

# Intro to Conjoint Experiments

## Session 5

Alberto Stefanelli

# Session 4: Recap

## 1 AMCE

- Effect decomposition
- Advantages
- Calculation
- Interpretation

## 2 Marginal Means

- Purpose and interpretation
- Relation with the AMCE

# Session 5: Outline

- ① Sub-group analysis
- ② Attributes interaction
- ③ Power Analysis for CJ
- ④ Restricted Randomization
- ⑤ Weighted Randomization
- ⑥ Conjoint mixture model

# Materials

- Lecture's PDF
- Lab
- Exercise
- Solutions

Where to find the material:

- On my [GitHub/conjoint\\_class](#)

# Sub-group analysis

- ① Subgroup analyses of conjoint experiments in order to discover preference heterogeneity.
- ② Increasingly common feature of experimental analysis
- ③ e.g., AMCE of male and female political candidates among male and female respondents ([Teele, Kalla, and Rosenbluth 2018](#))
- ④ **Technically:** Attribute's marginal effect conditional on the respondent characteristic of interest
- ⑤ 2 quantities of interest
  - Causal effects of profile features within each subgroup (Hainmueller, Hopkins, and Yamamoto ([2014](#)) term “conditional AMCEs”)
  - Difference between two conditional AMCEs across subgroups

# Common issues in sub-group analysis (1)

- ① Most of the time, difference-in-AMCEs is interpreted descriptively and not causally
  - Difference-in-AMCEs are used to descriptively interpret apparent differences in favorability toward objects with a given feature (e.g., immigrants from Syria) between the two groups (low and high ethnocentrism respondents) ([Hainmueller, Hopkins, and Yamamoto 2014](#))
  - However, differences-in-preferences (that is to say, the difference in degree of favorability toward profiles containing a given feature) are not directly reflected in differences-in-AMCEs.
  - That is, difference-in-AMCEs do not provide inference into difference between subgroups' favorability toward a conjoint feature

## Common issues in sub-group analysis (2)

- ② Reason: difference between subgroups (low and high ethnocentrism respondents) diverge in the reference category ([Leeper, Hobolt, and Tilley 2019](#))
  - Lack of differences when another (meaningful) baseline category is chosen
  - Smaller/Larger effect in difference-in-AMCEs for a given feature because a group has already a weak/strong preference for a given attribute
  - AMCE difference is valid only when preferences toward profiles with the reference category are equivalent across groups.
  - E.g. Political Experience (No Experience, 3 Years, 6 years) X PID (Republican, Democrats): Republican might experience a small effect because their preference towards candidate experience is already very high and as such, a large positive effect for Democrats occurs despite Democrat being less supportive in every experimental conditions
- ③ **Take Home:** Differences in the sign/size of causal effects should not be interpreted as differences in preferences

# Solutions for difference-in-AMCEs

- ① Unadjusted marginal means
  - AMCE is very sensitive to which reference category you choose (as we have seen in Lab 4?) but MM are not
  - They express favorability on the scale of the outcome over alternative values of each feature
- ② An omnibus F-test, measuring differences therein.
  - **Model 1:** Model estimating only marginal effects of a given attribute
  - **Model 2:** Same as Model 1 with additional interactions between the sub-grouping covariate and the features
  - F-test for the model comparison between two models
- ③ Other approaches? Check out Ratkovic and Tingley ([2017](#)) on Bayesian Lasso for sparse estimation and uncertainty in subgroup analysis



# Marginal Means: comparison with AMCE

Figure 4: Estimated Preference Differences between Inequity Averse and Non-Averse Respondents from Ballard-Rosa et al. (2016) Tax Preference Experiment for Each Possible Reference Category

