MA1500

Introduction to Probability Theory

READING MATERIAL

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Lecture 1 Set Theory

1.1 Elementary set theory

A set is a collection of distinct *elements*.

- If a is an element of the set A, we denote this by $a \in A$.
- If a is not an element of A, we denote this by $a \notin A$.
- The *cardinality* of a set is the number of elements it contains.
- The *empty set* contains no elements, and is denoted by \emptyset .

Algebra is the study of operations and relations.

- The basic relations of set algebra are set inclusion and set equality.
- The basic operations of set algebra are complementation, union and intersection.

1.1.1 Set relations

Definition 1.1

Let A and B be sets.

- (1) If every element of A is also an element of B, we say that A is a *subset* of B. This is denoted by $A \subseteq B$.
- (2) If every element of A is an element of B, and every element of B is an element of A, we say that A and B are equal.

This is denoted by A = B.

(3) If A is a subset of B, but A is not equal to B, we say that A is a proper subset of B. This is denoted by $A \subset B$.

Example 1.2

Let $A = \{a, b\}, B = \{a, b\}$ and $C = \{a, b, c\}$.

- A is a subset of B: $A \subseteq B$,
- A is also equal to B: A = B, and
- A is a proper subset of C: $A \subset C$.

1.1.2 Set operations

Definition 1.3

Let A, B and Ω be sets, with A, $B \subseteq \Omega$.

(1) The union of A and B is the set

$$A \cup B = \{a \in \Omega : a \in A \text{ or } a \in B\}.$$

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(2) The intersection of A and B is the set

$$A \cap B = \{a \in \Omega : a \in A \text{ and } a \in B\}.$$

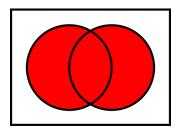
(3) The complement of A is the set

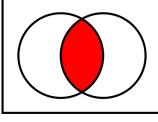
$$A^c = \{ a \in \Omega : a \notin A \}.$$

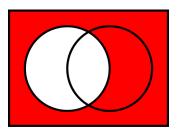
Example 1.4

Let $A = \{a, b\}, B = \{b, c\}$ and $\Omega = \{a, b, c, d\}$.

Then $A \cup B = \{a, b, c\}, A \cap B = \{b\} \text{ and } A^c = \{c, d\}.$







Union

Intersection

Complement

Set Theory		Logic		
Union	$A \cup B$	Disjunction	OR	V
Intersection	$A \cap B$	Conjunction	AND	\wedge
Complement	A^c	Negation	NOT	_

1.1.3 Set algebra

Definition 1.5

- (1) Commutative property.
 - $A \cup B = B \cup A$,
 - $\bullet \ A\cap B=B\cap A.$
- (2) Associative property.
 - $(A \cup B) \cup C = A \cup (B \cup C)$,
 - $(A \cap B) \cap C = A \cap (B \cap C)$.
- (3) Distributive property.
 - $A \cup (B \cap C) = (A \cup B) \cap (A \cup C)$,
 - $A \cap (B \cup C) = (A \cap B) \cup (A \cap C)$.

Remark 1.6

A statement such as $A \cup B \cap C$ is ambiguous.

1.2 De Morgan's laws

Union and intersection swap roles under complementation.

Theorem 1.7

- $(1) (A \cup B)^c = A^c \cap B^c.$
- $(2) (A \cap B)^c = A^c \cup B^c.$

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Proof:		

1.3 Set difference

Definition 1.8

Let A, B and Ω be sets, with $A, B \subseteq \Omega$.

(1) The set difference between A and B is the set

$$A \setminus B = \{ a \in \Omega : a \in A \text{ and } a \notin B \}.$$

(2) The symmetric difference between A and B is the set

$$A \triangle B = (A \setminus B) \cup (B \setminus A).$$

- $A \setminus B$ is the set of points that are in A but not in B.
- $A \triangle B$ is the set of points that are in either A or B, but not both.

Example 1.9

Let $A = \{a, b\}$ and $B = \{b, c\}$. Then

- $A \setminus B = \{a\}$
- $A \triangle B = \{a, c\}.$

1.4 Exercises

Exercise 1.1

- 1. Illustrate the basic set operations using Venn diagrams.
- 2. State and prove De Morgan's laws.

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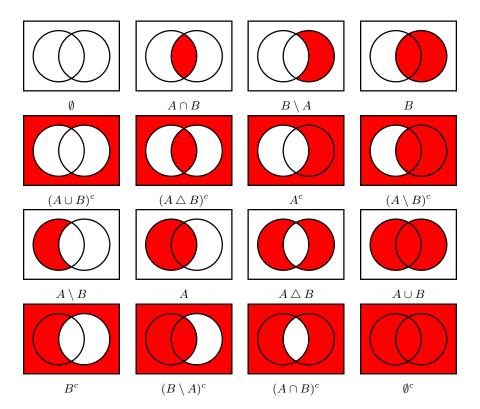


Figure 1.2: Set Operations (A on the left, B on the right).

Lecture 2 Events

A brief history of probability

Games of chance have been played since antiquity, but the mathematical principles of chance and uncertainty were first established only in the 17th century.

