Hypothesis testing - a review Statistical Learning - Modern Inference

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Darwin data

- Charles Darwin collected data on Zea mays plants
- The plants were descended from the same parents and planted at the same time.
- Half of the plants were self-fertilized, and half were cross-fertilized
- Planted in pairs in different pots
- Purpose of the experiment: compare heights

	Pot	Cross	Self
1	I	23.500	17.375
2	I	12.000	20.375
3	I	21.000	20.000
4	П	22.000	20.000
5	П	19.125	18.375
6	П	21.500	18.625
7	Ш	22.125	18.625
8	Ш	20.375	15.250
9	Ш	18.250	16.500
10	Ш	21.625	18.000
11	Ш	23.250	16.250
12	IV	21.000	18.000
13	IV	22.125	12.750
14	IV	23.000	15.500
15	IV	12.000	18.000

Questions

	type	average height
1	Cross	20.19
2	Self	17.57

- Is the difference in heights too large to have occurred by chance?
- **2** Can we estimate the height increase, and assess the uncertainty of our estimate?

Hypothesis testing

 ${\it H}_{\rm 0}$: There is no difference in height between cross-fertilized and self-fertilized plants

Hypothesis testing is a type of stochastic proof by contradiction

Deterministic proof by contradiction

- Assume a proposition, the opposite of what you think about, i.e. the opposite conclusion of your theorem
- 2 Write down a sequence of logical steps/math
- 3 Derive a contradiction
- Occide that the proposition is false (which implies that the theorem is true)

Stochastic proof by contradiction

- Set H_0 (the proposition)
- 2 Collect data (which is noisy)
- 3 Derive an apparent contradiction (i.e. if H_0 is true, then this data is very weird)
- **4** Hence we reject H_0 ; this is called a "discovery"

Hypothesis testing is stochastic because we might make errors:

Type I (false discoveries) and Type II (missed discoveries)

Galton model

Galton considered a model where the height of a self-fertilized plant is

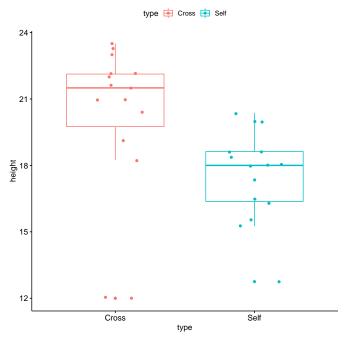
$$Y = \mu + \sigma \varepsilon$$

and of a cross-fertilized plant is

$$X = \mu + \theta + \sigma \epsilon$$

where μ , θ and σ are unknown parameters, and ε and ϵ are independent random variables with mean 0 and variance 1

Self-fertilized plants: Y_1, \ldots, Y_{15} i.i.d. as Y Cross-fertilized plants: X_1, \ldots, X_{15} i.i.d. as X.



Two-sample t-test

- **1** Is the average height increase $\theta \neq 0$?
- **2** Can we estimate θ , and assess the uncertainty of our estimate?

If we assume that ε and ϵ have a N(0,1) distribution, we can use

Two Sample t-test

```
data: height by type
t = 2.4371, df = 28, p-value = 0.02141
alternative hypothesis: true difference in means is not equ
95 percent confidence interval:
    0.4173433   4.8159900
sample estimates:
mean in group Cross mean in group Self
    20.19167    17.57500
```

Fisher model

Comparison of different pairs would involve differences in humidity, growing conditions, and lighting, which are not of interest, whereas *comparison within pairs* would depend only on the type of fertilization. Fisher considered the model

$$Y_i = \mu_i + \sigma \varepsilon_i, \quad X_i = \mu_i + \theta + \sigma \varepsilon_i, \quad i = 1, \dots, n$$

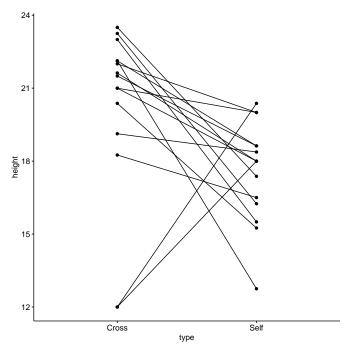
The parameter μ_i represents the effects of the planting conditions for the *i*th pair, and ε_i and ϵ_i are independent random variables with mean 0 and variance 1.