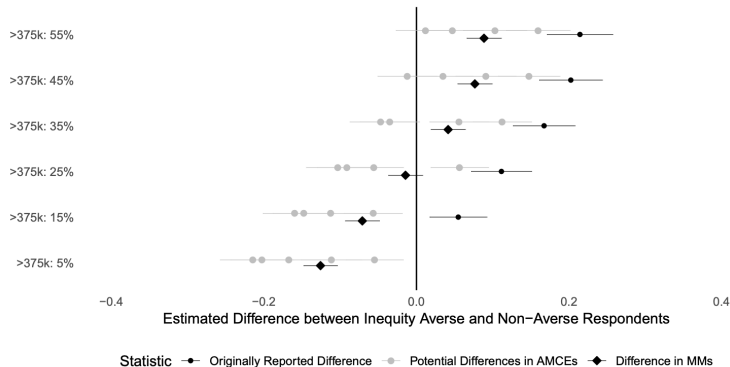


Figure 1: Conjointsdtd Attributes and Levels

# Interactions between attributes (ACIE)

- 1 The decision of choosing a particular profile could be the results of a particular combination of features.
- 2 The causal effect of one attribute (say candidate's income) may vary depending on what value another attribute (e.g., ideology) is held at.
- 3 We may want to quantify the magnitude of such interactions.
- 4 ACIE: Average component interaction effect
- 5 ACIE: Difference in the AMCEs of a given attribute between conditional on another attribute

# Interactions between Attributes: an example

- ① Voters might value the **programmatic coherence** of a candidate policy positions
  - ② A voter might be in favour of a reduction of governmental intervention in the economy and as such would support a candidate in favour of cutting social spending
  - ③ What would happen if a candidate is in favour of increasing social spending but – at the same time – is pro Medicare?
  - ④ ACIE: the percentage point difference in the AMCEs of cutting social spending between a pro Medicare candidate and a candidate against it
- ② Q: Can you come up with another example of attribute interaction? (**think theoretically**)

# AMCE with Restricted Randomization

- ❶ Restricted randomization means that we have constraints between profile features
- ❷ E.g., Doctor without an academic education
- ❸ AMCE is a weighted average across each combination of the constrained features
- ❹ The effect of being a doctor on candidate favourability is marginalized **only** across higher level of educations
- ❺ Recommendations
  - Check the frequencies to be sure that the conditional probability between two attributes/levels is actually 0
  - Analysing only the complete and comparable subset of the design. This needs to be implemented while analysing the CJ in R or other software
  - Be clear about what features are being marginalized over

# Non-uniform distribution

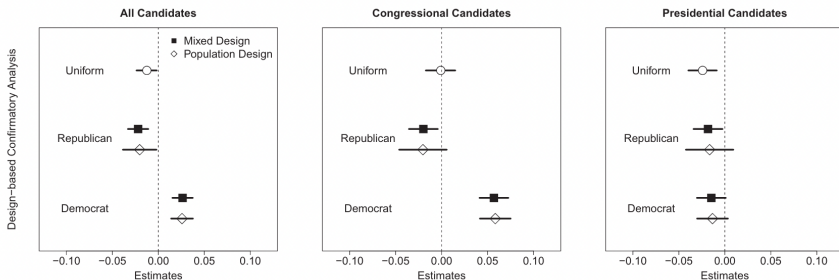
- ❶ The actual distribution of profiles in the real world and the distribution of theoretical interest are often far from being uniform.
- ❷ We should match the real world distribution
- ❸ It compromise the external validity of conjoint analysis.
  - ❶ Design-based confirmatory analysis: incorporates the target profile distribution in the design
  - ❷ Model-based exploratory analysis: takes into account the target profile distribution at the analysis stage, after randomizing profiles and collecting data
- ❹ **Paper to read:** Improving the External Validity of Conjoint Analysis: The Essential Role of Profile Distribution ([de la Cuesta, Egami, and Imai 2022](#))

