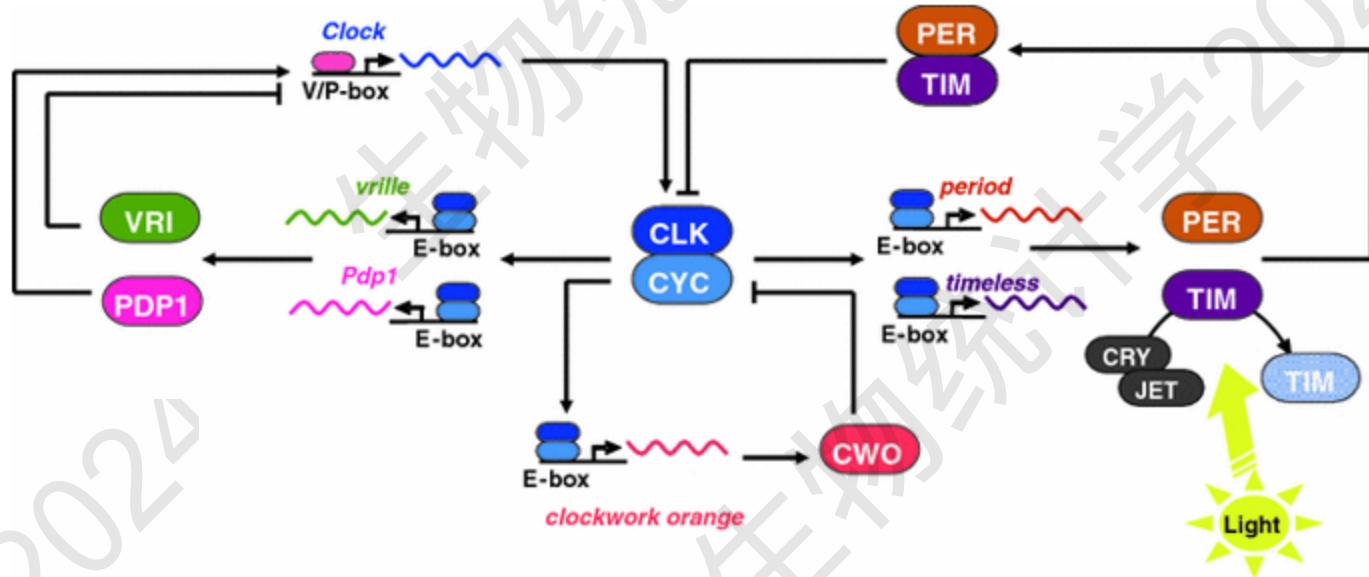
概率专家系统

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- N.Friedman. Learning bayesian networks with local structure.
- D.Heckman. Bayesian Networks for data mining. Data Mining and Knoledge Discovery **1**: 79-119, 1997.
- N.Friedman. Using bayesian networks to analyze expression data. J. Comp. Biol. 2002.
- A.Hartemink Combining location and expression data for principled discovery of genetic regulatory network models. *PSB2002*.
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- Novershtern N, Subramanian A, Lawton LN, Mak RH, Haining WN, McConkey ME, Habib N, Yosef N, Chang CY, Shay T, Frampton GM, Drake AC, Leskov I, Nilsson B, Preffer F, Dombkowski D, Evans JW, Liefeld T, Smutko JS, Chen J, Friedman N, Young RA, Golub TR, Regev A, Ebert BL. Densely interconnected transcriptional circuits control cell states in human hematopoiesis. *Cell*. 144(2):296-309. 2011.

