Structural Equation Modeling

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Lab Session 3

Objectives

- Multiple Group Analysis
 - MIMIC
 - MGCFA
 - MGSEM
- "Groups" can be...
 - Countries
 - Schools
 - Males / Females
 - Time points

Multiple Groups Research

- Substantive research questions
 - Focus on factor means, factor (co-)variances, regression coefficients
 - We hope to find some interesting differences
- But before we can compare:
 - Is the measurement instrument performing in the same way in each of the groups?
 - We need to establish "measurement equivalence" before we can make meaningful comparisons:
 - Same number of factors?
 - Same factor loadings?
 - Same intercepts / thresholds?
 - Same residual variances of the factor indicators?
 - We hope to find no differences

Multiple Groups Research - Levels of Measurement Equivalence

- Configural
- Metric
- Scalar

Multiple Groups Research - Approaches

- MGCFA
 - Simultaneously testing a CFA model in multiple groups
 - Possible to test equality of each parameter
 - Requires a larger sample size than the MIMIC approach
 - Requires more programming
 - Not very elegant for comparing many groups
- MIMIC
 - Group membership captured in dummy-coded variable(s)
 - Not possible to test equality of each parameter
 - Smaller sample sizes
 - Easier to program
 - Good approach when dealing with many groups

Sequence of tests

Testing equivalencies across groups: sets of parameters are tested in a logically ordered and increasingly restrictive manner:

- Factor loadings (metric equivalence)
- Intercepts (scalar equivalence)

And, in some cases (depending on the purpose of the research):

- Factor variances / covariances
- Structural regression paths
- Error variances / covariances
- Disturbance terms

Lavaan syntax

```
setwd("~/Documents/KUL/SEM/2015")
ds<-read.table("session3")
library(lavaan)
model<-'f =~ x1 + x2 + x3 + x4'
# configural equivalence
fit1<-cfa(model, data=ds, group="group")

# metric equivalence: set the factor loadings equal across groups
fit2<-cfa(model, data=ds, group="group", group. equal=c("loadings"))

# scalar equivalence: set the factor loadings and the intercepts equal across groups
fit3<-cfa(model, data=ds, group="group", group. equal=c("loadings", "intercepts"))</pre>
```

Lavaan alternative syntax

```
# configural equivalence
model.conf.eq<-'f = x1 + x2 + x3 + x4'
```

```
fit. conf. eq<-cfa (model. conf. eq, data=ds, group="group")

# metric equivalence: set the factor loadings equal across groups
model. metr. eq<-'f = c(11,11)*x1 + c(12,12)*x2 + c(13,13)*x3 + c(14,14)*x4'
fit. metr. eq<-cfa (model. metr. eq, data=ds, group="group")

# scalar equivalence: set the factor loadings and the intercepts equal across groups
model. scal. eq<-'f = c(11,11)*x1 + c(12,12)*x2 + c(13,13)*x3 + c(14,14)*x4

x1 c(t1,t1)*1
x2 c(t2,t2)*1
x3 c(t3,t3)*1
x4 c(t4,t4)*1
f c(a1,a2)*1
a1 == 0'
fit. scal. eq<-cfa (model. scal. eq, data=ds, group="group")
```

Model fit

```
anova (fit1, fit2)
## Chi Square Difference Test
##
       Df
            AIC
                  BIC
                        Chisq Chisq diff Df diff Pr(>Chisq)
## fit1 4 21729 21863 4.3422
                                               3 3.731e-14 ***
## fit2 7 21788 21906 69.9430
                                  65.601
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
anova (fit1, fit3)
## Chi Square Difference Test
##
       Df
            AIC
                  BIC
                         Chisq Chisq diff Df diff Pr(>Chisq)
                        4.3422
## fit1 4 21729 21863
## fit3 10 22066 22167 353.7448
                                    349.4
                                                6 < 2.2e-16 ***
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
```