The μ_i could be eliminated by using the differences

$$D_i = X_i - Y_i$$

which have mean θ and variance $2\sigma^2$.





Paired t-test

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If we assume that \varepsilon and \epsilon have a N(0,1) distribution, we can use
One Sample t-test
data: differences
t = 2.148, df = 14, p-value = 0.0497
alternative hypothesis: true mean is not equal to 0
95 percent confidence interval:
 0.003899165 5.229434169
sample estimates:
mean of x
 2.616667
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P-values

- Choose a test statistic T = t(Y), large values of which cast doubts on H_0
- Observe the data yobs, realization of Y

•

$$p_{\rm obs} = P_0(T \ge t_{\rm obs})$$

where $t_{\rm obs} = t(y_{obs})$ and P_0 is the probability under H_0

P-value null distribution

- $p_{\rm obs} = 1 F_0(t_{\rm obs})$, where F_0 is the null distribution function of T, supposed to be continuous and invertible.
- One interpretation of $p_{\rm obs}$ stems from the corresponding random variable $P=1-F_0(T)$
- The null distribution of P is Uniform(0,1): for any $u \in (0,1)$,

$$P_0(P \le u) = P_0(F_0^{-1}(1-u) \le T) = 1 - F_0(F_0^{-1}(1-u)) = u$$

Valid p-values

We have a *valid test* if its p-value is uniformly distributed under H_0 , i.e.

$$P_0(P \le u) = u \quad \forall u \in (0,1) \tag{1}$$

or more generally if the p-value is stochastically dominated by the uniform distribution under H_0 , i.e.

$$P_0(P \le u) \le u \quad \forall u \in (0,1) \tag{2}$$

One- and two-sided tests

- Suppose that we have a test statistic T (with continuous distribution), small and large values of which indicate a departure from H₀
- Calculate

$$p_{\mathrm{obs}}^- = \mathrm{P}_0(T \le t_{\mathrm{obs}}), \quad p_{\mathrm{obs}}^+ = \mathrm{P}_0(T \ge t_{\mathrm{obs}})$$

• The *p*-value is

$$p_{\mathrm{obs}} = 2 \min(p_{\mathrm{obs}}^-, p_{\mathrm{obs}}^+)$$

This follows because the null distribution of $Q = \min(P^-, P^+)$ is

$$Q = \min(1 - U(0,1), U(0,1)) = U(0,1/2)$$

thus the null distribution of 2Q is U(0,1)



Nonparametric tests

Consider a more general matched pair model for the Darwin data:

$$Y_i = \mu_i + \sigma_i \varepsilon_i, \quad X_i = \mu_i + \theta + \tau_i \varepsilon_i, \quad i = 1, \dots, n$$

The height differences may be written as

$$D_i = \theta + (\tau_i \epsilon_i - \sigma_i \epsilon_i)$$

If we assume that ε_i and ϵ_i are independent and symmetrically distributed around 0, then D_i is symmetrically distributed around θ , i.e.

$$D_i - \theta \stackrel{d}{=} \theta - D_i$$
 $i = 1, \dots, n$

Sign test

• If $H_0: \theta = 0$ is true, then the probability that D_i falls on either side of 0 is 1/2 (because D_i is symmetric around 0)

$$T = \sum_{i=1}^{n} \mathbb{1}\{D_i > 0\}$$

• Under H_0 , $T \sim \text{Binomial}(n, 1/2)$

$$p_{\mathrm{obs}}^+ = \mathrm{P}_0(T \geq t_{\mathrm{obs}}) = \sum_{k=-1}^n \ \binom{n}{k} \frac{1}{2^n}$$

$$p_{\text{obs}}^{-} = P_0(T \le t_{\text{obs}}) = \sum_{k=0}^{t_{\text{obs}}} \binom{n}{k} \frac{1}{2^n}$$

In a discrete problem $p_{\rm obs}$ is $q_{\rm obs} = \min(p_{\rm obs}^-, p_{\rm obs}^+)$ plus the achievable p-value from the other tail of the distribution nearest to but not exceeding $q_{\rm obs}$.

Exact binomial test

data: 13 and 15
number of successes = 13, number of trials = 15,
p-value = 0.007385
alternative hypothesis:
true probability of success is not equal to 0.5
95 percent confidence interval:
0.5953973 0.9834241
sample estimates:
probability of success
0.8666667