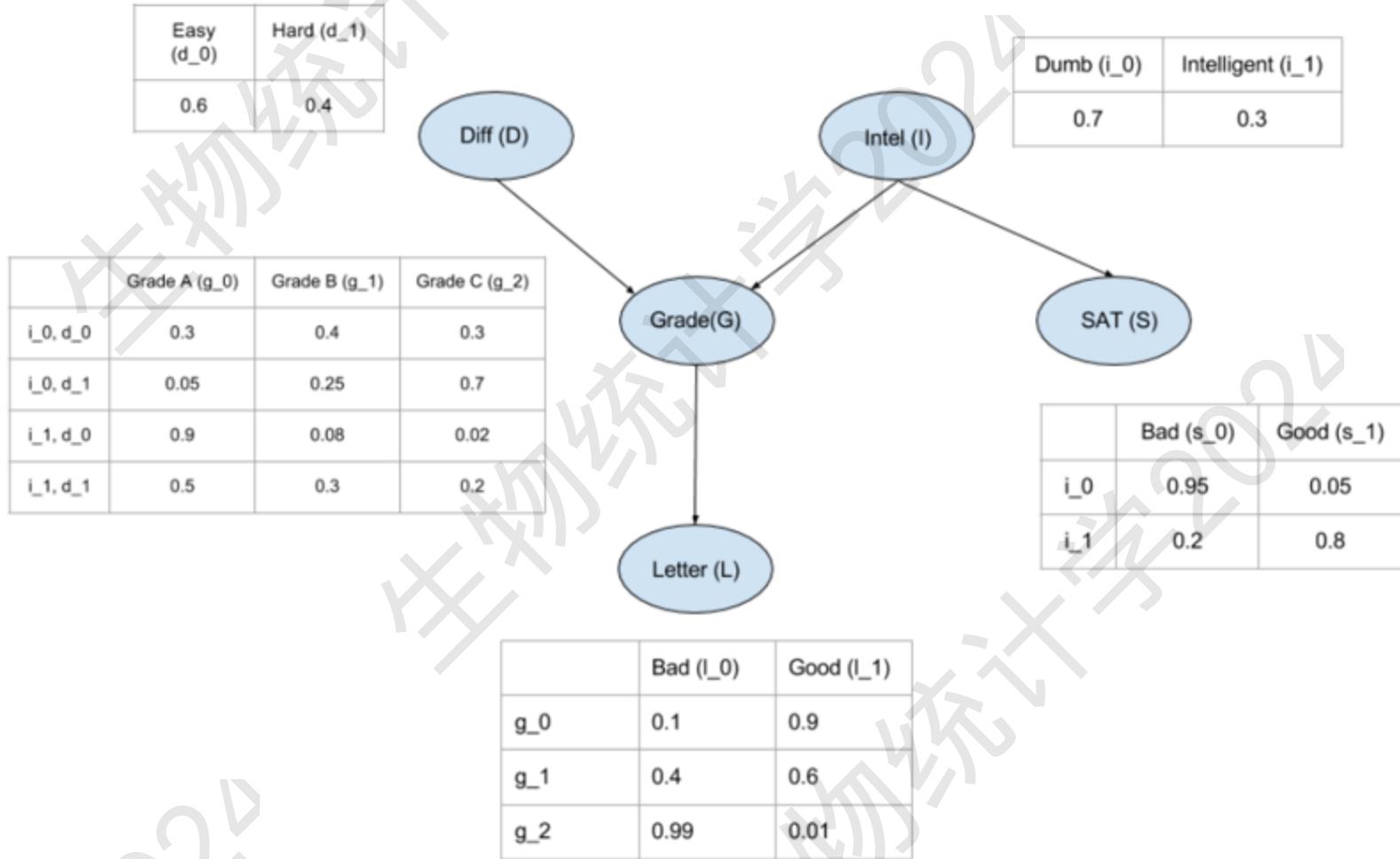
补充知识

Circadian clock regulation

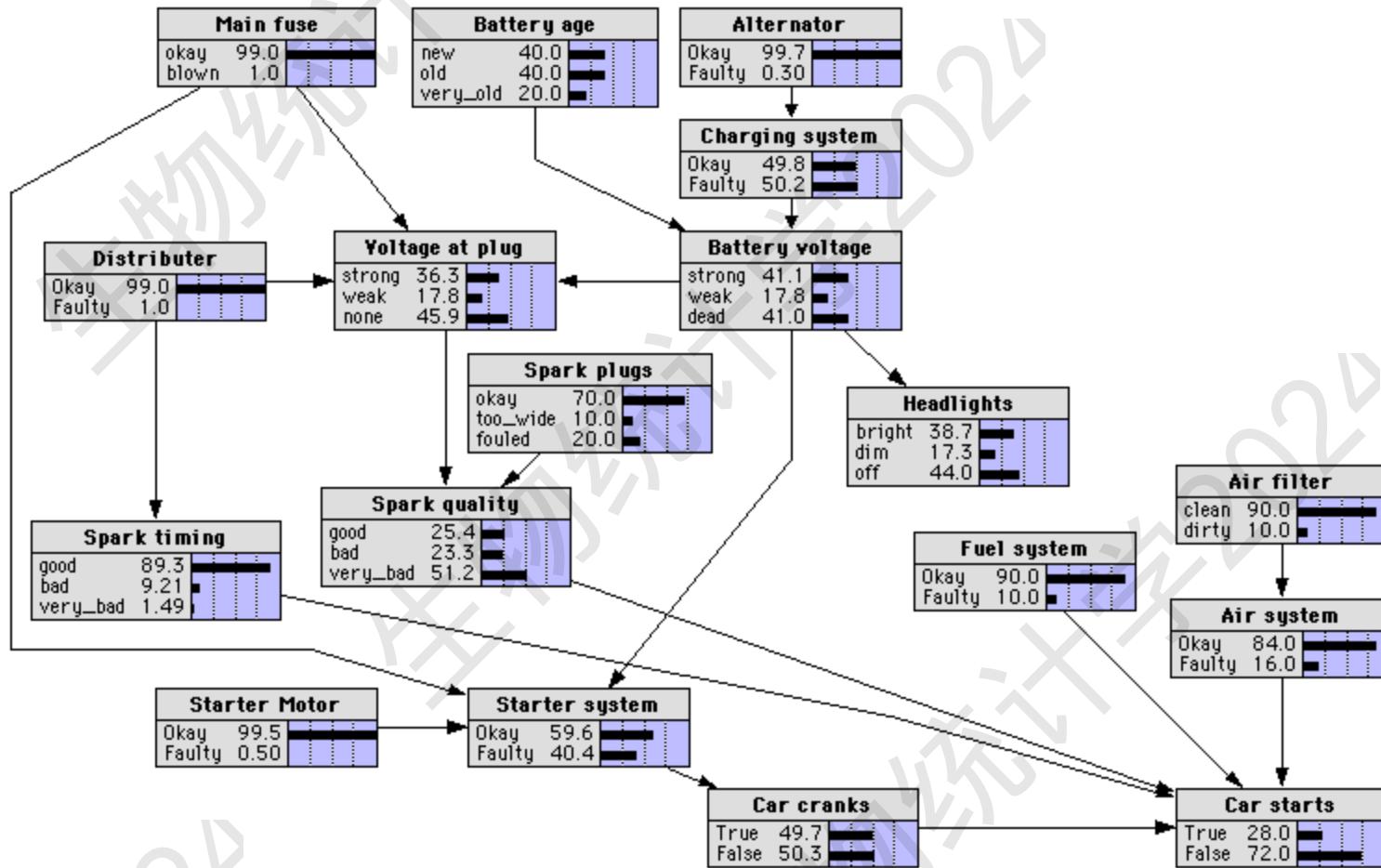


parameters

