Clustering market regimes using the Wasserstein distance [1]

Seminar for the Phd course in Quantitative Finance (SNS)

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Link for the code: https://github.com/alebati3



Clusters analysis

- Cluster analysis or Clustering is an unsupervised technique used to group objects into clusters.
- It's fundamental to define a way to quantify the similarity among objects.
 - It extends the concept of distance.
- "The definition of an optimal clustering is not well defined, and in the case of financial data, this is certainly true." [1]
- "Heuristically, we would like individual clusters to contain objects that are similar to each other whilst being distinct from objects in other clusters". [1]

k-means algorithm

• Suppose $X = \{(x_1, \dots, x_N) : x_i \in V\}$, where $(V, \|\cdot\|_V)$ is a normed vector space. We further assume that each $x_i = (x_i^1, \dots, x_i^d)$ has been standardized coordinate-wise, that is,

$$\mathbb{E}[(x_i^j)_{1 \leq i \leq N}] = 0 \quad \text{and} \quad \mathsf{Var}((x_i^j)_{1 \leq i \leq N}) = 1 \quad \text{for } j = 1, \dots, d.$$

• The *k-means clustering algorithm* is an unsupervised vector quantization method which assigns elements of X to k disjoint clusters. Each of these clusters is defined by central elements $\bar{x} = \{\bar{x}_j\}_{j=1,\dots,k}$ called *centroids*, which are initially sampled from X.

k-means algorithm

• At each step $n \in \mathbb{N}$ of the algorithm, one first calculates the *nearest* neighbours

$$C_I^n := \left\{ x_i \in X : \arg\min_{j=1,\dots,k} d(x_i, \bar{x}_j^{n-1}) = I \right\}$$

associated to each \bar{x}_{l}^{n-1} for l = 1, ..., k.

• Each set C_l^n is then aggregated into a new centroid x_l^n for $l=1,\ldots,k$ via a function $\alpha:2^V\to V$, so

$$\bar{x}_I^n := \alpha(C_I^n)$$
 for $I = 1, \dots, k$.

• In the classical k-means on \mathbb{R}^d , we take as new centroid the barycenter of C_l

$$\alpha(C_l) = \left(\frac{1}{|C_l|} \sum_{x_j \in C_l} x_j\right)_{1 < j < d}.$$

where, $|C_I|$ denotes the cardinality of the set C_I .

k-means algorithm

• For a given tolerance level $\epsilon>0$ and a loss function $I:V^k\times V^k\to [0,+\infty)$, the k-means algorithm terminates at step $n\in\mathbb{N}$ if the stopping condition

$$I(\bar{x}^n, \bar{x}^{n-1}) < \epsilon$$

is satisfied.

• The loss function / is given by

$$I(x,y) = \sum_{i=1}^{k} ||x_i - y_i||_V,$$

• At the end, the algorithm outputs the final clusters $C^* = \{C_l^n\}_{l=1,\dots,k}$ and the k-quantization $\bar{x}^n = \{x_l^n\}_{l=1,\dots,k}$.



The market regime clustering problem (MRCP)

Given the return series of a security price $\mathbf{r} = (r_0, r_1, \dots, r_N)$

• The MRCP is defined as the task of classifying segments of return series $(I_i)_{i=1}^M$, where

$$I_i = (r_i^1, \dots, r_i^n)$$
 for $n \in \mathbb{N}$

Any vector $l_i \in \mathbb{R}^n$ can be associated to an empirical probability measure

$$\mu_i = \frac{1}{n} \sum_{j=1}^n \delta_{r_i^j}$$

for i = 1, ..., M with n atoms. Thus the problem of classifying market regimes is equivalent to assigning a label to probability measures $(\mu_i)_{i=1}^M$.

Problem setting and notation

Given the return series $\mathbf{r} = (r_j)_{j=0}^{N-1}$, where $r_j = \log(s_{j+1}) - \log(s_j)$, we define the segments of the return series as follows:

• if $h_1, h_2 \in \mathbb{N}$ with $h_1 > h_2$ then

$$I_i = (r_{(h_1 - h_2)(i-1)}, \dots, r_{(h_1 - h_2)(i-1) + h_1})$$
 for $i = 1, \dots, M$

where M is the maximum number of partitions with length h_1+1 that can be extracted from the return series with sliding offset parameter h_2 .

• In general the different partitions l_i may overlap.