Lavaan syntax - partial equivalence

- Results suggest that setting all factor loadings equal across groups is too restrictive
- Use model modification indexes to determine which factor loadings should be freed
- Specify which part of the model should not be constrained equal across groups using the group.partial parameter in the cfa function

```
## Chi Square Difference Test
##

## Df AIC BIC Chisq Chisq diff Df diff Pr(>Chisq)
## fit1  4 21729 21863 4.3422
## fit2b 6 21728 21851 7.2372 2.895 2 0.2352
```

Lavaan syntax - partial equivalence

• Based on "model 2b", check whether intercepts are equal across groups

```
fit3b<-cfa(model, data=ds, group="group", group. equal=c("loadings", "intercepts"),</pre>
            group. partial=\mathbf{c} ("f = ^{\sim} x2"))
anova (fit1, fit3b)
## Chi Square Difference Test
##
                             Chisq Chisq diff Df diff Pr(>Chisq)
##
          4 21729 21863
## fit1
                            4.3422
## fit3b 9 22003 22109 288.3441
                                                       5 < 2.2e-16 ***
                                            284
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
fit3c<-cfa(model, data=ds, group="group", group. equal=c("loadings", "intercepts"),
            group. partial=\mathbf{c}("f = x2", "x3 \sim 1"))
anova(fit1, fit3c)
## Chi Square Difference Test
                            Chisq Chisq diff Df diff Pr(>Chisq)
##
         Df
               AIC
                      BIC
## fit1
          4 21729 21863
                           4.3422
## fit3c 8 21727 21839 10.2978
                                       5.9557
                                                            0.2025
                                                      4
```

Lavaan alternative syntax - partial equivalence

• Loadings and intercepts kept equal, except for loading of x2 and intercept of x3 ("fit3c")

```
model. partial. scal. eq<-'f = ^{\sim} c(11,11)*x1 + c(121,122)*x2 + c(13,13)*x3 + c(14,14)*x4 x1 ^{\sim} c(t1,t1)*1 x2 ^{\sim} c(t2,t2)*1 x3 ^{\sim} c(t31,t32)*1 x4 ^{\sim} c(t4,t4)*1 f ^{\sim} c(a1,a2)*1 a1 == 0' fit. partial. scal. eq<-cfa(model. partial. scal. eq, data=ds, group="group")
```

Comparing means

- Once (partial) measurement equivalence is demonstrated, the latent means can be compared
- Use the summary function for object fit3c to check difference in latent means

- The latent mean of the first group is set to 0; the latent mean of the second group is free to deviate.
- The part you need to look at is called "Intercepts", and you want to look at the intercept of f
- The latent mean is different in Group 2 than in Group 1 as the intercept of f in Group 2 equals 0.169 (p=2.44e-07)
- If we had not corrected for partial invariance, we would have estimated the difference in latent means as 0.044 (p=0.21)

Ex. 1.

- Use file "exS3_1" to perform a two-group CFA (2 factors with 4 indicators each)
- The variable names are x1-x8, with x1-x4 loading on factor 1 and x5-x8 loading on factor 2
- No cross-loadings are assumed
- 1.1 Statistically test for measurement equivalence and compare the latent means of both factors across the 2 groups; are they different?
- 1.2 Statistically test whether the variance of each factor is equal across groups (do this in the model that results from Ex. 1.1 so that any measurement non-equivalence is accounted for)
- 1.3 Statistically test whether the covariance between the 2 factors is equal across groups (based on the model from 1.2)
- 1.4 Build a MIMIC model in which the group variable is used as independent variable and test whether you reach the same conclusions as with the MGCFA approach regarding measurement equivalence and difference in means

```
ds<-read. table ("exS3_1")
model < -"f1 = x1 + x2 + x3 + x4
f2 = x5 + x6 + x7 + x8"
fit1<-cfa(model, data=ds, group="group")</pre>
fit2<-cfa(model, data=ds, group="group", group. equal=c("loadings"))</pre>
fit3<-cfa(model, data=ds, group="group", group. equal=c("loadings", "intercepts"))
anova (fit1, fit2)
## Chi Square Difference Test
##
##
                    BIC Chisq Chisq diff Df diff Pr(>Chisq)
        Df
             AIC
## fit1 38 33725 33993 48.982
## fit2 44 33723 33959 58.844
                                    9.8627
                                                 6
                                                        0.1306
anova (fit1, fit3)
## Chi Square Difference Test
##
                    BIC
                          Chisq Chisq diff Df diff Pr(>Chisq)
##
        Df
             AIC
## fit1 38 33725 33993 48.982
## fit3 50 33814 34018 161.703
                                    112.72
                                                 12 < 2.2e-16 ***
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

Ex. 1 - Solution 1.1