Bayesian Network

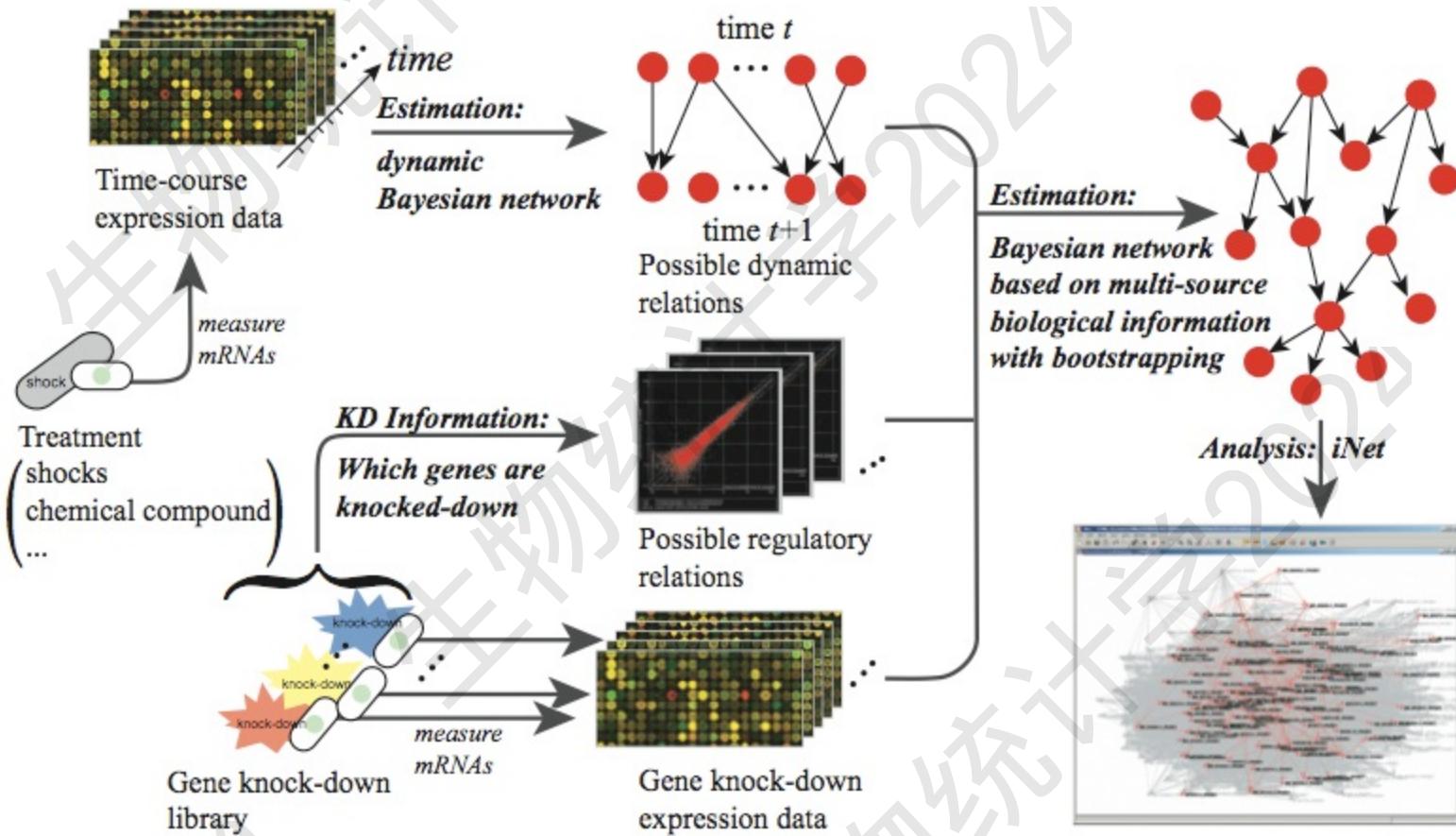


Bayesian Network

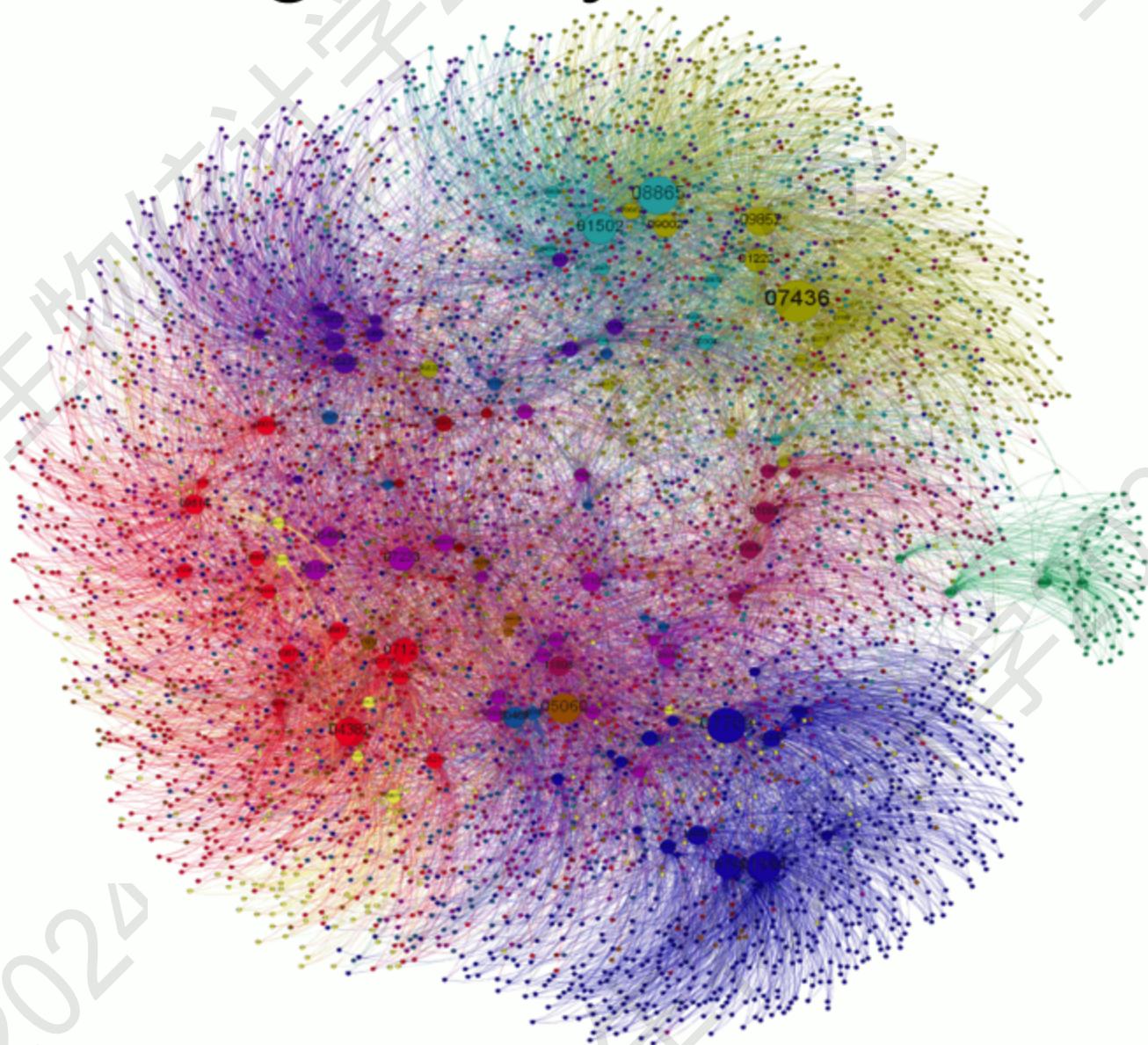


Bayesian Network in biology

