

CONTRIBUTEURS

Note The moments of a mixture, are the mixture of the moments.

Lesson 25 : Estimator Quality

Sample Statistics

Sample Mean Unbiased estimator of the true mean μ .

$$\bar{x} = \frac{\sum_{i=1}^n x_i}{n}$$

Sample Variance Unbiased estimator of the true variance σ^2 .

$$s^2 = \frac{\sum (x_i - \bar{x})^2}{n - 1}$$

Empirical Variance Biased estimator of the true variance σ^2 .

$$\hat{\sigma}^2 = \frac{\sum (x_i - \bar{x})^2}{n}$$

Lesson 27 : Method of Moments

Notation

μ'_k k^{th} moment centred around 0, $\mu'_k = E[X^k]$.

$\hat{=}$ Set equal.

Exponential Distribution There is only one parameter θ which is the mean, we set

$$\hat{\theta} \hat{=} \mu'_1.$$

Gamma Distribution We have :

$$E[X] = \alpha\theta \hat{=} \bar{x}$$

$$\therefore \hat{\theta} = \frac{\hat{\sigma}^2}{\bar{x}} = \frac{\hat{\mu}'_2 - \hat{\mu}'_1^2}{\hat{\mu}'_1}$$

Pareto Distribution

$$E[X] = \frac{\theta}{\alpha - 1} \hat{=} \hat{\mu}'_1$$

$$\therefore \hat{\alpha} = \frac{2(\hat{\mu}'_2 - \hat{\mu}'_1^2)}{(\hat{\mu}'_2 - 2\hat{\mu}'_1^2)}$$

Lognormal Distribution

$$\hat{\mu} = 2 \ln(\hat{\mu}'_1) - 0.5 \ln(\hat{\mu}'_2)$$

$$\text{Var}(X) = \alpha\theta^2 \hat{=} \hat{\sigma}^2$$

$$\hat{\alpha} = \frac{\bar{x}^2}{\hat{\sigma}^2} = \frac{\hat{\mu}'_1^2}{\hat{\mu}'_2 - \hat{\mu}'_1^2}$$

$$E[X^2] = \frac{2\theta^2}{(\alpha - 1)(\alpha - 2)} \hat{=} \hat{\mu}'_2$$

$$\hat{\theta} = \frac{\hat{\mu}'_1 \hat{\mu}'_2}{\hat{\mu}'_2 - 2\hat{\mu}'_1^2}$$

$$\hat{\sigma}^2 = \ln(\hat{\mu}'_2) - 2 \ln(\hat{\mu}'_1)$$

Uniform Distribution

$$E[X] = \frac{\theta}{2}$$

$$\therefore \hat{\theta} = 2\hat{\mu}'_1$$

Lesson 28 : Percentile Matching

Note : Exams don't typically ask a lot of percentile matching questions, thus it's not really worth memorizing each distribution's formulas as it can easily be about a random distribution.

Exponential Distribution

$$\hat{\theta} = \frac{-\pi_g}{\ln(1 - g)}$$

Weibull Distribution

$$\hat{\tau} = \frac{\ln(\ln(1 - g_1) / \ln(1 - g_2))}{\ln(\pi_{g_1} / \pi_{g_2})}$$

$$\hat{\theta} = \frac{\pi_{g_1}}{\sqrt[3]{-\ln(1 - g_1)}}$$

Lognormal Distribution (use the percentiles of a normal distribution z_p)

$$\hat{\sigma} = \frac{\ln(\pi_{g_2}) - \ln(\pi_{g_1})}{z_{g_2} - z_{g_1}}$$

$$\hat{\mu} = \ln(\pi_{g_1}) - z_{g_1} \hat{\sigma}$$

Truncated data

For $(X|X > d)$:

$$F_X(x|X > d) = \frac{F_X(x) - F_X(d)}{S_X(d)}$$

$$S_X(x|X > d) = \frac{S_X(x)}{S_X(d)}$$

Lesson 29 :

Lesson 30 : MLE Special Techniques

If the likelihood function is of the form $\mathcal{L}(\gamma) = \gamma^{-a} e^{-b/\gamma}$ then $\hat{\gamma}^{\text{MLE}} = \frac{b}{a}$.

If the likelihood function is of the form $\mathcal{L}(\lambda) = \lambda^a e^{-\lambda b}$ then $\hat{\lambda}^{\text{MLE}} = \frac{a}{b}$.

If the likelihood function is of the form $\mathcal{L}(\lambda) = \theta^a (1 - \theta)^b$ then $\hat{\theta}^{\text{MLE}} = \frac{a}{a+b}$.

Exponential

$$\hat{\theta} = \frac{\sum_{i=1}^{n+c} (x_i - d_i)}{n}$$

Weibull with a fixed τ

$$\hat{\theta} = \sqrt[n]{\frac{\sum_{i=1}^{n+c} x_i^\tau - \sum_{i=1}^{n+c} d_i^\tau}{\tau}}$$

Lognormal (use the percentiles of a normal distribution z_p)

$$\hat{\sigma} = \sqrt{\frac{\sum_{i=1}^n \ln^2 x_i}{n} - \hat{\mu}^2} \quad \hat{\mu} = \frac{\sum_{i=1}^n \ln x_i}{n}$$

Uniform $(0, \theta)$ for individual data

$$\hat{\theta} = \max x_i$$

Uniform $(0, \theta)$ for grouped data

c_j Upper bound of highest finite interval

n_j Number of observations below c_j

$$\hat{\theta} = c_j \left(\frac{n}{n_j} \right)$$

Inverse exponential

$$\hat{\theta} = \frac{n}{\sum_{i=1}^n (1/x_i)}$$

Two-parameter Pareto, fixed θ

$$\hat{\alpha} = -\frac{n}{K} \quad K = \sum_{i=1}^{n+c} \{\ln(\theta + d_i) - \ln(\theta + x_i)\}$$

Single-parameter Pareto, fixed θ

$$\hat{\alpha} = -\frac{n}{K} \quad K = \sum_{i=1}^{n+c} \{\ln \max(\theta, d_i) - \ln x_i\}$$

Beta, fixed $\theta, \beta = 1$

$$\hat{\alpha} = -\frac{n}{K} \quad K = \sum_{i=1}^{n+c} \{\ln(x_i)\} - n \ln \theta$$

Beta, fixed $\theta, \alpha = 1$

$$\hat{\beta} = -\frac{n}{K} \quad K = \sum_{i=1}^{n+c} \{\ln(\theta - x_i)\} - n \ln \theta$$