Problem setting and notation

• We can associate to each segment of data l_i the empirical measure μ_i for i = 1, ..., M. This gives us a family of measures:

$$\mathcal{K} = \{\mu_1, \dots, \mu_M\}.$$

It is this family ${\cal K}$ which will be the subject of our clustering algorithm.

• Let Q^j be the function which extracts the j-th order statistic of l_i , for $j=1,\ldots,h_1+1$. Then, the cumulative distribution function of the *empirical measure* μ_i associated to l_i is defined as:

$$\mu_i((-\infty,r]) = \frac{1}{h_1+1} \sum_{j=1}^{h_1+1} \chi(Q^j(l_i) \leq r),$$

where $\chi: \mathbb{R} \to [0,1]$ is the indicator function.



p-Wasserstein distance (W_p)

Main properties:

• W_p is a metric in the set of probability measures having the first p moments finite, denoted by $\mathcal{P}_p(\mathbb{R}^d)$.

• Convergence with respect to W_p is equivalent to the usual weak convergence of measures plus convergence of the first p-th moments.

• Intuition: connection with the optimal transport problem.

Connection with the optimal transport problem

- The optimal transport problem aims to transform the distribution of mass $\mu(x)$ into another distribution $\nu(y)$ on a space X, while minimizing the cost.
- The function $c(x, y) \ge 0$ represents the cost of moving mass from point x to point y. A transport plan $\gamma(x,y)$ defines how much mass is moved from x to y.
- The total cost of the transport plan γ is:

$$\int \int c(x,y)\gamma(x,y)\,dx\,dy = \int c(x,y)\,d\gamma(x,y).$$

• The optimal transport plan minimizes the total cost:

$$C = \min_{\gamma \in \Gamma(\mu, \nu)} \int c(x, y) \, d\gamma(x, y).$$

• If $c(x,y) = d(x,y)^p$, then C corresponds to the W_p .

W_p for empirical probability measure

• Suppose $\mu, \nu \in \mathcal{P}_p(\mathbb{R}^d)$ and let d=1. Moreover, suppose that μ, ν are absolutely continuous with respect to the Lebesgue measure on \mathbb{R} . Then, the p-Wasserstein distance $W_p(\mu, \nu)$ is given by

$$W_p(\mu,\nu) = \left(\int_0^1 \left|F_{\mu}^{-1}(z) - F_{\nu}^{-1}(z)\right|^p dz\right)^{1/p},$$

where the quantile function $F_{\mu}^{-1}:[0,1)
ightarrow\mathbb{R}$ is defined as

$$F_{\mu}^{-1}(z) = \inf\{x : F_{\mu}(x) \ge z\}.$$

• If μ, ν are empirical measures with equal numbers of atoms $N \in \mathbb{N}$, with $(\alpha_i)_{1 \leq i \leq N}$ and $(\beta_i)_{1 \leq i \leq N}$ their corresponding order statistics, then

$$W_{p}(\mu,\nu)^{p} = \frac{1}{N} \sum_{i=1}^{N} |\alpha_{i} - \beta_{i}|^{p}.$$

W_p for empirical probability measure

- Calculating the Wasserstein distance between two empirical measures can be done in linear time, assuming the atoms of each measure are already sorted ascending. If not, is an $\mathcal{O}(N \log N)$ operation, where N is the number of atoms.
- Suppose that $\{\mu_i\}_{1\leq i\leq M}$ are a family of empirical probability measures, each with order statistics $\{\alpha_j^i\}_{1\leq j\leq N}$. We can define the Wasserstein barycenter as the probability measure $\bar{\mu}$ that minimizes the sum of p-Wasserstein distances to each μ_i .
- ullet In particular $ar{\mu}$ is charaterazed by the following the order statistics

$$\mathbf{a}_j = \mathsf{Median}\left(lpha_j^1, \dots, lpha_j^M
ight) \quad \mathsf{for} \ j = 1, \dots, \mathit{N}.$$

Wasserstein k-means algorithm

Set of vectors
$$\longrightarrow$$
 $\mathcal{K} = \{\mu_1, \dots, \mu_M\}$ $\|\cdot\|_V \longrightarrow p$ -Wasserstein distance

Aggregation function ______ Wasserstein barycenter to update centroids

The last specification we need to make is regarding the loss function.

