## **EECS 126**

Albert Ye

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## 1 Probability Space

### 1.1 Definition

Essentially from 70. Events happen with some probability in a larger probability space that contains all events that can happen.

### 1.2 Axioms of Probability

**Proposition 1** (Axioms) 1. (Positivity)  $P(\omega > 0)$  for any event  $\omega$  in probability space  $\Omega$ .

- 2. (Totality) In any sample space  $\Omega$ ,  $P(\Omega) = 1$ .
- 3. (Additivity) If  $A_1, A_2, \ldots, A_n$  are independent, then

$$\sum_{i=1}^{n} A_i = \bigcup_{i=1}^{n} A_i.$$

From just this, we can get some useful information, such as the union bound.

Theorem 2 (Union Bound)

$$P\left(\bigcup_{i=1}^{n} A_i\right) = \sum_{i=1}^{n} P(A_i).$$

The proof is left as an exercise to the student, probably in the homework.

### 1.3 $\sigma$ -algebra

#### **Definition 3** ( $\sigma$ -algebra)

Given a sample space  $\Omega$ , a set  $\mathcal{F} \subseteq 2^{\Omega}$  is a  $\sigma$ -algebra if:

- 1.  $\Omega \in \mathcal{F}$
- 2. If any event A is in  $\mathcal{F}$ , then its complement  $\Omega \setminus A$  is also in  $\mathcal{F}$ .
- 3. For countably many events  $A_1, A_2, \ldots, A_n, \ldots \in \mathcal{F}$ , their union  $A = \bigcup_{i=1}^{\infty} A_i \in \mathcal{F}$ .

The biggest note is that  $\Omega$  must be in a  $\sigma$ -algebra in order for any of the axioms of probability to apply.

# 2 Conditional Probability

- 2.1 Definition
- 2.2 Total Probability
- 2.3 Bayes' Rule
- 2.4 Continuous Bayes

## 3 It Depends

## 3.1 Independence / (Un)correlation

## 3.2 Conditional Expectation

Notice that E[X|Y] is a random variable, but E[X|Y=y] is a number. We can call E[X|Y] a function g(Y), where then E[X|Y=y]=g(y) is just a value in the function.

## 3.3 Iterated Expectation

## 4 Distributions

## 4.1 Joint Distribution

**Definition 4** (Joint Distribution)

A joint distribution  $f_{X,Y}(x,y)$ 

## 4.2 Marginal Distribution

## 4.3 Derived Distribution

## 5 Random Variables

### 5.1 Discrete

### 5.1.1 Bernoulli

- PMF:  $p_X(k) = \begin{cases} p & k = 1\\ 1 = p & k = 0 \end{cases}$
- $\bullet$  Expected value: p
- Variance: p(1-p).

#### 5.1.2 Binomial

- PMF:  $p_X(k) = \binom{n}{k} p^k (1-p)^{n-k}$  over all  $k \in 0, 1, \dots, n$ .
- $\bullet$  Expected value: np
- Variance: np(1-p).

Run a Bernoulli test n times, find how many are positive.

#### 5.1.3 Geometric

- PMF:  $p_X(k) = (1-p)^{k-1}p$ , for k = 1, 2, ...
- Expected value:  $\frac{1}{p}$
- Variance:  $\frac{1-p}{p^2}$ .

Here, each trial has a p probability of success, and we want to find the # of trials until one success.

#### 5.1.4 Poisson

- PMF:  $p_X(k) = \frac{\lambda^k(e^{-\lambda})}{k!}$ .
- Expected value:  $\lambda$
- Variance:  $\lambda$

Used to simulate arrivals, I guess. More useful later, with Poisson processes.

#### 5.2 Continuous

#### 5.2.1 Uniform

#### 5.2.2 Exponential

### 5.2.3 Gaussian

#### 5.2.4 Joint Gaussian

The main tips for Joint Gaussian are to approach it as a sort of vectorized Gaussians over a certain number N of dimensions. Most of the addition / whatever operations in a Gaussian can be remodeled as a Joint Gaussian.

## 6 Moment Generating Functions

#### Definition 5

The **moment generating function** (also known as a transform) associated with a RV X, is a function  $M_X(s)$  of a scalar parameter s defined by  $M_X(s) = E(e^{sX})$ .

the simpler notation M(S) can be used whenever the underlying random variable X is clear from context. In more detail, when X is a discrete random variable, the corresponding MGF is given by

$$M(s) = \sum_{x} e^{sx} p_X(x).$$

Analogously, when continuous, we just replace the summation with an integral to get

$$M(s) = \int_{-\infty}^{\infty} e^{sx} f_X(x) dx.$$

Just an example so that I know what the reference is here:

### Example 6 (Discrete Example)

Let

$$p_X(x) = \begin{cases} \frac{1}{2} & x = 2\\ \frac{1}{6} & x = 3\\ \frac{1}{3} & x = 5. \end{cases}$$

Then the corresponding transform is

$$M(s) = E(e^{sx}) = \frac{1}{2} + \frac{1}{6}e^{3s} + \frac{1}{3}e^{5s}.$$

#### **Example 7** (Continuous Example)

Let X be an exponential RV with parameter  $\lambda$ :

$$f_X(x) = \lambda e^{-\lambda x}$$
  $x \ge 0$ .