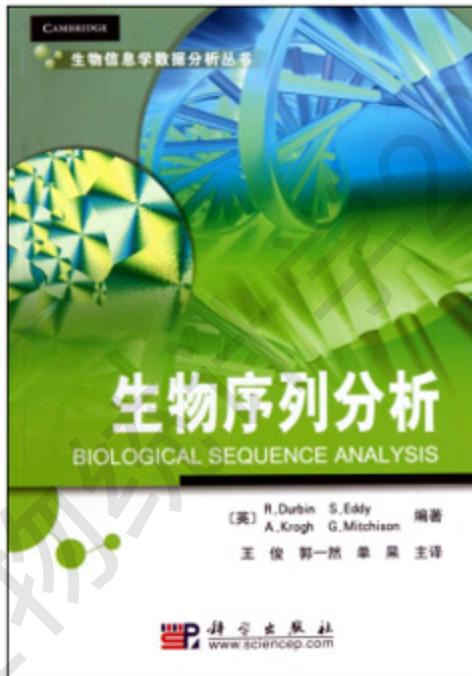
Bayesian Network in biology



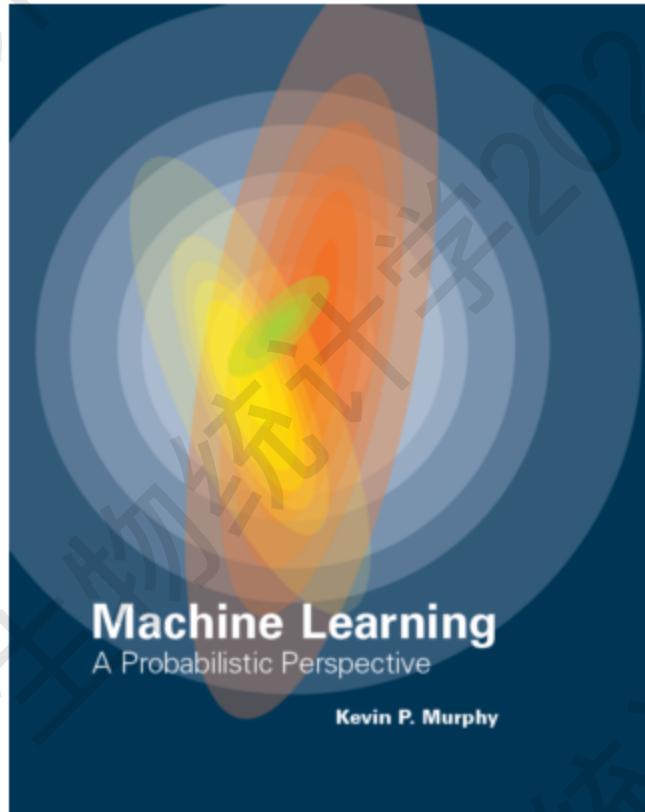
Regulatory Network



References



References



<https://github.com/probml/pyprobml>

Slides credits

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