 The most natural choice is to replace the distance induced by the norm on V with p-Wasserstein distance

$$I(\bar{\mu}^{n-1}, \bar{\mu}^n) = \sum_{i=1}^k W_p(\bar{\mu}_i^{n-1}, \bar{\mu}_i^n).$$

where $\bar{\mu}^n = (\bar{\mu}_i^n)_{1 \leq i \leq k}$ are the centroids obtained after step n of the Wasserstein k-means algorithm

Alternative clustering algorithms as benchmarks

- k-means with statistical moments (Moment k-means)
- Hidden Markov models
- Agglomerative Hierarchical clustering
 - W_p based (W-Hierchical clustering)
 - statistical moments based (Moment-Hierchical clustering)

Moment k-means

- A natural and more classical approach to clustering regimes may involve studying the first $p \in \mathbb{N}$ raw moments associated to each measure $\mu \in \mathcal{K}$
- each empirical probability measure μ_i is mapped in vector of \mathbb{R}^p , whose components are the corresponding first p moments

$$\varphi^{p}(\mu_{i}) = \left(\int_{\mathbb{R}} x^{n} \mu_{i}(dx)\right)_{1 \leq n \leq p},$$

• Thus, for a given $p \ge 1$, we obtain

$$\varphi^{p}(\mathcal{K}) = \{ \varphi^{p}(\mu_{1}), \dots, \varphi^{p}(\mu_{M}) : \varphi^{p}(\mu_{i}) \in \mathbb{R}^{p} \text{ for } i = 1, \dots, M \}.$$

• After standardising each element of $\varphi^p(\mathcal{K})$ component-wise, we obtain a clustering set on \mathbb{R}^p , which we can apply the standard k-means algorithm to.

Hidden Markov models

- A more classical approach to market regime clustering involves fitting a *Hidden Markov model (HMM)* to observed time series data **r**
 - We assume the existence of $k \in \mathbb{N}$ hidden latent states $\{1, \ldots, k\}$ which govern the dynamics of \mathbf{r} . The transitions between the latent states are assumed Markovian.
 - The aim is to determine the sequence of hidden states that most likely produced the observed data.
- HMM does not cluster segments of return series; instead, it associates
 a given latent state to a single return at time t.

Hidden Markov models

Model parameters:

- **Transition probabilities A**: probabilities of transitioning from one hidden state to another.
- Emission probability densities B: represented by $f(r_t|z_t)$, where z_t is the given latent state and r_t is the observed return.
 - $f(x|z_t)$ are assumed to be gaussians (Gaussian HMM)
- Initial state distribution π : the probabilities of starting in a particular hidden state.

The triplet $\lambda = (\mathbf{A}, \mathbf{B}, \pi)$ defines the Hidden Markov model.

Solution:

- find the best model λ^* by maximizing the Likelihood $\mathcal{L}(\lambda|\mathbf{r})$.
- Compute the most probable sequence of hidden states using the *Viterbi algorithm*.

Agglomerative Hierarchical clustering

- Start with the points as individual clusters;
- at each step, merge most similar pair of clusters until only one cluster (or k clusters) left;
- a key point is to define the Inter-Cluster Similarity.
 - Different approaches to defining the similarity between clusters disinguish the different algorithms
- The standard algorithm for Hierarchical Agglomerative clustering (HAC) has a complexity of $\mathcal{O}(N^3)$, where N is the number of objects.

Agglomerative Hierarchical clustering

Let's consider two different clusters, A and B, and a distance d to quantify the similarity among objects.

There are several popular way to define the Inter-Cluster distance:

Complete-linkage clustering:

$$\max_{a \in A, b \in B} d(a, b)$$

• Minimum or single-linkage clustering:

$$\min_{a \in A, b \in B} d(a, b)$$

Average linkage clustering:

$$\frac{1}{|A|\cdot|B|}\sum_{a\in A}\sum_{b\in B}d(a,b)$$

W and Moment Hierarchical

W Hierarchical

• We refer to the same setting of the W k-means.

Moment Hierarchical

- As for the Moment k-means each μ_i is mapped to a vector of \mathbb{R}^p , whose components are the first p moments. In order to quantify the similarity we take the Euclidean distance.
- In this case we define another criterion for merging clusters, the so called Ward's method, a variance-minimizing approach. In particular we merge the clusters that minimize the following quantity

$$\sum_{x \in A \cup B} \|x - \mu_{A \cup B}\|^2 - \sum_{x \in A} \|x - \mu_A\|^2 - \sum_{x \in B} \|x - \mu_B\|^2$$

where μ_C represents the centroid of the cluster C.