```
# Use summary function to detect largest MIs - x3 and x6 appear to need different
# intercepts across groups
anova(fit1, fit3b)
## Chi Square Difference Test
##
            AIC
                 BIC Chisq Chisq diff Df diff Pr(>Chisq)
## fit1
       38 33725 33993 48,982
## fit3b 48 33716 33931 60.033
                               11.051
                                         10
                                               0.3536
# Check overall model fit of specification 3b
c(inspect(fit3b, "fit")["chisq"], inspect(fit3b, "fit")["df"], inspect(fit3b, "fit")["pvalue"])
##
       chisq
                   df
                         pvalue
## 60.0326006 48.0000000 0.1140915
```

Ex. 1 - Solution 1.1

```
# Equal means? Check output you get from summary(fit3b) or use below statements
int <- inspect(fit3b, "est")[10]$alpha["f1", "intercept"] # gives intercept of f1
z <- inspect(fit3b, "est")[10]$alpha["f1", "intercept"] /
    inspect(fit3b, "se")[10]$alpha["f1", "intercept"] # gives z value
p<-(1-pnorm(abs(z)))*2 # gives 2 sided p-value
paste("Factor 1: alpha=",round(int,3),", z=",round(z,3),", p=",round(p,3), sep="")

## [1] "Factor 1: alpha=0.053, z=1.508, p=0.132"

int <- inspect(fit3b, "est")[10]$alpha["f2", "intercept"] # gives intercept of f2
z <- inspect(fit3b, "est")[10]$alpha["f2", "intercept"] /
    inspect(fit3b, "est")[10]$alpha["f2", "intercept"] # gives z value
p <- (1-pnorm(abs(z)))*2 # gives 2 sided p-value
paste("Factor 2: alpha=",round(int,3),", z=",round(z,3),", p=",round(p,3), sep="")

## [1] "Factor 2: alpha=0.209, z=6.512, p=0"</pre>
```

```
x4 \sim c(t4, t4) *1
x5 \sim c(t5, t5) *1
x6 \sim c(t61, t62) *1 \qquad \text{# free intercept for x6 across groups}
x7 \sim c(t7, t7) *1
x8 \sim c(t8, t8) *1
f1 \sim c(a11, a12) *1
f2 \sim c(a21, a22) *1
a11 == 0 \qquad \text{# set intercept of f1 to 0 in first group}
a21 == 0 \qquad \text{# set intercept of f2 to 0 in first group}
f1 \sim c(psi11, psi11) *f1 \qquad \text{# equal variance of factor 1 across groups}
f2 \sim c(psi22, psi22) *f2 \qquad \text{# equal variance of factor 2 across groups}
fit4 \leftarrow \text{-cfa(model3, data=ds, group="group")}
```

Ex. 1 - Solution 1.2

```
anova(fit3b, fit4)

## Chi Square Difference Test
##

## Df AIC BIC Chisq Chisq diff Df diff Pr(>Chisq)
## fit3b 48 33716 33931 60.033
## fit4 50 33713 33917 60.799 0.76638 2 0.6817
```

• The non-significant p-value indicates that we can constrain the factor variances to be equal across groups (Note that since we are testing 2 parameters at once, it is possible that this finding does not remain when testing one at the time)