Then,

$$M(s) = \lambda \int_0^\infty e^{sx} e^{-\lambda x} dx$$
$$= \lambda \int_0^\infty e^{(s-\lambda)x} dx$$
$$= \lambda \left( \frac{e^{(s-\lambda)x}}{s-\lambda} \Big|_0^\infty \right)$$
$$= \frac{\lambda}{\lambda - s}.$$

Notice, in above examples, that MGF is a **function** of parameter s, and not a number. We can also find MGF's for functions of X:

### Proposition 8 (MGF of Linear Function of RV)

Let Y = aX + b. Then,

$$M_Y(s) = E(e^{s(aX+b)}) = e^{sb}E(e^{saX}) = e^{sb}M_X(sa).$$

From our previous example, we see that  $M_X(s) = \frac{1}{1-s}$  where X is the exponential distribution

#### 6.1 Moments

Now that we've established what a moment generating function is, now it's time to understand what is being generated.

Let's do a generic MGF

$$M(s) = \int_{-\infty}^{\infty} e^{sx} f_X(x) dx.$$

Now, we take the derivative of this.

$$\frac{d}{ds}M(s) = \frac{d}{ds} \int_{-\infty}^{\infty} e^{sx} f_X(x) dx$$
$$= \int_{-\infty}^{\infty} \frac{d}{ds} e^{sx} f_X(x) dx$$
$$= \int_{-\infty}^{\infty} x e^{sx} f_X(x) dx.$$

When s=0, we have that this evaluates to  $\int_{-\infty}^{\infty} x f_X(x) dx = E(X)$ . If we differentiate n times, then we will get

$$\left. \left( \frac{d^n}{ds^n} M(s) \right|_{s=0} = \int_{-\infty}^{\infty} x^n f_X(x) dx = E(X^n).$$

#### 6.2 Inversion

**Proposition 9** (Inversion Property)

The MGF  $M_X(s)$  associated with an RV X uniquely determines the CDF of X, assuming that  $M_X(s)$  is finite for all s in some interval [-a, a] for positive a.

### 6.3 Sum of Independent Random Variables

#### **Proposition 10**

Addition of independent random variables corresponds to multiplication of transforms.

Proof. Let Z = X + Y.  $M_Z(s) = E(e^{sZ}) = E(e^{s(X+Y)}) = E(e^{sX}e^{sY})$ . Since X, Y are independent,  $e^{sX}$  and  $e^{sY}$  are independent random variables for any fixed s. Thus,  $E(e^{sX}e^{sY}) = E(e^{sX})E(e^{sY}) = M_X(s)M_Y(s)$ .

We can further extend this; if  $X_1, \ldots, X_n$  is a collection of independent random variables and  $Z = X_1 + \cdots + x_n$ , then  $M_Z(s) = M_{X_1}(s) \cdots M_{X_n}(s)$ .

## 7 Concentration Inequalities

Theorem 11 (Markov's Inequality)

$$P(X > a) = \frac{E(X)}{a}.$$

**Theorem 12** (Chebyshev's Inequality)

$$P(|X - E(X)| > a) = \frac{\operatorname{Var}(X)}{a^2}.$$

Used in lieu of confidence interval tests.

## 8 Modes of Convergence

#### 8.1 Pointwise

**Definition 13** (Pointwise Convergence)

Fix  $\omega \in \Omega$ ,  $\{X_n(\omega)\}_{n=1}^{\infty}$  converges **pointwise** if it becomes a real-valued sequence.

Usually, people don't use this because of reasons highlighted in 104.

### 8.2 Almost Sure

**Definition 14** (Almost Sure Convergence)

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\{x_n\}_{n=1}^{\infty} converges almost surely to X if P(\{\omega : \omega \in \Omega, \lim_{n \to \infty} X_n(\omega) = X(\omega)\}) = 1.
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This gets rid of  $\omega$  with probability 0. If you find an  $\omega$  such that convergence doesn't hold, it's fine as long sa  $P(\omega) = 0$ .

### 8.2.1 Checking for Almost Sure Convergence

There are a couple ways to check if some sequence converges almost surely.

### 8.3 In Probability

This is a weaker bound for convergence than almost sure convergence.