# Assumptions design with non-uniform distribution

**Table 3.** Data requirements and assumptions of design-based and model-based approaches.

Approach	Data requirement	Assumption	Note
<b>Design-based confirmatory analysis</b>			
• Joint population randomization	Joint distribution over all profile attributes	None	
• Marginal population randomization	Marginal distributions of each profile attribute	Absence of three-way or higher order interaction	Relax the assumption with partial joint distributions
• Mixed randomization	Joint distribution over control factors	None	Efficient when focus on one or two main factors
<b>Model-based exploratory analysis</b>			
• Linear probability model	Marginal distributions of each profile attribute	Absence of three-way or higher order interaction	Relax the assumption with partial joint distributions

# Non-uniform distribution example



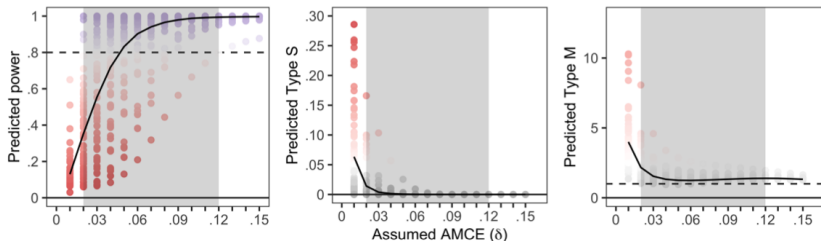
# Power Analysis for conjoint experiments

- ❶ Experiments must ensure that statistical power is be reasonably high
  - Type I error: Successfully rejecting the null hypothesis when it is false (false-positive)
  - Type II error: Fails to reject a null hypothesis that is actually false (false-negative)
  - Type S error: opposite direction of the true effect size, given that the statistic is statistically significant (sign) ([Gelman and Carlin 2014](#))
  - Type M error: magnitude of your effect is overestimated (size) ([Gelman and Carlin 2014](#))
- ❷ Calculating the required minimal sample size for a conjoint experiment is not a trivial exercise.
  - Multiple profiles
  - Multiple tasks
  - High numbers of attributes
  - High numbers of levels



# Retrospective statistical power, Type M and Type S error rates.

- Scholars think that conjoint designs “free us from the power constraints that limit traditional factorial experiments” (Kertzer, Renshon, and Yarhi-Milo 2019, 7)
- This results in conjoint studies that are under-powered and thus are likely to result in biased estimates, both in terms of direction and magnitude
- Stefanelli, A., & Lukac, M. (2020, November 18). Subjects, Trials, and Levels: Statistical Power in Conjoint Experiments.  
<https://doi.org/10.31235/osf.io/spkcy>



# Power calculation tool

- Shiny App for power calculations in CJ
- <https://mblukac.shinyapps.io/conjoints-power-shiny/>

## Conjoint Experiments: Power Analysis Tool

Respondents

1000

500 1,000 3,000

Tasks

3

1 3 9

Effect size (%)

0.05

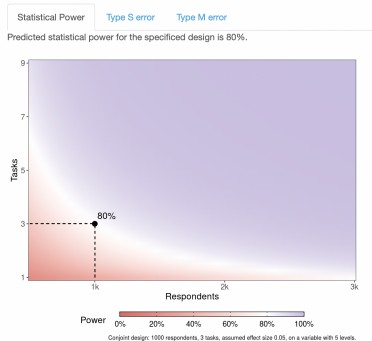
0.01 0.05 0.2

Variable levels

5

2 5 30

About Help



# Mixture Modelling for conjoint data: Limitation of the AMCE

- 1 The AMCE can be swayed both by **intensity** as well as **prevalence** of some preferences (Abramson et al., 2019).
- 2 Existence of subgroups with distinct preferences and different responses to the treatment.
- 3 Existence of unobserved subgroups with extreme preferences and their prevalence in one of the groups can seriously bias the results of a conjoint experiment
- 4 Presence of heterogeneity within observed subgroups and not only between subgroups (i.e. strong VS weak party identifiers).

# Mixture Modelling for conjoint data: Rationale

- We should first investigate the heterogeneity itself and then try to explain where it comes from.
  - Subgroups as unobserved (latent) classes in the population.
  - Investigate differences in the effects between unobserved groups.
  - The effects of the attributes are allowed to differ in magnitude or direction across the extracted unobserved groups.