1654	Classical principles	Blaise Pascal (1623–1662)
		Pierre de Fermat (1601–1665)
1657	De Ratiociniis in Ludo Aleae	Christiaan Huygens (1629–1695)
1713	$Ars\ Conject and i$	Jakob Bernoulli (1654–1705)
1718	The Doctrine of Chances	Abraham de Moivre (1667–1754)
1812	Theorie Analytique des Probabilites	Pierre de Laplace (1749-1827)
1919	Relative frequency	Richard von Mises (1883–1953)
1933	Modern axiomatic theory	Andrey Kolmogorov (1903–1987)

2.1 Sample spaces

Definition 2.1

- (1) Any process of observation or measurement will be called an *experiment* or *trial*.
- (2) Any experiment whose outcome is uncertain is called a random experiment.
- (3) A random experiment has a set of possible outcomes.
- (4) Each time a random experiment is performed, exactly one of its outcomes will occur.
- (5) The set of all possible outcomes is called the *sample space* of the experiment, denoted by Ω .
- (6) Outcomes are also called *elementary events*, and denoted by $\omega \in \Omega$.

Example 2.2

For any random experiment, the sample space is the set of all possible outcomes:

Experiment	Sample space
A coin is tossed once.	$\Omega = \{H, T\}$
A six-sided die is rolled once.	$\Omega = \{1, 2, 3, 4, 5, 6\}$
A coin is tossed repeatedly until a head occurs.	$\Omega = \{1, 2, 3, \ldots\}$
The height of a randomly chosen student is measured:	$\Omega = [0, \infty)$

2.2 Events

Definition 2.3

- An event A is a subset of the sample space, Ω .
- If outcome ω occurs, we say that event A occurs if and only if $\omega \in A$.
- Two events A and B with $A \cap B = \emptyset$ are called disjoint or mutually exclusive.
- The empty set \emptyset is called the *impossible event*.

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• The sample space Ω is called the *certain event*.

Remark 2.4

- If A occurs and $A \subseteq B$, then B must also occur.
- If A occurs and $A \cap B = \emptyset$, then B does not occur.

Example 2.5

A die is rolled once. The sample space can be represented by $\Omega = \{1, 2, 3, 4, 5, 6\}$.

We may be interested in whether or not the following events occur:

<u>Event</u>	$\underline{\text{Subset}}$
The outcome is the number 1.	$A = \{1\}$
The outcome is an even number.	$A = \{2, 4, 6\}$
The outcome is even but does not exceed 3.	$A = \{2, 4, 6\} \cap \{1, 2, 3\}$
The outcome is not even	$A = \Omega \setminus \{2, 4, 6\}$

2.3 Families of events

Definition 2.6

Let Ω be any set.

- (1) The set of all subsets of Ω is called its *power set*, which we denote by $\mathcal{P}(\Omega)$.
- (2) Any subset of $\mathcal{P}(\Omega)$ is called a family of sets over Ω .

Let Ω be the sample space of some random experiment. If we are interested in events A and B, we must also be interested in whether:

- event A occurs or event B occurs this is the event $A \cup B$,
- event A occurs and event B occurs this is the event $A \cap B$,
- event A does not occur this is the event A^c .

We cannot therefore use arbitrary families of sets over Ω as the basis for investigating random experiments. Instead, we allow only families that are *closed* under certain set operations.

Definition 2.7

A family of sets \mathcal{F} over Ω is said to be

- (1) closed under complementation if $A^c \in \mathcal{F}$ for every $A \in \mathcal{F}$,
- (2) closed under pairwise unions if $A \cup B \in \mathcal{F}$ for every $A, B \in \mathcal{F}$,
- (3) closed under finite unions if $\bigcup_{i=1}^n A_i \in \mathcal{F}$ for every $A_1, A_2, \dots A_n \in \mathcal{F}$,

Definition 2.8

A family of sets \mathcal{F} over Ω is called a *field of sets* over Ω if

- (1) $\Omega \in \mathcal{F}$,
- (2) \mathcal{F} is closed under complementation, and
- (3) \mathcal{F} is closed under pairwise unions.

Example 2.9

A six-sided die is rolled once, and the score is observed. A suitable sample space for this experiment is the set $\Omega = \{1, 2, 3, 4, 5, 6\}$. The power set of Ω will always provide a field of sets to work with. However, suppose we are only interested in whether or not the outcome is an even number. In this case, we need only consider the following family of events:

$$\mathcal{F} = \{\emptyset, \{1, 3, 5\}, \{2, 4, 6\}, \{1, 2, 3, 4, 5, 6\}\}.$$

We can see that \mathcal{F} is a field of sets over Ω , because

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- (1) it contains the sample space $\{1, 2, 3, 4, 5, 6\}$,
- (2) the complement of every set in \mathcal{F} is also contained in \mathcal{F} , and
- (3) the union of any two sets in \mathcal{F} is also contained in \mathcal{F} .

Theorem 2.10 (Properties of fields)

Let \mathcal{F} be a field over Ω . Then

- (1) $\emptyset \in \mathcal{F}$,
- (2) \mathcal{F} is closed under pairwise intersections,
- (3) \mathcal{F} is closed under set differences.