```
mode12 \leftarrow "f1 = c(11, 11) *x1 + c(12, 12) *x2 + c(13, 13) *x3 + c(14, 14) *x4
f2 = (15, 15) \times x5 + c(16, 16) \times x6 + c(17, 17) \times x7 + c(18, 18) \times x8
x1 \sim c(t1, t1)*1
x2^{\circ} c(t2, t2) *1
x3 ^
    c(t31, t32) *1
                          # free intercept for x3 across groups
x4 \sim c(t4, t4) *1
x5^{\circ} c(t5, t5) *1
x6 \sim c(t61, t62) *1
                          # free intercept for x6 across groups
x7 ~ c(t7, t7)*1
x8 ~ c(t8, t8)*1
f1 ~ c(a11, a12) *1
f2 ~ c(a21, a22) *1
a11 == 0
                          # set intercept of f1 to 0 in first group
a21 == 0
                          # set intercept of f2 to 0 in first group
f1 ~~ c(psi11,psi11)*f1 # equal variance of factor 1 across groups
f2 ~~ c(psi22, psi22)*f2 # equal variance of factor 2 across groups
f1 ^{\sim} c(psi1,psi1)*f2 # constrain the factor covariance to be equal"
fit5<-cfa (mode12, data=ds, group="group")
```

Ex. 1 - Solution 1.3

```
anova(fit4, fit5)

## Chi Square Difference Test
##
## Df AIC BIC Chisq Chisq diff Df diff Pr(>Chisq)
## fit4 50 33713 33917 60.799
## fit5 51 33711 33910 60.962 0.16322 1 0.6862
```

• The non-significant p-value indicates that we can constrain the factor covariance to be equal across groups

Ex. 1 - Solution 1.4

```
# Dummy-code group for ease of interpretation
ds$grpB <- ifelse(ds$group=="B", 1, 0)
# Test for measurement non-equivalence (only intercepts of indicators are checked)
model < -"f1 = x1 + x2 + x3 + x4
f2 = x5 + x6 + x7 + x8
f1 ~ grpB
f2 ~ grpB
   ~ grpB
~ 0*grpB
x1
    ~ 0*grpB
x2
x3 ~ 0*grpB
x4 ~ 0*grpB
    ~ 0*grpB
x5
x6
       0*grpB
       0*grpB
x7
x8 ~ 0*grpB"
fit.mimic <-cfa (model, data=ds)
mi<-inspect(fit.mimic, "mi") # You should review all modification indexes!
mi<-subset(mi, substr(lhs, 1, 1) == "x" & op == "~" & rhs == "grpB") # But here we will</pre>
                      # restrict to the direct effects of grpB on the x-variables
mi.sorted<-mi[order(-mi$mi),] # sort from high to low
mi.sorted[1:3,] # only display some large MI values
##
                                 epc sepc.lv sepc.all sepc.nox
      lhs op rhs
                         mi
## 1 x3
              grpB 67.375 0.391
                                        0.391
                                                  0.192
                                                             0.385
## 2 x6 ^{\sim}
## 2 x6 ^{\sim} grpB 32.116 -0.303 ## 3 x8 ^{\sim} grpB 8.955 0.175
                                      -0.303
                                                  -0.156
                                                             -0.311
                                                  0.088
                                                             0.175
                                       0.175
```

```
# Allow intercept of x3 and x6 to vary by group
model2<-"f1 = x1 + x2 + x3 + x4
f2 = x5 + x6 + x7 + x8
f1 grpB
```

```
f2 ~ grpB

x1 ~ 0*grpB

x2 ~ 0*grpB

x3 ~ grpB

x4 ~ 0*grpB

x5 ~ 0*grpB

x6 ~ grpB

x7 ~ 0*grpB

x8 ~ 0*grpB"

fit2. mimic <-cfa (model2, data=ds)
```

· Review the model output and verify you obtain similar substantive results as with MGCFA

Ex. 2. Major Depression - Gender analysis

- On page 272 of Brown's book "Confirmatory Factor Analysis for Applied Research" (2006), correlation matrices and mean and standard deviation vectors are given for 375 Females and 375 Males
- Nine variables (mdd1-mdd9) are supposed to measure a single latent variable "Major Depression". One error covariance between mdd1 and mdd2 is assumed for substantive reasons
- Read in the summary data and fit the 8 models that Brown presents on page 273
- Review the fit of each of the models and compare to Brown's results

Ex. 2 - Solution: reading in the data (Females)