# Mixture Modelling for conjoint data: Equation

The diagram shows the equation  $y_i | \mathbf{x}, k = \beta_{0k} + \sum_{p=1}^P \beta_{pk} x_{ip} + \epsilon_{ik}$  with several annotations:

- P attributes** ( $\mathbf{x} = (x_1, x_2, \dots, x_P)$ ) points to  $\mathbf{x}$ .
- Selection of a profile for each individual  $i$**  points to  $y_i$ .
- $\beta_{0k}$  is a class  $K$  specific intercept** points to  $\beta_{0k}$ .
- Latent membership is denoted by  $C$ , where  $C = 1, 2, \dots, K$ .** points to  $k$ .
- Values  $x_{ip}$  for subject  $i$  for attribute  $p$ .** points to  $x_{ip}$ .
- $\beta_{pk}$  is a class-specific effect of attribute  $p$  for class  $k$**  points to  $\beta_{pk}$ .
- and  $\epsilon_{ik} \sim N(0, \sigma_k^2)$  with  $\sigma_k^2$  as a residual variance/covariance matrix for class  $k$ .** points to  $\epsilon_{ik}$ .

$$y_i | \mathbf{x}, k = \beta_{0k} + \sum_{p=1}^P \beta_{pk} x_{ip} + \epsilon_{ik}$$

# Mixture Modelling for conjoint data: advantages

- 1 Subgroups are allowed to emerge even if they are not correlated to observed covariates.
- 2 Well-defined framework to test the presence of subgroups
- 3 Misclassification of the subjects within the unobserved subgroups can be investigated and taken into account

# Session 5: Recap

- ① Sub-group analysis
  - Issues
  - Solutions (MM)
- ② Attributes interaction
  - ACIE
- ③ Power Analysis for CJ
  - Underpowered studies
  - Shiny app
- ④ Restricted Randomization
  - Design-based approaches
  - Model-based approaches
- ⑤ Conjoint mixture model
  - Limitations AMCE
  - Rationale
  - Advantages

# References I

- Cuesta, Brandon de la, Naoki Egami, and Kosuke Imai. 2022. "Improving the External Validity of Conjoint Analysis: The Essential Role of Profile Distribution." *Political Analysis* 30 (1): 19–45. <https://doi.org/10.1017/pan.2020.40>.
- Gelman, Andrew, and John Carlin. 2014. "Beyond Power Calculations: Assessing Type S (Sign) and Type M (Magnitude) Errors." *Perspectives on Psychological Science* 9 (6): 641–51. <https://doi.org/b2h3>.
- Hainmueller, Jens, Daniel J. Hopkins, and Teppei Yamamoto. 2014. "Causal Inference in Conjoint Analysis: Understanding Multidimensional Choices via Stated Preference Experiments." *Political Analysis* 22 (1): 1–30. <https://doi.org/f5qzwp>.
- Kertzer, Joshua D, Jonathan Renshon, and Keren Yarhi-Milo. 2019. "How Do Observers Assess Resolve?" *British Journal of Political Science*, June, 1–23. <https://doi.org/gmqpck>.
- Leeper, Thomas J., Sara B. Hobolt, and James Tilley. 2019. "Measuring Subgroup Preferences in Conjoint Experiments." *Political Analysis*, August, 1–15. <https://doi.org/gh6p77>.



## References II

- Ratkovic, Marc, and Dustin Tingley. 2017. "Sparse Estimation and Uncertainty with Application to Subgroup Analysis." *Political Analysis* 25 (1): 1–40. <https://doi.org/10.1017/pan.2016.14>.
- Teele, Dawn Langan, Joshua Kalla, and Frances Rosenbluth. 2018. "The Ties That Double Bind: Social Roles and Women's Underrepresentation in Politics." *American Political Science Review* 112 (3): 525–41. <https://doi.org/gdwd55>.