Proof:			

2.4 Terminology

Notation	Set theory	Probability theory
Ω	Universal set	Sample space
$\omega \in \Omega$	Element of Ω	Elementary event, outcome
$A \subseteq \Omega$	Subset of Ω	Event A
$A \subseteq B$	Inclusion	If A occurs, then B occurs
$A \cup B$	Union	A or B occurs
$A \cap B$	Intersection	A and B occur
A^c	Complement of A	A does not occur
$A \setminus B$	Difference	A occurs, but B does not
$A \triangle B$	Symmetric difference	A or B occurs, but not both
Ø	Empty set	Impossible event
Ω	Universal set	Certain event

2.5 Exercises

Exercise 2.1

- 1. Identify a sample space, and the subset corresponding to event A, in each of the following scenarios:
 - (a) A coin is tossed three times. A is the event that at least two heads are obtained.

Answer:

 $\Omega = \{HHH, HHT, HTH, THH, HTT, THT, TTH, TTT\}$ and $A = \{HHH, HHT, HTH, THH\}$. Alternatively, if we are only interested in the number of heads, we could take $\Omega = \{0, 1, 2, 3\}$ and $A = \{2, 3\}$.

(b) A game of football is played. A is the event that the match ends in a draw.

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Answer: $\Omega = \{(a,b): a,b=0,1,2,\ldots\}$ and $A = \{(a,b): a=b\}$ where a and b are the numbers of goals scored by the first and second teams, respectively. Note that this is a (countably) infinite set.

Alternatively, we could take $\Omega = \{W, D, L\}$ and $A = \{D\}$ where W, D, L are respectively the events that the first team wins, draws or loses the game.

(c) A couple have two children. A is the event that both are girls.

Answer:
$$\Omega = \{GG, GB, BG, BB\}$$
 and $A = \{GG\}$.
Alternatively, we could take $\Omega = \{0, 1, 2\}$ and $A = \{2\}$.

(d) A shot hits a circular target of radius 10cm. A is the event that the shot hits within 3cm of the centre.

Answer:
$$\Omega = \{(x,y) : x^2 + y^2 \le 10^2\}$$
 and $A = \{(x,y) : x^2 + y^2 \le 3^2\}$.

- 2. A family of sets \mathcal{F} over Ω is said to be
 - closed under finite unions if $A_1 \cup A_2 \cup \ldots \cup A_n \in \mathcal{F}$ whenever $A_1, A_2, \ldots A_n \in \mathcal{F}$,
 - closed under finite intersections if $A_1 \cap A_2 \cap \ldots \cap A_n \in \mathcal{F}$ whenever $A_1, A_2, \ldots A_n \in \mathcal{F}$.

Suppose that \mathcal{F} is a field of sets over Ω . Show that

(a) \mathcal{F} is closed under finite unions, and that

Answer: Proof by induction. Suppose that \mathcal{F} is closed under unions of n sets (where $n \geq 2$). Let $A_1, A_2, \ldots, A_{n+1} \in \mathcal{F}$. By the inductive hypothesis, $\bigcup_{i=1}^n A_i \in \mathcal{F}$. Thus $\bigcup_{i=1}^{n+1} A_i = \left[\bigcup_{i=1}^n A_i\right] \cup A_{n+1} \in \mathcal{F}$, because \mathcal{F} is closed under pairwise unions.

(b) \mathcal{F} is closed under finite intersections.

Answer: Let $A_1, A_2, \ldots, A_n \in \mathcal{F}$. Then $\bigcap_{i=1}^n A_i = \left[\bigcup_{i=1}^n A_i^c\right]^c$ (De Morgan's laws). Hence $\bigcap_{i=1}^n A_i \in \mathcal{F}$ because \mathcal{F} is closed under complementation and finite unions.

Lecture 3 Probability

3.1 Probability measures

Probability is defined to be a function that assigns numerical value to random events.

Definition 3.1

Let Ω be the sample space of some random experiment, and let \mathcal{F} be a field of sets over Ω . A probability measure on (Ω, \mathcal{F}) is a function

$$\mathbb{P}: \quad \mathcal{F} \quad \to \quad [0,1]$$

$$A \quad \mapsto \quad \mathbb{P}(A)$$

such that $\mathbb{P}(\Omega) = 1$, and for any countable collection of pairwise disjoint events $\{A_1, A_2, \ldots\}$,

$$\mathbb{P}\left(\bigcup_{i=1}^{\infty} A_i\right) = \sum_{i=1}^{\infty} \mathbb{P}(A_i).$$

The triple $(\Omega, \mathcal{F}, \mathbb{P})$ is called a *probability space*.

Remark 3.2

- The second property is called *countable additivity*.
- The number $\mathbb{P}(A)$ is called the *probability* of event $A \in \mathcal{F}$.

Example 3.3

Consider a random experiment in which a fair six-sided die is rolled once.