```
# Females, n=375
lower1<-'1
0.616
0.315
        0.313
0.349
        0.332
                  0.261
                           1
                  0.27
0.314
         0.25
                           0.327
                  0.298
0.418
                           0.328
                                    0.317
                                             1
         0.416
0.322
         0.313
                  0.096
                           0.117
                                    0.13
                                             0.14
0.409
        0.415
                  0.189
                                    0.303
                                             0.281
                                                      0.233
                           0.314
                                                                1
        0.222
                                                                         1'
0.318
                  0.051
                                    0.14
                                                      0.217
                                                                0.222
                           0.115
                                             0.15
sds1<-c (1.717, 2.015, 2.096, 2.212, 2.132, 2.005, 2.062, 2.156, 1.791)
mean1 < -c (4.184, 3.725, 1.952, 3.589, 2.256, 3.955, 3.869, 3.595, 1.205)
cov1<-getCov(lower1, names=paste("mdd", 1:9, sep=''), sds=sds1)</pre>
```

Ex. 2 - Solution: reading in the data (Males)

```
# Males, n=375
lower2<-'1
0.689
        1
0.204
        0.218
0.335
        0.284
                 0.315
                         1
0.274
                 0.153
        0.32
                         0.265
0.333
        0.333
                 0.221
                         0.364
                                  0.268
```

```
0.258
        0. 211 0. 114 0. 139 0. 185 0. 132
0.319
        0.346
                 0.176
                         0.207
                                  0.231
                                           0.279
                                                    0.146
                                                                     1'
0.316
        0.269
                 0.111
                         0.14
                                  0.117
                                           0.131
                                                   0.263
                                                            0.163
sds2<-c(1.598, 2.018, 2.094, 2.232, 2.108, 2.113, 2.286, 2.174, 1.788)
mean2 < -c (4.171, 3.685, 1.739, 3.357, 2.235, 3.661, 3.421, 3.517, 1.259)
cov2<-getCov(lower2, names=paste("mdd", 1:9, sep=''), sds=sds2)
```

Ex. 2 - Solution: fitting the model in both groups separately

```
model1<-'f = mdd1 + mdd2 + mdd3 + mdd4 + mdd5 + mdd6 + mdd7 + mdd8 + mdd9
mdd1 mdd2'

# Single group analysis - Females
fit1. females<-cfa(model1, sample. cov=cov1, sample. mean=mean1, sample. nobs=375)

# Single group analysis - Males
fit1. males<-cfa(model1, sample. cov=cov2, sample. mean=mean2, sample. nobs=375)</pre>
```

Ex. 2 - Solution: Measurement Equivalence

Ex. 2 - Solution: Population heterogeneity

Ex. 2 - Solution: summary (1)

Single Group Solutions	Chi-squared	df	RMSEA	CFI	TLI
Women (n=375)	53.14		0.053	0.96	0.94
Men (n=375)	46.03		0.045	0.97	0.95

Ex. 2 - Solution: summary (2)

Measurement Invariance	Chi-squared	df	RMSEA	CFI	TLI
Equal form	99.17	52	0.049	0.96	0.95
Equal factor loadings	103.12	60	0.044	0.97	0.96
Equal ind. intercepts	115.63	68	0.043	0.96	0.96
Equal ind. error var.	125.33	77	0.041	0.96	0.96

Ex. 2 - Solution: summary (3)

Population heterogeneity	Chi-squared	df	RMSEA	CFI	TLI
Equal factor variance Equal latent mean	126.12 128.04	, 0	0.041 0.041	0.96 0.96	0.,,

Ex. 3. Holzinger-Swineford data

This exercise uses the HolzingerSwineford1939 dataset that comes with lavaan.

- 3.1. Fit the 3-factor model in each school separately (Grant-White and Pasteur). x1-x3 and x9 should load on the Spatial factor; x4-x6 on the Verbal factor; and x7-x9 on the Speed factor.
- 3.2. Fit the model in both groups simultaneously and verify that the sum of the Chi-squared values from 3.1. equals the Chi-squared value in 3.2. (the same should also hold for the df)
- 3.3. Test for metric equivalence; if necessary, build a model with partial metric equivalence.
- 3.4. Test for scalar equivalence; if necessary, build a model with partial scalar equivalence.
- 3.5a. Compare the latent means across the two groups: which are statistically different and which are not? 3.5b. If measurement non-equivalence is observed: what if measurement non-equivalence had not been accounted for? Would the same conclusions with regard to latent means have been attained?
- 3.6. Are all 3 factor variances equal across groups?

Ex. 3 - Solution 3.1.