Validation of the clustering algorithms

"The validation of clustering structures is the most difficult and frustrating part of cluster analysis.

Without a strong effort in this direction, cluster analysis will remain a black art accessible only to those true believers who have experience and great courage." [2]

Validation on synthetic data

Generating synthetic price path helps to define a validation procedure.

- We know both the underlying probabilistic structure and the regime change periods a priori.
- It's possible to evaluate both how accurately either algorithm is classifying sequences of returns into regimes, and how closely the centroids $\{\bar{\mu}_I\}$ of each cluster correspond to the true distributions $\{P_I\}$ associated to the synthetic data.

We assume 2 different regimes: a standard regime (regime-off) and the regime change (regime-on). They can be interpreted from an economic point of view as bull and bear regimes.

Validation on synthetic data

The methodology is as follows:

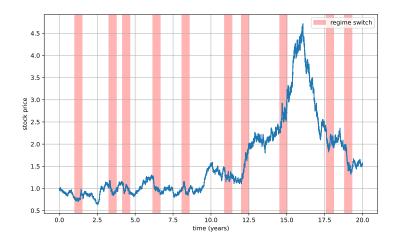
• For a given time interval [0, T] with $T \in \mathbb{N}$, we define a mesh so that each time increment roughly represents one market hour. That is, with $n := 252 \times 7$, we set

$$\Delta = \left\{ \left[\frac{i-1}{n}, \frac{i}{n} \right] : i = 1, 2, \dots, nT \right\}.$$

then $\Delta t = \frac{1}{n}$ represents 1 market hour.

- Next, we define the number of regime changes $r \in \mathbb{N}$ we wish to observe. We specify their starting points and the length of each interval.
- **Note**: we simulate a path over T=20 years with r=10 regime changes, randomly chosen with a duration of 1.5 years.

Validation with synthetic data



Validation on synthetic data

- For $i=0,\ldots,N-1$, associate to each log-return r_i the empirical measures $M_i=\{\mu_{j(i)},\ldots,\mu_{j(i)+v_i-1}\}$ it was a member of.
- $j(i) \in \mathbb{N}$ is the first measure that r_i is a member of. With $h_1 = 35$ and $h_1 = 28$, one has that $v_i \in [1, 6]$. Note that if the overlap parameter $h_2 = 0$, then $v_i = 1$.
- We then calculate which cluster each $\mu \in M_i$ is associated to, which gives us our predicted labels $\bar{y}^i = \{\bar{k}_1, \dots, \bar{k}_v\}$. We then aggregate these labels into the row vector

$$\bar{Y}^i = \left(\sum_{j=1}^{\nu} \chi_{\{x=l\}}(\bar{k}_j)\right)_{l=1}^k$$
 for $i = 0, \dots, N-1$,

where k=2 is the number of clusters. In what follows we assume the assignment $\bar{k}=1$ corresponds to the standard regime and $\bar{k}=2$ the regime change.

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Accuracy scores

For a given vector of log-returns \mathbf{r} and cluster assignments $C = \{C_l\}_{l=1}^2$, we define:

• regime-off accuracy score (ROFS)

$$\mathsf{ROFS}(\mathbf{r}, C) = \frac{\sum_{r_i \in \mathit{off}} \bar{Y}_1^i}{\sum_{r_i \in \mathit{off}} \sum_{k=1,2} \bar{Y}_k^i} \in [0,1]$$

• regime-on accuracy score (RONS)

$$\mathsf{RONS}(\mathbf{r},C) = \frac{\sum_{r_s^i \in on} \bar{Y}_2^i}{\sum_{r_i \in on} \sum_{k=1,2} \bar{Y}_k^i} \in [0,1]$$

total accuracy (TA)

$$\mathsf{TA}(\mathbf{r},C) = \frac{\sum_{r_i \in \textit{off}} \bar{Y}_1^i + \sum_{r_i \in \textit{on}} \bar{Y}_2^i}{\sum_{i=1}^{N-1} \sum_{k=1,2} \bar{Y}_k^i} \in [0,1]$$

Models for generating price paths

• Geometric Brownian motion (GBM)