- A suitable sample space for the experiment is $\Omega = \{1, 2, 3, 4, 5, 6\}$.
- A suitable field of events for the experiment is the power set, $\mathcal{F} = \mathcal{P}(\Omega)$.
- Because the die is fair, a suitable probability measure is given by the function

$$\begin{array}{cccc} \mathbb{P}: \mathcal{F} & \to & [0,1] \\ A & \mapsto & \frac{1}{6}|A|, & & \text{where } |A| \text{ denotes the cardinality of } A. \end{array}$$

	Event	Probability
The outcome is the number 1.	$A = \{1\}$	$\mathbb{P}(A) = 1/6$
The outcome is an even number.	$A = \{2, 4, 6\}$	$\mathbb{P}(A) = 3/6$
The outcome is even but does not exceed 3.	$A = \{2, 4, 6\} \cap \{1, 2, 3\}$	$\mathbb{P}(A) = 1/6$
The outcome is not even	$A = \Omega \setminus \{2, 4, 6\}$	$\mathbb{P}(A) = 3/6$

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Example 3.4

A fair six-sided die is rolled once. If we are only interested in whether the outcome is an odd or even number, we can take

- Sample space: $\Omega = \{1, 2, 3, 4, 5, 6\},\$
- Events: $\mathcal{F} = \{\emptyset, \{1, 3, 5\}, \{2, 4, 6\}, \{1, 2, 3, 4, 5, 6\}\}$
- Probability measure: $\mathbb{P}(\emptyset) = 0$, $\mathbb{P}(\{1,3,5\}) = 1/2$, $\mathbb{P}(\{2,4,6\}) = 1/2$, $\mathbb{P}(\{1,2,3,4,5,6\}) = 1$.

3.2 Properties of probability measures

Theorem 3.5 (Properties of probability measures)

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space, and let $A, B \in \mathcal{F}$.

- (1) Complementarity: $\mathbb{P}(A^c) = 1 \mathbb{P}(A)$.
- $(2) \mathbb{P}(\emptyset) = 0,$

Proof:

- (3) Monotonicity: if $A \subseteq B$ then $\mathbb{P}(A) \leq \mathbb{P}(B)$.
- (4) Addition rule: $\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) \mathbb{P}(A \cap B)$.

3.3 Exercises

Exercise 3.1

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1. What does it mean to say that \mathbb{P} is a probability measure over (Ω, \mathcal{F}) ?

Answer: Bookwork. The symbols Ω (sample space) and \mathcal{F} (field of events) should be defined before giving the definition of \mathbb{P} .

2. Show that $\mathbb{P}(A \cup B) \leq \mathbb{P}(A) + \mathbb{P}(B)$ for any two events A and B.

Answer: First we express A, B and $A \cup B$ as disjoint unions:

$$A = (A \cap B^c) \cup (A \cap B)$$

$$B = (B \cap A^c) \cup (A \cap B)$$

$$A \cup B = (A \cap B^c) \cup (A \cap B) \cup (B \cap A^c)$$

By the additivity property of probability measures,

$$\begin{split} \mathbb{P}(A) &= \mathbb{P}(A \cap B^c) + \mathbb{P}(A \cap B) \\ \mathbb{P}(B) &= \mathbb{P}(B \cap A^c) + \mathbb{P}(A \cap B) \\ \mathbb{P}(A \cup B) &= \mathbb{P}(A \cap B^c) + \mathbb{P}(A \cap B) + \mathbb{P}(B \cap A^c) \end{split}$$

From here, it follows that $\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B)$, and because $\mathbb{P}(A \cap B) \geq 0$ for any two events $A, B\mathcal{F}$, we see that $\mathbb{P}(A \cup B) \leq \mathbb{P}(A) + \mathbb{P}(B)$, as required.

- 3. Let A and B be events such that $\mathbb{P}(A) = 0.4$, $\mathbb{P}(B) = 0.5$ and $\mathbb{P}(A \cup B) = 0.8$. Compute the following probabilities:
 - (a) $\mathbb{P}(A \cap B)$.

Answer:
$$\mathbb{P}(A \cap B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cup B) = 0.4 + 0.5 - 0.8 = 0.1.$$

(b) $\mathbb{P}(A \cup B^c)$.

$$\textbf{Answer:} \ \mathbb{P}(A \cup B^c) = 1 - \mathbb{P}(B \setminus A) = 1 - \left[\mathbb{P}(B) - \mathbb{P}(A \cap B)\right] = 1 - 0.4 = 0.6.$$

- 4. Let A and B be random events, with probabilities $\mathbb{P}(A) = 1/2$ and $\mathbb{P}(B) = 3/4$.
 - (a) Show that $\frac{1}{4} \leq \mathbb{P}(A \cap B) \leq \frac{1}{2}$.

Answer: $A \cap B \subseteq A$ and $A \cap B \subseteq B$ means that:

$$\mathbb{P}(A \cap B) \le \min \{ \mathbb{P}(A), \mathbb{P}(B) \} = \frac{1}{2}.$$

Furthermore, $\mathbb{P}(A \cup B) \leq 1$ means that:

$$\mathbb{P}(A \cap B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cup B) \ge \frac{1}{4}.$$

(b) Show that $\frac{3}{4} \leq \mathbb{P}(A \cup B) \leq 1$.

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Answer: $A \subseteq A \cup B$ and $B \subseteq A \cup B$ means that:

$$\mathbb{P}(A \cup B) \geq \max \left[\mathbb{P}(A), \mathbb{P}(B) \right] = \frac{3}{4}.$$

Furthermore, $\mathbb{P}(A \cup B) \leq 1$ means that:

$$\mathbb{P}(A \cup B) \le \min\{1, \mathbb{P}(A) + \mathbb{P}(B)\} = 1.$$

Lecture 4 Conditional Probability

4.1 Conditional probability

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space, and let $A, B \in \mathcal{F}$ be any two events.

- If B occurs and $A \cap B = \emptyset$, then A cannot occur.
- If B occurs and $B \subseteq A$, then A is certain to occur.
- If B occurs, then A will also occur if and only if the event $A \cap B$ occurs.