```
model<-"spatial = x1 + x2 + x3 + x9
verbal = x4 + x5 + x6
speed = x7 + x8 + x9"
fit1.GW <- cfa(model, data=subset(HolzingerSwineford1939, school=="Grant-White"))
fit1.P <- cfa(model, data=subset(HolzingerSwineford1939, school=="Pasteur"))</pre>
```

- Request the summary for both model outputs and review them
- Note that the model fits well in het Grant-White school, but not well inthe Pasteur school.

Ex. 3 - Solution 3.2.

```
fit2 <- cfa (model, data=HolzingerSwineford1939, group="school")
fit2
## lavaan (0.5-18) converged normally after 56 iterations
##
##
     Number of observations per group
##
     Pasteur
                                                         156
##
     Grant-White
                                                         145
##
##
     Estimator
                                                          ML
##
     Minimum Function Test Statistic
                                                     81.547
##
     Degrees of freedom
                                                          46
     P-value (Chi-square)
##
                                                      0.001
##
## Chi-square for each group:
##
##
     Pasteur
                                                      53.253
##
     Grant-White
                                                      28.293
```

Ex. 3 - Solution 3.3.

• Metric equivalence seems to hold

Ex. 3 - Solution 3.4.

Scalar equivalence does not hold, review modification indexes to pin point the problematic item(s)

```
mi<-inspect(fit4, "mi") # You should review all modification indexes!
mi<-subset(mi, substr(lhs, 1, 1) == "x" & op == "^1" & rhs == "") # But here we will
# restrict to the direct effects of grpB on the x-variables
mi.sorted<-mi[order(-mi$mi),] # sort from high to low</pre>
```

Ex. 3 - Solution 3.4.

```
mi.sorted[1:5,] # only display some large MI values
##
     lhs op rhs group
                           mi
                                 epc sepc. lv sepc. all sepc. nox
## 1 x3 ~1
                    1 10.118 0.267
                                       0.267
                                                 0.224
                                                          0.224
## 2 x3 ~1
                    2 7.862 -0.207
                                      -0.207
                                                -0.196
                                                         -0.196
## 3 x7
         ^{\sim}1
                       7. 669 0. 220
                                      0.220
                                                 0.200
                                                          0.200
                    1
## 4 x7^{-1}
                    2 5.937 -0.170 -0.170
                                               -0.161
                                                         -0.161
## 5 x2 ^{\sim}1
                    1 3.238 -0.168 -0.168
                                               -0.135
                                                         -0.135
```

• Relax equality constraint on the intercept of x3

Ex. 3 - Solution 3.4.