Merton jump diffusion model (MJD)

Heston model

Validation on real data

- "Heuristically, we would like individual clusters to contain objects that are similar to each other whilst being distinct from objects in other clusters" [1]
- Evaluating derived k-means clusters is typically done by evaluating the final total cluster variation $TC(C^*)$.
 - Associate to each C_l its centroid \bar{x}_l for $l=1,\ldots,k$. Then, for a given C_l , the within-cluster variation is defined as

$$WC(C_I) = \sum_{x \in C_I} ||x - \bar{x}_I||_V^2 \text{ for } I = 1, ..., k.$$

• Then, the total cluster variation is given by

$$TC(C) = \sum_{i=1}^{k} WC(C_i)$$

• Since the total cluster variation depends on V, one cannot use it in evaluation between clusterings on different choices of V.

Maximum mean discrepancy (MMD)

Let (X, d) be a metric space and F be a class of functions
 f: X → ℝ. If μ, ν ∈ P(X) are Borel measures, the maximum mean discrepancy (MMD) between μ and ν is defined as

$$\mathsf{MMD}[\mathcal{F},\mu,\nu] := \sup_{f \in \mathcal{F}} \left(\mathbb{E}_{\mu}[f(x)] - \mathbb{E}_{\nu}[f(y)] \right).$$

ullet If as ${\mathcal F}$ we take the Gaussian kernel

$$\kappa_G: \mathbb{R}^d \times \mathbb{R}^d \to [0, +\infty), \quad \kappa_G(x, y) = \exp\left(-\frac{\|x - y\|_{\mathbb{R}^d}^2}{2\sigma^2}\right)$$

the MMD is a metric on $\mathcal{P}(\mathcal{X})$.

 This last property makes MMD the building block of the validation procedure.

Maximum Mean Discrepancy (MMD)

• If we draw samples $\mathbf{x} = (x_1, \dots, x_n)$ and $\mathbf{y} = (y_1, \dots, y_m)$ where $x_i \sim \mu$ for $i = 1, \dots, n$ and $y_j \sim \nu$ for $j = 1, \dots, m$, a biased empirical estimate of the previous MMD is given by:

$$MMD_{b}^{2}[\kappa_{G}, \mathbf{x}, \mathbf{y}] = \left[\frac{1}{n^{2}} \sum_{i,j=1}^{n} k(x_{i}, x_{j}) - \frac{2}{mn} \sum_{i,j=1}^{m,n} k(x_{i}, y_{j}) + \frac{1}{m^{2}} \sum_{i,j=1}^{m} k(y_{i}, y_{j}) \right].$$

Cluster validation via MMD

Between-cluster evaluation

Repeat the following 3 steps for each clustering algorithm:

- given the two cluster C_1 , C_2 , draw $n \in \mathbb{N}$ pairwise samples $(\mu_i, \nu_i) \in C_1 \times C_2$ for $i = 1, \dots, n$.
 - We represent each empirical measure $\mu_i, \nu_i \in \mathcal{P}_p(\mathbb{R})$ by its corresponding vector of log-returns $\mathbf{x}_i, \mathbf{y}_i$.
- For each pair we compute $MMD_b[\kappa_G, \mathbf{x}_i, \mathbf{y}_i]$, where we choose a gaussian kernel with $\sigma = 0.1$.
- Finally, compute the similarity score

$$\mathsf{Sim} = \mathsf{Median}\left(\left(\mathsf{MMD}^2_b[\kappa_G, \mathbf{x}_i, \mathbf{y}_i]\right)_{1 \leq i \leq n}\right),$$

Then,

 we compare the associated distribution of the MMD among the histograms generated from all the methods by reporting the similarity score.

Cluster validation via MMD

Within-cluster evaluation is performed much in the same way as the between-cluster case. Repeat the following steps for each clustering algorithm:

- for each cluster C_l , l=1,2, we draw $n \in \mathbb{N}$ pairwise samples $(\mu_l^1, \mu_l^2) \in C_l \times C_l$ and evaluate the biased MMD.
- We report the similarity score associated to the empirical distribution of each within-cluster MMD and plot the resulting histograms.

So, for each algorithm we get 1 between-cluster similarity score and 2 within-cluster similarity scores.

• "We stated that a clustering algorithm was successful if the *self-similarity* and *distinctness* of derived clusters were appropriately traded off against each other." [1]

References