Given that B occurs, the probability that A also occurs is $\mathbb{P}(A \cap B)$ expressed as a proportion of $\mathbb{P}(B)$.

Definition 4.1

If $\mathbb{P}(B) > 0$, the conditional probability of A given B is defined to be

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}$$

Remark 4.2

- $\mathbb{P}(A|B) = 0$ whenever $A \cap B = \emptyset$, and
- $\mathbb{P}(A|B) = 1$ whenever $B \subseteq A$.

Example 4.3

Let A and B be two events, with probabilities $\mathbb{P}(A) = 0.3$, $\mathbb{P}(B) = 0.8$ and $\mathbb{P}(A \cap B) = 0.2$. Find the probabilities $\mathbb{P}(A \cup B)$, $\mathbb{P}(A \cap B^c)$, $\mathbb{P}(A|B)$ and $\mathbb{P}(A|B^c)$.

Solution:			

Example 4.4 (The Second Child Paradox)

If we know that a man has two children, and that one of them is a boy, what is the probability that he has two boys?

Solution:

4.2 The partition theorem

Definition 4.5

A partition of a set B is a collection of non-empty sets $\{A_1, A_2, \ldots\}$ such that every element of B lies in exactly one of these sets, or equivalently,

- (1) $A_i \cap A_j = \emptyset$ for all $i \neq j$, and
- (2) $B \subseteq \bigcup_{i=1}^{\infty} A_i$.

Theorem 4.6 (The Partition Theorem)

If $\{A_1, A_2, \ldots\}$ is a partition of B, then

$$\mathbb{P}(B) = \sum_{i=1}^{\infty} \mathbb{P}(B \cap A_i) = \sum_{i=1}^{\infty} \mathbb{P}(B|A_i)\mathbb{P}(A_i)$$

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-	10	vi.

4.3 Bayes' theorem

Lemma 4.7

For any two events A and B such that $\mathbb{P}(B) > 0$,

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(B|A)\mathbb{P}(A)}{\mathbb{P}(B)}.$$

Proof:
Theorem 4.8 (Bayes' Theorem) Let $\{A_1, A_2,\}$ be a partition of an event B and suppose that $\mathbb{P}(B) > 0$. Then
$\mathbb{P}(A_i B) = \frac{\mathbb{P}(B A_i)\mathbb{P}(A_i)}{\sum_{j=1}^{\infty} \mathbb{P}(B A_j)\mathbb{P}(A_j)}$
Proof:
Example 4.9 Bob tries to buy a newspaper every day. He tries in the morning with probability $1/3$, in the evening with probability $1/2$ and forgets completely with probability $1/6$. The probability of successfully buying a newspaper in the morning is $9/10$ (plenty of copies left), and in the evening is $2/10$ (often sold out). If Bob buys a newspaper, what is the probability that he bought it in the morning?
Solution:

4.4 Exercises

Exercise 4.1

- 1. Let A and B be events such that $\mathbb{P}(A)=0.4$, $\mathbb{P}(B)=0.5$ and $\mathbb{P}(A\cup B)=0.8$. Compute the following probabilities:
 - (a) $\mathbb{P}(A \cap B)$.

Answer: $\mathbb{P}(A \cap B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cup B) = 0.4 + 0.5 - 0.8 = 0.1.$

(b) $\mathbb{P}(A \cup B^c)$.

Answer:
$$\mathbb{P}(A \cup B^c) = 1 - \mathbb{P}(B \setminus A) = 1 - \lceil \mathbb{P}(B) - \mathbb{P}(A \cap B) \rceil = 1 - 0.4 = 0.6.$$

(c) $\mathbb{P}(A \mid B)$.

Answer:
$$\mathbb{P}(A|B) = \mathbb{P}(A \cap B)/\mathbb{P}(B) = 0.1/0.5 = 0.2.$$

(d) $\mathbb{P}(A \mid A \cup B)$.

Answer:
$$\mathbb{P}(A|A \cup B) = \mathbb{P}(A)/\mathbb{P}(A \cup B) = 0.4/0.8 = 0.5.$$

2. Let A, B and C be events such that $\mathbb{P}(A) = 0.7$, $\mathbb{P}(B) = 0.6$, $\mathbb{P}(C) = 0.5$, $\mathbb{P}(A \cap B) = 0.4$, $\mathbb{P}(A \cap C) = 0.3$, $\mathbb{P}(B \cap C) = 0.2$ and $\mathbb{P}(A \cap B \cap C) = 0.1$.

Compute the following probabilities:

(a) $\mathbb{P}(A \cup B)$.

Answer:
$$\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B) = 0.7 + 0.6 - 0.4 = 0.9.$$

(b) $\mathbb{P}(A|B)$.

Answer:
$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)} = \frac{0.4}{0.6} = \frac{2}{3}$$
.

(c) $\mathbb{P}(A \mid A \cup B)$.

$$\textbf{Answer:} \ \mathbb{P}(A|A\cup B) = \frac{\mathbb{P}[A\cap (A\cup B)]}{\mathbb{P}(A\cup B)} = \frac{\mathbb{P}(A)}{\mathbb{P}(A\cup B)} = \frac{0.7}{0.9} = \frac{7}{9}.$$

(d) $\mathbb{P}(A \cup B \cup C)$.