```
mi<-inspect(fit4b, "mi") # You should review all modification indexes!
mi<-subset(mi, substr(lhs, 1, 1) == "x" & op== "~1" & rhs== "") # But here we will
                  # restrict to the direct effects of grpB on the x-variables
mi.sorted (-mi[order(-mi$mi),] # sort from high to low
mi.sorted[1:5,] # only display some large MI values
##
     lhs op rhs group
                                epc sepc. lv sepc. all sepc. nox
                          mi
## 1 x7 ^{\sim}1
                    1 6. 289 0. 198
                                      0.198
                                               0.181
                                                         0.181
## 2 x7
                    2 4.893 -0.154
                                     -0.154
                                               -0.145
                                                        -0.145
## 3 x8 ~1
                    1 3.781 -0.135
                                    -0.135
                                              -0.138
                                                        -0.138
## 4 x8 ~1
                    2 3, 436 0, 123
                                      0.123
                                               0.119
                                                        0.119
## 5 x2 ~1
                    1 1. 231 -0. 102 -0. 102
                                                       -0.083
                                              -0.083
```

Ex. 3 - Solution 3.4.

• Relax equality constraint on the intercept of x3 and x7

Ex. 3 - Solution 3.5a / 3.5b

Taking into account measurement non-equivalence (partial measurement equivalence):

Factor	Mean in Group 2	p-value
Spatial	0.032	0.8
Verbal	0.576	0
Speed	-0.079	0.404

• Not taking into account measurement non-equivalence:

Factor	Mean in Group 2	p-value
Spatial	-0.119	0.32
Verbal	0.576	0
Speed	-0.2	0.042

Ex. 3 - Solution 3.6

```
fit5 <- cfa (model, data=HolzingerSwineford1939, group="school",
              group. equal=c("loadings", "intercepts", "lv. variances"),
              group. partial=\mathbf{c}("x3^{\sim}1", "x7^{\sim}1"))
anova(fit4c, fit5)
## Chi Square Difference Test
##
                        BIC Chisq Chisq diff Df diff Pr(>Chisq)
##
                AIC
          Df
## fit4c 57 7445.6 7634.7 95.064
## fit5
         60 7442. 2 7620. 2 97. 704
                                          2.6392
                                                        3
                                                               0.4507
```

• The non-significant p-value reflects that we can set all 3 error variances equal across groups (Note: testing each factor variance at the time may give different results)

Ex. 4. MGSEM

This exercise builds on ex. 3

We might hypothesize that students' abilities increase as they move to the next grade. So we would expect to see a positive effect of "grade" on each of the three latent variables. But is this positive effect observed in both schools? And is the effect equal in both schools? Or would students in one school progress less quickly than students in the other school?

- 4.1. Extend the model from ex. 3 that takes into account any measurement non-equivalence by including the variable "grade". Regress the three latent variables on this covariate. Allow the regression parameters to be different across both schools.
- 4.2. Test whether it is possible to constrain each of the three regression parameters to be equal across the two schools. If this is not possible, develop a model which holds as many regression parameters equal across groups as possible.
- 4.3. To verify the robustness of findings from 4.2, test for measurement equivalence (loadings and intercepts of the indicator variables) of the measurement instrument by grade

Ex. 4 - Solution 4.1

• Review the output using the summary() function

Ex. 4 - Solution 4.2

• The non-significant p-value suggests that all regression coefficients can be held equal across the two schools. (Note: this is an omnibus test. Testing each regression separately may give different results)

Ex. 4 - Solution 4.3

Ex. 4 - Solution 4.3

```
anova (fit3. ce, fit3. me)
## Chi Square Difference Test
##
                 AIC
                         BIC Chisq Chisq diff Df diff Pr(>Chisq)
## fit3.ce 46 7469.2 7698.8 74.445
## fit3.me 53 7458.6 7662.3 77.908
                                         3.4623
                                                      7
                                                             0.8392
anova (fit3. ce, fit3. se)
## Chi Square Difference Test
##
##
                         BIC Chisq Chisq diff Df diff Pr(>Chisq)
                  AIC
## fit3.ce 46 7469.2 7698.8 74.445
## fit3.se 59 7455.4 7636.9 86.702
                                         12.256
                                                     13
                                                             0.5068
```

• Non-significant p-values suggest measurement equivalence, and so results from 4.2. should not be affected by measurement non-equivalence.