Answer: By the inclusion-exclusion principle,

$$\mathbb{P}(A \cup B \cup C) = [\mathbb{P}(A) + \mathbb{P}(B) + \mathbb{P}(C)] - [\mathbb{P}(A \cap B) + \mathbb{P}(A \cap C) + \mathbb{P}(B \cap C)] + \mathbb{P}(A \cap B \cap C)$$
$$= (0.7 + 0.6 + 0.5) - (0.4 + 0.3 + 0.2) + 0.1 = 1.$$

(e) $\mathbb{P}(A^c \cap B^c \cap C)$.

Answer: Because
$$A^c \cap B^c \cap C = (A \cup B)^c \cap C$$
, the sets $A \cup B$ and $A^c \cap B^c \cap C$ form a partition of $A \cup B \cup C$. Hence $\mathbb{P}(A^c \cap B^c \cap C) = \mathbb{P}(A \cup B \cup C) - \mathbb{P}(A \cup B) = 1.0 - 0.9 = 0.1$

(f) $\mathbb{P}(A^c \cap B^c \cap C | A \cup B)$.

Answer: Because A^c and A are disjoint, $\mathbb{P}(A^c \cap B^c \cap C | A \cup B) = 0$.

- 3. A student has three opportunities to pass an exam. The probability of failing the first attempt is 0.6; the probability of failing the second attempt, given that they have failed the first is 0.75, and the probability of failing the third attempt, given that they have failed the first and second is 0.4.
 - (a) What is the probability that the student eventually passes the exam.

Answer: Let F_i denote the event that the student fails at the *i*th attempt, so that

$$\mathbb{P}(F_1) = 0.6$$
, $\mathbb{P}(F_2|F_1) = 0.75$ and $\mathbb{P}(F_3|F_1 \cap F_2) = 0.4$

The probability that the student fails all three attempts is

$$\mathbb{P}(F_1 \cap F_2 \cap F_3) = \mathbb{P}(F_1)\mathbb{P}(F_2|F_1)\mathbb{P}(F_3|F_1 \cap F_2) = 0.6 \times 0.75 \times 0.4 = 0.18$$

Hence, the probability that the student eventually passes is 1 - 0.18 = 0.82.

(b) What are the respective probabilities of passing at the first, second and third attempts.

Answer: The probability that the student passes on the first attempt is $1-\mathbb{P}(F_1)=1-0.6=0.4$. The probability that the student takes a second test is $\mathbb{P}(F_1)=0.6$. If the second test is taken, the (conditional) probability that the student passes it is $1-\mathbb{P}(F_2|F_1)=1-0.75=0.25$. Hence, the probability that the student passes on the second attempt is $0.6\times0.25=0.15$. Similarly, the probability that the student takes the third test is $\mathbb{P}(F_1\cap F_2)=\mathbb{P}(F_1)\mathbb{P}(F_2|F_1)=0.6\times0.75=0.45$. If the third test is taken, the (conditional) probability that the student passes it is $1-\mathbb{P}(F_3|F_1\cap F_2)=1-0.4=0.6$. Hence, the probability that the student passes on the third attempt is $0.45\times0.6=0.27$. (The probability of eventually passing is 0.4+0.15+0.27=0.82, which agrees with the answer to part (a).)

4. An insurance company divides its customers into three categories: 60% of customers are classed as low-risk, 30% as moderate-risk and 10% as high-risk. The probabilities that low-risk customers, moderate-risk customers and high-risk customers make a claim in any given year are 0.01, 0.1 and 0.5 respectively. Given that a customer makes a claim this year, what is the probability that the customer is in the high-risk category?

Answer: Let L denote the event that the customer is low-risk, M the event that the customer is moderate-risk, and H the event that the customer is high-risk:

$$\mathbb{P}(L) = 0.6$$
, $\mathbb{P}(M) = 0.3$, $\mathbb{P}(H) = 0.1$.

Let C be the event that the customer makes a claim this year:

$$\mathbb{P}(C \mid L) = 0.01, \quad \mathbb{P}(C \mid M) = 0.1, \quad \mathbb{P}(C \mid H) = 0.5.$$

We need to find $\mathbb{P}(H \mid C)$:

$$\mathbb{P}(H \mid C) = \frac{\mathbb{P}(H \cap C)}{\mathbb{P}(C)} = \frac{\mathbb{P}(C \mid H)\mathbb{P}(H)}{\mathbb{P}(C)}.$$

The events $\{L, M, H\}$ form a partition of sample space (the set of all customers), so by the law of total probability,

$$\mathbb{P}(C) = \mathbb{P}(C \,|\, L) \mathbb{P}(L) + \mathbb{P}(C \,|\, M) \mathbb{P}(M) + \mathbb{P}(C \,|\, H) \mathbb{P}(H),$$

and hence

$$\mathbb{P}(H \mid C) = \frac{\mathbb{P}(C \mid H) \mathbb{P}(H)}{\mathbb{P}(C \mid L) \mathbb{P}(L) + \mathbb{P}(C \mid M) \mathbb{P}(M) + \mathbb{P}(C \mid H) \mathbb{P}(H)},$$

which is Bayes' theorem. The probability that a customer is in the high-risk category, given that the customer makes a claim this year, is

$$\mathbb{P}(H \mid C) = \frac{(0.5 \times 0.1)}{(0.01 \times 0.6) + (0.1 \times 0.3) + (0.5 \times 0.1)} = 0.5814 \quad \text{(approx.)}.$$

- 5. A horse has three opportunities to clear a fence. The probability that it fails at the first attempt is 0.4. The probability that it fails at the second attempt, given that it has failed at the first attempt, is 0.3. The probability that it fails at the third attempt, given that it has failed at the first and second attempts, is 0.8.
 - (a) What is the probability that the horse eventually clears the fence?
 - (b) What are the respective probabilities that the horse clears the fence at the first, second and third attempts.

Answer: Let A, B and C denote the events that the horse fails at the first, second and third attempts, respectively.

$$\mathbb{P}(A) = 0.4$$
, $\mathbb{P}(B|A) = 0.3$, $\mathbb{P}(C|A \cap B) = 0.8$.

(1) The probability that the horse fails all three attempts is

$$\mathbb{P}(A \cap B \cap C) = \mathbb{P}(A)\mathbb{P}(B|A)\mathbb{P}(C|A \cap B) = 0.4 \times 0.3 \times 0.8 = 0.096,$$

so the probability that the horse eventually succeeds is 1 - 0.096 = 0.904.

- (2) The probability that the horse succeeds at the first attempt is $1 \mathbb{P}(A) = 0.6$.
 - The probability that the horse has a second attempt is $\mathbb{P}(A) = 0.4$. In this case the (conditional) probability of success is $1 \mathbb{P}(B|A) = 1 0.3 = 0.7$. Hence the probability of success on the second attempt is $0.4 \times 0.7 = 0.28$.
 - The probability that the horse has a third attempt is $\mathbb{P}(A \cap B) = \mathbb{P}(B|A)\mathbb{P}(A) = 0.3 \times 0.4 = 0.12$. In this case the (conditional) probability of success is $1 \mathbb{P}(C|A \cap B) = 1 0.8 = 0.2$. Hence the probability of success on the third attempt is $0.12 \times 0.2 = 0.024$.

Check: the probability of success on either the first, second or third attempt is 0.6+0.28+0.024 = 0.904, which agrees with our answer to part (i).

Lecture 5 Independence

5.1 Independence

If the probability that event A occurs is not affected by whether or not event B occurs, then $\mathbb{P}(A|B) = \mathbb{P}(A)$. In such cases, we say that events A and B are independent:

Definition 5.1

Two events A and B are said to be independent if $\mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B)$.

Lemma 5.2

If A and B are independent, then A and B^c are also independent.

Proof:			

Example 5.3

A fair die is rolled once. Let A be the event that the outcome is an even number, and let B be the event that the outcome is divisible by 3. Are A and B independent?

Solution:			

5.2 Pairwise independence and total independence

Definition 5.4

A family of events $\{A_1, A_2, \ldots\}$ is said to be

- (1) pairwise independent if $\mathbb{P}(A_i \cap A_j) = \mathbb{P}(A_i)\mathbb{P}(A_j)$ for all $i \neq j$.
- (2) totally independent if, for every finite subset $\{B_1, B_2, \dots, B_m\} \subseteq \{A_1, A_2, \dots\}$, we have

$$\mathbb{P}(B_1 \cap B_2 \cap \ldots \cap B_m) = \mathbb{P}(B_1)\mathbb{P}(B_2)\cdots\mathbb{P}(B_m).$$

This can also be written as $\mathbb{P}\left(\bigcap_{j=1}^{m} B_j\right) = \prod_{j=1}^{m} \mathbb{P}(B_j)$.

Example 5.5 (de Méré's Paradox)

Show that you are more likely to obtain a six in 4 rolls of a single fair die, than to obtain a double-six in 24 rolls of two fair dice.

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Solution:		

Example 5.6

Consider a sample space $\Omega = \{1, 2, 3, 4\}$ where each outcome is equally likely. Let $A = \{1, 2\}$, $B = \{1, 3\}$ and $C = \{1, 4\}$. Show that $\{A, B, C\}$ is pairwise independent, but not totally independent.

Solution:		 	 	

5.3 Exercises

Exercise 5.1

1. A fair six-sided die is rolled repeatedly. How many times should it be rolled to ensure that the probability of getting a six is at least 0.8?

Answer: For convenience, let p = 1/6 denote the probability that the die shows a six.

Let A_1 be the event that we observe a six on the first roll, A_2 that we observe a six on the second roll, and so on.

Let A be the event that we observe at least one six in n rolls. The complementary event A^c is the event that we do not observe a six on any of the n rolls. This can be written as

$$A^c = A_1^c \cap A_2^c \cap \ldots \cap A_n^c$$

If we assume that the rolls are independent of each other, the set of events $\{A_1, A_2, \ldots, A_n\}$ is totally independent, and hence $\{A_1^c, A_2^c, \ldots, A_n^c\}$ is also totally independent. The probability of getting at least one six in n rolls is therefore given by

$$\mathbb{P}(A) = 1 - \mathbb{P}(A^c)
= 1 - \mathbb{P}(A_1^c \cap A_2^c \cap \dots \cap A_n^c)
= 1 - \mathbb{P}(A_1^c) \mathbb{P}(A_2^c) \cap \dots \cap A_n^c)
= 1 - [1 - \mathbb{P}(A_1)][1 - \mathbb{P}(A_2)] \cdots [1 - \mathbb{P}(A_n)]
= 1 - (1 - p)^n$$

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Let n denote the number of rolls. To find the required value of n, we need that $\mathbb{P}(A) \geq 0.8$, i.e.

$$1 - (1 - p)^n \ge 0.8 \Rightarrow (1 - p)^n \le 0.2$$

$$\Rightarrow n \log(1 - p) \le \log 0.2$$

$$\Rightarrow n \ge \frac{\log 0.2}{\log(1 - p)} = \frac{\log 0.2}{\log(5/6)} \approx 8.827$$

Therefore we should roll the die at least n = 9 times.

2. A multiple choice test has five questions, with each question having four alternative choices. At least three questions must be answered correctly to pass the test. If a candidate chooses her answers at random, what is the probability that she passes the test? State any assumptions you make.

Answer: Assumption: The choices are independent of each other. Let A_i be the event that exactly i questions are answered correctly (i = 0, 1, 2, 3, 4, 5), and let A be the event that the student passes the test. Then

$$\mathbb{P}(A) = \mathbb{P}(A_3 \cup A_4 \cup A_5)$$

$$= \mathbb{P}(A_3) + \mathbb{P}(A_4) + \mathbb{P}(A_5) \qquad \text{(because events } A_3, A_4 \text{ and } A_5 \text{ are disjoint)}\}$$

$$= 10(0.25)^3 (0.75)^2 + 5(0.25)^4 (0.75) + (0.25)^5$$

$$= 0.1035.$$

3. Two fair dice are rolled. Show that the event that their sum is 7 is independent of the score shown on the first die.

Answer: For every $j \in \{1, 2, 3, 4, 5, 6\}$,

- $\mathbb{P}(\text{first die shows } j \text{ and sum is } 7) = \frac{1}{36}$
- $\mathbb{P}(\text{first die shows } j)\mathbb{P}(\text{sum is } 7) = \frac{1}{6} \times \frac{1}{6} = \frac{1}{36}$

Alternatively, let A be the event that the sum is 7 and let B be the event that any score is shown on the first die. Then $\mathbb{P}(A) = 1/6$ and $\mathbb{P}(B) = 1$, while $\mathbb{P}(A \cap B) = 1/6$, so A and B are independent.

- 4. A fair die is rolled twice, each roll being independent of the other. Let A be the event that the first roll shows 3, let B be the event that the second roll shows 4, and let C be the event that the total of the two rolls is 7.
 - (a) Define a suitable sample space, and identify the subsets corresponding to events A, B and C.

Answer:

- $\Omega = \{(a, b) : 1 \le a, b \le 6\}$
- $A = \{(3, b) : 1 \le b \le 6\}$
- $B = \{(a,4) : 1 \le a \le 6\}$
- $C = \{(a,b) : 1 \le a, b \le 6 \text{ and } a+b=7\}$

The sample space Ω contains 36 outcomes, and the events A, B and C each contains 6 outcomes. Because each outcome is equally likely, $\mathbb{P}(A) = \mathbb{P}(B) = \mathbb{P}(C) = 1/6$.

(b) Show that $\{A, B, C\}$ is pairwise independent but not totally independent.

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Answer: To show that the set $\{A, B, C\}$ is pairwise independent, we need to show that any two events chosen from the set are independent of each other. For A and B, their intersection $A \cap B$ is the event $\{(3,4)\}$, consisting of the single outcome (3,4). This means that $\mathbb{P}(A \cap B) = 1/36$ and therefore

$$\mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B)$$

so A and B are pairwise independent. Similarly, $A \cap C = \{(3,4)\}$ which implies that $\mathbb{P}(A \cap C) = \mathbb{P}(A)\mathbb{P}(C)$, and $B \cap C = \{(3,4)\}$ implies that $\mathbb{P}(B \cap C) = \mathbb{P}(B)\mathbb{P}(C)$. Hence, the set $\{A,B,C\}$ is pairwise independent. However, the intersection of all three events is also $\{(3,4)\}$ so $\mathbb{P}(A \cap B \cap C) = 1/36$ and hence

$$\mathbb{P}(A \cap B \cap C) \neq \mathbb{P}(A)\mathbb{P}(B)\mathbb{P}(C)$$

so the set $\{A, B, C\}$ is *not* independent.

5. A coin has probability p of showing heads. Let q_n be the probability that in n independent tosses, a head is observed an even number of times (for this question, take 0 to be an even number). Using the partition theorem, show that

$$q_n = p(1 - q_{n-1}) + (1 - p)q_{n-1}$$
 for any $n \ge 1$.

Answer: Let $n \geq 1$, let A_1, \ldots, A_n be any sequence of tosses, and let $N(A_1, \ldots, A_n)$ be the number of heads in the sequence. Using the fact that A_n is independent of the previous tosses A_1, \ldots, A_{n-1} , it follows that the number $N(A_1, \ldots, A_n)$ is even if and only if

- A_n is heads and $N(A_1, \ldots, A_{n-1})$ is odd, which occurs with probability $p(1-q_{n-1})$, or
- A_n is tails and $N(A_1, \ldots, A_{n-1})$ is even, which occurs with probability $(1-p)q_{n-1}$

Hence

$$q_n = p(1 - q_{n-1}) + (1 - p)q_{n-1}$$

as required.